

# Likelihood Inference for Large Financial Systems

Justin A. Sirignano

Stanford University  
[www.stanford.edu/~jasirign](http://www.stanford.edu/~jasirign)

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Joint work with **Gustavo Schwenkler** (Boston University) and  
**Kay Giesecke** (Stanford University)

# Introduction

- Large interacting systems appear in many areas
- Finance: models of default and systemic risk for large credit pools and banking systems
  - Bush, Hambly, Haworth, Jin, and Reisinger (2011)
  - Cvitanic, Ma, and Zhang (2011)
  - Fouque and Ichiba (2012)
  - Garnier, Papanicolaou, and Yang (2012)
  - Giesecke, Spiliopoulos, Sowers, and Sirignano (2013)
  - Spiliopoulos, Sirignano, and Giesecke (2014)
- **Statistical estimation of model parameters is challenging** due to the size and complexity of the systems

# Our Approach

- We exploit weak convergence results for large systems to develop tractable parameter estimation.
- **Large system approximation** constructed using:
  - Law of large numbers
  - Central limit theorem
- Instead of parameter estimation using the true likelihood of large but finite systems, we estimate parameters using the **approximate likelihood** from the large system approximation.
- Approximate maximum likelihood estimators (MLEs) are computationally tractable and we show they are consistent and asymptotically normal.

We consider a class of large systems:

- Interacting particle system of jump-diffusion processes
- Each particle is driven by
  - an **idiosyncratic** process
  - a **systematic** random process
  - **interaction with other particles** through a mean field term
- Default and systemic risk models mentioned earlier fall into this class of systems.

We consider a system of interacting particles  $\{Y_t^n\}_{n=1}^N$  with a systematic process  $X$  and parameter vector  $\theta$ :

$$dX_t = a(X_t)dt + dW_t,$$

$$dY_t^n = \nu_\theta dt + \sigma_\theta^\top dW_t^n + \gamma_\theta dJ_t^N + \zeta_\theta^\top dX_t, \quad n = 1, \dots, N,$$

where the coefficient functions  $\nu, \sigma, \gamma, \zeta$  can depend upon  $(Y_t^n, X, \mu_t^N)$ . The particles interact through the empirical measure

$$\mu_t^N = \frac{1}{N} \sum_{n=1}^N \delta_{Y_t^n},$$

and the jump process  $J_t^N$ , whose intensity can also depend upon  $(Y_t^n, X, \mu_t^N)$ . **Our goal is to estimate  $\theta$ !**

# Data

- Suppose we have the observation  $D_{M,N} = (Z_{M,N}, X_M)$  where

$$Z_{M,N} = (Z_{t_1}^N, \dots, Z_{t_M}^N),$$

$$Z_t^N = \langle f, \mu_t^N \rangle,$$

$$X_M = (X_t)_{t \in [0, t_M]}.$$

- The goal is to estimate the true parameter  $\theta_0$  from the data  $D_{M,N}$ .

# Likelihood Function

- The true likelihood of the observed data is

$$\mathcal{L}_{M,N}(\theta) = p_{\theta}^N(Z_{t_1}^N, \dots, Z_{t_M}^N | \mathbf{X}_M)$$

- These likelihood functions are typically computationally intractable!
  - The transition density is usually only available via Monte Carlo simulation and is computationally expensive for large  $N$ .
  - Furthermore, computing  $p_{\theta}^N$  typically requires filtering. Here, there are **N particles to filter!** Consequently, maximum likelihood estimation is computationally intractable.

- Our approach: use the **large system approximation** to obtain a computationally tractable **approximate likelihood function**.
- Define the empirical fluctuation process

$$\Xi_t^N = \sqrt{N}(\mu_t^N - \bar{\mu}_t).$$

- If  $(\mu_t^N, \Xi_t^N, X) \xrightarrow{d} (\bar{\mu}_t, \Xi_t, X)$ , we have the second-order approximation

$$\mu_t^N \xrightarrow{d} \bar{\mu}_t + \frac{1}{\sqrt{N}} \Xi_t.$$

- Typically,  $\bar{\mu}_t$  satisfies a nonlinear stochastic evolution equation and  $\Xi_t$  is the conditionally Gaussian solution of a linear stochastic evolution equation.

# Stochastic Evolution Equations

- The limiting measures satisfy stochastic evolution equations:

$$\begin{aligned} d\bar{\mu}_s &= \mathcal{A}_{X_s, \bar{\mu}_s}^1 \bar{\mu}_s ds + \mathcal{A}_{X_s, \bar{\mu}_s}^2 \bar{\mu}_s dX_s, \\ d\bar{\Xi}_s &= \mathcal{G}_{X_s, \bar{\mu}_s}^1 \bar{\Xi}_s ds + \mathcal{G}_{X_s, \bar{\mu}_s}^2 \bar{\Xi}_s dX_s + d\bar{\mathcal{M}}_s, \end{aligned}$$

- Reduce  $\bar{\Xi}_s$  to a system of SDEs:

$$d\bar{\mathbf{v}}_t = A_1(X_t)\bar{\mathbf{v}}_t dt + A_2(X_t)\bar{\mathbf{v}}_t dX_t + Bd\mathbf{B}_t,$$

- Use the fundamental solution  $\Psi(s, t)$ !

$$\bar{\mathbf{v}}_t = \Psi(s, t)\bar{\mathbf{v}}_s + \Psi(s, t) \int_s^t \Psi^{-1}(s, u) Bd\mathbf{B}(u).$$

# Approximate Likelihood Function

- The approximate likelihood is

$$\mathcal{L}_{M,N}^A(\theta) = \bar{p}_\theta^N(Z_{t_1}^N, \dots, Z_{t_M}^N | \mathbf{X}_M),$$

where  $\bar{p}_\theta^N$  is the density under the second-order approximation  
 $\mu_t^N \stackrel{d}{\approx} \bar{\mu}_t + \frac{1}{\sqrt{N}} \Xi_t$ .

- The approximate estimator is

$$\hat{\theta}_{M,N}^A = \arg \max_{\theta \in \Theta} \mathcal{L}_{M,N}^A(\theta).$$

- $\bar{p}_\theta^N(Z_{M,N} | \mathbf{X}_M)$  can be calculated **semi-analytically** since the large system approximation is **conditionally Gaussian**!

## Theorem (Consistency of Approximate Estimator)

*Suppose the data  $D_{M,N}$  is produced by the true parameter  $\theta_0$ . Provided certain technical assumptions hold (see Appendix), then the approximate estimator is consistent:  $\hat{\theta}_{M,N}^A \xrightarrow{P} \theta_0$  as  $N \rightarrow \infty$ .*

## Theorem (Asymptotic Normality of Approximate Estimator)

*In addition, provided certain technical assumptions, the approximate estimator is asymptotically normal:*

$\sqrt{N}(\hat{\theta}_{M,N}^A - \theta_0) \xrightarrow{d} Q$ , where  $Q$  is conditionally Gaussian given  $X_M$ .

# Relationship with Least Squares

- The approximate estimator converges to a weighted least squares estimator as  $N \rightarrow \infty$ :

$$\theta_{M,\infty}^A \in \arg \max_{\theta \in \Theta} -\frac{1}{2}(\mathbf{Z}_{M,\infty} - m_\theta^M)^\top (\Sigma_\theta^M)^{-1} (\mathbf{Z}_{M,\infty} - m_\theta^M).$$

- $\Sigma_\theta^M$  is the covariance from the CLT with parameter  $\theta$  and  $m_\theta^M$  is the LLN with parameter  $\theta$ .
- In fact, the (unweighted) least squares estimator using just the LLN is a consistent estimator of  $\theta_0$ !

Consider the mean-field system:

$$\begin{aligned} dY_t^{N,n} &= \beta^C \langle y, \mu_t^N \rangle dt + \sqrt{Y_t^{N,n}} dW_t^{N,n} + \beta^S Y_t^{N,n} dX_t, \\ dX_t &= dW_t. \end{aligned} \tag{1}$$

We wish to estimate the parameters  $\theta = (\beta^C, \beta^S)$ .

N	Approx. Estimator for $\beta^C$	Approx. Estimator for $\beta^S$
100	.2397	.008325
500	.2635	.008660
1,000	.2429	.01122
10,000	.2492	.01035
100,000	.2495	.009931

Table : Approximate estimators for the system (1) with true parameter  $\theta_0 = (.25, .01)$ . N is the size of the system.

# Extensions

- Discrete observations of  $X$
- Continuous observation of  $Z_t^N$
- Simultaneous parameter estimation for the systematic process  $X$
- Discrete instead of continuous-time dynamics

# Conclusion

- Parameter estimation for a class of interacting systems is computationally intractable
- Approximate likelihood estimators are obtained via a large system approximation
- Approximate estimators are consistent and asymptotically normal