## Ridge Regression Explained

## James Taylor

In the multivariate regression setting, the estimated coefficients of correlated predictor variables are very unstable - a large positive coefficient for one variable can be offset by a large negative coefficient for a correlated variable. Ridge regression mitigates this problem by introducing an  $(L^2)$  regularization term which penalizes large values of the model's (non-intercept) coefficients.

**The setup.** Suppose we are in the traditional multivariate regression setting with one quantitative response Y and p predictors  $X_1, X_2, \ldots, X_p$ , with an (approximately) linear relationship between them:

$$Y \approx \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_n X_n$$

Say we have training data consisting of n observations. We want to use these data to find approximations  $b_i$  for  $\beta_i$ ; we will then be able to predict future values of Y given  $(X_1, \ldots, X_p) = (x_1, \ldots, x_p)$  via the formula

$$\hat{y} = b_0 + b_1 x_1 + \dots + b_p x_p$$

Our estimates for the model coefficients are going to be

$$\operatorname{argmin}_{b} \left[ RSS + \alpha \sum_{i=1}^{p} b_{i}^{2} \right]$$

where  $\alpha \geq 0$  is some user-supplied constant. For  $\alpha = 0$  ridge regression is just ordinary least squares regression. But as  $\alpha$  gets larger, the regularization term becomes an increasingly larger part of the optimization function, thus forcing the optimal  $(b_1, \ldots, b_p)$  to have smaller  $L^2$  norm. Clearly, as  $\alpha \to \infty$  the  $L^2$  norm of  $(b_1, \ldots, b_p)$  goes to 0.

Note: Traditionally, one centers the data, estimates  $\beta_0$  with  $b_0 = \frac{1}{n} \sum_{i=1}^n y_i$ , and the performs ridge regression without intercept to derive the equation  $b = (X^T X + \alpha I)^{-1} X^T y$ , where X is the  $n \times p$  matrix containing our observed predictor data. We are not going to go this route.

Minimizing the ridge regression equation. Letting our n observations be  $(x_i, y_i)$  for i = 1, ..., n where  $x_i = (x_{i1}, x_{i2}, ..., x_{in})$ , we want to minimize

$$RSS + \alpha \sum_{i=1}^{p} b_i^2 = \sum_{i=1}^{n} (y_i - \hat{y_i})^2 + \alpha \sum_{i=1}^{p} b_i^2$$
$$= \sum_{i=1}^{n} (y_i - (b_0 + b_1 x_{i1} + \dots + b_p x_{ip}))^2 + \alpha \sum_{i=1}^{p} b_i^2$$

We minimize this function by differentiating with respect to  $b_j$ , setting this expression equal to zero, and simplifying the resulting equation to be of the form  $\rho_j \cdot b = \gamma_j$ . We do this for each j, and then solve the following matrix equation for b:

$$\rho b = \gamma$$

where  $\rho$  is the  $(p+1)\times(p+1)$  matrix whose jth row is  $\rho_{j-1}$ , and  $\gamma$  is the (p+1)-vector with jth entry  $\gamma_{j-1}$ .

More explicitly, for j = 0 we have

$$\frac{\partial}{\partial b_0} \left( RSS + \alpha \sum_{i=1}^p b_i^2 \right) = \sum_{i=1}^n -2(y_i - (b_0 + b_1 x_{i1} + \dots + b_p x_{ip})) = 0$$

$$\implies \sum_{i=1}^n (y_i - (b_0 + b_1 x_{i1} + \dots + b_p x_{ip})) = 0$$

$$\implies nb_0 + b_1 \sum_{i=1}^n x_{i1} + \dots + b_p \sum_{i=1}^n x_{ip} = \sum_{i=1}^n y_i$$

and for  $j = 1, \ldots, p$  we have

$$\frac{\partial}{\partial b_{j}} \left( RSS + \alpha \sum_{i=1}^{p} b_{i}^{2} \right) = \sum_{i=1}^{n} -2x_{ij}(y_{i} - (b_{0} + b_{1}x_{i1} + \dots + b_{p}x_{ip})) + 2\alpha b_{j} = 0$$

$$\implies \sum_{i=1}^{n} x_{ij}(y_{i} - (b_{0} + b_{1}x_{i1} + \dots + b_{p}x_{ip})) - \alpha b_{j} = 0$$

$$\implies \alpha b_{j} + \sum_{i=1}^{n} x_{ij}(b_{0} + b_{1}x_{i1} + \dots + b_{p}x_{ip}) = \sum_{i=1}^{n} x_{ij}y_{i}$$

$$\implies b_{0} \sum_{i=1}^{n} x_{ij} + b_{1} \sum_{i=1}^{n} x_{ij}x_{i1} + \dots + b_{j}(\alpha + \sum_{i=1}^{n} x_{ij}^{2}) + \dots + b_{p} \sum_{i=1}^{n} x_{ij}x_{ip} = \sum_{i=1}^{n} x_{ij}y_{i}$$

These equations give rise to the following matrix equation:

$$\begin{pmatrix} n & \sum_{i=1}^{n} x_{i1} & \sum_{i=1}^{n} x_{i2} & \cdots & \sum_{i=1}^{n} x_{ip} \\ \sum_{i=1}^{n} x_{i1} & \alpha + \sum_{i=1}^{n} x_{i1}^{2} & \sum_{i=1}^{n} x_{i1} x_{i2} & \cdots & \sum_{i=1}^{n} x_{i1} x_{ip} \\ \sum_{i=1}^{n} x_{i2} & \sum_{i=1}^{n} x_{i2} x_{i1} & \alpha + \sum_{i=1}^{n} x_{i2}^{2} & \cdots & \sum_{i=1}^{n} x_{i2} x_{ip} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \sum_{i=1}^{n} x_{ip} & \sum_{i=1}^{n} x_{ip} x_{i1} & \sum_{i=1}^{n} x_{ip} x_{i2} & \cdots & \alpha + \sum_{i=1}^{n} x_{ip}^{2} \end{pmatrix} \begin{pmatrix} b_{0} \\ b_{1} \\ b_{2} \\ \vdots \\ b_{p} \end{pmatrix} = \begin{pmatrix} \sum_{i=1}^{n} y_{i} \\ \sum_{i=1}^{n} x_{i1} y_{i} \\ \sum_{i=1}^{n} x_{i2} y_{i} \\ \vdots \\ \sum_{i=1}^{n} x_{ip} y_{i} \end{pmatrix}$$

Let  $x_{*j}$  be the *n*-vector consisting of all *n* observations of the jth predictor  $X_j$ . If we set  $x_{*0}$  to be the *n*-vector containing all 1's, then this equation becomes

$$\begin{pmatrix} x_{*0} \cdot x_{*0} & x_{*0} \cdot x_{*1} & x_{*0} \cdot x_{*2} & \cdots & x_{*0} \cdot x_{*p} \\ x_{*1} \cdot x_{*0} & \alpha + x_{*1} \cdot x_{*1} & x_{*1} \cdot x_{*2} & \cdots & x_{*1} \cdot x_{*p} \\ x_{*2} \cdot x_{*0} & x_{*2} \cdot x_{*1} & \alpha + x_{*2} \cdot x_{*2} & \cdots & x_{*2} \cdot x_{*p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ x_{*p} \cdot x_{*0} & x_{*p} \cdot x_{*1} & x_{*p} \cdot x_{*2} & \cdots & \alpha + x_{*p} \cdot x_{*p} \end{pmatrix} b = \begin{pmatrix} x_{*0} \cdot y \\ x_{*1} \cdot y \\ x_{*2} \cdot y \\ \vdots \\ x_{*p} \cdot y \end{pmatrix}$$

Notice that the first matrix is  $X^TX + diag(0, \alpha, \dots, \alpha)$ , and the RHS is  $X^Ty$ , where X is the  $n \times (p+1)$  matrix

$$\begin{pmatrix} 1 & x_{11} & \cdots & x_{1p} \\ 1 & x_{21} & \cdots & x_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & \cdots & x_{np} \end{pmatrix}$$

So our matrix equation is now

$$(X^TX + diag(0, \alpha, \dots, \alpha))b = X^Ty$$

and thus

$$b = (X^T X + diag(0, \alpha, \dots, \alpha))^{-1} X^T y$$

Note that, depending on how large X is, inverting  $X^TX + diag(0,\alpha,\ldots,\alpha)$  could be extremely resource intensive. A way around this is to compute the QR decomposition of  $X^TX + diag(0,\alpha,\ldots,\alpha)$ ; that is, we find an orthogonal matrix Q and an upper triangular matrix R such that  $X^TX + diag(0,\alpha,\ldots,\alpha) = QR$ . (Note: A QR decomposition exists for every matrix.) We'll still need to solve  $QRb = X^Ty$ , but since Q is orthogonal we have  $Q^{-1} = Q^T$ , and thus we only need to solve

$$b = R^{-1}Q^T X^T y = R^{-1}(XQ)^T y$$

So now the only matrix we need to invert is the triangular matrix R, which is a much simpler task.