

# Presentation on CSIR UGC NET (June 2015) Q. 106

Jatin Tarachandani

CS20BTECH11021

# Some Definitions

## Definition

The pmf for a binomial distribution with parameters  $n$  and  $p$  is

$$\Pr(X = x; n, p) = {}^nC_x p^x (1 - p)^{n-x} \quad (1)$$

## Definition

The pmf for a Poisson distribution with parameter  $\lambda$  is:

$$\Pr(Y = y; \lambda) = \frac{\lambda^y}{y!} e^{-\lambda} \quad (2)$$

# Unbiased Estimators

## Definition

Suppose we have an estimator  $T(X, Y)$ , acting on a set of 2 RVs  $X, Y$ ; which estimates a parameter  $\theta$ , then if

$$E(T(X, Y)) = \theta \quad (3)$$

the estimator is unbiased.

## Problem statement

Let  $X, Y$ , have the joint discrete distribution such that  $X|Y = y \sim \text{Binomial}(y, 0.5)$  and  $Y \sim \text{Poisson}(\lambda)$ ,  $\lambda > 0$ , where  $\lambda$  is an unknown parameter. Let  $T = T(X, Y)$  be any unbiased estimator of  $\lambda$ . Then

- ①  $\text{Var}(T) \leq \text{Var}(Y)$  for all  $\lambda$
- ②  $\text{Var}(T) \geq \text{Var}(Y)$  for all  $\lambda$
- ③  $\text{Var}(T) \geq \lambda$  for all  $\lambda$
- ④  $\text{Var}(T) = \text{Var}(Y)$  for all  $\lambda$

## Solution

We know that since  $Y$  has a Poisson distribution with parameter  $\lambda$ ,

$$\text{Var}(Y) = \lambda \quad (4)$$

Since  $X, Y$  are discrete, we define  $p_{XY}(X, Y)$  as the joint discrete distribution of  $X$  and  $Y$ . From Bayes Theorem,

$$p_{XY}(X, Y) = \Pr(X = x | Y = y) \Pr(Y = y) \quad (5)$$

$$= {}^y C_x \frac{1}{2^y} \frac{\lambda^y}{y!} e^{-\lambda} \quad (6)$$

# Cramer Rao lower bound

## Definition

If  $T(X, Y)$  is an unbiased estimator of a parameter  $\lambda$ , the Cramer-Rao bound states that:

$$\text{Var}(T(X, Y)) \geq -\frac{1}{E\left(\frac{\partial^2 \ln(p_{XY}(X, Y))}{\partial \lambda^2}\right)} \quad (7)$$

Since  $\lambda$  is a continuous parameter, we can apply this inequality to the logarithm of the joint discrete distribution seen in (6).

## Lemma

$$\frac{\partial^2 \ln(p_{XY}(X, Y))}{\partial \lambda^2} = -\frac{Y}{\lambda^2} \quad (8)$$

## Proof.

$$p_{XY}(X, Y) = {}^Y C_X \frac{1}{2^Y} \frac{\lambda^Y}{Y!} e^{-\lambda} \quad (9)$$

$$\frac{\partial^2 \ln(p_{XY}(X, Y))}{\partial \lambda^2} = \frac{\partial^2 \left( \ln {}^Y C_X - Y \ln 2^Y + Y \ln \lambda - \ln Y! - \lambda \right)}{\partial \lambda^2} \quad (10)$$

$$= \frac{\partial \left( \frac{Y}{\lambda} - 1 \right)}{\partial \lambda} \quad (11)$$

$$= -\frac{Y}{\lambda^2} \quad (12)$$

which is the required result. □

From using (7) on  $T(X, Y)$ ,

$$\text{Var}(T(X, Y)) \geq -\frac{1}{E\left(\frac{\partial^2 \ln(p_{XY}(X, Y))}{\partial \lambda^2}\right)} \quad (13)$$

From (8),

$$\text{Var}(T(X, Y)) \geq -\frac{1}{E\left(-\frac{Y}{\lambda^2}\right)} \quad (14)$$

$$\geq \frac{\lambda^2}{E(Y)} \quad (15)$$

$$\text{Var}(T(X, Y)) \geq \lambda \quad (16)$$

because the expectation value of a Poisson R.V with parameter  $\lambda$  is  $\lambda$ .



The correct options are options (2) and (3), since by (16), we see that

$$\text{Var}(T) \geq \lambda = \text{Var}(Y)$$

(from (4)). (1) and (4) do not hold for all  $\lambda$ .