ST1131

AY21/22 Sem 2

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01. Exploratory Data Analysis

- Quantitative Variable: Discrete vs. Continuous
- Categorical Variable: Ordinal vs. Nominal
- Difference: Is distance between 2 points meaningful?

Single Variable

Frequency Table - Categorical

- ullet Proportion aka relative frequency. $\frac{\# \text{ of obs. in } 1 \text{ cat.}}{Total\#ofobs.}$
- **Modal Frequency** Category with highest frequency
- Summarizing: Modal category and its proportion

Bar Plots - Categorical

 Summarizing: Modal category and its proportion, Cat. with high/low proportions, Mention trends if ordinal

Histogram - Quantitative

- Skewed left/right: Left/right tail is longer
- Summarizing: Unimodal/Bimodal/Multimodal, Skewness, Outlier

Describing Center

- Mean $\bar{X} = \frac{1}{n} \sum_{i=1}^{n} x_i$
- Linear Transformation: $\hat{Y} = b\hat{X} + a$
- Sensitive to outliers, unlike median
- Median X_(0.5)
- \bullet If $\bar{X} > X_{(0,5)}$, skew right. If $\bar{X} < X_{(0,5)}$, skew left.

Describing Variability

- Range Sensitive to outliers
- Variance $S^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i \bar{x})^2$
- Standard deviation $sd = \sqrt{S^2}$
- Linear Transformation: $S_y^2 = b^2 s_x^2 \ S_y = |b| s_x$
- Inter-quartile Range (IQR) $Q_3 Q_1$
- ullet Quantile (q_p) 100p% of observations are below q_p
- ullet Lower quartile (Q_1) , Median (Q_2) , Upper quartile (Q_3)
- \bullet Symm. \rightarrow Mean, Variance. Skewed \rightarrow Median, IQR.

Boxplot - Variability

- Includes: Min, Q_1 , Q_2 , Q_3 , Max
- Outliers $< Q_1 1.5IQR$ or $> Q_3 + 1.5IQR$
- Max/min Whisker Reach Boundary of outliers
- **Upper/lower Whisker** Min/max obs. exc. outliers
- If unimodal, can show skewness.
- \bullet Summarizing: Median, Outliers, Compare medians and IQRs if >1 boxplots

Two Variables

• Response Variable vs. Explanatory Variable

Bar Plots - 2 categorical

Contingency Table - 2 categorical

- Conditional Percentage % out of total
- Join Percentage % out of some group. Use explanatory as group.
- Be careful of phrasing (Eg. Ppl w/o cancer of PMH users vs. PMH users of those w/o cancer)
- Relative Risk Ratio of 2 percentages. (Eg. % of cancer in PMH users is 1.24 times the % of cancer in non-PMH users)

2 Boxplots - 1 Categorical and 1 Quantitative

Scatter Plot - 2 Quantitative Variables

 Summarizing: Pos./neg. association, Linear, Constant variability, Outliers

Correlation - $r \in [-1, 1]$

ullet $r=\pm 1 o$ Correlation is linear

02. Data Collection

- Confounding Variable Related to exp. and resp. variable. Confounds their association. Observed.
- Lurking Variable Unobserved
- Experimental Study Assign subjects (or experimental units) to treatments and observe response variable
- Observational Study Explanatory and response variable observed for subjects. No treatments.

Sample Survey

- 1. Compile sampling frame Where sample is from
- Sampling design How to choose subjects from sampling frame
- Simple Random Sample Each sample has same chance of being chosen

Sources of Bias in Sample Survey:

- Sampling Bias Sample not random or undercoverage
- Non-response Bias No response from subject
- Response Bias Subject does not answer truthfully

Elements of Good Experimental Study:

- Control comparison group
- Randomization: Eliminate lurking variables
- Blinding the study: Placebo

03. Probability

- ullet Sample space (S) Set of all possible outcomes
- ullet **Event** (E) Subset of sample space
- $P(A) = \frac{\# \text{ of outcomes in A}}{\text{Total } \# \text{ of possible outcomes}}$

Axioms of Probability

- 1. $0 \le P(A) \le 1$
- 2. P(S) = 1
- 3. If A and B are mutually exclusive (or disjoint), then $P(A \cup B) = P(A) + P(B)$ and $P(A \cap B) = 0$
- **4.** $P(A \cup B \cup C) = P(A) + P(B) + P(C) P(A \cap B) P(A \cap C) P(B \cap C) + P(A \cap B \cap C)$

For any events A and B:

- $P(A^c) = 1 P(A)$
- $P(A \cup B) = P(A) + P(B) P(A \cap B)$
- $P(A) = P(A \cap B) + P(A \cap B^c)$
- A and B are independent if $P(A \cap B) = P(A)P(B)$

Conditional Probability

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Law of Total Probability

 $P(A) = P(A \cap B_1) + \dots + P(A \cap B_n)$

Bayes' Theorem

$$P(B_i|A) = \frac{P(A|B_i)P(B_i)}{P(A|B_1)P(B_1) + \dots + P(A|B_n)P(B_n)}$$

Epidemiological Terms:

- Sensitivity Given has disease, prob. of positive test
- Specificty Given has no disease, prob. of neg. test
- Prevalence # of people with disease
 Total population
- Pos. Predictive Val. Given pos. test, prob. of disease

04. Random Variables

- Random variable Unpredictable outcome of exp.
- Probability distribution of random variable values and their probabilities

Discrete Random Variables

• **Probability distribution**: (P_x) Prob. for each possible x. Sum of all $P_x = 1$

Mean

- aka expected value. $\mu = \sum_x x P_x$
- Linear transformation: $E(Y)=a\mu+b$ and $E(a_1x_1+\ldots+a_nx_n)=a_1\mu_1+\ldots+a_n\mu_n$
- If $x_1,...,x_n$ have same prob. distri., mean of these variables (\bar{X}) is a random variable where $E(\bar{X})=\frac{1}{n}\sum_{i=1}^n \mu_i=\mu$

Variance

- ullet $\sigma^2 = \sum_x P_x (x \mu)^2$ and sd: σ
- Linear transformation: $Var(Y) = b^2 Var(X) = b^2 \sigma^2$ and $Var(a_1x_1 + ... + a_nx_n) = a_1^2 \sigma_1^2 + ... + a_n^2 \sigma_n^2$
- Likewise, $Var(\bar{X}) = \frac{\sigma^2}{n}$

Continuous Random Variables

- Probability distribution: Represented by probability density function. Area under curve = 1.
- Mean and variance have same properties as discrete

Binomial Distribution

- Combinations $-\binom{n}{k} = \frac{n!}{k!(n-k)!}$
- 3 Conditions for binomial distribution:
- 1. n trials with 2 outcomes
- 2. Each trial has probability of p to succeed
- 3. All trials are independent

Binomial Random Variable - # of successes in n trials

- Follows distribution: Bin(n, p)
- Bernoulli Distri. Bin(1, p). Sum of Ber. \sim Bin.

Binomial Formula: Suppose $X \sim Bin(n, p)$

•
$$P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$$

- \bullet E(X) = np
- $\bullet \ Var(X) = np(1-p)$

Normal Distribution

- Symmetric about μ , bell-shaped, $X \sim N(\mu, \sigma^2)$
- Note: In R, use sd. Else, use variance.
- Linear transformation of 2 Normal Random Variables: $aX + bY \sim N(a\mu_x + b\mu_y, a^2\sigma_x^2 + b^2\sigma_y^2)$
- Approximate binomial distribution using N(np, np(1-p)) when $np(1-p) \geq 5$

Z-score -
$$z = \frac{x-\mu}{\sigma} \sim N(0,1)$$

• Outlier: Any observation with z-score of > 3 or < -3

QQ Plot - Check normality

- Right tail below/above line → Longer/shorter
- ullet Left tail below/above line o Shorter/longer

05. Sampling Distribution

- Data Distribution Distribution of some observations
- Sampling Distribution Distribution of \bar{X} and \hat{y}
- Central Limit Theorem Suppose there are independent observations that form a distribution (not necessarily normal) with mean μ and variance σ^2 and sample size n is large, then sample mean $\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$

Sample Proportion \hat{p}

- Population Proportion p that we want to estimate
- Population Distribution Ber(p) where $\mu = p$ and $\sigma^2 = p(1-p)$
- • When $np(1-p) \geq 5$, $\hat{p} \sim N(p, \frac{p(1-p)}{n})$ appxorimately by CLT

Sample Mean $ar{X}$

when population distribution is normal:

- $E(\bar{X}) = \mu$ and $Var(\bar{X}) = \frac{\sigma^2}{n}$
- $\bar{X} \sim N(\mu, \frac{\sigma^2}{\sigma})$ exactly

when population distribution is not normal:

- $E(\bar{X}) = \mu$ and $Var(\bar{X}) = \frac{\sigma^2}{n}$
- ullet When $n\geq 30$, $ar{X}\sim N(\mu,rac{\sigma^2}{n})$ approximately by CLT

06. Confidence Intervals

 In long run, 95% of intervals will contain population proportion/mean.

Point Estimate

- ullet $ar{X}
 ightarrow \mu$ and $\hat{p}
 ightarrow p$
- Does not show how close they are to true value

Confidence Interval for Proportion

- $CI = Point estimate \pm Margin of error$
- Standard Error (se) Estimated sd of sampling distri-

Find CI given confidence lvl. (x):

- 1. Find \hat{p} and check $n\hat{p}(1-\hat{p}) \geq 5$
- 2. Let $\alpha = 1 x$
- 3. $CI = \hat{p} \pm q_{1-\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

Determine sample size (n) before study:

- 1. Decide confidence level (x) and width of CI (D)
- 2. $n \ge (\frac{2q_{1-\frac{\alpha}{2}}}{D})^2 p(1-p)$ where $p = \frac{1}{2}$

Confidence Interval for Mean

• **t-distribution** - (t_{df}) Approaches N(0,1) as $df \uparrow$

Find CI given confidence vI. (x):

- Assumptions: Sample is random (not robust crucial), Data distribution symmetric (or n is big)
- 2. $CI = \bar{X} \pm t_{n-1;1-\frac{\alpha}{2}} \frac{s}{\sqrt{n}}$

Determine sample size (n) before study:

- 1. Decide confidence level (x) and width of CI (D)
- 2. $n \ge (\frac{2sq_{1-\frac{\alpha}{2}}}{D})^2$
- 3. For s, look for similar studies. Ensure $n \ge 30$.

07. Hypothesis Testing

- Null hypothesis (H_0) vs. Alternative hypothesis (H_1)
- Test statistic How far point estimate falls from guess
- Null distribution Distribution of test stat. under H_0
- p-Value How unlikely observed value is, if H_0 is true
- Significance level (α) Reject H_0 if p-Val $\leq \alpha$
- Test is statistically significant when we reject H_0
- Test is statistically significant which we reject II
- Type I Error Reject H_0 , but H_0 is true
- Type II Error Do not reject H_0 , but H_0 is false
- Increase sample size to reduce both errors

One sample, Proportion

- 1. Assumptions: Categorical, Random, $np_0(1-p_0) \geq 5$
- 2. Hypothesis: $H_0: p = p_0$ and $H_1: p > \neq p_0$
- 3. Test statistic: $z=\frac{\hat{p}-p_0}{\sqrt{p_0(1-p_0)}}$ and $z\sim N(0,1)$
- 4. p-Value: If right sided, $P(z \geq \text{Test stat.}|z \sim N(0,1))$. If 2-sided, $2*P(z \geq \text{Test stat.}|z \sim N(0,1))$
- 5. Conclusion: Reject H_0 if p-Val $\leq \alpha$. Else, cannot reject

One sample, Mean

- 1. Assumptions: Quantitative, Random, Data distri. is approx. normal (or $n \geq 30$)
- 2. Hypothesis: $H_0: \mu = \mu_0$ and $H_1: \mu > < \neq \mu_0$
- 3. Test statistic: $T = \frac{\bar{X} \mu_0}{\frac{S}{\sqrt{n}}}$ and $T \sim t_{n-1}(0,1)$
- 4. p-Value and Conclusion: Same as proportion
- \bullet Result of 2-sided test for mean is same as using CI

Two sample, Independent, Equal variance

- 1. Assumptions: Quantitative, Random, Independent samples, Pop. distri. is approx. normal (or n is large enough), **Equal variance test** > 0.05
- 2. Hypothesis: $H_0: \mu_1 = \mu_2$ and $H_1: \mu_1 > < \neq \mu_2$
- 3. Test statistic: $T=\frac{\bar{X}-\bar{Y}}{se}$ where $se=s_p\sqrt{\frac{1}{n_1}+\frac{1}{n_2}}$ and $S_p^2=\frac{(n_1-1)s_1^2+(n_2-1)s_2^2}{n_1+n_2-2}$ (Pooled estimate of common variance) and $T\sim t_{n_1+n_2-2}$

Two sample, Independent, Unequal variance

- $1. \ \, \mathsf{Assumptions} \colon \mathsf{Same}, \ \mathsf{except} \ \mathsf{pop}. \ \mathsf{var}. \ \mathsf{is} \ \mathsf{different}$
- 2. Hypothesis: $H_0: \mu_1 = \mu_2$ and $H_1: \mu_1 > < \neq \mu_2$
- 3. Test statistic: $T=rac{ar{X}-ar{Y}}{se}$ where $se=\sqrt{rac{s_1^2}{n_1}+rac{s_2^2}{n_2}}$ and $T\sim t_{df}$ where df needs R

Two sample, Dependent

- \bullet 2 samples are dependent \leftrightarrow Each obs. has matching pair (Eg. Before and after)
- Take difference of matched observations and compare mean of difference with 0. Similar to 1 sample test.

08. Linear Regression

Simple Linear Regression

- $\bullet \ Y = \beta_0 + \beta_1 x + \epsilon$
- Assumptions: Random data, Relationship is linear, $\epsilon \sim N(0,\sigma^2)$ which implies $Y \sim N(\beta_0 + \beta_1 x,\sigma^2)$ (Check if response var. is symmetric)
- \bullet $\hat{\sigma} = \text{Residual standard error in model summary}$
- ullet Ordinary Least Square Estimate sum of square residuals. $ss_{Res} = \sum_{i=1}^n e_i^2$ where $e_i = y_i \hat{y}_i$
- Interpolation Estimate not observed. Within range
- Extrapolation Estimate that's outside range. Avoid!

T-test and F-test

- T-test Test significance of 1 coefficient
- F-test Test significance of entire model
- If coeff. are not significant, intercept model $\hat{Y} = \hat{\beta_0}$
- 1. Assumptions: Same as building model
- 2. Hypothesis for T-test: $H_0: \beta_1 = 0$ and $H_1: \beta_1 \neq 0$
- 3. Hypothesis for F-test: H_0 : All coeff. (except β_0) = 0 and H_1 : At least 1 coeff. is non-zero.
- 4. t-stat: Check from summary. Null distri: t_{n # of coeff
- 5. F-stat: For simple lin. reg., $F=t^2$. Null Distri: $F_\#$ of coeff; n # of coeff

Regression Diagnostics

Plots and what to look out for:

- Scatter plot $r_i vs. X_i/\hat{Y}_i$ and $Xvs.\hat{Y} \rightarrow$ Linearity (Curved band), Constant variance (Funnel shape), Normality (Many points outside (-3,3))
- ullet QQ Plot/Histogram of $r_i
 ightarrow$ Normality

How to fix:

- Constant variance: Transform response variable
- Linearity: Add higher order term

Coefficient of Determination - (R^2) Goodness of fit

- $|Cor(x,y)| = \sqrt{R^2} = R$
- ullet Weakness: More variables $ightarrow R^2 \uparrow$
- Thus, can use **Adjusted** R^2

Multivariable Linear Regression Regression Function with Categorical Var:

- Indicator Variable 1 if cat. is observed. 0 otherwise.
- Reference Category The category not in equation
- Eg. $Y = \beta_0 + \beta_1 x_1 + \beta_2 I(x_2 = Auto) + \epsilon$ • Auto: $Y = \beta_0 + \beta_1 x_1 + \beta_2 + \epsilon$
- Manual: $Y = \beta_0 + \beta_1 x_1 + \epsilon$

Interaction between variables: $Y = \beta_0 + \beta_1 x_1 + \beta_2 I(x_2 = Auto) + \beta_3 x_1 I(x_2 = Auto) + \epsilon$

09. R Code

```
matrix(c(1:6), nrow=2, ncol=3, byrow=T)
rbind(m, c(1,2,3))
data = read.csv("./crab.txt")
names(data) = c("Subject", "Gender")
rownames(mat) # colnames(mat)
data$Subject # Or attach()
data[1:8,]
data[Gender == "M" & HW == "A",]
# Replace elements based on condition
ifelse(Gender == "0", "F", "M")
# Return indices that match condition
which(flat == "3 ROOM")
for (i in 1:100) {...}
choose(6, 3)
summary(marks)
# Frequency Table
table(data)
prop.table(table(data))
# Bar Plot
barplot(table(data))
# Contingency Table
tab = table(bbd, pmh) # (r, c)
prop.table(tab) # Joint probabililty
prop.table(tab, "pmh") # Conditional
    probability on pmh groups
# Bar Plot with 2 variables
barplot(proptab)
# Boxplot
bp = boxplot(age cancer) # quan. cat.
```

```
bp$out # Values of outliers
grp = bp$group # Outliers in each group
which(grp == 1) # Outliers in group 1
bp$out[which(grp == 1)]
# Histogram
hist(flatPrice)
# Scatter Plot, Correlation
plot(size, price) # (x-axis, y-axis)
cor(size, price)
# QQ plot
qqnorm(data)
qqline(data)
# Generate vector of 10 IID samples
rbinom(n=10, size=100, prob=0.5)
rnorm(n=10, mean=100, sd=15)
rexp(n=10, rate=1/500)
\# P(X >= 70)
pbinom(70, 100, 0.5, lower.tail=T)
pnorm (70, 100, 15)
# Find q0.9 where area on left = 0.9
qbinom(0.9, 100, 0.5)
qnorm(0.9, 100, 15)
qt(p=0.9, df=12)
# Generate N samples with 10 obs.
m = matrix(rnorm(10*N, 70, 10), N, 10)
sampleMeans = rowMeans(m)
# 1 Sample T-test
t.test(x=data, mu=38, alternative="less
    ", conf.level=0.95)
var.test(d1, d2) # Equal var. if >0.05
shapiro.test(d) # Test normality.
   Normal if >0.1
wilcox.test(d, 4) # If data not normal
anova(m1) # Test sig. of var. with >2
   categories. Add at end.
# 2 Sample, Independent
t.test(data1, data2, alternative="less"
    , var.equal=T, conf.level=0.95)
# 2 Sample, Dependent
t.test(d1 - d2, mu=0)
# Linear Regression
m1 = lm(price area) # y x
m2 = lm(price~area+type+area*type)
summary (m1)
confint(m1, level=0.95) # CI of coeff.
abline(m1) # Add fitted model to plot
# Predict
new1 = data.frame(area = c(20, 40))
predict(m1, new1)
predict(m1, new1, interval="confidence"
    , level=0.95) # with CI
rawRes = m1$res # Raw residuals
SR = rstandard(m1) # Standard residuals
which (SR > 3 | SR < -3) # Outliers
which (cooks.distance(m1)>1) #Influen.
```