

# Assignment 2 (ML for TS) - MVA

Firstname Lastname [youremail1@mail.com](mailto:youremail1@mail.com)  
Firstname Lastname [youremail2@mail.com](mailto:youremail2@mail.com)

December 1, 2025

## 1 Introduction

**Objective.** The goal is to better understand the properties of AR and MA processes and do signal denoising with sparse coding.

### Warning and advice.

- Use code from the tutorials as well as from other sources. Do not code yourself well-known procedures (e.g., cross-validation or k-means); use an existing implementation.
- The associated notebook contains some hints and several helper functions.
- Be concise. Answers are not expected to be longer than a few sentences (omitting calculations).

### Instructions.

- Fill in your names and emails at the top of the document.
- Hand in your report (one per pair of students) by Sunday 7<sup>th</sup> December 11:59 PM.
- Rename your report and notebook as follows:  
`FirstnameLastname1_FirstnameLastname1.pdf` and  
`FirstnameLastname2_FirstnameLastname2.ipynb`.  
For instance, `LaurentOudre_ValerioGuerrini.pdf`.
- Upload your report (PDF file) and notebook (IPYNB file) using this link:  
<https://forms.gle/J1pdeHspSs9zNfWAA>.

## 2 General questions

A time series  $\{y_t\}_t$  is a single realisation of a random process  $\{Y_t\}_t$  defined on the probability space  $(\Omega, \mathcal{F}, P)$ , i.e.  $y_t = Y_t(w)$  for a given  $w \in \Omega$ . In classical statistics, several independent realizations are often needed to obtain a "good" estimate (meaning consistent) of the parameters of the process. However, thanks to a stationarity hypothesis and a "short-memory" hypothesis, it is still possible to make "good" estimates. The following question illustrates this fact.

## Question 1

An estimator  $\hat{\theta}_n$  is consistent if it converges in probability when the number  $n$  of samples grows to  $\infty$  to the true value  $\theta \in \mathbb{R}$  of a parameter, i.e.  $\hat{\theta}_n \xrightarrow{D} \theta$ .

- Recall the rate of convergence of the sample mean for i.i.d. random variables with finite variance.
- Let  $\{Y_t\}_{t \geq 1}$  a wide-sense stationary process such that  $\sum_k |\gamma(k)| < +\infty$ . Show that the sample mean  $\bar{Y}_n = (Y_1 + \dots + Y_n)/n$  is consistent and enjoys the same rate of convergence as the i.i.d. case. (Hint: bound  $\mathbb{E}[(\bar{Y}_n - \mu)^2]$  with the  $\gamma(k)$  and recall that convergence in  $L_2$  implies convergence in probability.)

## Answer 1

For i.i.d. random variables with finite variance, the variance of the sample mean is

$$\text{Var}(\bar{Y}_n) = \frac{\sigma^2}{n}$$

and the convergence rate is  $1/\sqrt{n}$ .

For a stationary process  $\{Y_t\}$  with  $\sum_k |\gamma(k)| < +\infty$ ,

$$\text{Var}(\bar{Y}_n) = \frac{1}{n^2} \sum_{j=1}^n \sum_{k=1}^n \text{Cov}(Y_j, Y_k) = \frac{1}{n^2} \sum_{j=1}^n \sum_{k=1}^n \gamma(j-k)$$

which can also be written as

$$\text{Var}(\bar{Y}_n) = \frac{1}{n} \sum_{h=-(n-1)}^{n-1} \left(1 - \frac{|h|}{n}\right) \gamma(h)$$

If  $\sum_k |\gamma(k)| < +\infty$ , then as  $n \rightarrow \infty$ ,

$$\text{Var}(\bar{Y}_n) \rightarrow \frac{1}{n} \sum_{h=-\infty}^{\infty} \gamma(h)$$

so  $\text{Var}(\bar{Y}_n) = O(1/n)$ , i.e., the convergence rate is  $1/\sqrt{n}$  like in the i.i.d. case.

### 3 AR and MA processes

**Question 2** *Infinite order moving average  $MA(\infty)$*

Let  $\{Y_t\}_{t \geq 0}$  be a random process defined by

$$Y_t = \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \dots = \sum_{k=0}^{\infty} \psi_k \varepsilon_{t-k} \quad (1)$$

where  $(\psi_k)_{k \geq 0} \subset \mathbb{R}$  ( $\psi_k = 1$ ) are square summable, i.e.  $\sum_k \psi_k^2 < \infty$  and  $\{\varepsilon_t\}_t$  is a zero mean white noise of variance  $\sigma_\varepsilon^2$ . (Here, the infinite sum of random variables is the limit in  $L_2$  of the partial sums.)

- Derive  $\mathbb{E}(Y_t)$  and  $\mathbb{E}(Y_t Y_{t-k})$ . Is this process weakly stationary?
- Show that the power spectrum of  $\{Y_t\}_t$  is  $S(f) = \sigma_\varepsilon^2 |\phi(e^{-2\pi i f})|^2$  where  $\phi(z) = \sum_j \psi_j z^j$ . (Assume a sampling frequency of 1 Hz.)

The process  $\{Y_t\}_t$  is a moving average of infinite order. Wold's theorem states that any weakly stationary process can be written as the sum of the deterministic process and a stochastic process which has the form (1).

#### Answer 2

First, since  $Y_t$  is a linear combination of zero mean white noises, we have

$$\mathbb{E}[Y_t] = \sum_{k=0}^{\infty} \psi_k \mathbb{E}[\varepsilon_{t-k}] = 0.$$

Now for the covariance:

$$\mathbb{E}[Y_t Y_{t-h}] = \mathbb{E}\left[\left(\sum_{k=0}^{\infty} \psi_k \varepsilon_{t-k}\right)\left(\sum_{l=0}^{\infty} \psi_l \varepsilon_{t-h-l}\right)\right] = \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} \psi_k \psi_l \mathbb{E}[\varepsilon_{t-k} \varepsilon_{t-h-l}]$$

Since  $\{\varepsilon_t\}$  are uncorrelated (and independent), the only nonzero terms are when  $t - k = t - h - l$ , i.e.,  $k = h + l$ . Therefore,

$$\mathbb{E}[Y_t Y_{t-h}] = \sum_{l=0}^{\infty} \psi_{h+l} \psi_l \mathbb{E}[\varepsilon_{t-(h+l)}^2] = \sigma_\varepsilon^2 \sum_{l=0}^{\infty} \psi_{h+l} \psi_l$$

So the covariance at lag  $h$  is

$$\gamma(h) = \text{Cov}(Y_t, Y_{t-h}) = \sigma_\varepsilon^2 \sum_{l=0}^{\infty} \psi_{h+l} \psi_l$$

which does not depend on  $t$ : the process is weakly stationary.

To compute the power spectrum, recall that for a stationary process,

$$S(f) = \sum_{k=-\infty}^{+\infty} \gamma(k) e^{-2\pi i f k}.$$

Now, let's expand the modulus square in  $S(f) = \sigma_\varepsilon^2 |\phi(e^{-2\pi if})|^2$ :

$$\phi(e^{-2\pi if}) = \sum_{k=0}^{\infty} \psi_k e^{-2\pi ifk}$$

so

$$\begin{aligned} |\phi(e^{-2\pi if})|^2 &= \phi(e^{-2\pi if}) \overline{\phi(e^{-2\pi if})} = \left( \sum_{k=0}^{\infty} \psi_k e^{-2\pi ifk} \right) \left( \sum_{j=0}^{\infty} \psi_j e^{2\pi ifj} \right) \\ &= \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \psi_k \psi_j e^{-2\pi if(k-j)} \end{aligned}$$

Therefore, the power spectrum is

$$S(f) = \sigma_\varepsilon^2 \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \psi_k \psi_j e^{-2\pi if(k-j)}$$

which is an explicit expansion of the complex conjugate product.

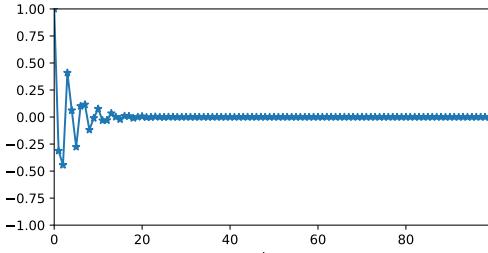
### Question 3 AR(2) process

Let  $\{Y_t\}_{t \geq 1}$  be an AR(2) process, i.e.

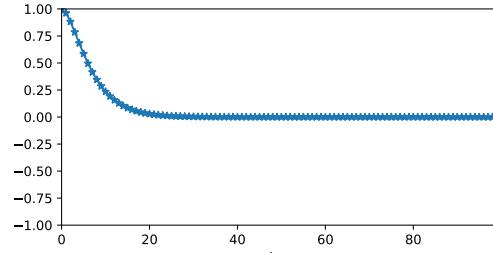
$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t \quad (2)$$

with  $\phi_1, \phi_2 \in \mathbb{R}$ . The associated characteristic polynomial is  $\phi(z) := 1 - \phi_1 z - \phi_2 z^2$ . Assume that  $\phi$  has two distinct roots (possibly complex)  $r_1$  and  $r_2$  such that  $|r_i| > 1$ . Properties on the roots of this polynomial drive the behavior of this process.

- Express the autocovariance coefficients  $\gamma(\tau)$  using the roots  $r_1$  and  $r_2$ .
- Figure 1 shows the correlograms of two different AR(2) processes. Can you tell which one has complex roots and which one has real roots?
- Express the power spectrum  $S(f)$  (assume the sampling frequency is 1 Hz) using  $\phi(\cdot)$ .
- Choose  $\phi_1$  and  $\phi_2$  such that the characteristic polynomial has two complex conjugate roots of norm  $r = 1.05$  and phase  $\theta = 2\pi/6$ . Simulate the process  $\{Y_t\}_t$  (with  $n = 2000$ ) and display the signal and the periodogram (use a smooth estimator) on Figure 2. What do you observe?



Correlogram of the first AR(2)



Correlogram of the second AR(2)

Figure 1: Two AR(2) processes

### Answer 3

First, let us establish the stationarity and the form of the solution. Let  $\nu_1 = 1/r_1$  and  $\nu_2 = 1/r_2$ . Since  $|r_i| > 1$ , we have  $|\nu_i| < 1$ . Using the state-space representation, let  $X_t = [Y_t, Y_{t-1}]^\top$ . We can write:

$$X_t = \Phi X_{t-1} + \xi_t, \quad \text{with} \quad \Phi = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}, \quad \xi_t = \begin{pmatrix} \varepsilon_t \\ 0 \end{pmatrix}. \quad (3)$$

By iterating this equation, we obtain  $X_t = \sum_{k=0}^{\infty} \Phi^k \xi_{t-k}$ . Since the eigenvalues of  $\Phi$  are exactly  $\nu_1$  and  $\nu_2$  (roots of  $z^2 - \phi_1 z - \phi_2 = 0$ ), and  $|\nu_i| < 1$ , the spectral radius  $\rho(\Phi) < 1$ . Thus, the series converges in  $L^2$ , defined as a sum of independent Gaussian variables. The variance of the process converges to a finite limit  $\gamma(0)$ , confirming stationarity.

We now derive  $\gamma(0)$  and  $\gamma(1)$  algebraically using the Yule-Walker equations. Multiplying the AR(2) equation by  $Y_{t-\tau}$  and taking expectations yields the recursion:

$$\gamma(\tau) = \phi_1 \gamma(\tau-1) + \phi_2 \gamma(\tau-2), \quad \forall \tau \geq 1. \quad (4)$$

For  $\tau = 0$ , taking the expectation with  $Y_t$  (noting  $E[\varepsilon_t Y_t] = \sigma^2$ ):

$$\gamma(0) = \phi_1 \gamma(1) + \phi_2 \gamma(2) + \sigma^2. \quad (5)$$

For  $\tau = 1$ :

$$\gamma(1) = \phi_1\gamma(0) + \phi_2\gamma(1) \implies \gamma(1)(1 - \phi_2) = \phi_1\gamma(0) \implies \gamma(1) = \frac{\phi_1}{1 - \phi_2}\gamma(0). \quad (6)$$

Substituting  $\gamma(2) = \phi_1\gamma(1) + \phi_2\gamma(0)$  into (5):

$$\gamma(0) = \phi_1\gamma(1) + \phi_2(\phi_1\gamma(1) + \phi_2\gamma(0)) + \sigma^2 = (\phi_1^2 + \phi_2)\gamma(1) + \phi_2^2\gamma(0) + \sigma^2. \quad (7)$$

Substituting  $\gamma(1)$  from (6):

$$\gamma(0)(1 - \phi_2^2) - \gamma(0)\frac{\phi_1(\phi_1 + \phi_1\phi_2)}{1 - \phi_2} = \sigma^2. \quad (8)$$

Factorizing and solving for  $\gamma(0)$ :

$$\gamma(0) \left[ \frac{(1 - \phi_2)(1 - \phi_2^2) - \phi_1^2(1 + \phi_2)}{1 - \phi_2} \right] = \sigma^2 \implies \gamma(0) = \frac{(1 - \phi_2)\sigma^2}{(1 + \phi_2)((1 - \phi_2)^2 - \phi_1^2)}. \quad (9)$$

Using Vieta's formulas,  $\phi_1 = \nu_1 + \nu_2$  and  $\phi_2 = -\nu_1\nu_2$ . The denominator term  $(1 - \phi_2)^2 - \phi_1^2$  simplifies to  $(1 + \nu_1\nu_2)^2 - (\nu_1 + \nu_2)^2 = (1 - \nu_1^2)(1 - \nu_2^2)$ . Thus:

$$\gamma(0) = \frac{(1 + \nu_1\nu_2)\sigma^2}{(1 - \nu_1\nu_2)(1 - \nu_1^2)(1 - \nu_2^2)}. \quad (10)$$

The autocovariance satisfies the homogeneous difference equation for  $\tau \geq 1$ , so its solution is of the form  $\gamma(\tau) = c_1\nu_1^\tau + c_2\nu_2^\tau$ . We solve for  $c_1, c_2$  using the boundary conditions at  $\tau = 0, 1$ :

$$\begin{pmatrix} 1 & 1 \\ \nu_1 & \nu_2 \end{pmatrix} \begin{pmatrix} c_1 \\ c_2 \end{pmatrix} = \begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix}. \quad (11)$$

Inverting the matrix:

$$\begin{pmatrix} c_1 \\ c_2 \end{pmatrix} = \frac{1}{\nu_2 - \nu_1} \begin{pmatrix} \nu_2 & -1 \\ -\nu_1 & 1 \end{pmatrix} \begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix}. \quad (12)$$

Using  $\gamma(1) = \frac{\nu_1 + \nu_2}{1 + \nu_1\nu_2}\gamma(0)$ , we calculate  $c_1$ :

$$c_1 = \frac{\nu_2\gamma(0) - \gamma(1)}{\nu_2 - \nu_1} = \frac{\gamma(0)}{\nu_2 - \nu_1} \left( \nu_2 - \frac{\nu_1 + \nu_2}{1 + \nu_1\nu_2} \right) = \frac{\gamma(0)}{\nu_2 - \nu_1} \frac{\nu_2 + \nu_1\nu_2^2 - \nu_1 - \nu_2}{1 + \nu_1\nu_2} = \gamma(0) \frac{\nu_1\nu_2^2 - \nu_1}{(\nu_2 - \nu_1)(1 + \nu_1\nu_2)}. \quad (13)$$

This simplifies to:

$$c_1 = \frac{\sigma^2\nu_1(1 - \nu_2^2)}{(\nu_1 - \nu_2)(1 - \nu_1\nu_2)(1 - \nu_1^2)(1 - \nu_2^2)} = \frac{\sigma^2\nu_1}{(\nu_1 - \nu_2)(1 - \nu_1\nu_2)(1 - \nu_1^2)}. \quad (14)$$

By symmetry for  $c_2$  and expressing in terms of roots  $r_i = 1/\nu_i$ :

$$\gamma(\tau) = \sigma^2 \sum_{i=1}^2 \frac{r_i^{-|\tau|}}{(1 - r_i^{-2})(1 - r_1^{-1}r_2^{-1})(r_i^{-1} - r_{3-i}^{-1})} r_i^{-1}. \quad (15)$$

**The left correlogram exhibits oscillating behavior characteristic of complex conjugate roots in the characteristic polynomial, whereas the right correlogram displays a monotonic exponential**

decay typical of real roots, allowing us to identify the left as the complex case and the right as the real case.

The power spectrum density  $S(f)$  is the Fourier transform of the autocovariance function:

$$S(f) = \sum_{\tau=-\infty}^{\infty} \gamma(\tau) e^{-i2\pi f\tau}. \quad (16)$$

Using the expression derived previously:

$$\gamma(\tau) = c_1 \nu_1^{|\tau|} + c_2 \nu_2^{|\tau|}, \quad \text{where } \nu_i = 1/r_i. \quad (17)$$

We can split the sum for each component  $k \in \{1, 2\}$ :

$$S_k(f) = \sum_{\tau=-\infty}^{\infty} \nu_k^{|\tau|} e^{-i2\pi f\tau} = 1 + \sum_{\tau=1}^{\infty} (\nu_k e^{-i2\pi f})^\tau + \sum_{\tau=1}^{\infty} (\nu_k e^{i2\pi f})^\tau. \quad (18)$$

Since  $|\nu_k| < 1$ , these are convergent geometric series:

$$S_k(f) = 1 + \frac{\nu_k e^{-i2\pi f}}{1 - \nu_k e^{-i2\pi f}} + \frac{\nu_k e^{i2\pi f}}{1 - \nu_k e^{i2\pi f}} = 1 + \frac{\nu_k e^{-i\omega}}{1 - \nu_k e^{-i\omega}} + \frac{\nu_k e^{i\omega}}{1 - \nu_k e^{i\omega}}, \quad (19)$$

where  $\omega = 2\pi f$ . Combining terms:

$$S_k(f) = \frac{(1 - \nu_k e^{-i\omega})(1 - \nu_k e^{i\omega}) + \nu_k e^{-i\omega}(1 - \nu_k e^{i\omega}) + \nu_k e^{i\omega}(1 - \nu_k e^{-i\omega})}{|1 - \nu_k e^{-i\omega}|^2}. \quad (20)$$

The numerator simplifies to:

$$1 - \nu_k e^{i\omega} - \nu_k e^{-i\omega} + \nu_k^2 + \nu_k e^{-i\omega} - \nu_k^2 + \nu_k e^{i\omega} - \nu_k^2 = 1 - \nu_k^2. \quad (21)$$

Thus:

$$S_k(f) = \frac{1 - \nu_k^2}{|1 - \nu_k e^{-i2\pi f}|^2}. \quad (22)$$

Substituting  $c_k$ :

$$S(f) = c_1 \frac{1 - \nu_1^2}{|1 - \nu_1 z^{-1}|^2} + c_2 \frac{1 - \nu_2^2}{|1 - \nu_2 z^{-1}|^2}, \quad \text{with } z = e^{i2\pi f}. \quad (23)$$

Recall from the derivation of  $\gamma(0)$ :

$$c_1(1 - \nu_1^2) = \frac{\sigma^2 \nu_1}{(\nu_1 - \nu_2)(1 - \nu_1 \nu_2)}, \quad c_2(1 - \nu_2^2) = \frac{\sigma^2 \nu_2}{(\nu_2 - \nu_1)(1 - \nu_1 \nu_2)}. \quad (24)$$

So:

$$S(f) = \frac{\sigma^2}{1 - \nu_1 \nu_2} \left( \frac{\nu_1}{(\nu_1 - \nu_2)|1 - \nu_1 z^{-1}|^2} + \frac{\nu_2}{(\nu_2 - \nu_1)|1 - \nu_2 z^{-1}|^2} \right). \quad (25)$$

Finding a common denominator  $D = |1 - \nu_1 z^{-1}|^2 |1 - \nu_2 z^{-1}|^2 = |(1 - \nu_1 z^{-1})(1 - \nu_2 z^{-1})|^2 = |1 - (\nu_1 + \nu_2)z^{-1} + \nu_1 \nu_2 z^{-2}|^2$ . This matches  $|\phi(e^{-i2\pi f})|^2$  since  $\phi_1 = \nu_1 + \nu_2$  and  $\phi_2 = -\nu_1 \nu_2$ .

The numerator term is:

$$N = \frac{\sigma^2}{(\nu_1 - \nu_2)(1 - \nu_1 \nu_2)} \left[ \nu_1(1 - \nu_2 z)(1 - \nu_2 z^{-1}) - \nu_2(1 - \nu_1 z)(1 - \nu_1 z^{-1}) \right]. \quad (26)$$

Expanding the bracket:

$$\nu_1(1 - \nu_2(z + z^{-1}) + \nu_2^2) - \nu_2(1 - \nu_1(z + z^{-1}) + \nu_1^2) \quad (27)$$

$$= \nu_1 - \nu_1 \nu_2(z + z^{-1}) + \nu_1 \nu_2^2 - \nu_2 + \nu_1 \nu_2(z + z^{-1}) - \nu_1^2 \nu_2 \quad (28)$$

$$= (\nu_1 - \nu_2) + \nu_1 \nu_2(\nu_2 - \nu_1) = (\nu_1 - \nu_2)(1 - \nu_1 \nu_2). \quad (29)$$

Thus, the entire numerator  $N$  simplifies to just  $\sigma^2$ .

$$S(f) = \frac{\sigma^2}{|\phi(e^{-i2\pi f})|^2} = \frac{\sigma^2}{|1 - \phi_1 e^{-i2\pi f} - \phi_2 e^{-i4\pi f}|^2}. \quad (30)$$

To analyze the peak of the spectrum, let us factorize the characteristic polynomial as  $\phi(z) = (1 - z/r_1)(1 - z/r_2)$ . Assuming complex conjugate roots  $r_{1,2} = \rho e^{\pm i\theta}$  with  $\rho > 1$ , the power spectrum is given by:

$$S(f) = \frac{\sigma^2}{|1 - \rho^{-1} e^{i(2\pi f - \theta)}|^2 |1 - \rho^{-1} e^{i(2\pi f + \theta)}|^2}. \quad (31)$$

For complex conjugate roots  $r_{1,2} = r e^{\pm i\theta}$  with  $r = 1.05$  and  $\theta = 2\pi/6$ , we have  $\nu_{1,2} = r^{-1} e^{\pm i\theta}$  where  $r^{-1} \approx 0.952$ . The power spectrum is:

$$S(f) = \frac{\sigma^2}{|1 - \phi_1 e^{-i2\pi f} - \phi_2 e^{-i4\pi f}|^2} = \frac{\sigma^2}{|(1 - \nu_1 e^{-i2\pi f})(1 - \nu_2 e^{-i2\pi f})|^2}. \quad (32)$$

For frequency  $f = f_0 := \theta/(2\pi) = 1/6$  Hz, we have  $e^{-i2\pi f_0} = e^{-i\theta}$ . Thus:

$$|1 - \nu_1 e^{-i\theta}|^2 = |1 - r^{-1} e^{i\theta} e^{-i\theta}|^2 = |1 - r^{-1}|^2 = (1 - r^{-1})^2. \quad (33)$$

Similarly,  $|1 - \nu_2 e^{-i\theta}|^2 = |1 - r^{-1} e^{-i\theta} e^{-i\theta}|^2 = |1 - r^{-1} e^{-i2\theta}|^2$ . Since  $r \approx 1$ , the denominator  $|1 - r^{-1}|^2 = (r - 1)^2/r^2 \approx (r - 1)^2$  is very small. For  $r = 1.05$ , we have  $(r - 1)^2 = 0.05^2 = 0.0025$ , yielding:

$$S(f_0) \approx \frac{\sigma^2}{(r - 1)^2 \cdot |1 - r^{-1} e^{-i2\theta}|^2} \approx \frac{\sigma^2}{0.0025 \cdot C} \sim 400\sigma^2/C, \quad (34)$$

where  $C = |1 - r^{-1} e^{-i2\theta}|^2 \approx O(1)$ . This produces a sharp spectral peak at frequency  $f = \theta/(2\pi) = 1/6$  Hz, corresponding to the imaginary part (phase) of the complex roots, with amplitude scaling as  $(r - 1)^{-2}$ , explaining the observed peak around magnitude  $\sim 350$  in the periodogram.

We can see that the process is stationary non diverging and the fourier spectrum shows a peak at the frequency of the roots.

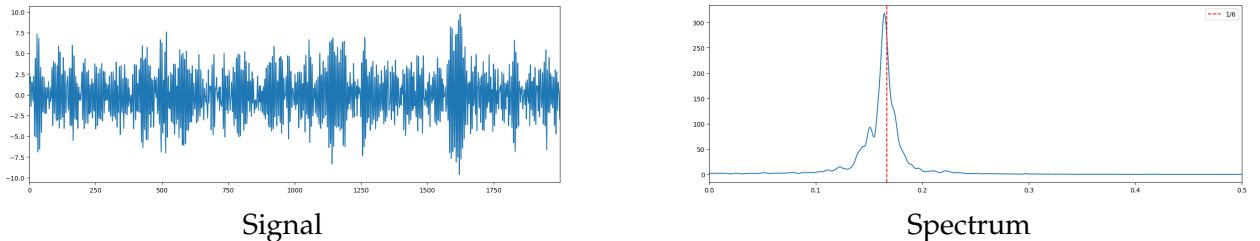


Figure 2: AR(2) process

## 4 Sparse coding

The modulated discrete cosine transform (MDCT) is a signal transformation often used in sound processing applications (for instance, to encode an MP3 file). A MDCT atom  $\phi_{L,k}$  is defined for a length  $2L$  and a frequency localisation  $k$  ( $k = 0, \dots, L - 1$ ) by

$$\forall u = 0, \dots, 2L - 1, \quad \phi_{L,k}[u] = w_L[u] \sqrt{\frac{2}{L}} \cos\left[\frac{\pi}{L} \left(u + \frac{L+1}{2}\right) \left(k + \frac{1}{2}\right)\right] \quad (35)$$

where  $w_L$  is a modulating window given by

$$w_L[u] = \sin\left[\frac{\pi}{2L} \left(u + \frac{1}{2}\right)\right]. \quad (36)$$

### Question 4 Sparse coding with OMP

For the signal provided in the notebook, learn a sparse representation with MDCT atoms. The dictionary is defined as the concatenation of all shifted MDCT atoms for scales  $L$  in  $[32, 64, 128, 256, 512, 1024]$ .

- For the sparse coding, implement the Orthogonal Matching Pursuit (OMP). (Use convolutions to compute the correlation coefficients.)
- Display the norm of the successive residuals and the reconstructed signal with 10 atoms.

### Answer 4

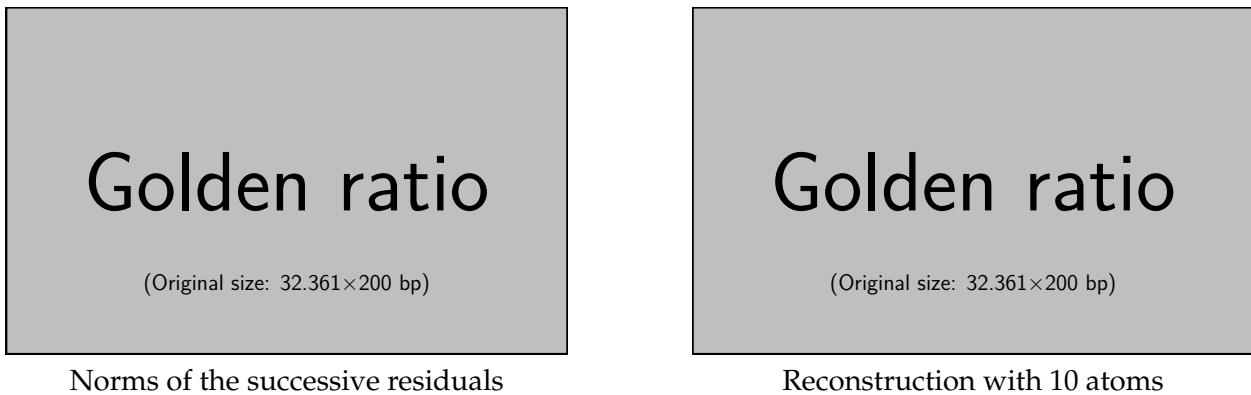


Figure 3: Question 4