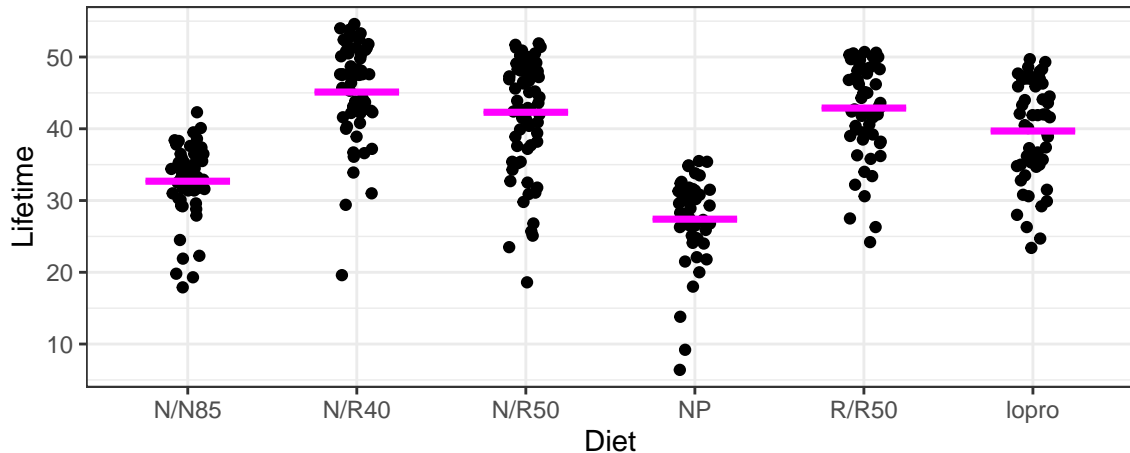


R07 - Contrasts

STAT 5870 (Engineering)
Iowa State University

November 22, 2024

Diet Effect on Mice Lifetimes



ANOVA and Regression Models

ANOVA model:

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$

with Y_{ij} being the lifetime for the i th mouse on the j th diet for $j = 0, 1, 2, 3, 4, 5$.

Regression model:

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \dots + \beta_p X_{i,p}, \sigma^2)$$

where Y_i is the lifetime for the i th mouse and $X_{i,j}$ is an indicator for the i th mouse being on the j th diet.

Reparameterized model since

$$\mu_0 = \beta_0 \quad \text{and} \quad \mu_j = \beta_0 + \beta_j$$

for $j > 0$.

Scientific questions

Here are a few example scientific questions:

1. What is the effect of pre-wean calorie restriction on mean lifetimes?
2. What is the difference in mean lifetimes between mice on a 40 kcal diet compared to those on a 50 kcal diet?
3. What is the effect of high calorie vs low calorie diets on mean lifetimes?

We can compute **contrasts**:

$$\gamma_1 = \mu_{R/R50} - \mu_{N/R50}$$

$$\gamma_2 = \mu_{N/R40} - \frac{1}{2}(\mu_{N/R50} + \mu_{R/R50})$$

$$\gamma_3 = \frac{1}{4}(\mu_{N/R50} + \mu_{R/R50} + \mu_{N/R40} + \mu_{lopro}) - \frac{1}{2}(\mu_{NP} + \mu_{N/N85})$$

Contrasts

A **linear combination** of group means has the form

$$\gamma = C_1\mu_1 + C_2\mu_2 + \dots + C_J\mu_J$$

where C_j are known coefficients and μ_j are the unknown population means.

A linear combination with $C_1 + C_2 + \dots + C_J = 0$ is a **contrast**.

Contrast interpretation is usually best if $|C_1| + |C_2| + \dots + |C_J| = 2$, i.e. the positive coefficients sum to 1 and the negative coefficients sum to -1.

Inference on Contrasts

Contrast

$$\gamma = C_1\mu_1 + C_2\mu_2 + \cdots + C_J\mu_J \quad \text{with} \quad \hat{\gamma} = C_1\bar{Y}_1 + C_2\bar{Y}_2 + \cdots + C_J\bar{Y}_J$$

with standard error

$$SE(\hat{\gamma}) = \hat{\sigma} \sqrt{\frac{C_1^2}{n_1} + \frac{C_2^2}{n_2} + \cdots + \frac{C_J^2}{n_J}}.$$

p -values for $H_0 : \gamma = g_0$ vs $H_A : \gamma \neq g_0$ and posterior probabilities (i.e. $2P(\gamma > 0|y)$ or $2P(\gamma < 0|y)$):

$$t = \frac{g - g_0}{SE(g)}, \quad p = 2P(T_{n-J} < -|t|).$$

Two-sided equal-tail $100(1 - \alpha)\%$ confidence/credible intervals:

$$g \pm t_{n-J, 1-\alpha/2} SE(g).$$

Contrasts for mice lifetime dataset

For these contrasts:

1. Difference in mean lifetimes for N/R50 v R/R50 diet
2. Difference in mean lifetimes for N/R40 v N/R50 and R/R50 combined
3. Difference in mean lifetimes for high calorie (NP and N/N85) diets v low calorie (others) diets

$$H_0 : \gamma = 0 \quad H_A : \gamma \neq 0 :$$

$$\gamma_1 = \mu_{R/R50} - \mu_{N/R50}$$

$$\gamma_2 = \mu_{N/R40} - \frac{1}{2}(\mu_{N/R50} + \mu_{R/R50})$$

$$\gamma_3 = \frac{1}{4}(\mu_{N/R50} + \mu_{R/R50} + \mu_{N/R40} + \mu_{lopro}) - \frac{1}{2}(\mu_{NP} + \mu_{N/N85})$$

	N/N85	N/R40	N/R50	NP	R/R50	lopro
early rest - none @ 50kcal	0.00	0.00	-1.00	0.00	1.00	0.00
40kcal/week - 50kcal/week	0.00	1.00	-0.50	0.00	-0.50	0.00
lo cal - hi cal	-0.50	0.25	0.25	-0.50	0.25	0.25

Fit the Multiple Regression Model

```
m <- lm(Lifetime ~ Diet, data = Sleuth3::case0501)
summary(m)
```

Call:
lm(formula = Lifetime ~ Diet, data = Sleuth3::case0501)

Residuals:

	Min	1Q	Median	3Q	Max
	-25.5167	-3.3857	0.8143	5.1833	10.0143

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	32.6912	0.8846	36.958	< 2e-16 ***
DietN/R40	12.4254	1.2352	10.059	< 2e-16 ***
DietN/R50	9.6060	1.1877	8.088	1.06e-14 ***
DietNP	-5.2892	1.3010	-4.065	5.95e-05 ***
DietR/R50	10.1945	1.2565	8.113	8.88e-15 ***
Dietlopro	6.9945	1.2565	5.567	5.25e-08 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6.678 on 343 degrees of freedom
Multiple R-squared: 0.4543, Adjusted R-squared: 0.4463
F-statistic: 57.1 on 5 and 343 DF, p-value: < 2.2e-16

Estimate Group Means

```
library("emmeans")  
em <- emmeans(m, ~ Diet)  
em
```

Diet	emmean	SE	df	lower.CL	upper.CL
N/N85	32.7	0.885	343	31.0	34.4
N/R40	45.1	0.862	343	43.4	46.8
N/R50	42.3	0.793	343	40.7	43.9
NP	27.4	0.954	343	25.5	29.3
R/R50	42.9	0.892	343	41.1	44.6
lopro	39.7	0.892	343	37.9	41.4

Confidence level used: 0.95

```
K_list
```

```
$`early rest - none @ 50kcal`
```

```
[1] 0 0 -1 0 1 0
```

```
$`40kcal/week - 50kcal/week`
```

```
[1] 0.0 1.0 -0.5 0.0 -0.5 0.0
```

```
$`lo cal - hi cal`
```

```
[1] -0.50 0.25 0.25 -0.50 0.25 0.25
```

```
co <- contrast(em, K_list)
```

```
# p-values (and posterior tail probabilities)
```

```
co
```

contrast	estimate	SE	df	t.ratio	p.value
early rest - none @ 50kcal	0.589	1.19	343	0.493	0.6223
40kcal/week - 50kcal/week	2.525	1.05	343	2.408	0.0166
lo cal - hi cal	12.450	0.78	343	15.961	<.0001

```
# confidence/credible intervals
```

```
confint(co)
```

contrast	estimate	SE	df	lower.CL	upper.CL
early rest - none @ 50kcal	0.589	1.19	343	-1.759	2.94
40kcal/week - 50kcal/week	2.525	1.05	343	0.463	4.59
lo cal - hi cal	12.450	0.78	343	10.915	13.98

```
Confidence level used: 0.95
```

Summary

- Contrasts are linear combinations of means where the coefficients sum to zero
- t-test tools are used to calculate pvalues and confidence intervals

Sulfur effect on scab disease in potatoes

The experiment was conducted to investigate the effect of sulfur on controlling scab disease in potatoes. There were seven treatments: control, plus spring and fall application of 300, 600, 1200 lbs/acre of sulfur. The response variable was percentage of the potato surface area covered with scab averaged over 100 random selected potatoes. A completely randomized design was used with 8 replications of the control and 4 replications of the other treatments.

Cochran and Cox. (1957) Experimental Design (2nd ed). pg96 and Agron. J. 80:712-718 (1988)

Scientific questions:

- Does sulfur have any impact at all?
- What is the difference between spring and fall application of sulfur?
- What is the effect of increased sulfur application?

Data

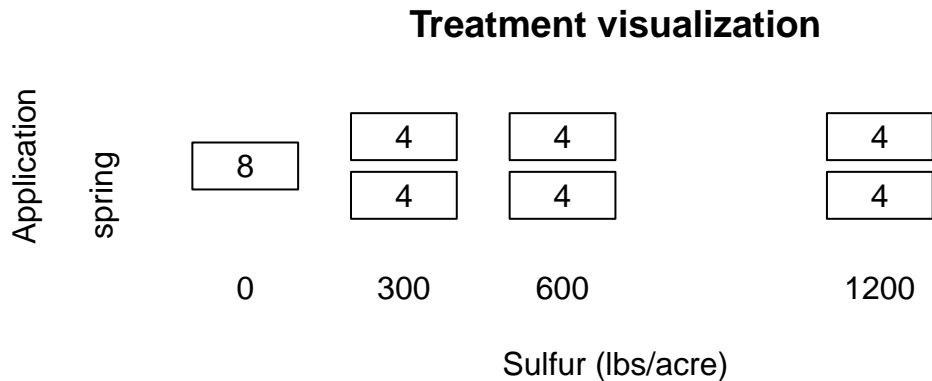
	inf	trt	row	col	sulfur	application	treatment
1	9	F3	4	1	300	fall	F3
2	12	0	4	2	0	not applicable	0
3	18	S6	4	3	600	spring	S6
4	10	F12	4	4	1200	fall	F12
5	24	S6	4	5	600	spring	S6
6	17	S12	4	6	1200	spring	S12
7	30	S3	4	7	300	spring	S3
8	16	F6	4	8	600	fall	F6
9	10	0	3	1	0	not applicable	0
10	7	S3	3	2	300	spring	S3
11	4	F12	3	3	1200	fall	F12
12	10	F6	3	4	600	fall	F6
13	21	S3	3	5	300	spring	S3
14	24	0	3	6	0	not applicable	0
15	29	0	3	7	0	not applicable	0
16	12	S6	3	8	600	spring	S6
17	9	F3	2	1	300	fall	F3
18	7	S12	2	2	1200	spring	S12
19	18	F6	2	3	600	fall	F6
20	30	0	2	4	0	not applicable	0
21	18	F6	2	5	600	fall	F6
22	16	S12	2	6	1200	spring	S12
23	16	F3	2	7	300	fall	F3
24	4	F12	2	8	1200	fall	F12
25	9	S3	1	1	300	spring	S3
26	18	0	1	2	0	not applicable	0
27	17	S12	1	3	1200	spring	S12

Design

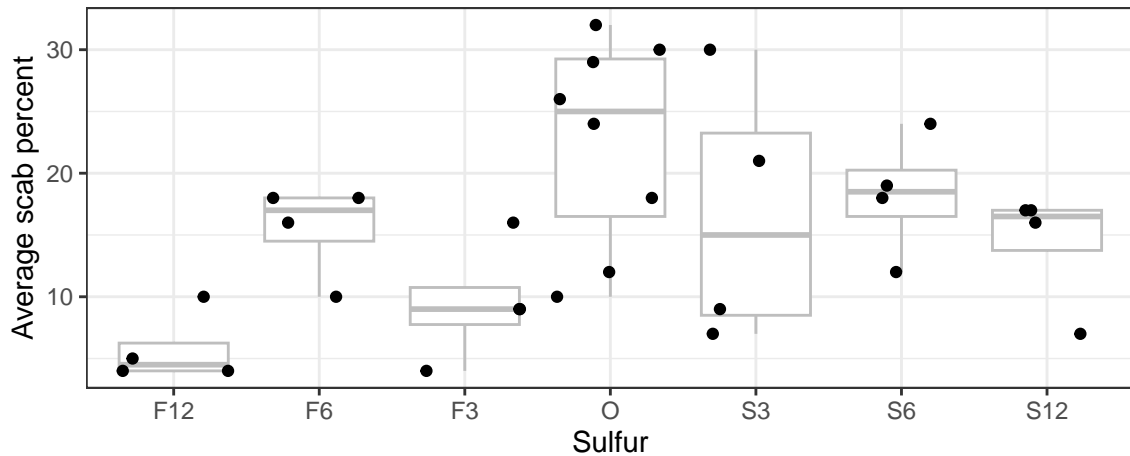
Completely randomized design potato scab experiment

row	3 1	F3	O	S6	F12	S6	S12	S3	F6
		O	S3	F12	F6	S3	O	O	S6
		F3	S12	F6	O	F6	S12	F3	F12
		S3	O	S12	S6	O	F12	O	F3
		1	2	3	4	5	6	7	8
		col							

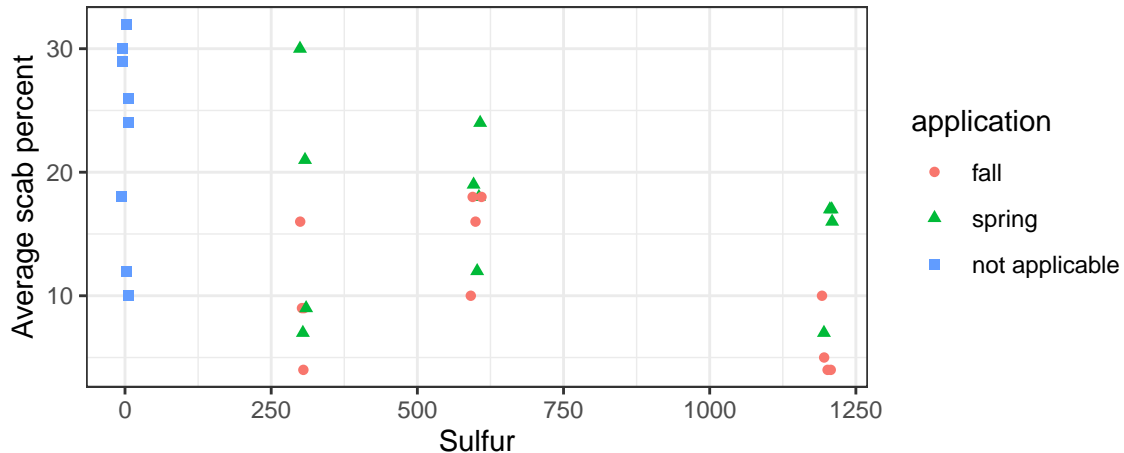
Design



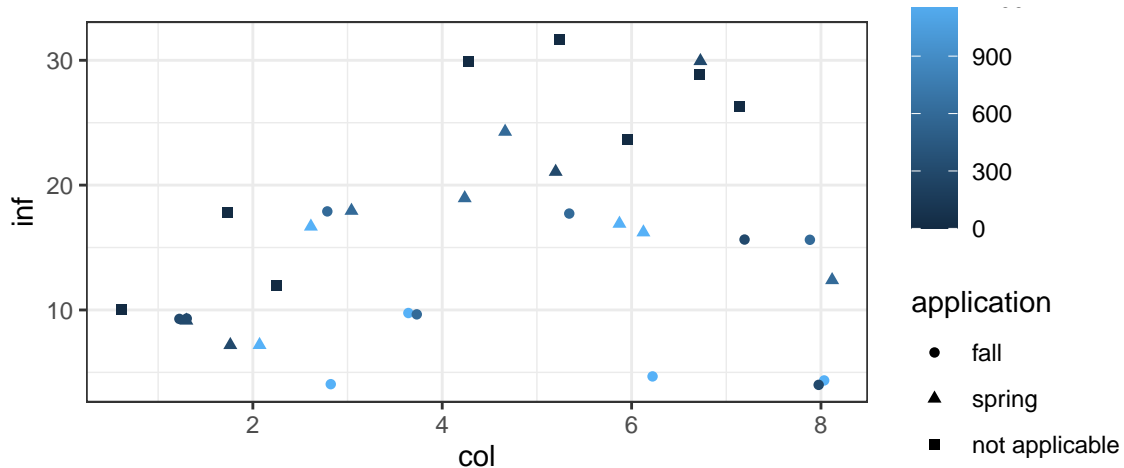
Data



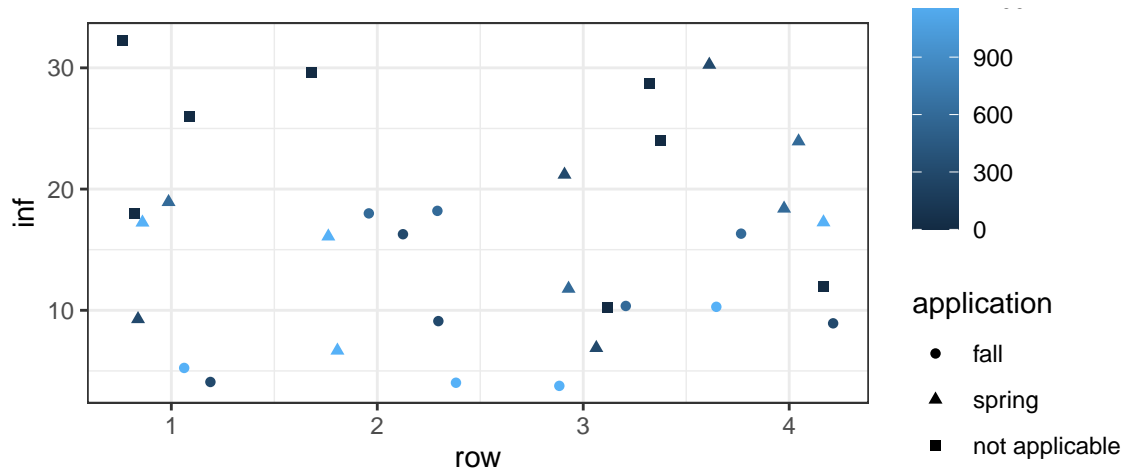
Data



Data



Data



Model

Y_{ij} : avg % of surface area covered with scab for plot i in treatment j for $j = 1, \dots, 7$.

Assume $Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$.

Hypotheses:

- Difference amongst any means:
One-way ANOVA F-test
- *Any effect*:
Contrast: control vs sulfur
- *Fall vs spring*:
Contrast: fall vs spring applications
- *Sulfur level*:
Contrast: linear trend

Contrasts

- *Sulfur effect*: Any sulfur vs none

$$\begin{aligned}\gamma &= \frac{1}{6}(\mu_{F12} + \mu_{F6} + \mu_{F3} + \mu_{S3} + \mu_{S6} + \mu_{S12}) - \mu_O \\ &= \frac{1}{6}(\mu_{F12} + \mu_{F6} + \mu_{F3} + \mu_{S3} + \mu_{S6} + \mu_{S12} - 6\mu_O)\end{aligned}$$

- *Fall vs spring*: Contrast comparing fall vs spring applications

$$\begin{aligned}\gamma &= \frac{1}{3}(\mu_{F12} + \mu_{F6} + \mu_{F3}) + 0\mu_O - \frac{1}{3}(\mu_{S3} + \mu_{S6} + \mu_{S12}) \\ &= \frac{1}{3}[1\mu_{F12} + 1\mu_{F6} + 1\mu_{F3} + 0\mu_O - 1\mu_{S3} - 1\mu_{S6} - 1\mu_{S12}]\end{aligned}$$

Contrasts (cont.)

- Sulfur linear trend

- The group sulfur levels (X_j) are 12, 6, 3, 0, 3, 6, and 12 (100 lbs/acre)
- and a linear trend contrast is $X_j - \bar{X}$

X_i	12	6	3	0	3	6	12
$X_i - \bar{X}$	6	0	-3	-6	-3	0	6

$$\gamma = 6\mu_{F12} + 0\mu_{F6} - 3\mu_{F3} - 6\mu_O - 3\mu_{S3} + 0\mu_{S6} + 6\mu_{S12}$$

Trt	F12	F6	F3	O	S3	S6	S12	Div
Sulfur v control	1	1	1	-6	1	1	1	6
Fall v Spring	1	1	1	0	-1	-1	-1	3
Linear Trend	6	0	-3	-6	-3	0	6	1

```
K <-
#
#               F12 F6 F3  O S3 S6 S12
list("sulfur - control" = c( 1, 1, 1,-6, 1, 1,  1)/6,
     "fall - spring"    = c( 1, 1, 1, 0,-1,-1, -1)/3,
     "linear trend"     = c( 6, 0,-3,-6,-3, 0,  6)/1)

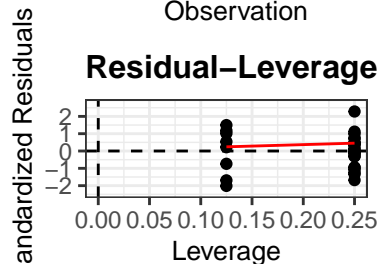
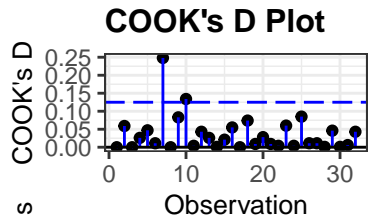
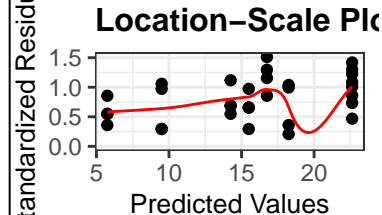
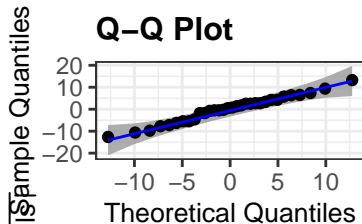
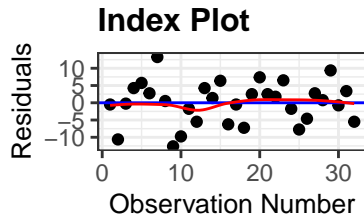
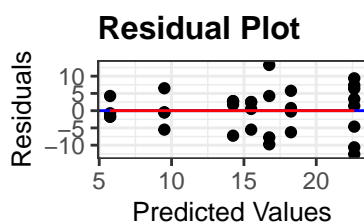
m <- lm(inf ~ treatment, data = d)
anova(m)
```

Analysis of Variance Table

Response: inf

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
treatment	6	972.34	162.057	3.6081	0.01026 *
Residuals	25	1122.88	44.915		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1




```
em <- emmeans(m, ~treatment); em
```

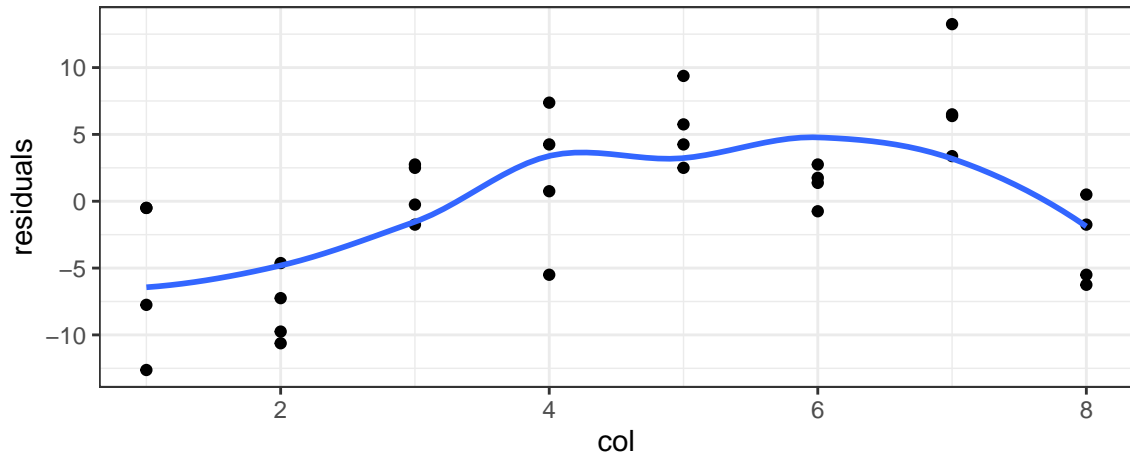
treatment	emmean	SE	df	lower.CL	upper.CL
F12	5.75	3.35	25	-1.15	12.7
F6	15.50	3.35	25	8.60	22.4
F3	9.50	3.35	25	2.60	16.4
0	22.62	2.37	25	17.74	27.5
S3	16.75	3.35	25	9.85	23.7
S6	18.25	3.35	25	11.35	25.2
S12	14.25	3.35	25	7.35	21.2

Confidence level used: 0.95

```
co <- contrast(em, K)
confint(co)
```

contrast	estimate	SE	df	lower.CL	upper.CL
sulfur - control	-9.29	2.74	25	-14.9	-3.657
fall - spring	-6.17	2.74	25	-11.8	-0.532
linear trend	-94.50	34.82	25	-166.2	-22.779

Confidence level used: 0.95



Summary

For this particular data analysis

- Significant differences in means between the groups (ANOVA $F_{6,25} = 3.61$ $p=0.01$)
- Having sulfur was associated with a reduction in scab % of 9 (4,15) compared to no sulfur
- Fall application reduced scab % by 6 (0.5,12) compared to spring application
- Linear trend in sulfur was significant ($p=0.01$)

- Concerned about spatial correlation among columns
- Consider a logarithm of the response
 - CI for F12 (-1.2, 12.7)
 - Non-constant variance (residuals vs predicted, sulfur, application)