

18.650 – Fundamentals of Statistics

4. Parametric hypothesis testing

Objectives

At the end of this lecture, you will be able to do the following:

- Reformulate experimental questions in terms of a **hypothesis test** by specifying an appropriate **null hypothesis** and an **alternative hypothesis** .
- **Design** statistical **tests** to decide between a null and alternative hypothesis.
- Understand the types of error of a test
- Compute the power function of a test
- Design statistical tests with a specific level or asymptotic level.
- Apply a test to a given sample to determine whether or not the null hypothesis should be rejected.
- Compute a test from a confidence interval
- Compute and interpret the p-value associated to a statistical test.

Goals

Recall: waiting time in the ER

$$H_0 : \mu \leq 30$$

$$H_1 : \mu > 30$$



How to perform this test based on data?

- ▶ test statistic
- ▶ rejection region
- ▶ p-value

How to measure the performance of a test?

- ▶ Type I and type II errors
- ▶ level
- ▶ power

Construct PARAMETRIC tests:

$$H_0 : \mu \leq 30$$

$$H_1 : \mu > 30$$

▶ Wald test

▶ T-Test

Waiting time in the ER

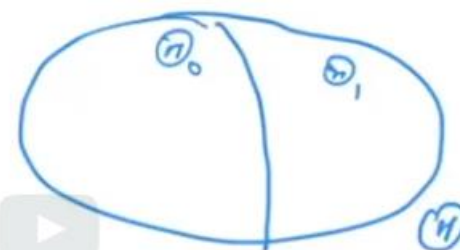
- ▶ The average waiting time in the Emergency Room (ER) in the US is 30 minutes according to the CDC
- ▶ Some patients claim that the new Princeton-Plainsboro hospital has a longer waiting time. Is it true?
- ▶ Collect a sample: X_1, \dots, X_n (waiting time in minutes for n random patients) with unknown expected value $\mathbb{E}[X_1] = \mu$.
- ▶ We want to know if $\mu > 30$.

$$H_0: \mu \leq 30$$
$$\rightarrow H_1: \mu > 30$$



Statistical formulation

- ▶ Consider a sample X_1, \dots, X_n of i.i.d. random variables and a statistical model $(E, (\mathbb{P}_\theta)_{\theta \in \Theta})$.
- ▶ Let Θ_0 and Θ_1 be a *partition* of Θ .
- ▶ Consider the two hypotheses:
$$\begin{cases} H_0 : & \theta \in \Theta_0 \\ H_1 : & \theta \in \Theta_1 \end{cases}$$
- ▶ H_0 is the *null hypothesis*, H_1 is the *alternative hypothesis*.
- ▶ We say that we *test* H_0 against H_1 .



Testing lexicon

H_k

- ▶ For $k = 0$ (H_0) or $k = 1$ (H_1), we say that
 - ▶ Θ_k is a simple hypothesis if $\Theta_k = \{\theta_k\}$
 - ▶ Θ_k is a composite hypothesis if Θ_k is of the following three forms

$$\Theta_k = \{\theta : \theta > \theta_k\}, \quad \Theta_k = \{\theta : \theta < \theta_k\}, \quad \Theta_k = \{\theta : \theta \neq \theta_k\}$$

- ▶ A test is typically either *one-sided* or *two-sided*

Two-sided

$$\begin{cases} H_0: \theta = \theta_0 \\ H_1: \theta \neq \theta_0 \end{cases}$$

One-sided

$$\begin{cases} H_0: \theta \leq \theta_0 \\ H_1: \theta > \theta_0 \end{cases} \quad \text{or} \quad \begin{cases} H_0: \theta \geq \theta_0 \\ H_1: \theta < \theta_0 \end{cases}$$

Examples

1. Waiting time in the ER

$$H_0 : \mu \leq 30$$

$$H_1 : \mu > 30$$

One-sided test.

composite

composite

2. In the Kiss example, we want to test

$$H_0 : p = .5$$

$$H_1 : p \neq .5$$

simple

$$\Theta_0 = \{.5\}$$

composite

Clinical trials

- ▶ Pharmaceutical companies use hypothesis testing to test if a new drug is efficient.
- ▶ To do so, they administer a drug to a group of patients (*test* group) and a placebo to another group (*control* group).
- ▶ We consider testing a drug that is supposed to lower LDL (low-density lipoprotein), a.k.a "bad cholesterol" among patients with a high level of LDL (above 200 mg/dL)

Notation and modelling

- ▶ Let $\mu_d > 0$ denote the expected decrease of LDL level (in mg/dL) for a patient that has used the drug.
- ▶ Let $\mu_c > 0$ denote the expected decrease of LDL level (in mg/dL) for a patient that has used the placebo.
- ▶ Hypothesis testing problem:

$$H_0 : \mu_d \leq \mu_c$$
$$H_1 : \mu_d > \mu_c$$



- ▶ We observe two independent samples:
 - ▶ $X_1, \dots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu_d, \sigma_d^2)$ from the *test* group and
 - ▶ $Y_1, \dots, Y_m \stackrel{iid}{\sim} \mathcal{N}(\mu_c, \sigma_c^2)$ from the *control* group.
- ▶ This is a *two-sample* test: these are very common (A/B testing).

Asymmetry in the hypotheses

$$H_0: \mu_d \leq \mu_c$$

$$H_1: \mu_d > \mu_c$$

- ▶ We want to decide whether to *reject* H_0 (look for evidence against H_0 in the data).
- ▶ H_0 and H_1 do not play a symmetric role: the data is only used to try to disprove H_0

$$H_0: \text{status quo}$$

$$H_1: \text{a (scientific) discovery}$$

- ▶ In particular lack of evidence, does not mean that H_0 is true (“innocent until proven guilty”)

Examples

1. Waiting time in the ER

$$H_0 : \mu \leq 30 \rightarrow \text{status quo}$$

$$H_1 : \mu > 30$$

Status quo: CDC statement. We collect data to show that Princeton-Plainsboro is different

2. Kiss

$$H_0 : p = .5$$

$$H_1 : p \neq .5$$

Status quo: our intuition tells us there should be no preference. We collect data to show that there is one.

3. Clinical trials

$$H_0 : \mu_d \leq \mu_c$$

$$H_1 : \mu_d > \mu_c$$

Status quo: The drug is not more effective than a placebo. We collect data to prove that the drug is effective.

What is a test?

- ▶ A *test* is a statistic $\psi \in \{0, 1\}$ that does not depend on unknown quantities and such that:

- ▶ If $\psi = 0$, H_0 is not rejected;

- ▶ If $\psi = 1$, H_0 is rejected.

Important remark: Can always write $\psi = \mathbb{1}\{R\}$, where R is an event called *rejection region*.

$$\psi = \mathbb{1}(\psi = 1)$$

- ▶ Waiting time in the ER:

$$H_0 : \mu \leq 30$$

$$H_1 : \mu > 30$$

$$\psi = \mathbb{1}\{\bar{X}_n > c\}$$

- ▶ Kiss:

$$H_0 : p = .5$$

$$H_1 : p \neq .5$$

$$\psi = \mathbb{1}\{|\bar{X}_n - \frac{1}{2}| > c\}$$

- ▶ Clinical trials

$$H_0 : \mu_d \leq \mu_c$$

$$H_1 : \mu_d > \mu_c$$

$$\psi = \mathbb{1}\{\bar{X}_n - \bar{Y}_m > c\}$$

Errors

A test can make two types of errors:

	Fail to reject Null	Reject Null
H_0 true ($\theta \in \Theta_0$)	✓	type 1
H_1 true ($\theta \in \Theta_1$)	type 2	✓

Both errors can be computed from the *power function*

$$\beta(\theta) = \mathbb{P}_\theta[\psi = 1]$$

► If $\theta \in \Theta_0$,

$$\beta(\theta) = \mathbb{P}_\theta[\psi \text{ makes an error of type } 1]$$

We want $\beta(\theta)$ to be *small*

► If $\theta \in \Theta_1$,

$$\beta(\theta) = 1 - \mathbb{P}_\theta[\psi \text{ makes an error of type } 2]$$

We want $\beta(\theta)$ to be *large*

The Neyman-Pearson paradigm

Recall the waiting time in the ER example

$$\begin{aligned} H_0 : \mu &\leq 30 \\ H_1 : \mu &> 30 \end{aligned} \quad \psi = \mathbb{I}\{\bar{X}_n > C\}$$

How to choose C ?

We are facing a dilemma: both errors should be small!

► To make Type I error $\rightarrow 0$, take $C \rightarrow +\infty$

► To make Type II error $\rightarrow 0$, take $C \rightarrow -\infty$ (enough to have $C = \infty$)

Cannot make both small at the same time.

The *Neyman-Pearson paradigm*:

- Make sure that $\mathbb{P}[\text{Type I error}] \leq \alpha$ (e.g., $\alpha = 5\%, 1\%, \dots$)
- Minimize $\mathbb{P}[\text{Type II error}]$ subject to this constraint

Level

The value of $\alpha \in [0, 1]$ chosen in the Neyman-pearson paradigm is called *level* of a test

For which $\theta \in \Theta_0$ should we compute $\mathbb{P}_\theta[\psi = 1]$ (probability of Type 1 error)?

- ▶ A test ψ has *level* α if

$$\mathbb{P}_\theta[\psi = 1] \leq \alpha, \quad \forall \theta \in \Theta_0.$$
$$\Leftrightarrow \max_{\theta \in \Theta_0} \mathbb{P}_\theta[\psi = 1] \leq \alpha$$

- ▶ A test $\psi = \psi_n$ has asymptotic level α if

$$\lim_{n \rightarrow \infty} \max_{\theta \in \Theta_0} \mathbb{P}_\theta[\psi_n = 1] \leq \alpha,$$

Building a test from a confidence interval

Given a confidence interval, we can often build a test (and vice versa).

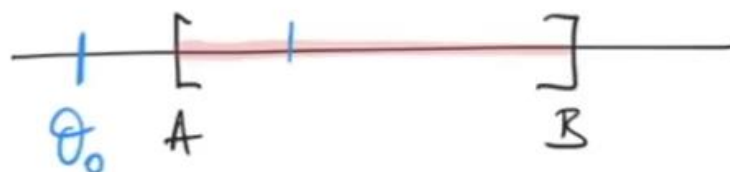
- ▶ Let $I = [A, B]$ be a confidence interval at level $1 - \alpha$ for a parameter θ :

$$\mathbb{P}_{\theta}(\theta \in [A, B]) \geq 1 - \alpha$$

- ▶ We want to use this I to build a test at level α for

$$H_0 : \theta = \theta_0$$

$$H_1 : \theta \neq \theta_0$$



- ▶ Natural candidate:

$$\psi = \mathbb{I} \{ \theta_0 \notin [A, B] \}$$

- ▶ Level of test:

$$\mathbb{P}_{\theta_0}[\psi = 1] = \mathbb{P}_{\theta_0}[\theta_0 \notin I] = 1 - \mathbb{P}_{\theta_0}[\theta_0 \in I] \leq 1 - (1 - \alpha) = \alpha$$

- ▶ Therefore ψ is a test with level

A test for the Kiss example

We want to test:

$$H_0 : p = 0.5$$

$$H_1 : p \neq 0.5$$

We observe $R_1, \dots, R_n \stackrel{iid}{\sim} \text{Ber}(p)$.

► Recall that

$$\mathcal{I}_{\text{conserv}} = \left[\bar{R}_n - \frac{1.96}{2\sqrt{n}}, \bar{R}_n + \frac{1.96}{2\sqrt{n}} \right]$$

is a confidence interval of asymptotic level $1 - \alpha$ for p .

► Consider the test:

$$\psi = \mathbb{I} \{ 0.5 \notin \mathcal{I}_{\text{conserv}} \}$$

$$\text{IP}_\alpha = 95\%$$

$$\mathcal{I}_{\text{conserv}} = [0.56, 0.73]$$

↳ reject.

► We have

$$\lim_{n \rightarrow \infty} \mathbb{P}_{.5}[\psi = 1] = 1 - \lim_{n \rightarrow \infty} \mathbb{P}_{.5} [.5 \in \mathcal{I}_{\text{conserv}}] \leq 1 - (1 - \alpha) = \alpha$$

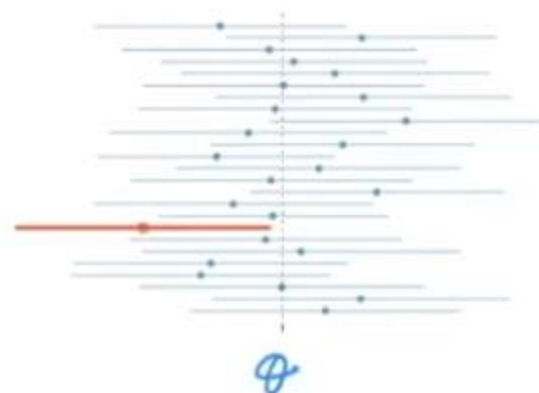
► Therefore ψ is a test with asymptotic level α

Meaning of the level

- Recall that

\mathcal{I} is a CI at level 95% for θ

means that if we repeat the experiment many times, at least 95% confidence intervals will contain the true parameter θ .



- Similarly:

ψ is a test at level 5% for H_0 vs H_1

means that if we repeat the experiment many times, at most 5% of the tests will make an error of type 1

What if we change the level?

$$\mathcal{I}_{\text{conserv}} = \left[\bar{R}_n \pm \frac{q_{\alpha/2}}{2\sqrt{n}} \right]$$

↙ level 95%

With our data $\mathcal{I}_{\text{conserv}} = [0.56, 0.73]$ so we *reject H_0 at level 5%*

α	$q_{\alpha/2}$	$\mathcal{I}_{\text{conserv}}$	decision
10%	1.64	[0.57, 0.72]	<i>Reject</i>
<i>5%</i>	<i>1.96</i>	<i>[0.56, 0.73]</i>	<i>Reject</i>
1%	2.76	[0.52, 0.77]	<i>Reject</i>
.1%	3.29	[0.497, 0.79]	<i>Fail to reject</i>
.01%	3.89	[0.47, 0.82]	<i>Fail to reject</i>

The value of α across which we switch from "reject" to "fail to reject" is called the .

p-value

Definition

The (asymptotic) *p-value* of a test ψ is the smallest (asymptotic) level α at which ψ rejects H_0 .

Golden rule

$\text{p-value} \leq \alpha \Leftrightarrow H_0$ is rejected by ψ , at the (asymptotic) level α .

p-value > $\alpha \Leftrightarrow H_0$ is not rejected by ψ , at the (asymptotic) level α .

Kiss example: we need to find α_0 such that $\bar{R}_n - \frac{q_{\alpha_0/2}}{2\sqrt{n}} = 0.5$

If $\bar{R}_n = .645$, $n = 124$ we get $q_{\alpha_0/2} = 3.23$. To find α_0 :

$$\frac{\alpha_0}{2} = P\left[Z > q_{\frac{\alpha_0}{2}}\right] = P[Z > 3.23] = 1 - .9994 = 0.0006 \Rightarrow \alpha_0 = 0.12\%$$

where $Z \sim \mathcal{N}(0, 1)$ and $\mathbb{P}(Z \leq 3.24) = 0.9994$ (read from table).

The evidence scale

- ▶ Statisticians, and more generally researchers, are used to communicating directly in terms of p-values rather than "reject/fail to reject at level..."
- ▶ The mental conversion is as follows:

p-value	evidence against H_0
$> 10\%$	almost none
$[5\%, 10\%]$	weak
$[1\%, 5\%]$	strong
$[\cdot 1\%, 1\%]$	very strong
$< \cdot 1\%$	undisputable