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# Genome Analysis

# Dryhic: Increasing Reproducibility of Hi-C Samples with Abnormal Karyotypes

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#### **Abstract**

**Motivation:** Several strategies have been proposed to remove biases from Hi-C experiments. However, samples with abnormal karyotypes (such as cancer cell lines) present challenges not completely addressed by existing methods. We introduce an alternative designed to model the total number of contacts per genomic loci, thus taking into account possible chromosomal abnormalities.

**Results:** The proposed method increases reproducibility among replicas of Hi-C samples with abnormal karyotype and obtained using different protocols. The implementation is fast and scales well in terms of computing resources for resolutions up to 1 Kbp.

**Availability:** The strategy proposed is implemented as an R package available at http://www.github.com/qenvio/dryhic.

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Supplementary information: Supplementary data are available at Bioinformatics online.

## 1 Introduction

High throughput chromosomal conformation capture (HiC) experiments allow us to explore the spatial organization of chromatin in cells (Lieberman-Aiden *et al.*, 2009). Contrary to other 3C techniques, HiC is unbiased in the sense that all possible pairwise interactions between restriction fragments might be interrogated. However, this does not guarantee absence of bias. For instance, genome features as CG content, restriction enzyme site distribution and sequence uniqueness (related to mappability) have been described to modify the result in such experiments (Yaffe and Tanay, 2011).

Several strategies have been proposed to remove biases in HiC experiments (Schmitt *et al.*, 2016), either modeling explicitly a set of genomic features (Yaffe and Tanay, 2011; Hu *et al.*, 2012) or using algorithms to implicitly correct for unknown biases (Imakaev *et al.*, 2012), the latter often being preferred due to its fast implementation. These methods are suitable for normal cells but they were not designed for cells with abnormal karyotypes. Although there have been some proposals that consider aneuploidy (Wu and Michor, 2016), to our knowledge, there

is a lack of methodology dealing with aberrant karyotypes (common in cancer cell lines) that present abnormalities more complex than entire chromosome duplications.

Here we propose oned, an alternative designed to model total number of contacts per genomic loci, thus taking into account possible chromosomal abnormalities. It explicitly models the contribution of known biases via a generalized additive model (GAM). The resulting data present better overall reproducibility between replicas and across different protocols. Besides, the implementation is fast and scales up to 1 Kbp resolution with reasonable computing resources.

## 2 Methods

#### 2.1 Model

A common representation for the data yielded in a HiC experiment is a contact matrix, obtained slicing the genome in contiguous bins of fixed length (resolution) and computing the number of contacts between each pair of bins, corresponding to a cell in the contact matrix  $(x_{ij})$ .

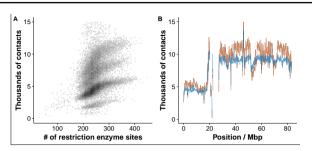
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**Fig. 1.** A Non linear relationship between a genomic feature (number of restriction sites per bin) and the total number of contacts per bin for a T47D sample. B Total number of contacts per bin in chromosome 17 of a T47D sample for raw data (in brown, noisier) and corrected data (in blue, preciser).

We assumed that the total number of contacts per bin  $(t_i)$  could be approximated by a negative binomial distribution with every known genomic feature contributing independently to its mean  $(\lambda_i)$ . Given that this relationship might not be linear (Figure 1A), we allowed a smooth representation using thin plate penalized regression splines (Wood, 2003) in a generalized additive model (Wood, 2011). The model can be parametrized as follows:

$$t_i = \sum_i x_{ij}$$
$$t_i \sim NB(\lambda_i, \theta)$$
$$log(\lambda_i) \propto \sum_K f_k(z_k)$$

where  $x_{ij}$  are the number of contacts between bins i and j, and there is an additive term for each  $z_k$  genomic feature. The smooth functions  $\{f_k(z_k)\}$  are estimated jointly with the negative binomial dispersion parameter  $\theta$  using R (R Core Team, 2017) mgcv package (Wood, 2011).

Once the model of the total counts has been estimated, we can further rescale  $\{\lambda_i\}$  to obtain a correction vector  $\{\lambda_i'\}$  that can be used to compute the corrected counts  $(c_{ij})$  or, alternatively, it can be included in downstream analysis:

$$\lambda_i' = \frac{\lambda_i}{\sum_j^n \lambda_j / n}$$
$$c_{ij} = \frac{x_{ij}}{\sqrt{\lambda_i' \lambda_j'}}$$

Following previous works (Yaffe and Tanay, 2011; Hu *et al.*, 2012) we included mappability, CG content and number of restriction sites (comparable to effective length) as genomic features, but the model and implementation allow easy extension with different genomic features.

## 2.2 Data sources

As an environment to test the bias correction, we gather a set of published (Le Dily et al., 2014; Consortium et al., 2012; Rao et al., 2014; Stadhouders et al., 2017; Lin et al., 2012; Dixon et al., 2012) and unpublished data from HiC experiments of different cell types and organisms. The details can be found in Table ??. Briefly, we used several experiments comprising T47D breast cancer cell lines (7 samples), K562 leukemia cell lines (4 samples), both with aberrant karyotypes and mouse primary B (6 samples) and ES (7 samples) cells, both with normal karyotypes. Experiments used different protocols (i.e. original HiC (Lieberman-Aiden et al., 2009) and in-situ HiC (Rao et al., 2014)), different restriction enzymes (DpnII, HindIII, MboI and NcoI) and produced in distinct laboratories.

We also make use of array-based copy-number segmentation of the two cell lines obtained from COSMIC database (Forbes *et al.*, 2010).

#### 2.3 Data processing

All data were processed though a in-house pipeline based on TADBit (Serra et al., 2016). For more details, see methods of Stadhouders et al. (2017). The pipeline is available, with minor modifications, in Quilez et al. (2017) material. Briefly, after fastq quality control, paired reads were mapped against the corresponding reference genome (hg38 and mm10) taking into account the restriction enzyme motifs. Non-informative contacts were removed applying the following TADBit (Serra et al., 2016) filters: Self-circle, dangling-ends, error, extra dangling-end, duplicated and random-breaks. Then, we developed the routines contained in the dryhic R package to efficiently create sparse representations of contact matrices (at 100 Kb resolution) and further apply vanilla, ice and oned corrections. HiTC (Servant et al., 2012) and HiCapp (Wu and Michor, 2016) were used to get lgf and caib corrections respectively.

#### 2.4 Total number of contacts

In non-diploid samples, we expect the coverage to be directly related to the number of copies (a duplicated region will show, on average, twice the number of contacts). Given that our model is based on total number of contacts, which in turn is directly related to the number of copies, we reasoned that a preliminary test would be to check if the corrected number of contacts per bin  $(ct_i = \sum_i c_{ij})$  reflects the copy number (as measured by an independent technique such as array-based copy-number segmentation).

#### 2.5 Sample comparison

In order to test performance, we relied on pair-wise comparisons of contact matrices. However, there is little consensus on how to compare two HiC matrices. The simplest metric is the Spearman correlation applied to cis (intra-chromosomal) contacts up to a given distance (i.e. 5 Mb). Though, the decay in the number of contacts with increasing genomic distance is the main driver of contact abundance and it could mask real differences between samples. This distance decay can be taken into account, for instance, computing the expected number of interactions given the genomic distance between any pair of loci. Thus, we can compare matrices in terms of observed over expected contacts, via the Pearson correlation (again focusing on a given distance range). Alternatively, we can compute the correlation per distance stratum and then obtain a stratumadjusted correlation coefficient (SCC) as defined in Yang et al. (2017). Additionally, Yan et al. (2017) proposed a reproducibility score based on the correlation between a set of the Laplacian matrix's eigenvectors (corresponding to the smallest eigenvalues) for the two experiments.

We defined three sets of samples: One with all T47D experiments plus two K562 samples, another with all K562 experiments plus two T47D samples and a third one comprising all mm10 samples (see Table ?? for details). Given a set of experiments and a metric, we first computed all pair-wise combinations between experiments and then classified the comparisons according to the characteristics of each pair (*i.e.* same / different cell types, same / different protocols, etc ...).

To benchmark our method (*oned*), we also test the performance of other alternatives. Three of them were based on the same implicit iterative correction (Imakaev *et al.*, 2012): One computing one iteration (*vanilla*), another iterating until convergence (*ice*) and a third one using the chromosome adjusted iterative correction (*caib*) (Wu and Michor, 2016). The fourth one was inspired on the explicit bias correction originally presented by Yaffe and Tanay (2011), but based on the local genomic features (*lgf*) approach introduced by Hu *et al.* (2012) as implemented in Servant *et al.* (2012).

Besides, we compared the raw matrices and used these comparisons as a baseline to present the gain / loss of correlation relative to the uncorrected









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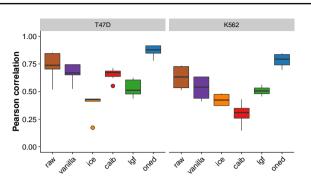


Fig. 2. Pearson correlation between total number of contacts per bin and independent copy number estimation (COSMIC) for each of the methods compared. Left panel T47D breast cancer cell line, right panel K562 leukemia cell line. The new proposal (in blue) outperforms the rest of alternatives.

data, thus allowing us to verify the effect of the bias correction procedure. The differences relative to the raw comparisons were estimated using a linear mixed model with the method as fixed effect and the chromosome as a random effect, and using the lmer function of the lme4 package (Bates *et al.*, 2015).

In an attempt to summarize the results, we reasoned that the correlation between unrelated samples (*i.e* different cell types) should be lower than between related samples (*i.e* same cell types), so we could re-interpret each correlation metric as a classifier score. Thus, we computed the receiver operating characteristic (ROC) curves and corresponding area under the curve (AUC) using the ROCR package (Sing *et al.*, 2005).

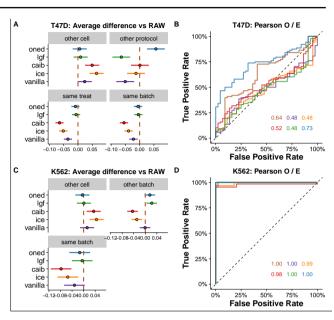
## 3 Results

## 3.1 Copy number

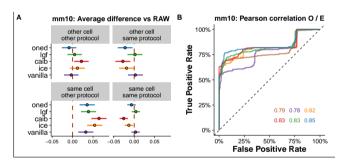
Figure 1B shows the total number of corrected contacts along chromosme 17 of a T47D sample. Our correction approach presents a more stable and precise estimation of the copy number. Quantitatively (Figure 2), Pearson correlation with independent copy number data is high for both T47D and K562 cell lines (averages of 0.87 and 0.78 respectively), not only compared with raw data (averages of 0.74 and 0.63 respectively) but also with any of the alternative methods (averages ranging 0.39-0.68 and 0.30-0.53 respectively). Similar results were obtained using Spearman correlation (Supplementary Figure 2).

### 3.2 Samples with aberrant karyotypes

Focusing on the Pearson correlation of observed over expected counts, methods based on the iterative correction (mainly *ice* and *caib* but also *vanilla*) present the undesirable feature of higher correlations between unrelated samples than using raw data (top left panels of Figures 3A and 3C, Supplementary Figure 3). Among the comparisons on cells of the same type, *oned* is the only method outperforming uncorrected data when comparing experiments using different protocols (top right panel of Figure 3A) and it is among the top performant when comparing experiments using the same protocol (rest of panels of Figures 3A and 3C). Consequently, our proposal is the one that better discriminates unrelated samples form similar ones (as it can be seen in the ROC curves and AUCs of Figures 3B and D). Complementary analysis using other metrics (Supplementary Figures 3, 4 and 5) yielded similar results. Notably, Spearman correlation of contacts presents the worst performance for all scenarios and seems to be a poor choice as a metric to compare HiC matrices.



**Fig. 3.** Results of the comparison between samples with aberrant karyotype. Left panels: Average difference between each correction method (Y axis) and the raw data (brown dashed line) with 95% CI for T47D (A) and K562 (C) sets. Right panels: Corresponding ROC curves and AUCs for T47D (B) and K562 (D) sets. All results in this figure are based on Pearson correlations between the observed over expected counts. Our method increases reproducibility of similar samples while being able to discriminate different ones.



**Fig. 4.** Results of the comparison between samples with normal karyotype. A Average difference between each correction method (Y axis) and the raw data (brown dashed line) with 95% CI for the mm10 sets. B Corresponding ROC curves and AUCs. The introduced strategy do not under perform compared to the alternatives.

#### 3.3 Samples with normal karyotypes

Regarding cells with diploid genomes, we confirmed that our bias correction strategy also achieves good results in terms separation between unrelated and similar samples (Figure 4 and Supplementary Figures 6 and 7), increasing the reproducibility across different experimental protocols for similar primary B and ES mouse cell samples. Again, Spearman correlation of contacts revealed as the worst metric.

# 3.4 Computing resources

One of the main reasons for the broad usage of the iterative correction (Imakaev *et al.*, 2012) is the easy and fast implementation, moreover compared to existing explicit methods (Servant *et al.*, 2012). Our proposal uses only the total number of contacts per bin instead the full matrix for the model estimation. Thus, it achieves similar computing times to the fastest approaches (Figure 5 A). Yet it scales with increasing resolutions (decreassing bin size) (Figure 5 B).

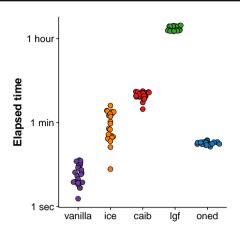








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**Fig. 5.** Computing time (Y axis, log scale) of all the bias methods used. Each dot corresponds to one sample. The only method faster than ours under performs in all sample comparisons.

# 4 Discussion and Conclusions Acknowledgements Funding

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