

STOR 455

STATISTICAL METHODS I

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Diagnostics for residuals

- Model: $Y_i = \beta_0 + \beta_1 X_i + \xi_i$
- Predicted values: $\hat{Y}_i = b_0 + b_1 X_i$
- Residuals: $e_i = Y_i - \hat{Y}_i$
 - The book recommends to standardize the residuals.

$$h_{i,i} = \frac{1}{n} + \frac{(x_i - \bar{x})^2}{SSX}$$

$$r_i = \frac{\hat{e}_i}{\hat{\sigma} \sqrt{1 - h_{i,i}}}$$

Residuals

- The e_i should be similar to the ξ_i
 - The model assumes ξ_i iid $N(0, \sigma)$
- Similarly the r_i should be similar to the ξ_i/σ
 - The model assumes ξ_i/σ iid $N(0, 1)$

What do we learn?

- Is the relationship linear?
- Is the variance a constant?
- Are there outliers?
- Are the errors normal?
- Are the errors dependent?

Is the Relationship Linear?

- Scatter plot of Y vs X
- Scatter plot of r vs X
- Scatter plot of r vs \hat{Y}
 - Plot of r vs ... emphasize deviations from linear pattern
 - Never plot r vs Y

Do it in SAS

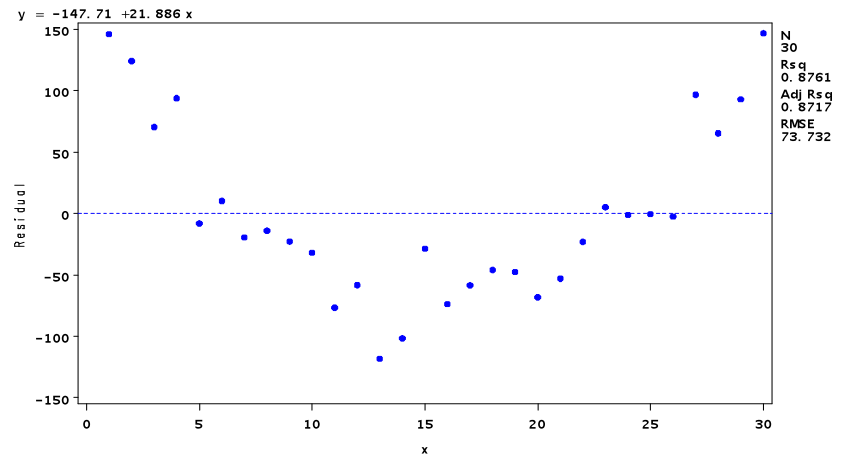
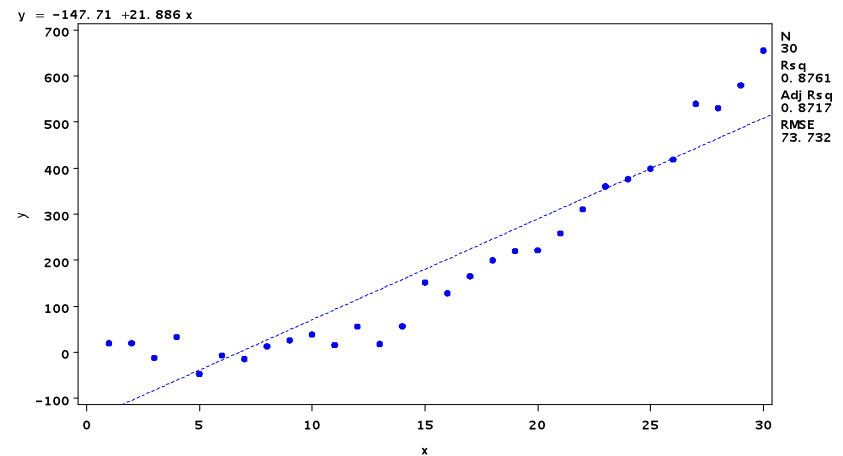
```
/* Residual plots */
symbol1 v=dot h=.8 c=blue;

/* simulated data, is it linear?
*/
Data resid;
  do x=1 to 30;
    y=x*x-10*x+30+25*normal
      (0);
    output;
  end;
proc print data=resid;
run;
```

Obs	x	y
1	1	20.268
2	2	20.292
3	3	-11.644
4	4	33.722
5	5	-46.357
6	6	-6.092
7	7	-13.902
8	8	13.392
9	9	26.473
10	10	39.391
11	11	16.491
12	12	56.712
13	13	18.588
14	14	57.087
...		

Do it in SAS

```
proc reg data=resid  
  noprint;  
  model y=x;  
  plot y*x student.*x  
        student.*p.;  
run;
```



Does the variance depend on X ?

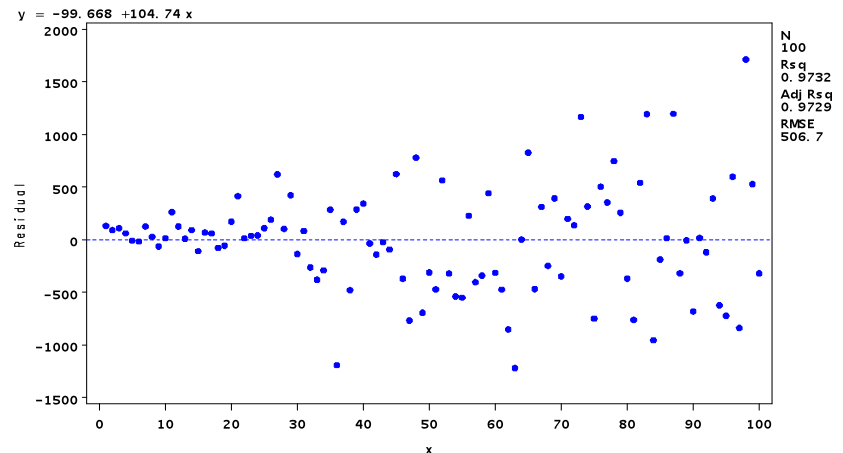
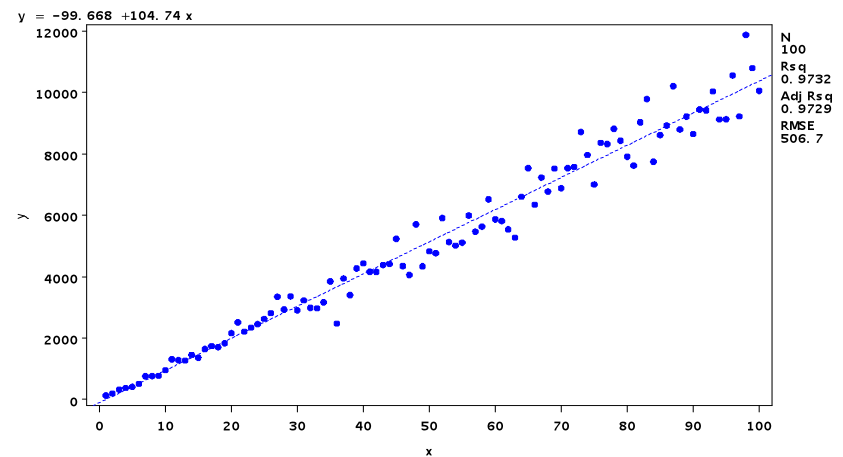
- Plot Y vs X
- Plot r vs X
 - Plot of r vs X will emphasize problems with the variance assumption

Do it in SAS

```
/* simulated data, is  
variance constant? */
```

```
Data resid2;  
  do x=1 to 100;  
    y=100*x  
    +30+10*x*normal(0);  
    output;  
  end;  
run;
```

```
proc reg data=resid2  
  noprint;  
model y=x;  
plot y*x student.*x;  
run;
```



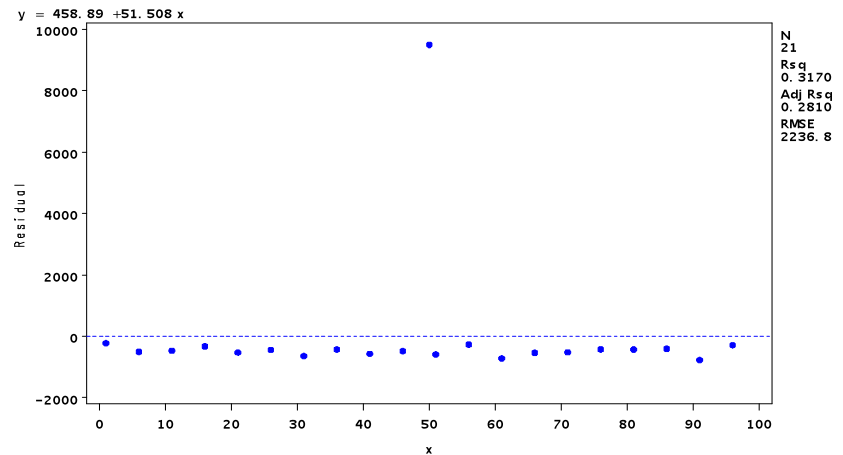
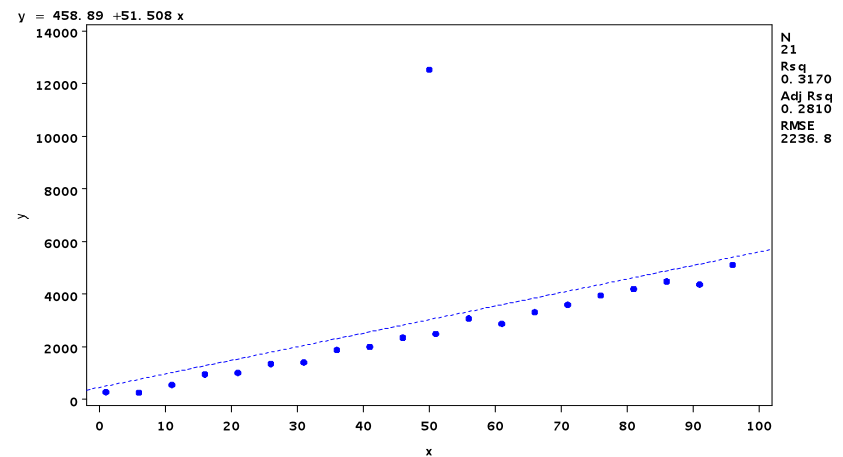
Outliers and Influential Cases

- Outliers inflate the variance and decrease our chances of finding statistically significant results
- Outliers may or may not be influential
- An outlier can be *influential* for some model parameters, and not influential for others
- Influential cases are usually outliers

Do it in SAS

```
/* simulated data, outlier */  
Data outlier;  
  do x=1 to 100 by 5;  
    y=30+50*x+200*normal(0);  
    output;  
  end;  
  x=50; y=30+50*50 +10000;  
  d='out'; output;  
run;  
proc print data=outlier;  
run;
```

```
proc reg data=outlier;  
  model y=x;  
  where d ne 'out';  
run;  
proc reg data=outlier;  
  model y=x;  
  plot y*x student.*x;  
run;
```

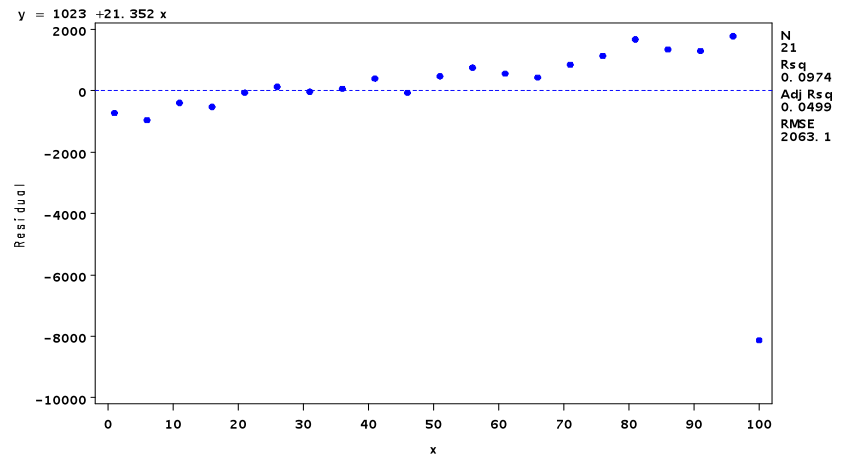
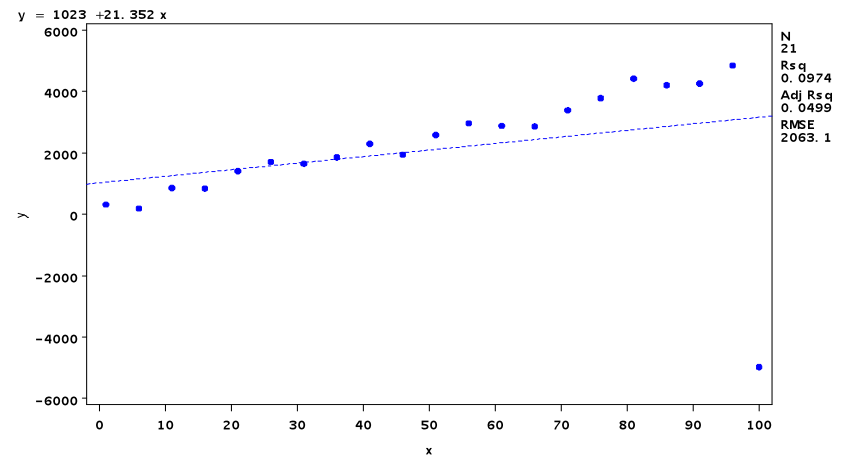


Are there outliers?

- Plot Y vs X
- Plot of r vs X should emphasize an outlier

Do it in SAS

```
/* Simulated data, inf. case */  
Data outlier2;  
  do x=1 to 100 by 5;  
    y=30+50*x+200*normal(0);  
    output;  
  end;  
  x=100; y=30+50*100 -10000;  
  d='out'; output;  
run;  
  
proc reg data=outlier2;  
  model y=x;  
  plot y*x student.*x;  
run;
```



Are the errors normal?

- The *real* question is whether the distribution of the errors is far enough away from normal to invalidate our confidence intervals and significance tests
- Look at the distribution of the residuals
- Use a normal quantile plot

Normal Quantile plots (Rankit plot)

- Consider $n=5$ observations iid $N(0,1)$
- From normal table, we find
 - $P(z \leq -.84) = .20$
 - $P(-.84 < z \leq -.25) = .20$
 - $P(-.25 < z \leq .25) = .20$
 - $P(.25 < z \leq .84) = .20$
 - $P(.84 < z) = .20$

Normal Quantile plots (2)

- So we *expect*
 - One observation ≤ -0.84
 - One observation in $(-0.84, -0.25]$
 - One observation in $(-0.25, 0.25]$
 - One observation in $(0.25, 0.84]$
 - One observation > 0.84

Normal Quantile plots (3)

- We use some theory to pick an appropriate value in the interval
(The book has a nice idea).
- $Z_{\text{norm}_i} = \Phi^{-1}((i-.375)/(n+.25)), i=1 \text{ to } n$
- Plot the order statistics $X_{(i)}$ versus Z_{norm_i}

Normal Quantile plots (4)

- The standardized X variable is

$$z = (X - \bar{X})/s$$

- So, $X = \bar{X} + s * z$
- If the data are approximately normal, the relationship will be approximately linear with slope s and intercept \bar{X} .

Do it in SAS

```
/*lot size example revisited */
```

```
data lot2;
```

```
set lot;
```

```
id = _n_;
```

```
group=1;
```

```
run;
```

```
/* scatter plot and QQ plot of  
residual */
```

```
symbol1 v=dot h=.8 c=blue i=none;
```

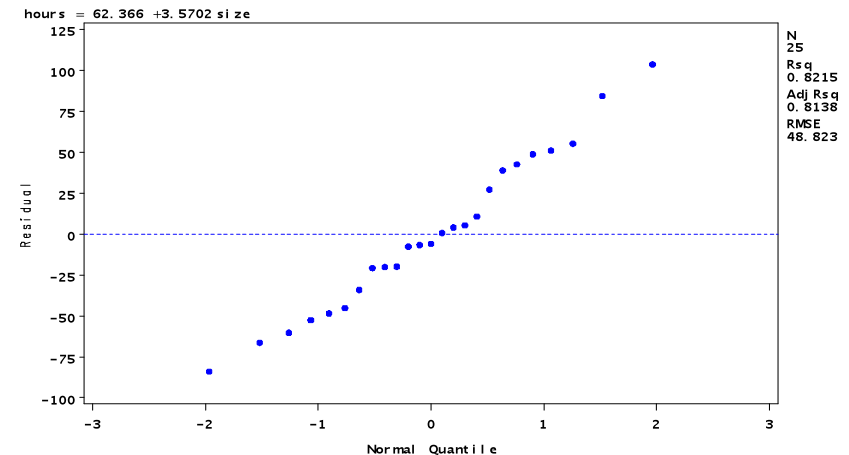
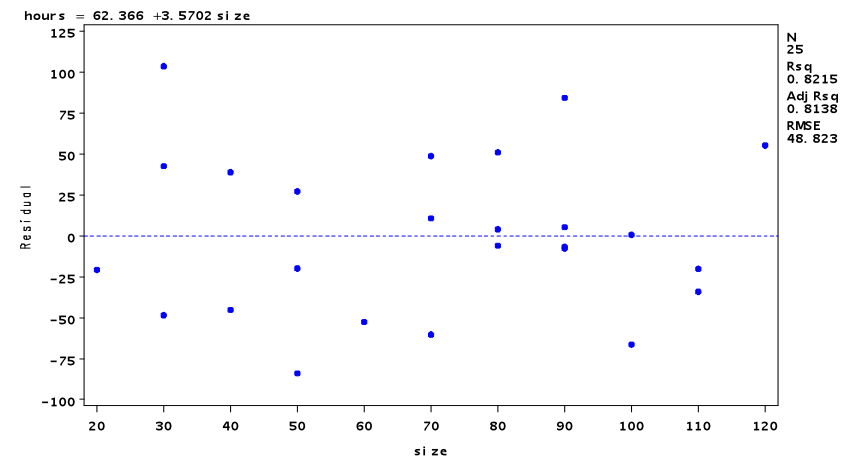
```
proc reg data = lot2 noprint;
```

```
model hours = size;
```

```
output out=temp student=r;
```

```
plot hours*size student.*size  
student.*nqq.;
```

```
run;
```



Significance tests for normality

- Many choices for a significance testing procedure
- Proc univariate with the normal option (`proc univariate normal;`) provides four
- Shapiro-Wilk is a good choice

Do it in SAS

```
/* test for normality */  
proc univariate normal data=temp;  
var r;  
run;
```

Tests for Normality

Test	--Statistic---	-----p Value-----
Shapiro-Wilk	W	0.978904
Kolmogorov-Smirnov	D	0.09572
Cramer-von Mises	W-Sq	0.033263
Anderson-Darling	A-Sq	0.207142

Pr < W	0.8626
Pr > D	>0.1500
Pr > W-Sq	>0.2500
Pr > A-Sq	>0.2500