

Worksheet 02

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```
library(tidyverse)
library(rjags)
library(coda)
library(bayesmeta)
library(pCalibrate)

## Warning in .recacheSubclasses(def@className, def, env): undefined subclass
## "ndiMatrix" of class "replValueSp"; definition not updated

library(ggthemes)

## Warning: package 'ggthemes' was built under R version 4.3.3

library(DescTools)

## Warning: package 'DescTools' was built under R version 4.3.3

theme_set(theme_solarized_2())
```

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Exercise 1

```
dat <- matrix(data = c(14, 9, 1, 5), ncol = 2, byrow = T)
rownames(dat) <- c("Secukinumab", "Placebo")
colnames(dat) <- c("Responder", "Not Responder")
dat |> addmargins()
```

```
##           Responder Not Responder Sum
## Secukinumab         14           9 23
## Placebo             1           5  6
## Sum                 15          14 29
```

a)

```
glue::glue("P[Responder|Secukinumab]")
BinomCI(dat[1,1], sum(dat[1,]))
glue::glue("")
```

```
glue::glue("P[Responder|Placebo]")
BinomCI(dat[2,1], sum(dat[2,]))
```

```
## P[Responder|Secukinumab]
##      est    lwr.ci    upr.ci
## [1,] 0.6086957 0.4078552 0.7784238
##
## P[Responder|Placebo]
##      est    lwr.ci    upr.ci
```

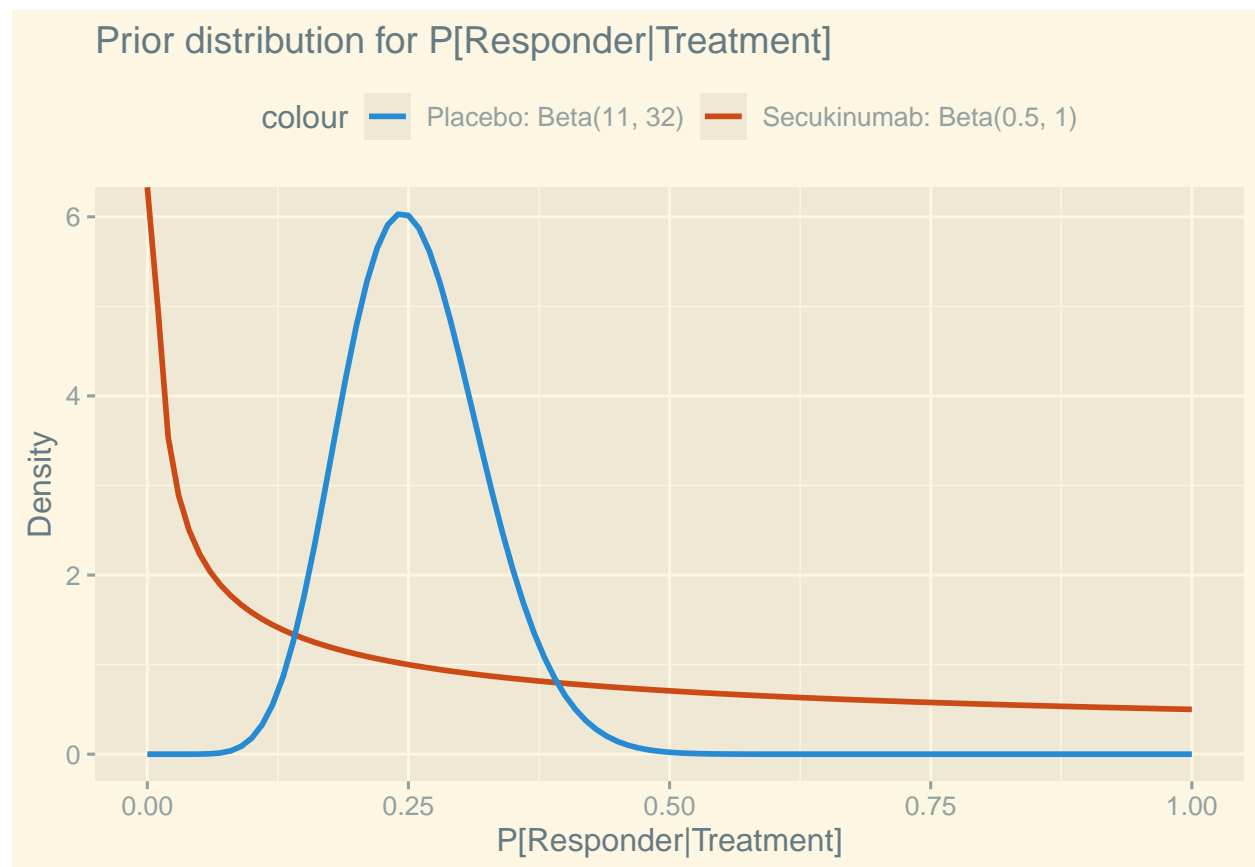
At this stage we explicitly use the interpretation:
Confidence interval
For repeated random samples from a distribution with unknown parameter θ , a $\gamma \cdot 100\%$ confidence interval will cover θ in $\gamma \cdot 100\%$ of all cases.

```
## [1,] 0.1666667 0.03005337 0.5635028
```

Given only the data in the table, classical statistics suggest that the probability of a patient responding to Secukinumab is probably between 0.41 and 0.78, and the probability of a patient responding to placebo is between 0.03 and 0.56. These intervals overlap, so we cannot conclude that Secukinumab is more effective than placebo.

b)

```
ggplot(data.frame(x = c(0, 1)), aes(x)) +  
  stat_function(  
    fun = dbeta,  
    args = list(shape1 = 0.5, shape2 = 1),  
    linewidth = 1,  
    aes(colour = "Secukinumab: Beta(0.5, 1)")) +  
  stat_function(  
    fun = dbeta,  
    args = list(shape1 = 11, shape2 = 32),  
    linewidth = 1,  
    aes(colour = "Placebo: Beta(11, 32)")) +  
  theme(legend.position = "top") +  
  labs(  
    title = "Prior distribution for P[Responder|Treatment]",  
    x = "P[Responder|Treatment]",  
    y = "Density") +  
  scale_colour_solarized()
```



```

glue::glue("Secukinumab prior: Beta(0.5, 1)")
alpha1 <- 0.5
beta1 <- 1
mean1 <- alpha1/(alpha1 + beta1)
median1 <- qbeta(0.5, alpha1, beta1)
ci1 <- qbeta(c(0.025, 0.975), alpha1, beta1)
glue::glue("Mean: {mean1}")
glue::glue("Median: {median1}")
glue::glue("95% CrI: {ci1[1]} - {ci1[2]}")
glue::glue("")

```

```

glue::glue("Placebo prior: Beta(11, 32)")
alpha2 <- 11
beta2 <- 32
mean2 <- alpha2/(alpha2 + beta2)
median2 <- qbeta(0.5, alpha2, beta2)
ci2 <- qbeta(c(0.025, 0.975), alpha2, beta2)
glue::glue("Mean: {mean2}")
glue::glue("Median: {median2}")
glue::glue("95% CrI: {ci2[1]} - {ci2[2]}")

```

```

## Secukinumab prior: Beta(0.5, 1)
## Mean: 0.333333333333333
## Median: 0.25
## 95% CrI: 0.000625 - 0.950625
##
## Placebo prior: Beta(11, 32)
## Mean: 0.255813953488372
## Median: 0.252000315740326
## 95% CrI: 0.13861013208847 - 0.394502429021576

```

The prior for Secukinumab is weak and its credible interval covers pretty much the whole range of possible values. Since the distribution is heavily skewed, I prefer to use the median as a measure of central tendency. It is 0.25 in this case, but I wouldn't put too much weight on it since the prior is so weak ($\alpha + \beta = 0.5 + 1 = 1.5$). The prior for placebo is much stronger and its confidence interval is much narrower ($\alpha + \beta = 11 + 32 = 43$). It is bell-shaped and thus the mean and median are very close. The true probability of a patient responding to placebo is probably around 0.25 and between 0.14 and 0.39.

c)

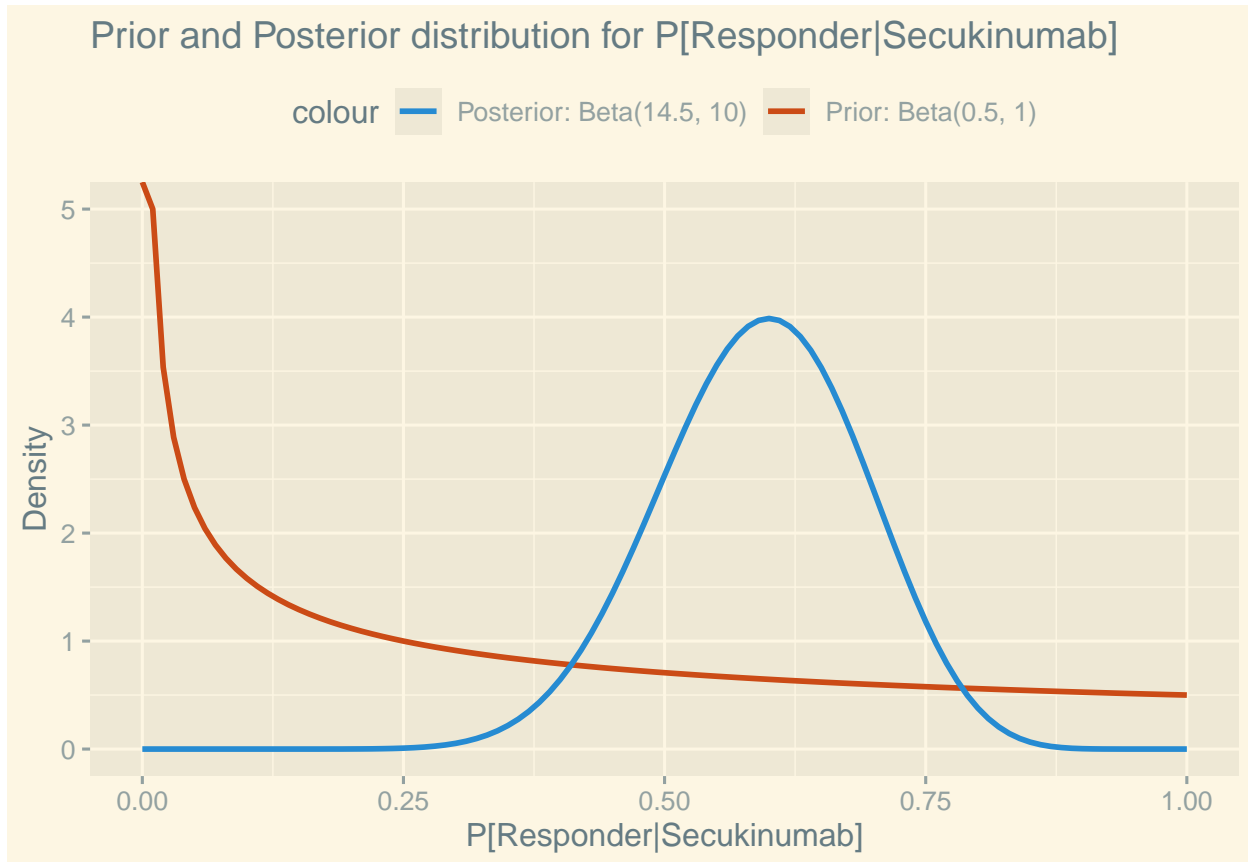
```

alpha1.post <- alpha1 + dat[1,1]
beta1.post <- beta1 + dat[1,2]

ggplot(data.frame(x = c(0, 1)), aes(x)) +
  stat_function(
    fun = dbeta,
    args = list(shape1 = alpha1, shape2 = beta1),
    linewidth = 1,
    aes(colour = "Prior: Beta(0.5, 1)")) +
  stat_function(
    fun = dbeta, args = list(shape1 = alpha1.post, shape2 = beta1.post),
    linewidth = 1,
    aes(colour = "Posterior: Beta(14.5, 10)")) +
  theme(legend.position = "top") +

```

```
labs(
  title = "Prior and Posterior distribution for P[Responder|Secukinumab]",
  x = "P[Responder|Secukinumab]",
  y = "Density") +
scale_colour_solarized()
```



```
glue::glue(glue::glue("Secukinumab posterior: Beta({alpha1.post}, {beta1.post}"))
mean1.post <- alpha1.post/(alpha1.post + beta1.post)
median1.post <- qbeta(0.5, alpha1.post, beta1.post)
ci1.post <- qbeta(c(0.025, 0.975), alpha1.post, beta1.post)
glue::glue("Mean: {mean1.post}")
glue::glue("Median: {median1.post}")
glue::glue("95% CrI: {ci1.post[1]} - {ci1.post[2]}")
```

```
## Secukinumab posterior: Beta(14.5, 10)
## Mean: 0.591836734693878
## Median: 0.594375005547927
## 95% CrI: 0.395840118720512 - 0.773625318723245
```

The posterior is much stronger than the prior and the credible interval is much narrower ($\alpha + \beta = 14.5 + 10 = 24.5$). Given our new data, the true probability of a patient responding to Secukinumab is probably around 0.59 and between 0.40 and 0.77.

given the data observed and the prior

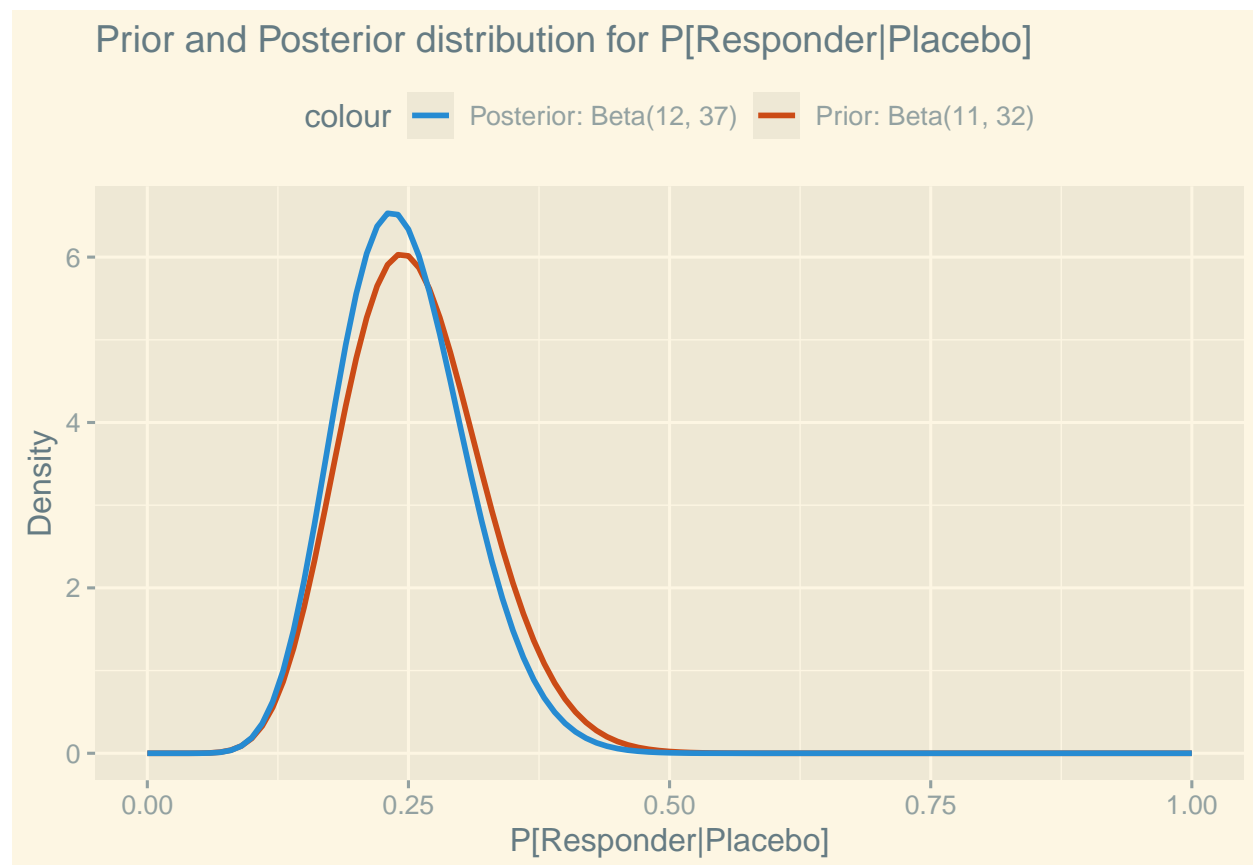
d)

```

alpha2.post <- alpha2 + dat[2,1]
beta2.post <- beta2 + dat[2,2]

ggplot(data.frame(x = c(0, 1)), aes(x)) +
  stat_function(
    fun = dbeta,
    args = list(shape1 = alpha2, shape2 = beta2),
    linewidth = 1,
    aes(colour = "Prior: Beta(11, 32)")) +
  stat_function(
    fun = dbeta,
    args = list(shape1 = alpha2.post, shape2 = beta2.post),
    linewidth = 1,
    aes(colour = "Posterior: Beta(12, 37)")) +
  theme(legend.position = "top") +
  labs(
    title = "Prior and Posterior distribution for P[Responder|Placebo]",
    x = "P[Responder|Placebo]",
    y = "Density") +
  scale_colour_solarized()

```



```

glue::glue(glue::glue("Placebo posterior: Beta({alpha2.post}, {beta2.post}"))
mean2.post <- alpha2.post/(alpha2.post + beta2.post)
median2.post <- qbeta(0.5, alpha2.post, beta2.post)
ci2.post <- qbeta(c(0.025, 0.975), alpha2.post, beta2.post)
glue::glue("Mean: {mean2.post}")

```

```
glue::glue("Median: {median2.post}")
glue::glue("95% CrI: {ci2.post[1]} - {ci2.post[2]}")
```

```
## Placebo posterior: Beta(12, 37)
## Mean: 0.244897959183673
## Median: 0.241405434669185
## 95% CrI: 0.1363722774206 - 0.37312024731586
```

Our prior was already fairly strong and the new data is in line with it. Thus, our posterior is not much different from our prior. The true probability of a patient responding to placebo is probably around 0.24 and between 0.14 and 0.37.

given data observed and the prior

Exercise 2

a)

See pictures below.

for $\text{Beta}(\alpha, \beta)$

$$\textcircled{A} \quad \mu = \frac{\alpha}{\alpha + \beta}$$

$$\textcircled{B} \quad \sigma^2 = \frac{\alpha \beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$$

$$\textcircled{A} \quad \mu = \frac{\alpha}{\alpha + \beta}$$

$$\Rightarrow \alpha \mu + \beta \mu = \alpha$$

$$\Rightarrow \beta \mu = \alpha - \alpha \mu$$

$$\Rightarrow \beta = \frac{\alpha (1 - \mu)}{\mu}$$

$$q := \frac{1 - \mu}{\mu} \Rightarrow \beta = \alpha q \quad \text{plug into } \textcircled{B}$$

$$\sigma^2 = \frac{\alpha \cdot \alpha q}{(\alpha + \alpha q)^2 (\alpha + \alpha q + 1)}$$

$$= \frac{\alpha^2 q}{(\alpha (1 + q))^2 (\alpha (1 + q) + 1)}$$

$$= \frac{\alpha^2 q}{\alpha^2 (1 + q)^2 (\alpha (1 + q) + 1)}$$

$$= \frac{q}{(1 + q)^2 (\alpha (1 + q) + 1)}$$

Figure 1: Exercise 2 part 1

$$\begin{aligned}
 &= \frac{4}{\alpha(1+q)^3 + (1+q)^2} \\
 &\Rightarrow q = \sigma^2 \left(\alpha \left(\frac{1}{\mu} \right)^3 + \left(\frac{1}{\mu} \right)^2 \right) \\
 &\text{using } q = \frac{1-\mu}{\mu} \text{ and } 1+q = \frac{1}{\mu} \\
 &\Rightarrow \frac{1-\mu}{\mu} = \sigma^2 \left(\alpha \left(\frac{1}{\mu} \right)^3 + \left(\frac{1}{\mu} \right)^2 \right) \\
 &= \sigma^2 \left(\frac{\alpha}{\mu^3} + \frac{1}{\mu^2} \right)
 \end{aligned}$$

$$\begin{aligned}
 &\Rightarrow \frac{1-\mu}{\mu\sigma^2} = \frac{\alpha}{\mu^3} + \frac{1}{\mu^2} \\
 &\Rightarrow \frac{\mu^2(1-\mu)}{\mu\sigma^2} = \alpha + \mu \\
 &\Rightarrow \frac{\mu^2(1-\mu)}{\sigma^2} - \mu = \alpha \\
 &\Rightarrow \alpha = \mu^2 \left(\frac{1-\mu}{\sigma^2} - \frac{1}{\mu} \right)
 \end{aligned}$$

from earlier:

$$\beta = \frac{\alpha(1-\mu)}{\mu}$$

Figure 2: Exercise 2 part 2

b)

```

estimate_beta_shapes <- function(mean, var) {
  alpha <- mean^2 * ((1 - mean)/var - 1/mean);
  beta <- alpha * (1 - mean)/mean;
  return(list(alpha = alpha, beta = beta));
}

```

c)

```

estimate_beta_shapes(mean = 0.255814, var = 0.004326663)

```

```

## $alpha
## [1] 11
##
## $beta
## [1] 32

```

$\alpha = 11$ and $\beta = 32$. These are the values of our prior for the placebo group.