

ASEN 5044, Fall 2018

# Statistical Estimation for Dynamical Systems

Lecture 25 [Special Topic #5]:  
Bayesian Recursive Updating;  
The General Bayes Filter

Prof. Nisar Ahmed ([Nisar.Ahmed@Colorado.edu](mailto:Nisar.Ahmed@Colorado.edu))

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# Overview

- Recursive Bayesian updating and estimation
- General probabilistic forward prediction/propagation of dynamical states
  - **The Chapman-Kolmogorov equation**
- General probabilistic measurement updating of dynamical states
  - **The general Bayes filter**

# Recursive Bayesian Updating and Estimation

- Consider static  $x$  (no dynamics or process noise) with prior pdf  $p(x)$
- Given  $T$  noisy independent measurements  $y_{1:T}$  with known observation likelihoods  $p(y_k|x)$ , for steps  $k=1,\dots,T$ .
- **To get Bayesian estimate of  $x$  at step  $\underline{k}$ :** need to minimize expected value of some desired cost fxn with respect to posterior pdf (see Lecture 20/Special topic #4)
- e.g. using Minimum Mean Squared Error cost function:

$$\hat{x}_{k, \text{MMSE}} = \arg \min_{x \in \mathbb{R}^n} E[(x - \hat{x}_k)^T (x - \hat{x}_k) \mid y_{1:\underline{k}}]$$

$$\rightarrow \hat{x}_{k, \text{MMSE}} = E[x \mid y_{1:\underline{k}}] \rightarrow \text{what is the mean}$$

$$p(x \mid y_1, \dots, y_{\underline{k}}), \underline{k} \leq T$$

→ Need to find  $p(x \mid y_{1:k}) = p(x \mid y_1, \dots, y_k)$

if meas cond.  $\Downarrow$   
of each other  
given  $x$

$$\propto p(x) \cdot p(y_1:k \mid x) \leftrightarrow \text{"batch"}$$

$$\hookrightarrow p(x) \cdot \prod_{i=1}^k p(y_i \mid x) = p(x) \cdot p(y_1 \mid x) \cdot p(y_2 \mid x) \cdots p(y_k \mid x) \leftrightarrow \text{"recursive basis"}$$

# Recursive Bayesian Updating and Estimation

- Alternative: what if data  $y_1, \dots, y_T$  absorbed sequentially, rather than in batch?
- As it turns out, we can use the posterior  $p(x|y_1)$  as the prior for  $p(x|y_{1:2})$ ...  
...and then use the posterior  $p(x|y_{1:2})$  as the prior for  $p(x|y_{1:3})$ , etc.

for  $k=1$  :  $p(x|y_1) \propto p(x) \cdot p(y_1|x) \rightarrow p(x|y_1) = \frac{p(x) \cdot p(y_1|x)}{\int_{-\infty}^{\infty} p(x) \cdot p(y_1|x) dx} \left( = \frac{p(x) \cdot p(y_1|x)}{p(y_1)} \right)$

for  $k=2$  :  $p(x|y_1, y_2) \propto \underbrace{p(x) p(y_1|x) \cdot p(y_2|x)}_{p(x|y_1)} \cdot p(y_2|x) = \frac{p(x) p(y_1|x) \cdot p(y_2|x)}{\int_{-\infty}^{\infty} p(x|y_1) \cdot p(y_2|x) dx} = \frac{p(x) p(y_1|x)}{p(y_1)} \cdot \frac{p(y_2|x)}{p(y_2, y_1)}$

$\rightarrow p(x|y_1, y_2) = \frac{p(x|y_1) \cdot p(y_2|x)}{\int_{-\infty}^{\infty} p(x|y_1) \cdot p(y_2|x) dx} = \frac{\cancel{p(x|y_1)} \cdot p(y_2|x)}{\cancel{p(y_2, y_1)}} \left( = \frac{p(x) \cdot p(y_1|x) \cdot p(y_2|x)}{p(y_2, y_1)} \right)$

For Step  $k=j$  :

Recursive Bayesian updating  $\otimes$

$p(x|y_1, y_2, \dots, y_j) = \frac{p(x|y_1, \dots, y_{j-1}) \cdot p(y_j|x)}{\int_{-\infty}^{\infty} p(x|y_1, \dots, y_{j-1}) \cdot p(y_j|x) dx} \propto p(x|y_{1:j-1}) \cdot p(y_j|x)$

*"has memory" of previous update  $\rightarrow$  posterior encodes all past info*

so  $\boxed{\text{Mathematically equivalent to batch, Bayes rule update!}}$

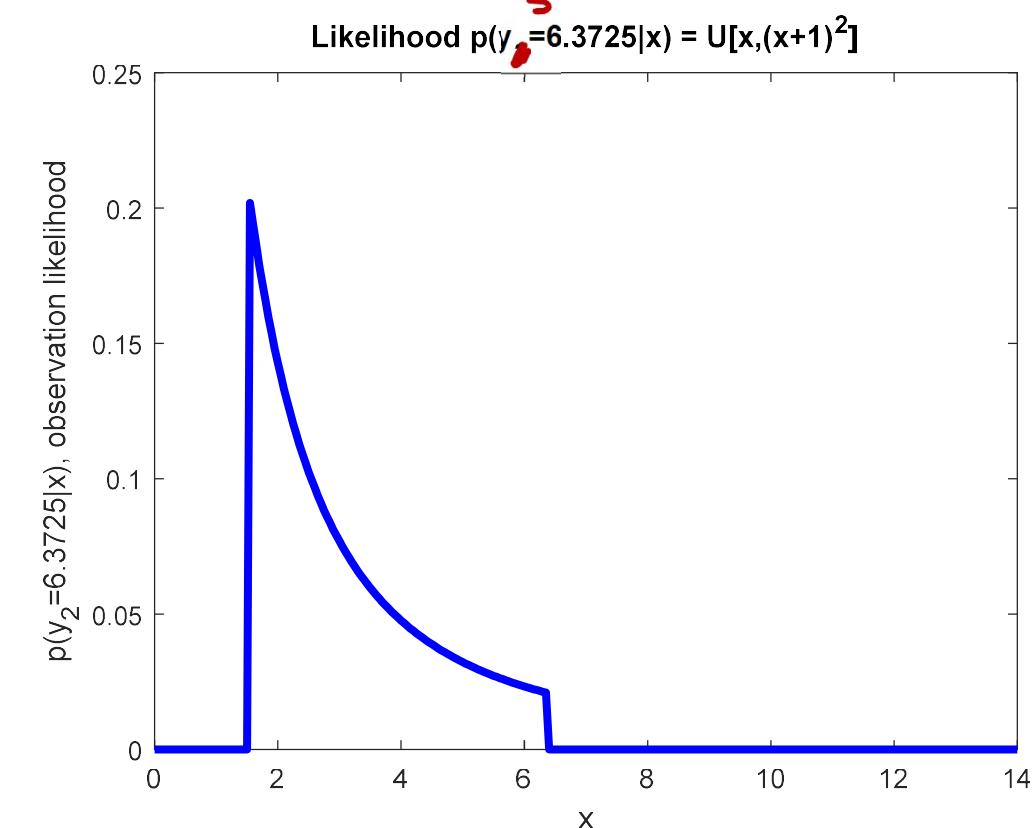
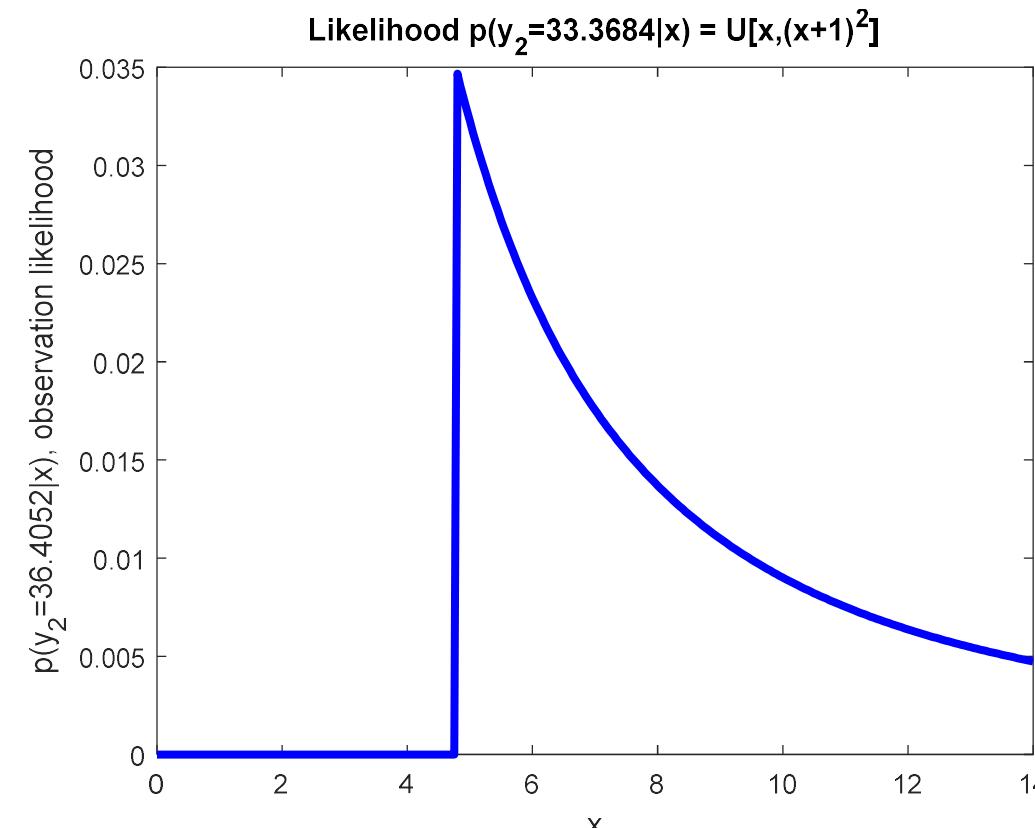
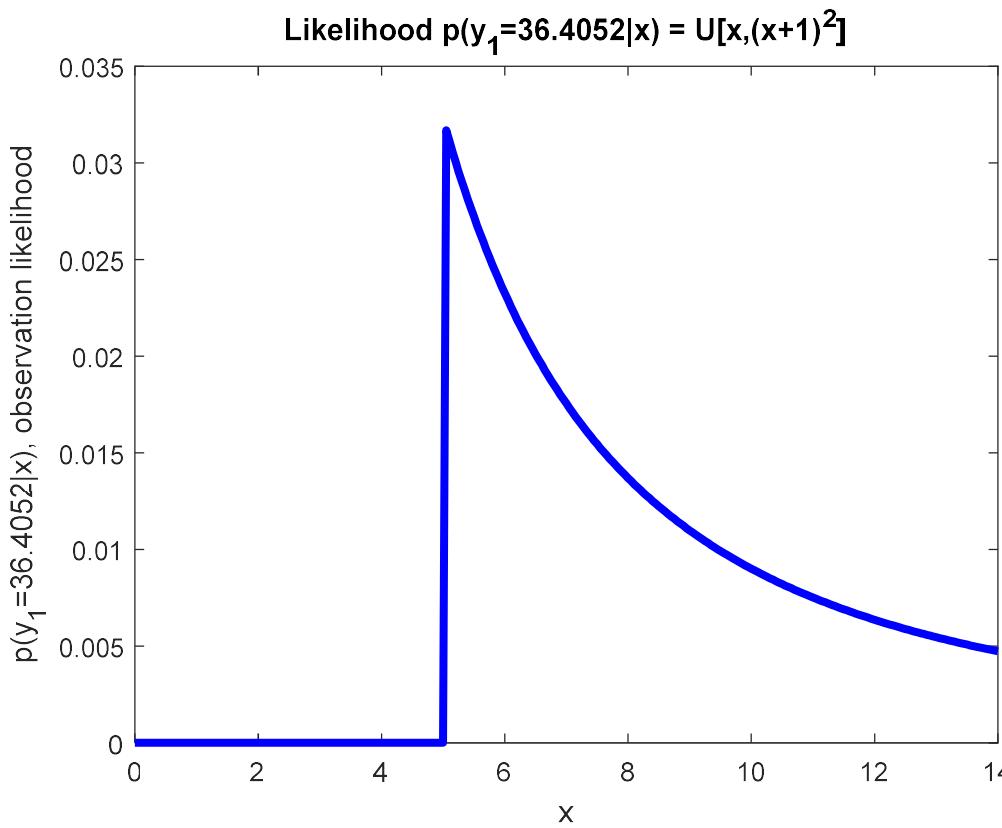
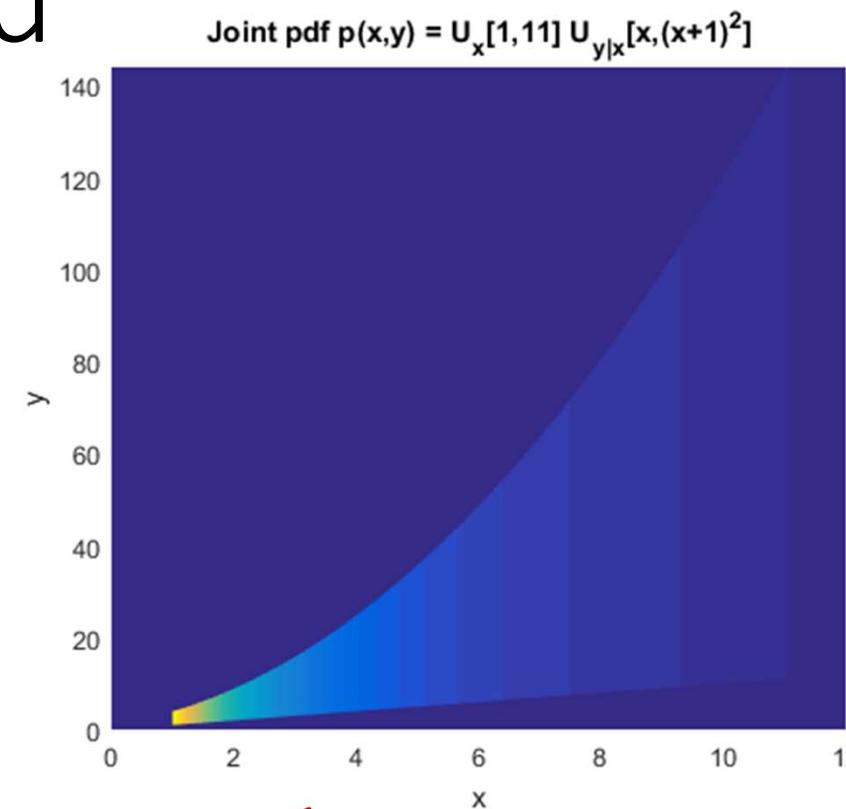
# Example: Conditional Uniform PDF Revisited

## (Lec 14, Special Topic #2)

Given joint pdf  $p(x,y) = p(x) p(y|x) = U_x[1,11] U_{y|x}[x, (x+1)^2]$

- Suppose true (unknown)  $x \sim p(x) = U_x[1,11] \rightarrow x = 5.0723$
- 3 indep. observations in sequence,  $y_k \sim p(y_k|x) = U_{y|x}[x, (x+1)^2]$ :

$$y_1 = 36.4052, y_2 = 33.3684, y_3 = 6.3725$$

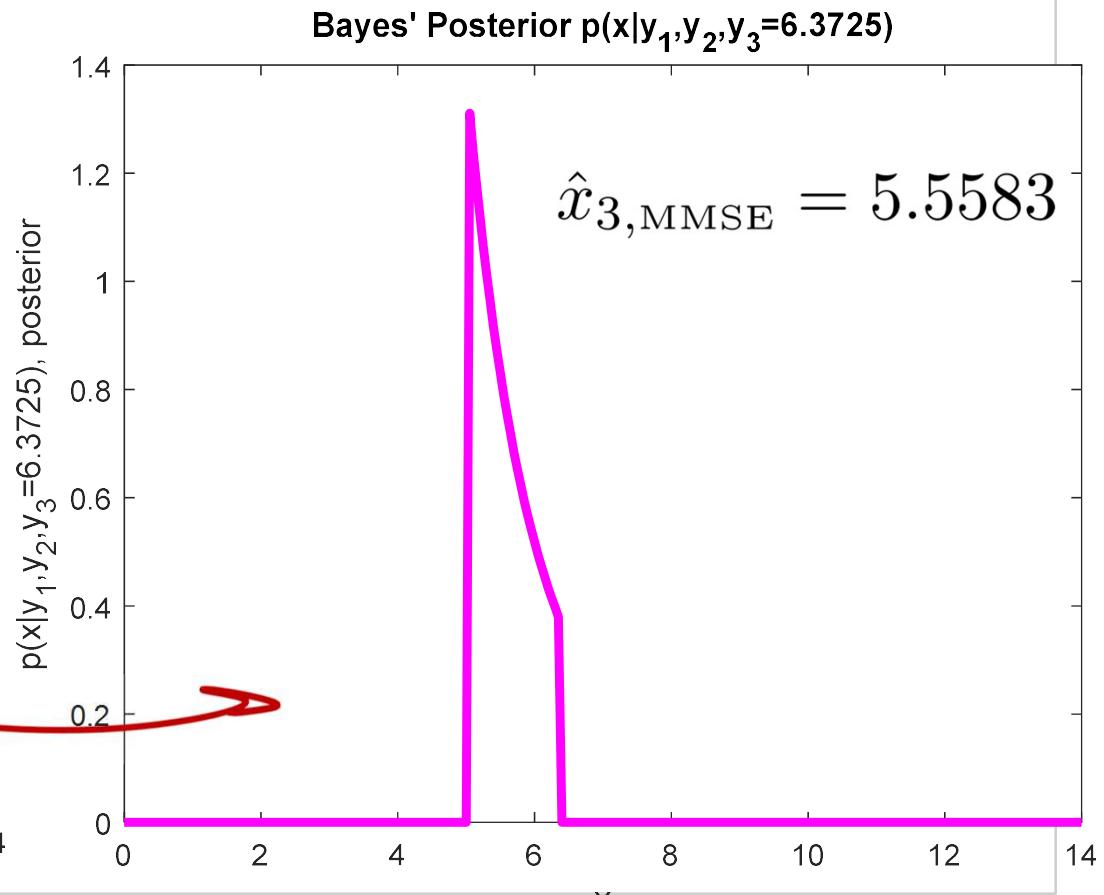
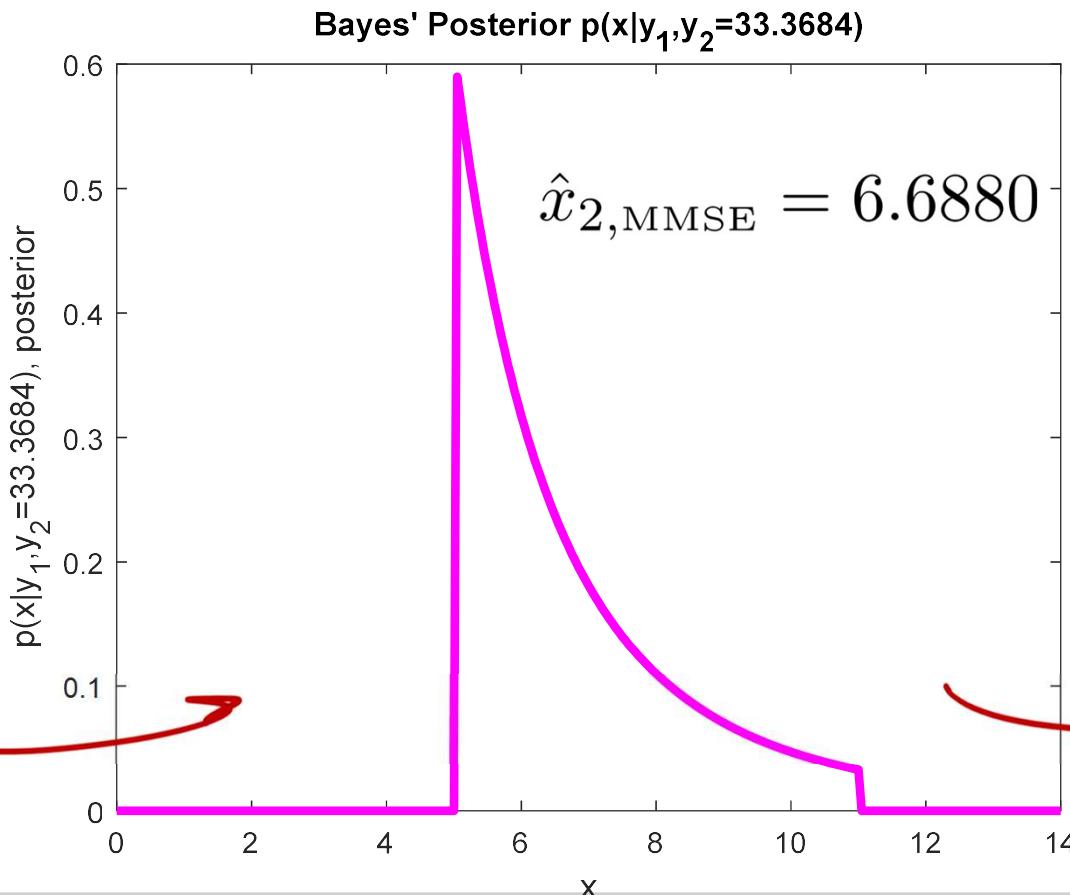
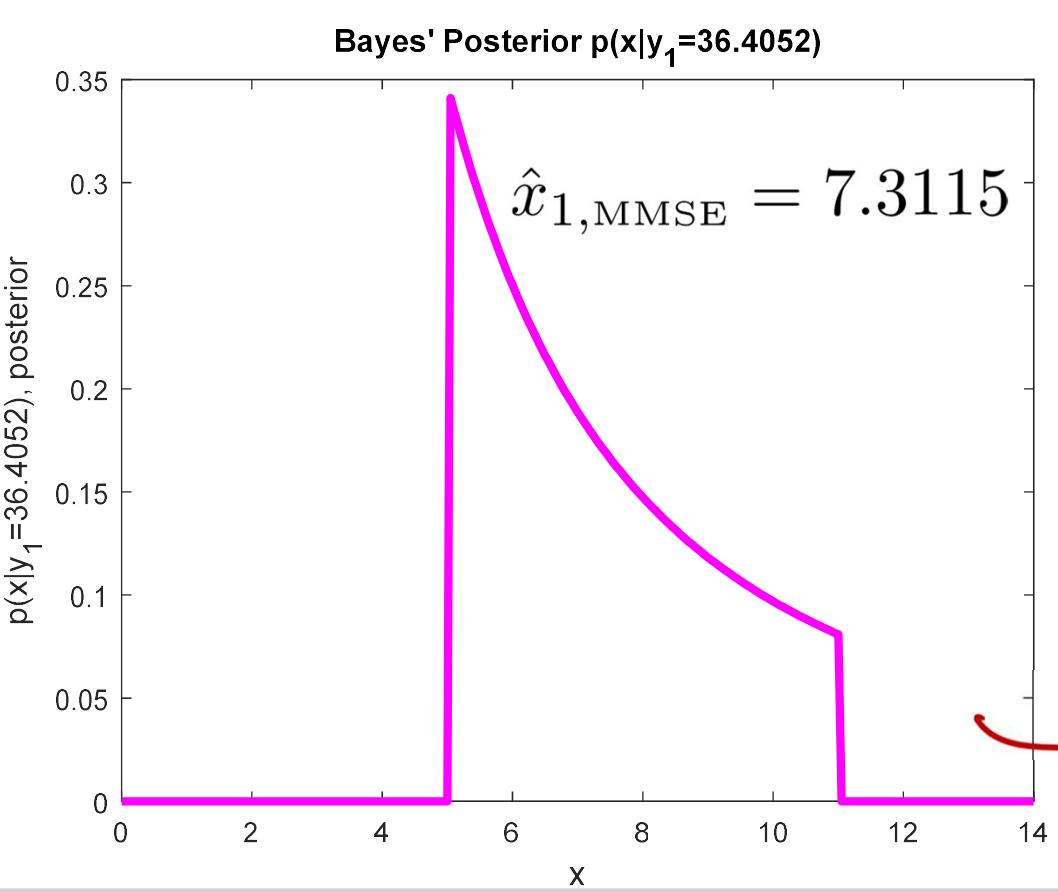
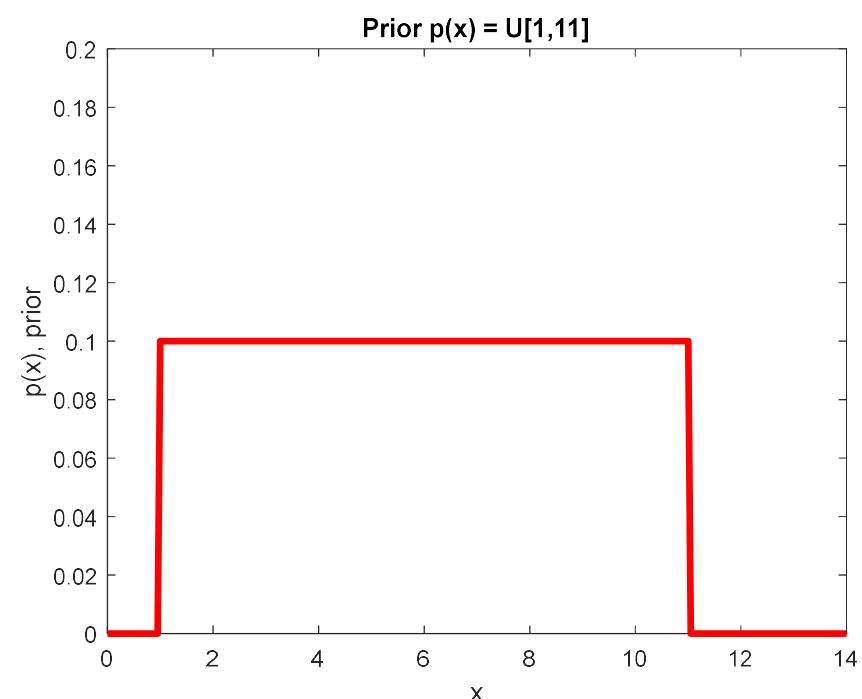


# Example: Conditional Uniform PDF

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# General probabilistic forward state prediction

- Suppose now  $x$  is NOT a static random vector, but has dynamics and process noise
- How to deal with general (non-linear/non-Gaussian) forward dynamical state propagation in a probabilistic (Bayesian) way?

Assume • suppose DT dynamics are  $x_{k+1} = f(x_k, u_k, w_k, k)$ , where  $w_k \sim p(w_k)$   
1st order markov process for  $x_k \rightarrow x_{k+1}$

• Can generally define the conditional state transition PDF:  $x_{k+1} | x_k \sim p(x_{k+1} | x_k)$

- Suppose  $x_k$  has known  $p(x_k)$   $\rightarrow$  what is  $p(x_{k+1})$ ?

$\rightarrow$  using law of total probability, we obtain the Chapman-Kolmogorov Eqn.

$$\textcircled{*} \quad p(x_{k+1}) = \int_{-\infty}^{\infty} p(x_{k+1} | x_k) \cdot p(x_k) dx_k$$

The basis of probabilistic recursive prediction:  $p(x_{k+2}) = \int_{-\infty}^{\infty} p(x_{k+2} | x_{k+1}) \cdot p(x_{k+1}) dx_{k+1} = \int_{-\infty}^{\infty} p(x_{k+2} | x_{k+1}) \cdot \left[ \int_{-\infty}^{\infty} p(x_{k+1} | x_k) \cdot p(x_k) dx_k \right] dx_{k+1}$

# The General Probabilistic Filtering Problem

- Given
  - Set of measurements (online)  $y_1, \dots, y_k$
  - & initial state PDF  $P(x_0)$  & state transition pdfs  $P(x_{k+1} | x_k)$
- Find:  $P(x_k | y_1, \dots, y_k)$  for each time step  $k=1, 2, 3, \dots$
- & estimate  $x_k$  as  $\hat{x}_k = \underset{x \in \mathbb{R}^n}{\operatorname{argmin}} E[\mathcal{C}(x, \hat{x}_k)] P(x_k | y_1, \dots, y_k)$

# The General Bayes Filter (Recursive!)

Initialized with some initial guess of state vector pdf at k=0:

Given:  $p(x_0)$  ( & also know:  $p(x_{k|k} | x_k)$  for all  $k=0,1,2,\dots$  &  $p(y_k | x)$ , &  $y_i \neq y_j$  given  $x$ )

Consists of two key steps for subsequent time steps  $k=1,2,3,\dots$ :

1. Dynamic Prediction Step (Chapman-Kolmogorov)

$$@ k=1: p(x_1) = \sum_{-\infty}^{\infty} p(x_1 | x_0) \cdot p(x_0) \cdot dx_0$$

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$$@ k \geq 2: p(x_k | y_1:k-1) = \sum_{-\infty}^{\infty} p(x_k | x_{k-1}) \cdot p(x_{k-1} | y_1:k-1) \cdot dx_{k-1}$$

2. Bayesian Measurement update step

$$p(x_k | y_k) = \frac{p(x_k) \cdot p(y_k | x_k)}{\sum_{-\infty}^{\infty} p(x_k) \cdot p(y_k | x_k) dx_k}$$

$$p(x_k | y_1, \dots, y_{k-1}) = \frac{p(x_k | y_1:k-1) \cdot p(y_k | x_k)}{\sum_{-\infty}^{\infty} p(x_k | y_1:k-1) \cdot p(y_k | x_k) dx_k}$$