

DeepLearning HW1

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1 Problem a

See `hw1a.py`

2 Problem b

See `hw1b.py`

3 Problem c

3.1 i

Let $y = g(x)$, since g is monotonic, we have

$$f_Y(y) = \left| \frac{d}{dy}(g^{-1}(y)) \right| \cdot f_X(g^{-1}(y))$$

Take $g(x) = -\frac{1}{\lambda} \ln(x)$ into the formula

$$f_Y(y) = |\lambda| e^{-\lambda y} f_X(e^{-\lambda y})$$

If $\lambda > 0$:

$$f_Y(y) = \begin{cases} \lambda e^{-y} & y \geq 0 \\ 0 & otherwise \end{cases}$$

else

$$f_Y(y) = \begin{cases} -\lambda e^{-y} & y \leq 0 \\ 0 & otherwise \end{cases}$$

3.2 ii

$$\begin{aligned} p(X = x) &= \int p(X = x, Y = y) dy \\ &= \int_0^1 3(xy^2 + yx^2) dy \\ &= x + \frac{3x^2}{2}, 0 \leq x \leq 1 \end{aligned}$$

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$$\begin{aligned} E(x) &= \int x \cdot p(x) dx \\ &= \int_0^1 x^2 + \frac{3}{2}x^3 dx \\ &= \frac{17}{24} \end{aligned}$$

$$\begin{aligned} E(y) &= \int y \cdot p(y) dy \\ &= \int_0^1 y^2 + \frac{3}{2}y^3 dy \\ &= \frac{17}{24} \end{aligned}$$

$$\begin{aligned} E(xy) &= \int_0^1 \int_0^1 xy \cdot 3(xy^2 + yx^2) dx dy \\ &= \frac{1}{2} \end{aligned}$$

4 Problem d

4.1 i

Likelihood:

$$\begin{aligned} p(X|\mu, \Sigma) &= p(x_1, x_2, \dots, x_m|\mu, \Sigma) \\ &= \prod_{i=1}^m p(x_i|\mu, \Sigma) \end{aligned}$$

Convert to log-likelihood:

$$L(X|\mu, \Sigma) = \sum_{i=1}^m \log\left(\frac{1}{(2\pi)^{n/2} \sqrt{|\Sigma|}}\right) \exp\left(-\frac{1}{2}(x_i - \mu)^T \Sigma^{-1} (x_i - \mu)\right)$$

Optimize L over μ is equivalent to solve the equation:

$$\frac{\partial}{\partial \mu} \sum_{i=1}^m -\frac{1}{2}(x_i - \mu)^T \Sigma^{-1}(x_i - \mu) = 0$$

which can be simplified as:

$$\frac{\partial}{\partial \mu} \sum_{i=1}^m -\frac{1}{2}(x_i^T \Sigma^{-1} x_i - 2x_i^T \Sigma^{-1} \mu + \mu^T \Sigma^{-1} \mu) = 0$$

Using relationship

$$\frac{\partial X^T A X}{\partial X} = X^T (A + A^T)$$

The equation can be further simplified as

$$\sum_{i=1}^m (x_i^T - \mu^T) \Sigma^{-1} = 0$$

So

$$\mu = \frac{\sum_{i=1}^m x_i}{m}$$

Now calculate Σ . Simplify again over log-likelihood w.r.t Σ :

$$L = -\frac{mn}{2} \log 2\pi + \frac{m}{2} \log |A| - \frac{1}{2} \sum_{i=1}^m \text{tr}[(x - \mu)(x - \mu)^T A]$$

Using relationship

$$\begin{aligned} \frac{\partial \log |A|}{\partial A} &= (A^{-1})^T \\ \frac{\partial \log |BA|}{\partial A} &= B^T \end{aligned}$$

to derive the derivative w.r.t Σ

$$\frac{\partial L}{\partial \Sigma} = \frac{m}{2} \Sigma - \frac{1}{2} \sum_{i=1}^m (x - \mu)(x - \mu)^T = 0$$

Thus

$$\Sigma = \frac{1}{m} \sum_{i=1}^m (x - \mu)(x - \mu)^T$$

4.2 ii

Check the biasness of μ_{ML} :

$$\begin{aligned} E(\mu_{ML}) - \mu &= \frac{1}{m} \sum_{n=1}^m E(x_i) - \mu \\ &= \frac{1}{m} \sum_{n=1}^m \mu - \mu \\ &= 0 \end{aligned}$$

Therefore, μ_{ML} is unbiased.

Now check the biasness of Σ_{ML} :

$$\begin{aligned} E(\Sigma_{ML}) &= \frac{1}{m} \sum_{i=1}^m E[(x_i - \frac{1}{m} \sum_{j=1}^m x_j)(x_i^T - \frac{1}{m} \sum_{j=1}^m x_j^T)] \\ &= \frac{1}{m} \left\{ \sum_{i=1}^m E[x_i x_i^T] - \sum_{i=1}^m E[\frac{1}{m} x_i \sum_{k=1}^m x_k^T] - \sum_{i=1}^m E[\frac{1}{m} \sum_{j=1}^m x_j x_i^T] + \sum_{i=1}^m E[\frac{1}{m^2} \sum_{j=1}^m \sum_{k=1}^m x_j x_k^T] \right\} \\ &= \frac{1}{m} \{ m(\mu\mu^T + \Sigma) - (m\mu\mu^T + \Sigma) \} \\ &= \frac{m-1}{m} \Sigma \end{aligned}$$

Clearly, Σ_{ML} is biased.