

# Stochastic dynamic models for low count observations (and forecasting from them)

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# Overview

- 1 Poisson-binomial state-space model
- 2 Inference and forecasting
- 3 Forecasting with noisy observations
- 4 Forecasting with inference on parameters

# S→I→R stochastic compartment model

Focus on Susceptible (S) - Infectious (I) - Recovered (R) model with Poisson transitions, i.e.

$$\vec{SI} \sim Po(\lambda_{S \rightarrow I} SI/N) \quad \vec{IR} \sim Po(\lambda_{I \rightarrow R} I)$$

updating the states, i.e.

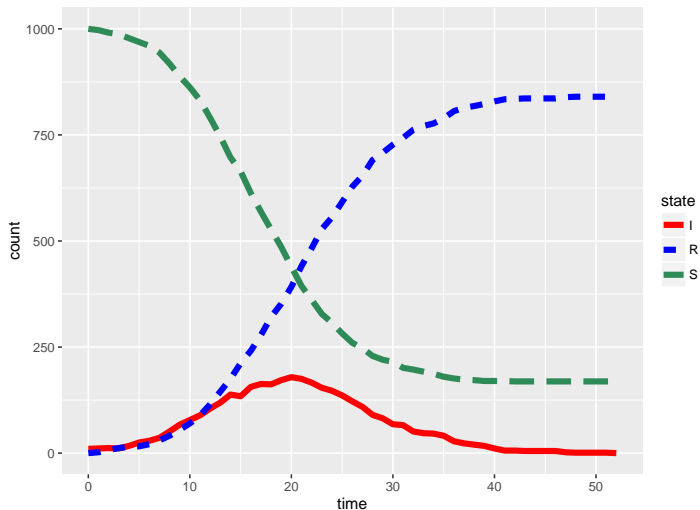
$$S_{t+1} = S_t - \vec{SI}, \quad I_{t+1} = I_t + \vec{SI} - \vec{IR}, \quad R_{t+1} = R_t + \vec{IR}$$

and binomial observations

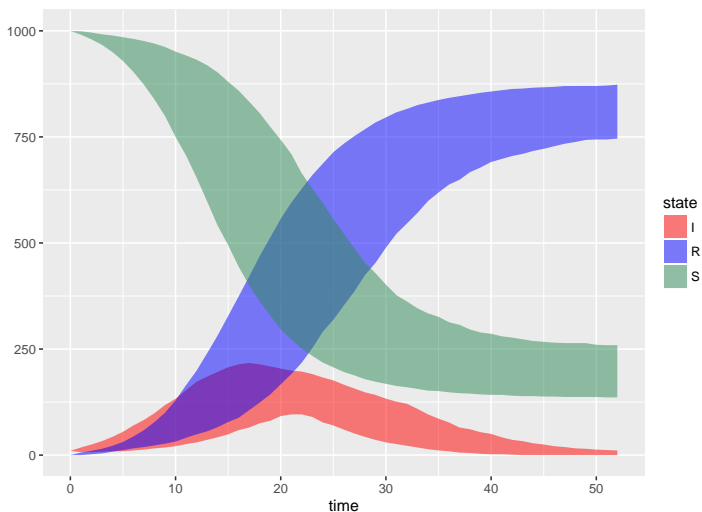
$$Y_{S \rightarrow I} \sim Bin(\vec{SI}, \theta_{S \rightarrow I}) \quad Y_{I \rightarrow R} \sim Bin(\vec{IR}, \theta_{I \rightarrow R})$$

A more general stochastic dynamic modeling structure can be used to extended to geographical regions, subpopulations, etc.

# SIR modeling simulations



# SIR modeling simulations

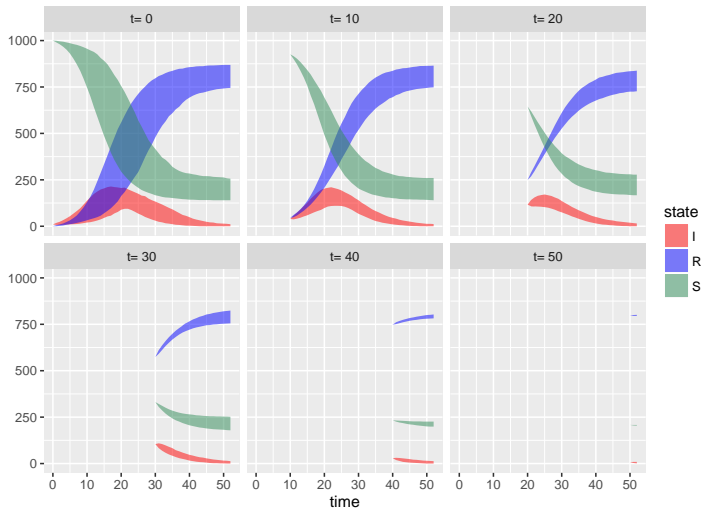


# Forecasting with perfect information

Suppose you know transition rates  $\lambda$ , observation probabilities  $\theta = 1$ , and the states  $X_{0:t}$  and your only goal is to forecast the future states  $X_{t+1:T}$ , i.e.

$$p(X_{t+1:T}|\theta, \lambda, X_{0:t}) = p(X_{t+1:T}|\theta, \lambda, X_t)$$

this distribution is estimated via Monte Carlo simulation.



# Delay in data analysis

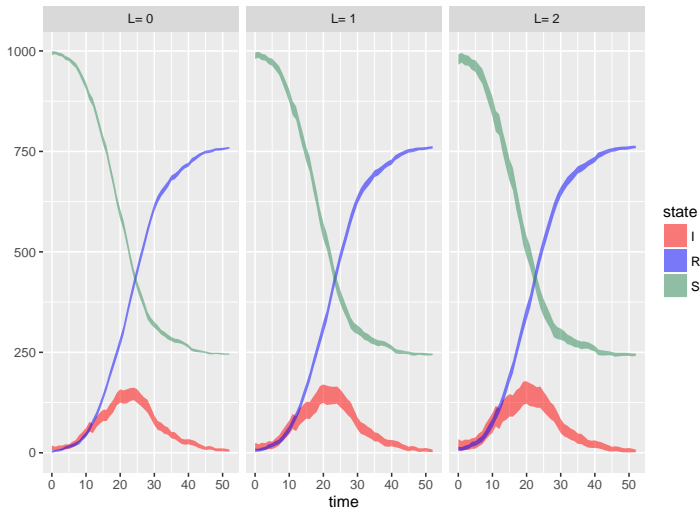
Suppose you have a one or two week delay in collecting, processing, and analyzing so that when trying to forecast future states you are using “old” data, i.e.

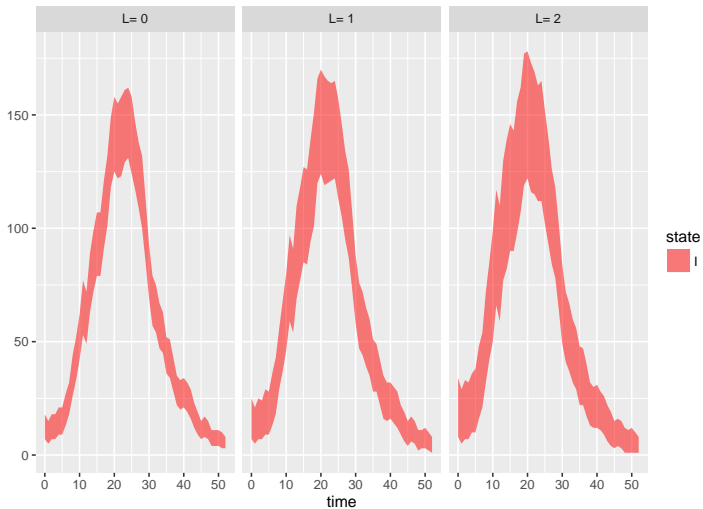
$$p(X_{t+1:T} | \theta, \lambda, X_{t-L})$$

where

- $L = 0$  indicates up-to-date data
- $L = 1$  indicates one-week old data
- $L = 2$  indicates two-week old data







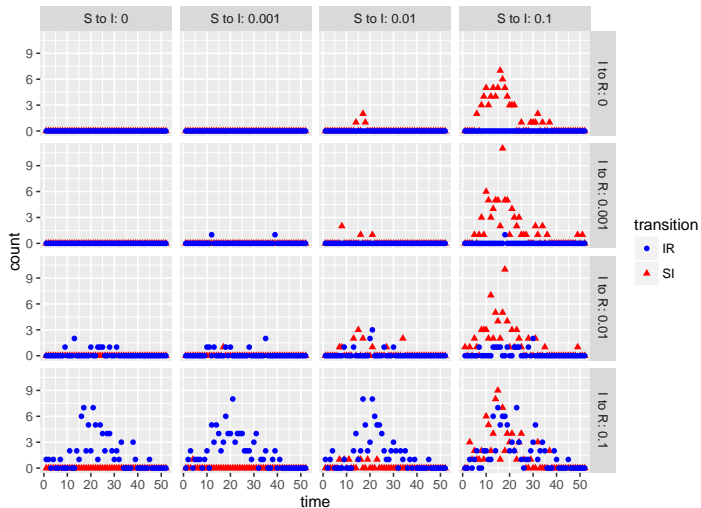
# Forecasting with noisy observations

Suppose, we know the transition rates ( $\lambda$ ) and the observation probabilities ( $\theta$ ), but we only observe a noisy version of the state transitions, i.e. .

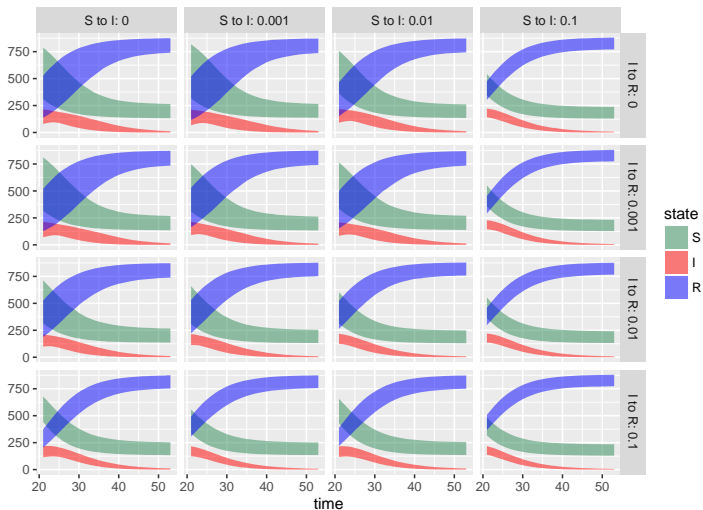
$$Y_{S \rightarrow I} \sim \text{Bin}(\vec{SI}, \theta_{S \rightarrow I}) \quad Y_{I \rightarrow R} \sim \text{Bin}(\vec{IR}, \theta_{I \rightarrow R})$$

Now the forecast distribution we need is

$$p(X_{t+1:T} | \lambda, \theta, y_{0:t}) = \int p(X_{t+1:T}, \lambda, \theta | X_t) p(X_t | \lambda, \theta, y_{0:t}) dX_t.$$



# Noisily observed state



# Forecasting with inference on parameters

In reality, we don't know the transition rates ( $\lambda$ ) and the observation probabilities ( $\theta$ ), and we only observe a noisy version of the state transitions.

Now the forecast distribution we need is

$$p(X_{t+1:T}|y_{0:t}) = \int \int \int p(X_{t+1:T}|\lambda, \theta, X_t) p(X_t, \lambda, \theta|y_{0:t}) d\lambda d\theta dX_t.$$

# Prior distributions

In order to calculate (or approximate) the integral

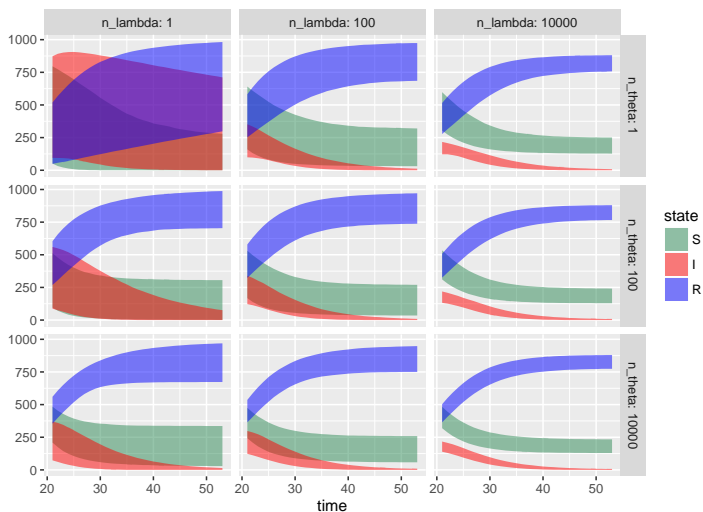
$$\int \int \int p(X_{t+1:T}, \lambda, \theta | X_t) p(X_t, \lambda, \theta | y_{0:t}) d\lambda d\theta dX_t$$

we need to assign priors to  $\lambda$ ,  $\theta$ , and  $X_0$ . Suppose, we assume

$$\begin{aligned}\theta_k &\overset{ind}{\sim} Be(n_\theta p_k, n_\theta [1 - p_k]) \\ \lambda_k &\overset{ind}{\sim} Ga(n_\lambda c_k, n_\lambda) \\ X_0 &\sim Mult(N; z_1, \dots, z_S)\end{aligned}$$

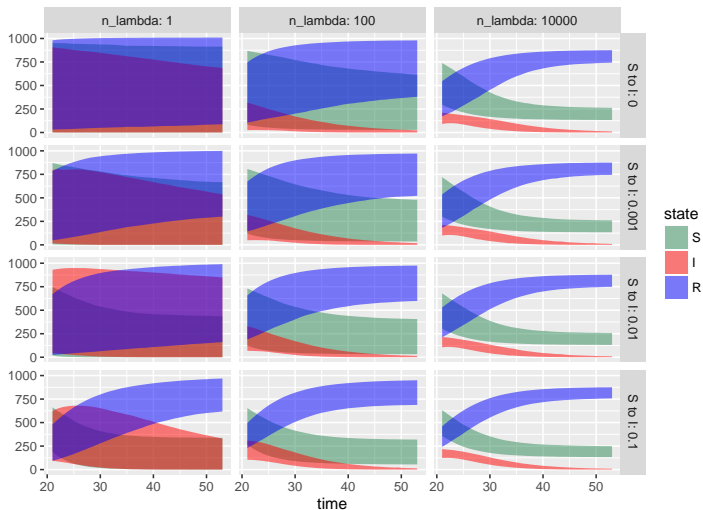
We can control how informative the priors are with  $n_\theta$  and  $n_\lambda$ .

# Informative priors





# Balance priors and data



# We need information

Information can come from

- Data
- Priors

We can quantitatively assess the impact of better information, i.e.

- increasing prior information, e.g.  $\lambda \sim Ga(n_\lambda c_k, n_\lambda)$  by increasing  $n_\lambda$ ,
- increasing surveillance, e.g.  $Y \sim Bin(S \rightarrow I, \theta)$  by increasing  $\theta$ , and
- increasing timeliness, e.g.  $p(X_{t+1:T} | y_{1:t-L})$  by decreasing  $L$ .

Then, we can discuss how to assign resources depending on the costs associated with each impact above.