

# Stochastic dynamic models for low count observations (and forecasting from them)

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April 19, 2016

# Overview

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  - Simulations
- 2 Inference and forecasting
- 3 Forecasting with noisy observations
- 4 Forecasting with inference on parameters

# Poisson-binomial state-space model

Let  $X_{tk}$  be the count of the number of individuals in state  $k$  and time  $t$ . We model state transitions using Poisson distributions, i.e.

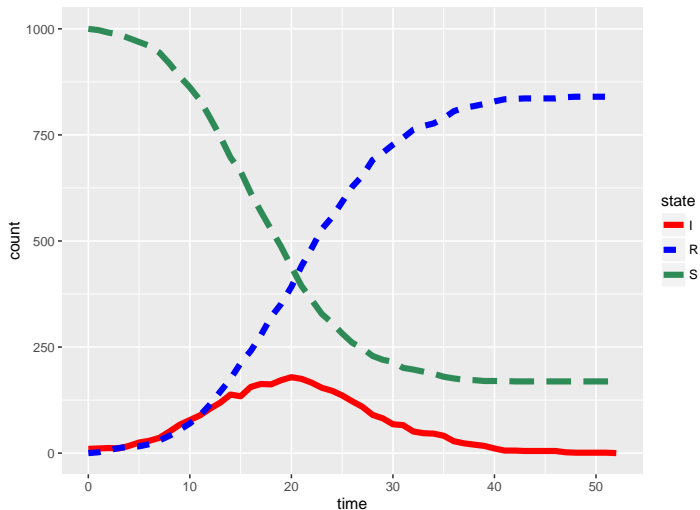
$$\begin{aligned}\Delta X_{tk} &\overset{\text{ind}}{\sim} \text{Po}(\lambda_k f_k(X_{t-1})), \\ X_{tm} &= X_{t-1,m} + \sum_{k=1}^K v_{mk} \Delta X_{tk}, \quad m = 1, \dots, M\end{aligned}$$

where  $f_k(\cdot)$  are known functions and  $v_{mk}$  is a known *stoichiometry*.

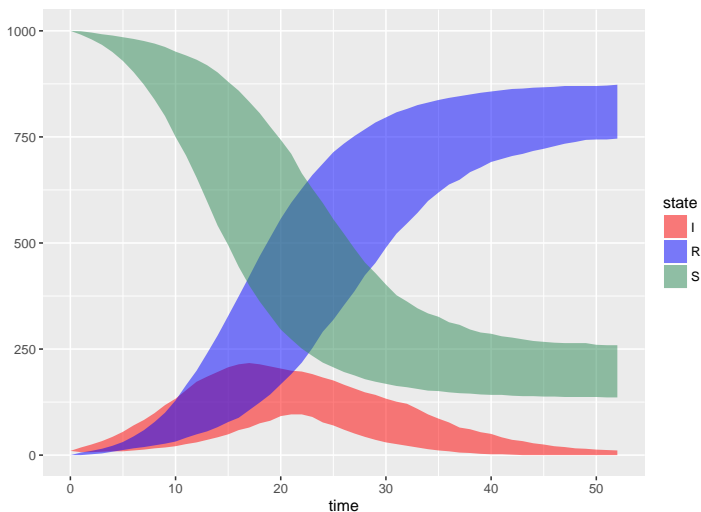
We assume the transitions  $\Delta X_{tk}$  are partially observed through a binomial distribution, i.e.

$$Y_{tk} \overset{\text{ind}}{\sim} \text{Bin}(\Delta X_{tk}, \theta_k).$$

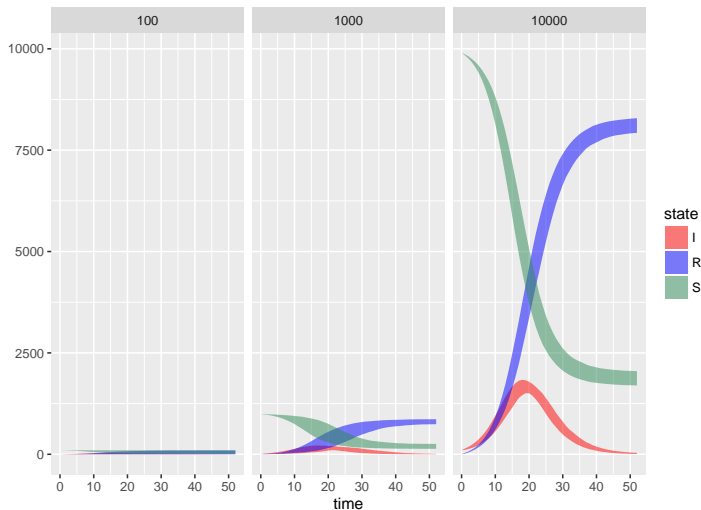
# SIR modeling simulations



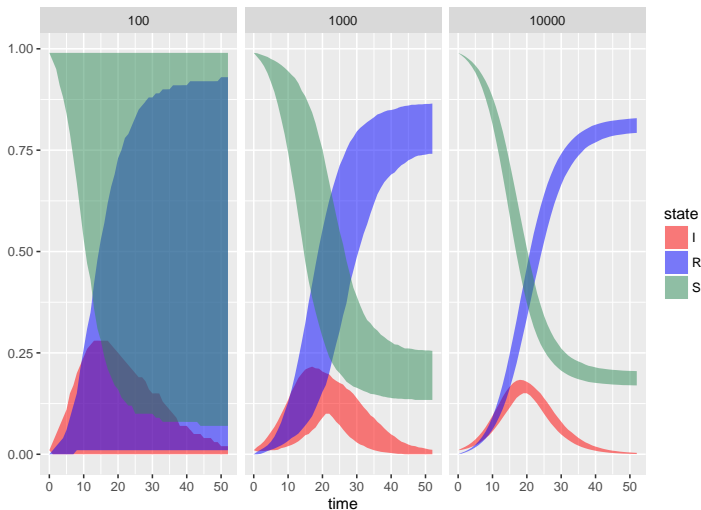
# SIR modeling simulations



# Variability as a function of population size



# Forecasts for proportion of population



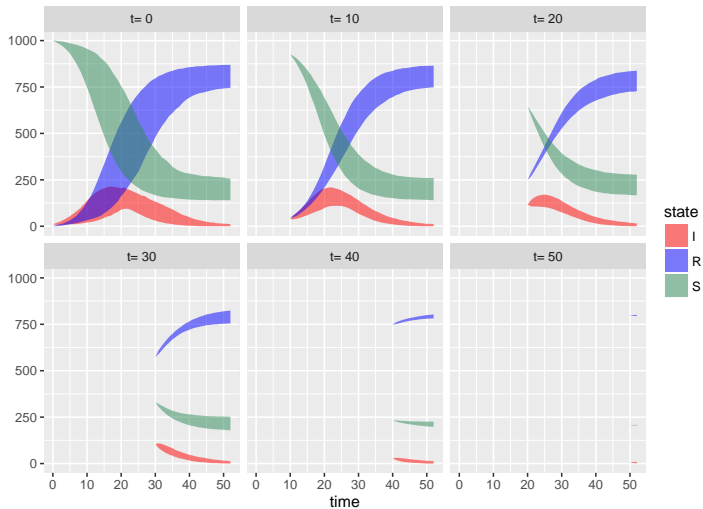
# Forecasting with perfect information

Suppose you know transition rates  $\lambda$ , observation probabilities  $\theta$ , and the states  $X_{0:t}$  and your only goal is to forecast the future states  $X_{t+1:T}$ , i.e.

$$p(X_{t+1:T}|\theta, \lambda, X_{0:t}) = p(X_{t+1:T}|\theta, \lambda, X_t)$$

this distribution is estimated via Monte Carlo simulation.





# Delay in data analysis

Suppose you have a one or two week delay in collecting, processing, and analyzing so that when trying to forecast future states you are using “old” data, i.e.

Real-time:

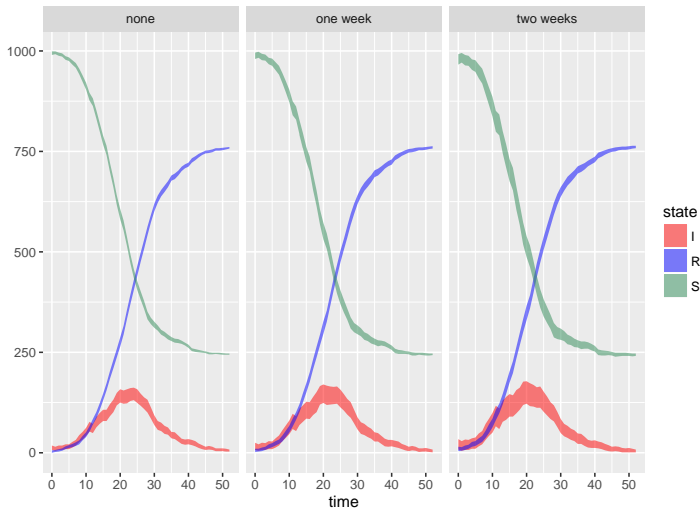
$$p(X_{t+1:T}|\theta, \lambda, X_t)$$

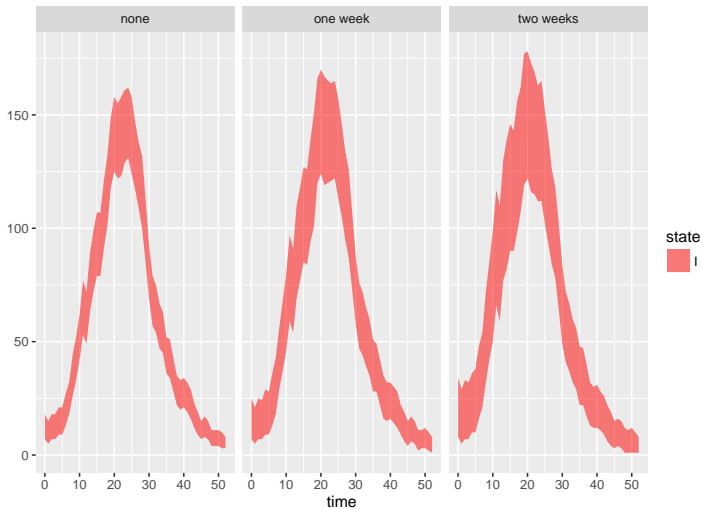
One-week delay:

$$p(X_{t+1:T}|\theta, \lambda, X_{t-1})$$

Two-week delay:

$$p(X_{t+1:T}|\theta, \lambda, X_{t-2})$$





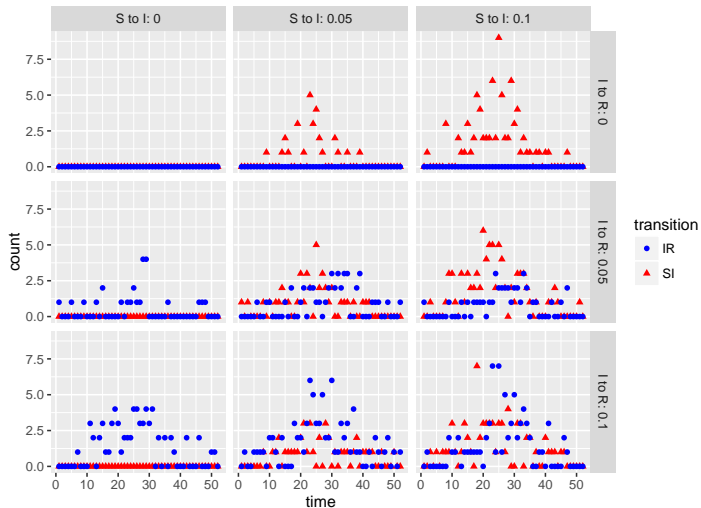
# Forecasting with noisy observations

Suppose, we know the transition rates ( $\lambda$ ) and the observation probabilities ( $\theta$ ), but we only observe a noisy version of the state transitions, i.e. .

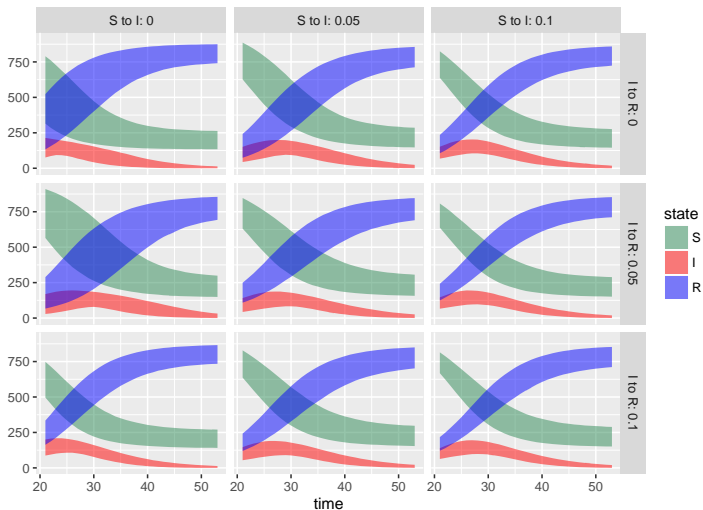
$$Y_{tk} \stackrel{ind}{\sim} \text{Bin}(\Delta X_{tk}, \theta_k).$$

Now the forecast distribution we need is

$$p(X_{t+1:T} | \lambda, \theta, y_{0:t}) = \int p(X_{t+1:T}, \lambda, \theta | X_t) p(X_t | \lambda, \theta, y_{0:t}) dX_t.$$



# Noisily observed state



# Forecasting with inference on parameters

In reality, we don't know the transition rates ( $\lambda$ ) and the observation probabilities ( $\theta$ ), and we only observe a noisy version of the state transitions.

Now the forecast distribution we need is

$$p(X_{t+1:T}|y_{0:t}) = \int \int \int p(X_{t+1:T}, \lambda, \theta | X_t) p(X_t, \lambda, \theta | y_{0:t}) d\lambda d\theta dX_t.$$



# Prior distributions

In order to calculate (or approximate) the integral

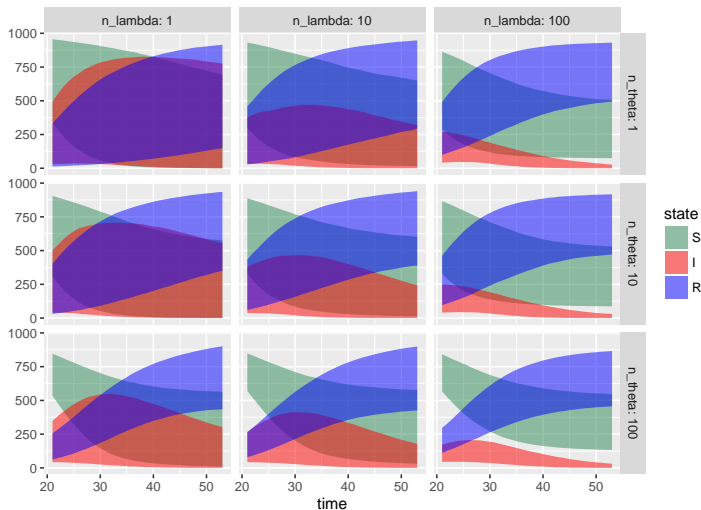
$$\int \int \int p(X_{t+1:T}, \lambda, \theta | X_t) p(X_t, \lambda, \theta | y_{0:t}) d\lambda d\theta dX_t$$

we need to assign priors to  $\lambda$ ,  $\theta$ , and  $X_0$ . Suppose, we assume

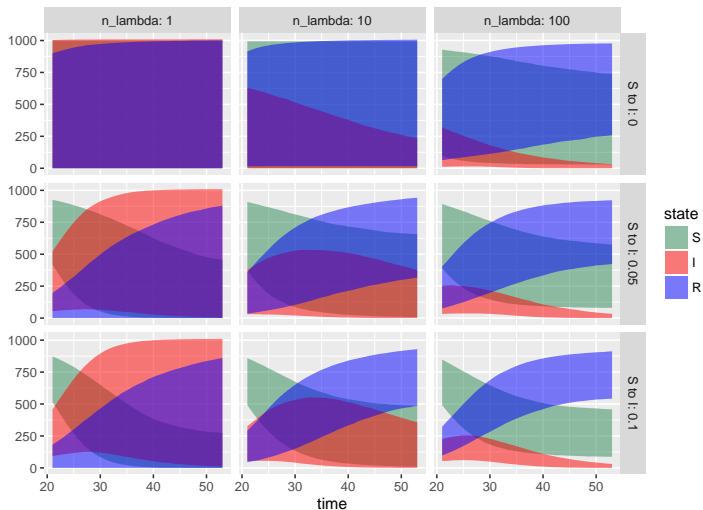
$$\begin{aligned}\theta_k &\overset{ind}{\sim} Be(n_\theta p_k, n_\theta [1 - p_k]) \\ \lambda_k &\overset{ind}{\sim} Ga(n_\lambda c_k, n_\lambda) \\ X_0 &\sim Mult(N; z_1, \dots, z_S)\end{aligned}$$

We can control how informative the priors are with  $n_\theta$  and  $n_\lambda$ .

# Informative priors



# Balance priors and data



# We need information

Information can come from

- Data
- Priors

We can quantitatively assess the impact of better information, i.e.

- Increasing prior information, e.g.  $\lambda \sim Ga(n_\lambda c_k, n_\lambda)$  by increasing  $n_\lambda$ .
- Increasing surveillance, e.g.  $Y_t \sim Bin(S \rightarrow I, \theta)$  by increasing  $\theta$ .
- Increasing timeliness, e.g.  $p(X_{t+1:T} | y_{1:t-L})$  by decreasing  $L$ .

Then, we can discuss how to assign resources depending on the costs associated with each impact above.