

## P4 - Central Limit Theorem

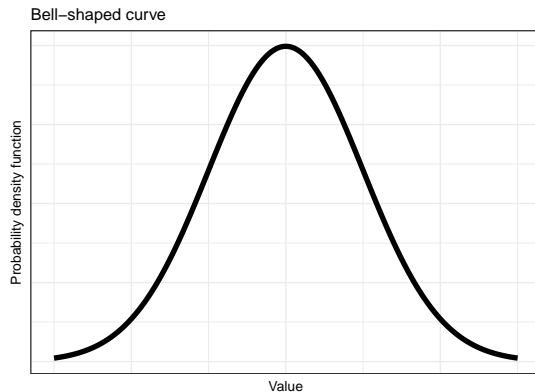
STAT 587 (Engineering)  
Iowa State University

March 30, 2021

**Main Idea:** Sums and averages of iid random variables from **any distribution** have approximate normal distributions for sufficiently large sample sizes.

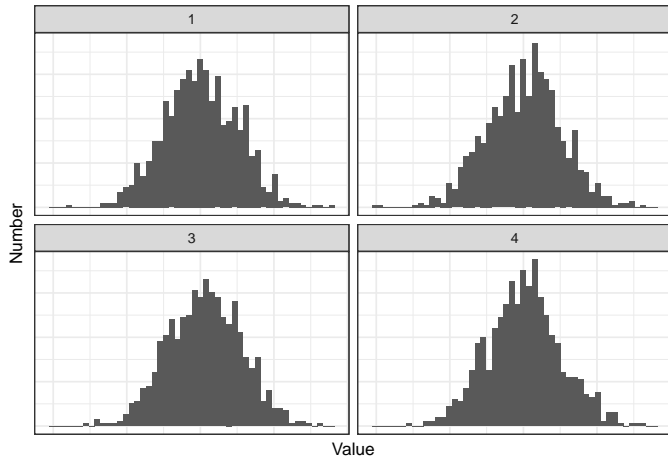
# Bell-shaped curve

The term **bell-shaped curve** typically refers to the probability density function for a normal random variable:



# Histograms of samples from bell-shaped curves

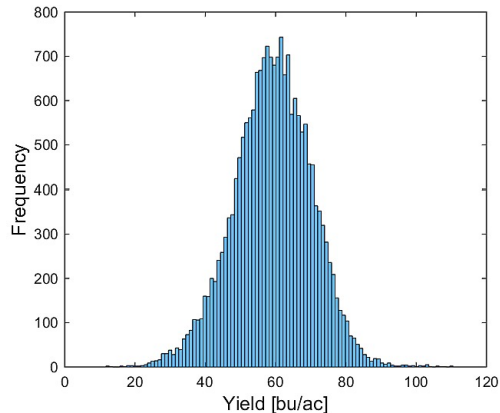
Histograms of 1,000 standard normal random variables



# Yield

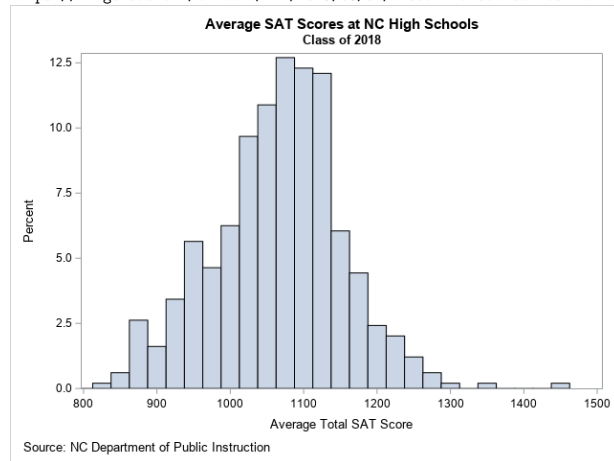
<https://journals.plos.org/plosone/article?id=10.1371/journal.pone.0184198>

**S1 Fig. Histogram of yield.**



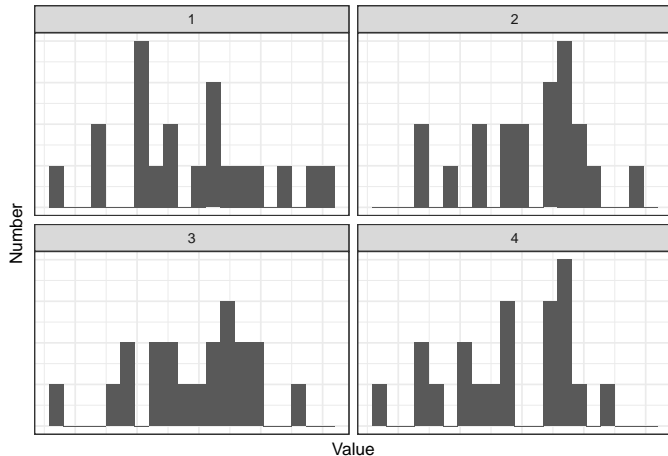
# SAT scores

<https://blogs.sas.com/content/iml/2019/03/04/visualize-sat-scores-nc.html>



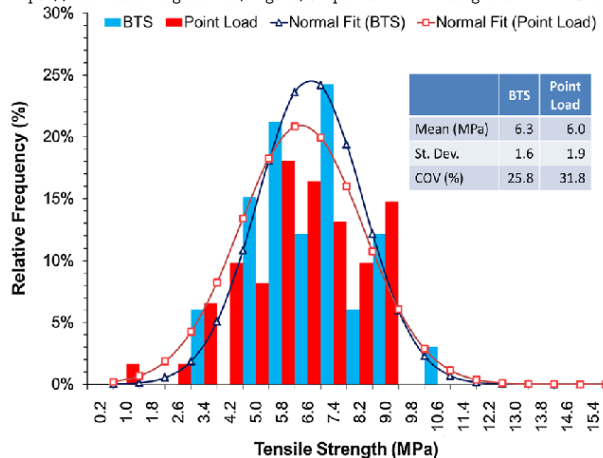
# Histograms of samples from bell-shaped curves

Histograms of 20 standard normal random variables



# Tensile strength

[https://www.researchgate.net/figure/Comparison-of-histograms-for-BTS-and-tensile-strength-estimated-from-point-load\\_fig5\\_260617256](https://www.researchgate.net/figure/Comparison-of-histograms-for-BTS-and-tensile-strength-estimated-from-point-load_fig5_260617256)



# Sums and averages of iid random variables

Suppose  $X_1, X_2, \dots$  are iid random variables with

$$E[X_i] = \mu \quad \text{Var}[X_i] = \sigma^2.$$

Define

$$\begin{aligned} \text{Sample Sum: } S_n &= X_1 + X_2 + \dots + X_n \\ \text{Sample Average: } \bar{X}_n &= S_n/n. \end{aligned}$$

For  $S_n$ , we know

$$E[S_n] = n\mu, \quad \text{Var}[S_n] = n\sigma^2, \quad \text{and} \quad SD[S_n] = \sqrt{n}\sigma.$$

For  $\bar{X}_n$ , we know

$$E[\bar{X}_n] = \mu, \quad \text{Var}[\bar{X}_n] = \sigma^2/n, \quad \text{and} \quad SD[\bar{X}_n] = \sigma/\sqrt{n}.$$



# Central Limit Theorem (CLT)

Suppose  $X_1, X_2, \dots$  are iid random variables with

$$E[X_i] = \mu \quad \text{Var}[X_i] = \sigma^2.$$

Define

$$\begin{aligned} \text{Sample Sum: } S_n &= X_1 + X_2 + \dots + X_n \\ \text{Sample Average: } \bar{X}_n &= S_n/n. \end{aligned}$$

Then the **Central Limit Theorem** says

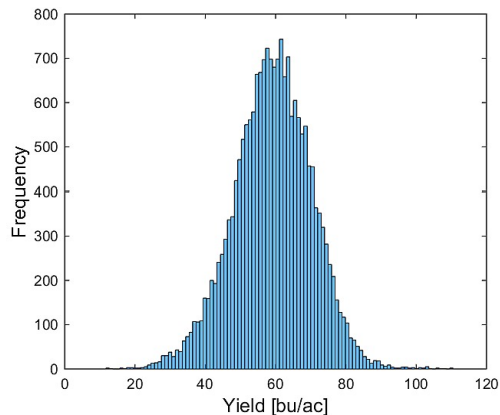
$$\lim_{n \rightarrow \infty} \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{d} N(0, 1) \quad \text{and} \quad \lim_{n \rightarrow \infty} \frac{S_n - n\mu}{\sqrt{n}\sigma} \xrightarrow{d} N(0, 1).$$

**Main Idea:** Sums and averages of iid random variables from **any distribution** have approximate normal distributions for sufficiently large sample sizes.

# Yield

<https://journals.plos.org/plosone/article?id=10.1371/journal.pone.0184198>

**S1 Fig. Histogram of yield.**



# Approximating distributions

Rather than considering the limit, I typically think of the following approximations as  $n$  gets large.

For the sample average,

$$\overline{X}_n \dot{\sim} N(\mu, \sigma^2/n).$$

where  $\dot{\sim}$  indicates *approximately distributed* because

$$E[\overline{X}_n] = \mu \quad \text{and} \quad \text{Var}[\overline{X}_n] = \sigma^2/n.$$

For the sample sum,

$$S_n \dot{\sim} N(n\mu, n\sigma^2)$$

because

$$\begin{aligned} E[S_n] &= n\mu \\ \text{Var}[S_n] &= n\sigma^2. \end{aligned}$$

# Averages and sums of uniforms

Let  $X_i \stackrel{ind}{\sim} Unif(0, 1)$ . Then

$$\mu = E[X_i] = \frac{1}{2} \quad \text{and} \quad \sigma^2 = Var[X_i] = \frac{1}{12}.$$

Thus

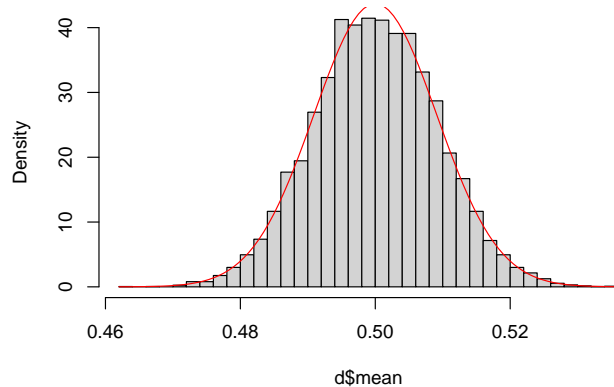
$$\overline{X}_n \dot{\sim} N\left(\frac{1}{2}, \frac{1}{12n}\right)$$

and

$$S_n \dot{\sim} N\left(\frac{n}{2}, \frac{n}{12}\right).$$

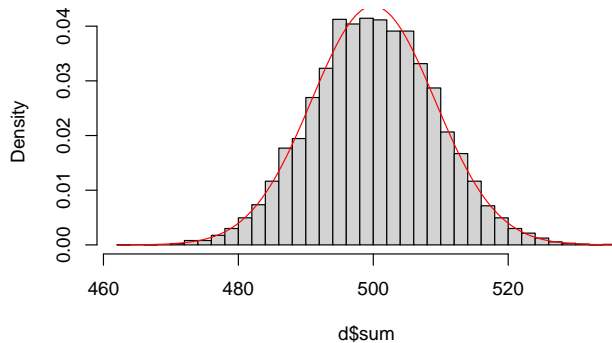
# Averages of uniforms

Histogram of d\$mean



# Sums of uniforms

Histogram of d\$sum



# Normal approximation to a binomial

Recall if  $Y_n = \sum_{i=1}^n X_i$  where  $X_i \stackrel{ind}{\sim} Ber(p)$ , then

$$Y_n \sim Bin(n, p).$$

For a binomial random variable, we have

$$E[Y_n] = np \quad \text{and} \quad Var[Y_n] = np(1 - p).$$

By the CLT,

$$\lim_{n \rightarrow \infty} \frac{Y_n - np}{\sqrt{np(1 - p)}} \rightarrow N(0, 1),$$

if  $n$  is large,

$$Y_n \dot{\sim} N(np, np[1 - p]).$$

# Roulette example

A European roulette wheel has 39 slots: one green, 19 black, and 19 red. If I play black every time, what is the probability that I will have won more than I lost after 99 spins of the wheel?

<https://isorepublic.com/photo/roulette-wheel/>





## Roulette example

A European roulette wheel has 39 slots: one green, 19 black, and 19 red. If I play black every time, what is the probability that I will have won more than I lost after 99 spins of the wheel?

Let  $Y$  indicate the total number of wins and assume  $Y \sim \text{Bin}(n, p)$  with  $n = 99$  and  $p = 19/39$ . The desired probability is  $P(Y \geq 50)$ . Then

$$P(Y \geq 50) = 1 - P(Y < 50) = 1 - P(Y \leq 49)$$

```
n = 99
p = 19/39
1-pbinom(49, n, p)

[1] 0.399048
```

## Roulette example

A European roulette wheel has 39 slots: one green, 19 black, and 19 red. If I play black every time, what is the probability that I will have won more than I lost after 99 spins of the wheel?

Let  $Y$  indicate the total number of wins. We can approximate  $Y$  using  $X \sim N(np, np(1-p))$ .

$$P(Y \geq 50) \approx 1 - P(X < 50)$$

```
1-pnorm(50, n*p, sqrt(n*p*(1-p)))
```

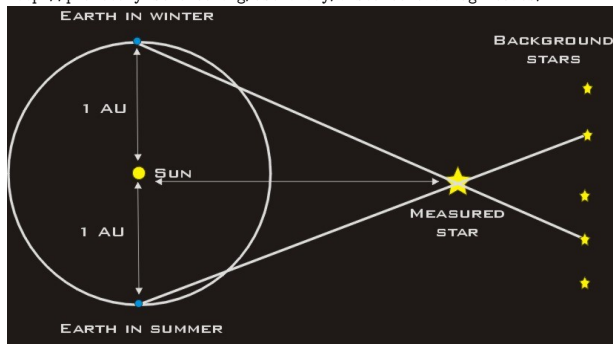
```
[1] 0.3610155
```

A better approximation can be found using a continuity correction.

## Astronomy example

An astronomer wants to measure the distance,  $d$ , from Earth to a star. Suppose the procedure has a known standard deviation of 2 parsecs. The astronomer takes 30 iid measurements and finds the average of these measurements to be 29.4 parsecs. What is the probability the average is within 0.5 parsecs?

<http://planetary-science.org/astrometry/distance-and-magnitudes/>



## Astronomy example

Let  $X_i$  be the  $i^{th}$  measurement. The astronomer assumes that  $X_1, X_2, \dots, X_n$  are iid with  $E[X_i] = d$  and  $Var[X_i] = \sigma^2 = 2^2$ . The estimate of  $d$  is

$$\bar{X}_n = \frac{(X_1 + X_2 + \dots + X_n)}{n} = 29.4.$$

and, by the Central Limit Theorem,  $\bar{X}_n \sim N(d, \sigma^2/n)$  where  $n = 30$ . We want to find

$$\begin{aligned} P(|\bar{X}_n - d| < 0.5) &= P(-0.5 < \bar{X}_n - d < 0.5) \\ &= P\left(\frac{-0.5}{2/\sqrt{30}} < \frac{\bar{X}_n - d}{\sigma/\sqrt{n}} < \frac{0.5}{2/\sqrt{30}}\right) \\ &\approx P(-1.37 < Z < 1.37) \end{aligned}$$

```
diff(pnorm(c(-1.37,1.37)))
```

```
[1] 0.8293131
```

## Astronomy example - sample size

Suppose the astronomer wants to be within 0.5 parsecs with at least 95% probability. How many more samples would she need to take?

We solve

$$\begin{aligned}
 0.95 \leq P(|\bar{X}_n - d| < .5) &= P(-0.5 < \bar{X}_n - d < 0.5) \\
 &= P\left(\frac{-0.5}{2/\sqrt{n}} < \frac{\bar{X}_n - d}{\sigma/\sqrt{n}} < \frac{0.5}{2/\sqrt{n}}\right) \\
 &= P(-z < Z < z) & z = 0.5/(2/\sqrt{n}) \\
 &= 1 - [P(Z < -z) + P(Z > z)] \\
 &= 1 - 2P(Z < -z)
 \end{aligned}$$

where  $z = 1.96$  since

```
1-2*pnorm(-1.96)
```

```
[1] 0.9500042
```

# Summary

- Central Limit Theorem
  - Sums
  - Averages
- Examples
  - Uniforms
  - Binomial
    - Roulette
- Sample size
  - Astronomy