The Effect of Employment Protection on Firms' Worker Selection

Jan Sauermann (SOFI, Stockholm University) with Sebastian Butschek (University of Cologne)

IFAU Seminar, February 17, 2020



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- ullet Anticipated firing costs reduce hiring but also firing (Bertola, 1999) o ambiguous effect on employment levels (OECD, 2013)
- Not much is known about how EPL affects which types of workers are hired
 - Distributional effects?
 - Hiring standards?
- Theoretical predictions
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 - Greater variance implies option value for high-productivity workers → private information about (match-specific) quality
 - ullet More stringent EPL o "safer" hires, i.e. firms reduce spread of their hires' ability distribution
 - Pries and Rogerson (2005):
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 - ullet argument that unemployed individuals are harder to screen o reform makes treated firms screen less
- Effects on labor turnover
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- Research question: does employment protection legislation (EPL) affect hiring standards?
 - We exploit the 2001 last-in-first-out (LIFO turordningsregler) reform in Sweden
 - Reform granted more exemptions for small firms
 - Did small firms change their hiring behavior, compared to larger firms?
 - Firm size threshold used for difference-in-differences framework
- We are not the first to study the 2001 reform:
 - Firm-level outcomes: firm productivity ↑ (Bjuggren, 2018); labor turnover ↑ employment → (von Below and Thoursie, 2010); firm growth → (Bornhäll et al., 2017)
 - Worker-level outcomes: sickness absence ↓ (Lindbeck et al., 2006; Olsson, 2009);
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- Preview of our findings
 - Less stringent EPL ⇒ establishments decrease their hiring standard
 - Hiring standard: minimum worker ability that is hired
 - Changes in screening activities likely underlying mechanism
- Contribution to the literature
 - Little direct evidence on how EPL affects who is hired and why this is → findings consistent with related papers (Marinescu, 2009; Bjuggren and Skedinger, 2018)
 - More evidence on the determinants of selectivity in worker-firm matching (e.g., Balsvik and Haller, 2015: foreign ownership; Bender et al., 2018: management quality; Hensvik and Skans, 2016: networks)
- And why is this important?
 - Efficiency perspective: firms' selectiveness affects allocation of workers
 - Distributional implications: disadvantaged groups increasingly hired?
 - How should we optimally design EPL policies? → current policy discussion in Sweden regarding labor market policies (LAS)

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Remainder of the presentation

Institutional background

Empirical strategy and data

Main results

Threats to identification

Conclusion and discussion

Institutional background

- Swedish EPL (*Lagen om anställningsskydd, LAS*) is relatively strict: Dismissals allowed in case of
 - misconduct
 - economic necessity → Last-in-first-out (LIFO) rule is applied
- LIFO first introduced in 1974
- We study the 2001 LIFO reform: introduction of exemption for small firms
 - \bullet Pre-reform (< 2001): LIFO applied to all firms \rightarrow last hire is laid off first
 - Post-reform(≥ 2001): small firms can choose between three most recent hires
 - How are small firms defined? ≤ 10 employees

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• To which degree are firms affected by the reform?

firm size	Pre-reform		Post-reform		Share protected
	Protected	Unprotected	Protected	Unprotected	becoming unp.
2	1	1	0	2	-100%
3	2	1	0	3	-100%
4	3	1	1	3	-67%
5	4	1	2	3	-50%
6	5	1	3	3	-40%
7	6	1	4	3	-34%
8	7	1	5	3	-29%
9	8	1	6	3	-25%
10	9	1	7	3	-22%
11+	10+	1	10+	1	0%

- Implementation of the 2001 reform:
 - Feb 2000: first proposal in parliament
 - Oct 2000:
 - Written into law
 - "Unlikely coalition": liberal-conservatives and greens against then-governing social democrats
 - Jan 2001: introduction
- Possible anticipation effects in 2000?
 - Expectation that hirings can be reversed

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Aim: estimate the effect of the 2001 relaxation of Sweden's LIFO for small establishment on hiring standards

• Difference-in-differences approach

$$y_{jt} = \alpha + \beta TR * POST_{jt} + \gamma_t + \delta TR_j + \epsilon_{jt}.$$

- y_{it}: hiring standard
- TR:
 - dummy for small firms (≤ 10 employees)
- POST:
 - 0=before (1993-1999)
 - 1=after (2001-2004)
 - (Possible anticipation effects in 2000)
- β: estimate of interest

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- Key assumption 1: absent of reform treatment and control firms would have followed similar trend ($\mathbb{E} \epsilon_{jt} | TR_j * POST_t = 0$)
- Key assumption 2: no major reforms at same time *and* with same treatment definition
 - ullet 1994: temporary exemption from LIFO o diff. year; no size threshold
 - ullet 1997: reform of temporary contracts o diff. year; no size threshold
 - ullet 2001: gender equality act o same year and similar threshold
- Treatment definition:
 - firm had \leq 10 employees in 1999
 - assumption that reform was not anticipated in 1999 (first draft 02/2000) cf. Bjuggren (2018)
- What do we estimate?
 - 1. Estimates intention-to-treat effect (ITT) time-invariant firm size (1999)
 - 2. Endogenous sorting of firms \rightarrow additionally estimate LATE with 1999 size as instrument for current firm size

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- How do we measure minimum hire quality $y_{jt} = min_{jt} \{ability_i\}$?
 - 1. Main measure: estimated AKM worker fixed effects
 - Years 1986-1992 used for estimation details
 - Birth cohorts 1922 to 1976
 - 2. Military draft cognitive and psychological test scores
 - Birth cohorts 1951 to 1991
 - 3. High school GPA age 15
 - Birth cohorts 1973 to 1982
- Why focus on AKM?
 - Male and female hires
 - Relatively long coverage

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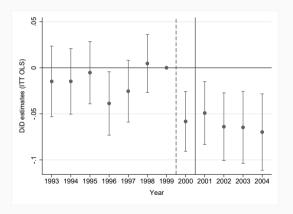
Individual-level Swedish register data (Statistics Sweden: SCB)

- Employment spell information (JOBB): timing of formation/dissolution of worker-firm matches, industry, location, firm age, public/private sector, firm size
 - Estimation of AKM worker fixed effects (1986-1992)
 - Information on hires (1993-2004)
- Education information (LISA)
- Grade 9 GPA (årskurs-9)
- Armed forces' database: military test scores

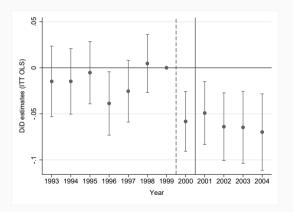
- Sample size and definitions
 - 1986-1992: sample used only for AKM estimation
 - 1993-2004
 - 1993-1999: pre-period
 - 2000 excluded: potential anticipation effects (excluded)
 - 2001-2004: post-period
- Firm level restrictions
 - Firm growth: exclude outliers with extreme firm growth (.5%)
 - Firms: keep only firms that exist in 1999 (N = 275,731)
- Firm size: keep only firms of size 2-15. (N = 129, 187)

Firm characteristics by treatment status

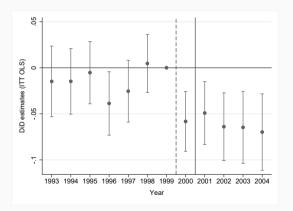
A: Continuous characteristics				
	Treated		Control	
	mean	sd	mean	sd
Head count (rounded)	4.857	2.453	12.688	1.388
Firm age (years)	8.171	4.548	8.859	4.407
Mean worker age (years)	40.158	9.311	38.644	8.532
Female worker share	0.409	0.297	0.373	0.267
Mean worker years of schooling	11.418	1.450	11.356	1.222
Mean worker monthly wage (100 SEK 1980)	45.080	29.996	45.506	29.199
Mean worker AKM person effect (std, 1986-92)	-0.007	0.998	-0.022	0.817
Estimated AKM firm effect (std, 1986-92)	-0.007	0.992	0.038	0.995
Observations	87,441		13,947	



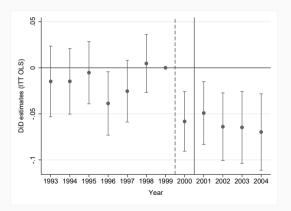
- Minimum hiring quality consistently lower in the post-period
- Less stringent EPL → significantly lower hiring standard
- 2000: likely anticipation effect (excluded from regressions)
- Results similar for military test scores and GPA details



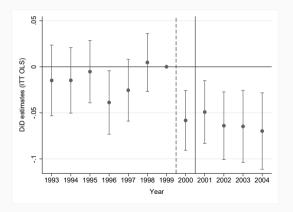
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• Quantify the effect of loosening EPL for different outcomes (standardized)

	(1)	(2)	(3)	(4)
	AKM	COG	NON-COG	GPA
DiD estimate (Treated*Post)=1	-0.0484***	-0.0603***	-0.0488***	-0.0724***
	(0.0096)	(0.0126)	(0.0124)	(0.0150)
Observations	314,144	193,025	189,923	132,281
Firms	101,388	79,923	79,210	63,909
Adjusted R ²	0.0027	0.0062	0.0058	0.0071

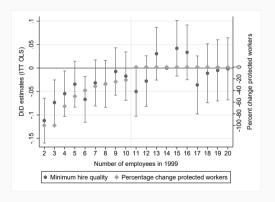
- Consistent results across outcome measures
- Similar results when including controls and firm FE Indicate of the state of the

4. Main results: heterogeneous effects

- Heterogeneous effects: effects by firm size (reform bite)
 - Smallest treatment group firms (2-3 workers): 100% of protected workers become unprotected
 - Largest treatment group firms (10 employees): 22% of protected workers become unprotected
- Smallest firms should be affected the strongest

4. Main results: heterogeneous effects

EPL effect heterogeneity by firm size (AKM)



- Dark gray dots show firm size-specific DiD estimates
- Light gray dots show change in share of protected workers becoming unprotected
- Results show that effects are much stronger for smallest (=most affected) firms

- Affected firms react to the reform by hiring more workers with lower ability
- Why do firms adjust their hiring standard?
- What are potential mechanisms behind this result?
 - Mechanism 1: mechanical labor turnover effect. Increased labor turnover implies
 more draws from the pool of applicants → this mechanically increases variation in
 hired individuals.
 - Mechanism 2: screening effect. When laying off is easier (cheaper) → firms can afford to reduce their screening intensity.
 - Mechanism 3: risky hiring. Increased flexibility in firing allows firms to take more risks in learning about individuals productivity

 implies greater variance in hiring

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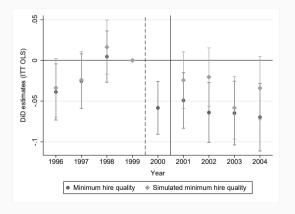
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- Mechanism 1: number of hires effect
 - Reform has been shown to increase labor turnover (von Below and Skogman Thoursie 2010)
 - Firms make more draws from the pool of potential hires
 - Simulation exercise:
 - Assume that hiring standard has not changed and that only the number of hires has changed
 - Randomly reshuffle hires within firm size groups (1996-2004)
 - ullet If purely mechanical o results based on simulated data should fully explain our findings

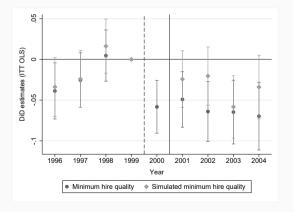
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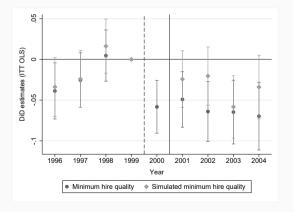
- Actual effects (dark dots) vs. simulated effects (light gray)
- Simulated effects more attenuated
- Actual effects about twice the size of simulated effects (-.048*** vs -.025***)
- Suggests that labor turnover does not fully explain our main finding

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- Actual effects about twice the size of simulated effects (-.048*** vs -.025***)
- Suggests that labor turnover does not fully explain our main finding

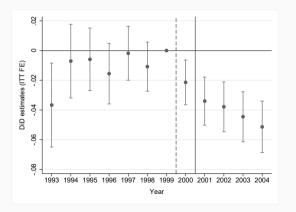
- Mechanism 2: Screening effect \rightarrow firms decrease their screening efforts
- Pries & Rogerson (2005): less severe consequences makes firms more likely to hire somebody with lower quality
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 - Nordström Skans and Hensvik (2016) have provided evidence that firms' use for network based hiring for screening purposes
 - We use a measure of the links between new hires and incumbent workers as ar indication of firm's screening intensity
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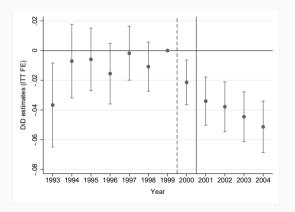
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EPL effect on network-based hiring



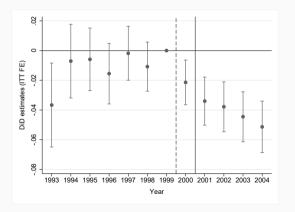
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- Result suggests that treated firms decreased their screening activities
- Caveat:
 - Without fixed effects, pre-trends not parallel
 - Outlier 1993

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- Mechanism 3: risky hires channel
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 - Productivity of individuals at both ends of the ability distribution is more difficult to assess
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	Minimum		Maximum	
	Actual	Simulated	Actual	Simulated
DiD estimate (Treated*Post)=1	-0.0478***	-0.0253**	0.0309***	0.0390***
	(0.0102)	(0.0099)	(0.0104)	(0.0104)
Observations	232,707	244,459	232,707	244,459
Firms	92,164	95,473	92,164	95,473
Adjusted R ²	0.0028	0.0017	0.0026	0.0023

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4. Main results: summary of findings

- Summary main results
 - Firms react to change in EPL by adjusting hiring standard
 - Results partially explained by purely mechanical effects (labor turnover)
 - Suggestive evidence for screening explanation

Threats to identification

5. Threats to identification

- Gender equality act: coinciding reform with similar size threshold details
- Demand vs. supply: sorting of workers into treatment and control firms details
- Non-classical measurement error: estimation error in AKM estimation and sorting of workers into treatment and control firms details
- Selective worker ability measurement: availability of AKM estimate related to firms' treatment status
- Time-invariant firm heterogeneity: changes in firm composition over time details
- Endogenous firm size and overcorrection: IV estimation to account for changes in firm size details
- Non-parallel trends: placebo tests details

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 - Mechanical effects are important
 - Importance of mechanical effects caused by labor turnover
 - Evidence for a screening-based mechanism
- Robustness tests suggest that these effects can be interpreted as causal

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• How do our results relate to previous research?

- Bjuggren and Skedinger (2018): same reform increased hiring from previously unemployed individuals → consistent with our finding that *lower* end of the hiring standard is affected
- Bjuggren (2018): reform increased firms' labour productivity
 - Inconsistent with our results of hiring more lower-ability individuals?
 - No! Reform also increased labor turnover (von Below and Skogman Thoursie 2010)
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 - Firms hire more widely and may retain the best workers
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 - Following 2018 GE: plans on reforming employment protection (LAS)
 - Expanding exemption rules of turordningsregler
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Measurement: AKM worker fixed effects

 We follow Abowd et al. (1999) and Card et al. (2013) and estimate two-way fixed-effects regression:

$$\ln(w_{ijt}) = \alpha_i + \psi_j + \gamma_t + x'_{it}\beta + r_{ijt},$$

- $\ln(w_{ijt})$ is the natural logarithm of individual i's hourly wage at firm j in year t
- Additive fixed effects for individuals (α_i) and firms (ψ_j)
- Year dummies (γ_t)
- Vector of time-varying individual-level controls (x_{it})
 - Age squared and age cubed
 - Education categories interacted with the year dummies, age squared and age cubed

Measurement: AKM worker fixed effects



- Use individual-level spell data on employment from JOBB
- Obtain full time-equivalent (FTE) monthly wages from the Wage Survey Statistics (WSS – survey based)
 - CPI deflated
 - Winsorize at 0.5% and 99.5% of the annual real monthly FTE wage distribution
 - WSS survey based: covers only stratified random sample of smaller firms
- Sample definition (1986 to 1992)
 - Individuals aged 18-65 with a November spell with FTE wage data
 - 45% of person-year observations have FTE information
 - 64% of all individuals have at last one FTE observation
 - ullet Education information available from 1990 o imputed from first year available

Measurement: military draft test scores



Military test scores

- Individual measurement: nearly all men at age 18-19.
- 9-point scale.
- Approximately normally distributed.
- Hiring standard: use minimum hire ability at firm-year level.

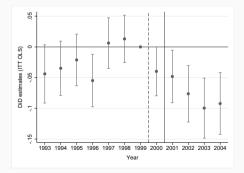
Measurement: military draft test scores

- Cognitive ability: 4 sub-scores from written tests
 - Inductive skill/reasoning
 - Verbal comprehension
 - Spatial ability
 - Technical understanding.
- Non-cognitive ability: 4 sub-scores from behavioral interview with psychologist
 - Social maturity
 - Psychological energy (focus and perseverance)
 - Intensity (e.g., activation without pressure)
 - Emotional stability (e.g., tolerance to stress).

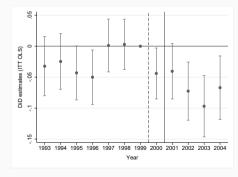
DiD estimates for other outcomes

• Dynamic effects for (1) cognitive test scores (military); (2) psychological test scores (military); (3) GPA test scores age 15

(1) cognitive test scores (military)



(2) psychological test scores (military)

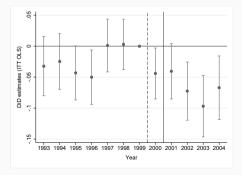


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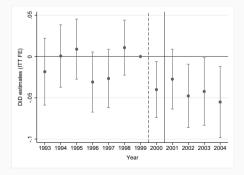
(3) GPA test score age 15



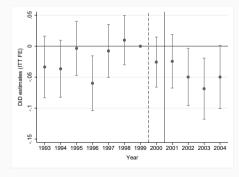
DiD estimates for other outcomes (with FE)

• Dynamic effects for (1) AKM; (2) cognitive test scores (military); (3) psychological test scores (military); (4) GPA test scores age 15

(1) AKM



(2) cognitive test scores (military)

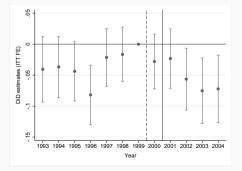


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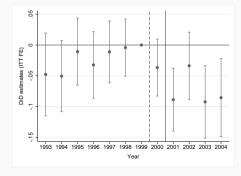


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(4) GPA test score age 15



Threat to identification: Gender equality act



- Reform Gender Equality Act (GEA) effective 2001 (SFS, 2000)
 - GEA introduced in 1992 (SFS, 1991): firms with \geq 10 employees are required to publish document specifying measure to promote gender equality
 - Reform made wording more concrete, but no sanctions or penalties
 - No econometric evaluation of the reform available
 - GEA-treatment group (10-15) \approx LIFO control group (11-15)
- ullet Hypothesis: GEA treatment group o favoring women to "comply" with GEA
 - LIFO control group hires more women
 - \bullet Labor market for LIFO treatment group tighter $\to \beta$ biased towards zero

	Women	Men
DiD estimate (Treated*Post)=1	-0.0353***	-0.0522***
	(0.0128)	(0.0122)
Observations (firms)	180,194 (76,453)	192,387 (78,853)
Adjusted R ²	0.0012	0.0019

Threat to identification: Demand vs. supply



- Implicit assumption that self-selection of workers into treatment / control firms not affected by the reform
- Ambiguous effects
 - Short-run: not most recent hire needs to be laid off in treated firms
 - Medium-run: seniority capital lower since lower protection in treated firms
- Information on applicants? No
- ullet Voluntary job-job transitions o sorting patterns systematic to the reform?

	E-E hires
DiD estimate (Treated*Post)=1	0.0109
	(0.0084)
Observations	314,144
Firms	101,388
Adjusted R ²	0.0022

Threat to identification: Non-classical measurement error



- Measurement error in estimated worker fixed effects (AKM)
 - AKMs are estimated using period before pre-period starts
 - ullet As long as treated and control firms are equally affected by measurement error ullet will be differenced out
 - ullet But: women facing wage discrimination o AKMs understate their ability
 - What if women are increasingly hired in treated firms?
 - Lowering hiring threshold could be due to underestimation of ability

	Share Women
DiD estimate (Treated*Post)=1	-0.0040
	(0.0085)
Observations	314,144
Firms	101,388
Adjusted R ²	0.0014

Threat to identification: Selective worker ability measurement



- AKM worker fixed effects cannot be estimated for the entire population (55% of all new hires)
- Problematic?
 - Yes if AKM measurement is positively related to the reform
 - This can create mechanical effects

	Share AKM
DiD estimate (Treated*Post)=1	-0.0196**
	(0.0089)
Observations	314,144
Firms	101,388
Adjusted R ²	0.0292

Threat to identification: Time-invariant firm heterogeneity



- Does firm composition change over time?
- Example: business cycles creating many new small firms (but not large ones)
- Problematic?
 - We condition on firms having existed in 1999 (when we measure firm size)
 - Firm fixed effects estimation (Figures: details)

	OLS	FE
DiD estimate (Treated*Post)=1	-0.0484***	-0.0359***
	(0.0096)	(0.0104)
Observations	314,144	314,144
Firms	101,388	101,388
Adjusted R ²	0.0027	0.0010 .

Threat to identification: Endogenous firm size and overcorrection



 Local average treatment effect (LATE) estimation to allow for time-varying firm size

• Instrument: 1999 firm size

	OLS	IV
DiD estimate (Treated*Post)=1	-0.0484***	-0.1162***
	(0.0096)	(0.0207)
Observations	314,144	314,144
Firms	101,388	101,388
Adjusted R ²	0.0027	

Threat to identification: Non-parallel trends



- Would treated firms changed hiring behavior also in absence of reform?
- \bullet Placebo treatment at actual threshold +20, 30 and 40

Placebo firm size cut-off	20	30	40
DiD estimate (Treated*Post)=1	0.0016	-0.0061	-0.0012
	(0.0029)	(0.0040)	(0.0050)
Observations	109,735	54,044	33,130
Firms	22,331	9,063	4,902
Adjusted R ²	0.0051	0.0064	0.0088