

# Extending Correlation and Regression from Multivariate to Functional Data

By He, Müller & Wang

Presented by Javier Zapata

# Preliminaries

- Let  $X \in \mathbb{R}^p$  and  $Y \in \mathbb{R}^q$

A multivariate linear regression can be defined as

$$Y = \alpha + \beta_0^T X + \epsilon, \quad (2.1)$$

where  $\epsilon \in \mathbb{R}^q$ , with  $E[\epsilon] = 0$ , and  $\beta_0 \in \mathbb{R}^{p \times q}$  is the parameter matrix.

In multivariate analysis, we seek the solution of a linear regression model (2.1) by finding the parameter matrix  $\beta_0^* \in \mathbb{R}^{p \times q}$  which minimizes the squared distance  $E \|Y - \beta X\|^2$ . When the covariance matrix of  $X$  is invertible, by classical least squares theory (see, *e.g.*, Anderson, 1984), the unique minimizer can be found as

$$\beta_0^* = R_{XX}^{-1} R_{XY}. \quad (2.10)$$

## $L_2$ -processes

For a stochastic process with support  $T$ , on a probability space  $\Omega$ ,  $X(t) = \{X(t, \omega); \omega \in \Omega, t \in T\}$ , it holds that  $X \in L_2(T)$ , if  $E \int_T X^2(t) dt < \infty$ . For convenience, we will always assume that  $T$  and  $T_1, T_2$  below are compact intervals. We note that  $L_2(T)$  is a Hilbert space if equipped with the inner product  $\langle f, g \rangle = \int_T f(t)g(t)dt$ , for  $f, g \in L_2(T)$ , where  $dt$  is the Lebesgue measure. The results can be easily extended to cover more general measures  $\mu$  and scalar products in spaces  $L_2(T; \mu)$ .

# Covariance Operators and Functions

the infinite-dimensional case, covariance matrices are generalized to *covariance operators*. Specifically, the covariance operator  $R_{XX}: L_2(T_1) \rightarrow L_2(T_1)$ , is given by

$$R_{XX}u(s) = \int_{T_1} r_{XX}(s, t)u(t)dt, \quad u \in L_2(T_1). \quad (2.4)$$

where

$$r_{XX}(s, t) = \text{Cov}[X(s), X(t)], \quad s, t \in T_1,$$

is the covariance function of process  $X$ . Similarly we can define covariance functions

$$r_{YY}(s, t) = \text{cov}[Y(s), Y(t)], \quad s, t \in T_2,$$

$$\text{and } r_{XY}(s, t) = \text{Cov}[X(s), Y(t)], \quad s \in T_1, t \in T_2.$$

The covariance operators  $R_{YY}: L_2(T_2) \rightarrow L_2(T_2)$ ,  $R_{XY}: L_2(T_2) \rightarrow L_2(T_1)$ , and  $R_{YX}: L_2(T_1) \rightarrow L_2(T_2)$  are defined in complete analogy to  $R_{XX}$ . We

# Karhunen-Loève decomposition for $X(s)$ and $Y(t)$

Using the Karhunen-Loève decomposition,  $X$  and  $Y$  may be expanded as

$$\begin{aligned} X(s) &= E[X(s)] + \sum_{i=1}^{\infty} \xi_i \theta_i(s), \quad s \in T_1, \\ Y(t) &= E[Y(t)] + \sum_{i=1}^{\infty} \zeta_i \phi_i(t), \quad t \in T_2, \end{aligned} \tag{2.5}$$

with a sequence of uncorrelated random variables  $\xi_i$  with  $E(\xi_i) = 0$ , and a sequence of uncorrelated random variables  $\zeta_i$  with  $E(\zeta_i) = 0$ . Here,  $\lambda_{Xi} = E[\xi_i^2]$ ,  $\lambda_{Yi} = E[\zeta_i^2]$ ,  $\sum_{i=1}^{\infty} \lambda_{Xi} < \infty$ ,  $\sum_{i=1}^{\infty} \lambda_{Yi} < \infty$ , and  $\{(\lambda_i, \theta_i)\}$ ,  $\{(\zeta_j, \phi_j)\}$  are the eigenvalues and eigenfunctions of the covariance operators

- We can see that:

$$r_{XY}(s, t) = \text{Cov}(X(s), Y(t)) = \sum_{i,j \geq 1} E[\xi_i \zeta_j] \theta_i(s) \phi_j(t)$$

## The functional linear model for $L_2$ -processes

*Consider  $L_2$ -processes  $X \in L_2(T_1)$ ,  $Y \in L_2(T_2)$ . The functional linear regression model is defined as*

$$Y(t) = \alpha(t) + \int_{T_1} X(s)\beta_0(s, t)ds + \epsilon(t), \quad (2.7)$$

*where  $\beta_0 \in L_2(T_1 \times T_2)$  is a parameter function,  $\alpha \in L_2(T_2)$  is an intercept function, and  $\epsilon \in L_2(T_2)$  is a random error process, with the assumption that  $X$  and  $\epsilon$  are uncorrelated, and that  $E[\epsilon(t)] = 0$ , for all  $t$ .*

By assuming, without loss of generality, that  $EX(t) = 0$  and  $EY(s) = 0$ , for all  $t, s$ , one may simplify the linear model (2.7) to

$$Y(t) = \int_{T_1} X(s)\beta_0(s, t)ds + \epsilon(t). \quad (2.8)$$

## Integral Operator $\mathcal{L}_X$ and its adjoint $\mathcal{L}_X^*$

Define a random integral operator  $\mathcal{L}_X: L_2(T_1 \times T_2) \rightarrow L_2(T_2)$  by

$$(\mathcal{L}_X \beta)(t) = \int_{T_1} X(s) \beta(s, t) ds, \quad \text{for } \beta \in L_2(T_1 \times T_2).$$

It is easy to see that the adjoint operator of  $\mathcal{L}_X$  is  $\mathcal{L}_X^*: L_2(T_2) \rightarrow L_2(T_1 \times T_2)$ , defined by

$$(\mathcal{L}_X^* z)(s, t) = X(s) z(t), \quad \text{for all } z \in L_2(T_2).$$

Note that (2.8) can be rewritten as

$$Y(t) = (\mathcal{L}_X \beta_0)(t) + \epsilon(t). \tag{2.9}$$

# How to achieve a normal equation?

- Let  $X \in \mathbb{R}^p$  and  $Y \in \mathbb{R}^q$

A multivariate linear regression can be defined as

$$Y = \alpha + \beta_0^T X + \epsilon, \quad (2.1)$$

where  $\epsilon \in \mathbb{R}^q$ , with  $E[\epsilon] = 0$ , and  $\beta_0 \in \mathbb{R}^{p \times q}$  is the parameter matrix.

- For class LS, if the covariance matrix of  $X$  is invertible:

$$\beta_0^* = R_{XX}^{-1} R_{XY}. \quad (2.10)$$

- For the functional linear model, we would like an operator  $\Gamma_{XX}$  to obtain:

$$\beta_0^* = \Gamma_{XX}^{-1} r_{XY}$$



- The model:

$$Y(t) = (\mathcal{L}_X \beta_0)(t) + \epsilon(t). \quad (2.9)$$

For the functional linear model (2.9), we seek a parameter function  $\beta_0^*$  such that

$$\beta_0^* = \arg \min_{\beta \in L_2(T_1 \times T_2)} E \|Y - \mathcal{L}_X \beta\|^2. \quad (2.11)$$

PROPOSITION 4.1. *Let  $\beta_0$  be a solution of the linear regression model (2.9). Then*

$$\beta_0 \in \arg \min_{\beta \in L_2(T_1 \times T_2)} E \|Y - \mathcal{L}_X \beta\|^2.$$

- Apply  $E[\mathcal{L}_X^*(\cdot)]$  on both sides of (2.9) to get:

$$\Gamma_{XX} \beta = r_{XY} \quad \text{for } \beta \in L_2(T_1 \times T_2)$$

- What is  $\Gamma_{XX}$  ? Can we invert it?

# Integral Operator $\Gamma_{XX}$

Motivated by the form of the least squares solution (2.10) for the multivariate linear model, we define a linear integral operator  $\Gamma_{XX}: L_2(T_1 \times T_2) \rightarrow L_2(T_1 \times T_2)$  as

$$(\Gamma_{XX}\beta)(s, t) = \int_{T_1} r_{XX}(s, w)\beta(w, t)dw.$$

- Some properties of  $\Gamma_{XX}$ 
  - i.  $\Gamma_{XX} = E[\mathcal{L}_X^* \mathcal{L}_X]$
  - ii.  $\Gamma_{XX}$  is self-adjoint
  - iii.  $\Gamma_{XX}$  is nonnegative Hilbert-Schmidt operator (a class of compact operators)

**PROPOSITION 4.2.** *Let  $\beta \in L_2(T_1 \times T_2)$ . Then*

$$\beta \in \arg \min_{\beta \in L_2(T_1 \times T_2)} E\|Y - \mathcal{L}_X \beta\|^2 \text{ if and only if } P_{R(\Gamma_{XX})}\beta = P_{R(\Gamma_{XX})}\beta_0,$$

where  $P_{R(\Gamma_{XX})}$  is the projection from  $L_2(T_1 \times T_2)$  to  $R(\Gamma_{XX})$ .

# Functional normal equation

- The model:  $Y(t) = (\mathcal{L}_X \beta_0)(t) + \epsilon(t). \quad (2.9)$
- Apply  $E[\mathcal{L}_X^*(\cdot)]$  on both sides of (2.9) to get:

$$\Gamma_{XX} \beta = r_{XY} \quad \text{for } \beta \in L_2(T_1 \times T_2)$$

- What is  $\Gamma_{XX}$  ? Can we invert it?
- To obtain an inverse  $\Gamma_{XX}^{-1}$ , we impose further structure on the problem:

*Condition 2.2.  $L_2$ -processes  $X$  and  $Y$  with the expansion (2.5) satisfy*

$$\sum_{i,j=1}^{\infty} \frac{E^2[\xi_i \zeta_i]}{\lambda_{Xi}^2} < \infty. \quad (2.12)$$

- More details in Conway (1990), Corollary 5.4

# Main result

**THEOREM 4.3.** *Let  $X$  and  $Y$  be  $L_2$ -processes with the expansion (2.5) which satisfy Condition 2.2. Then,*

- (a)  $\beta_0^* = \Gamma_{XX}^{-1} r_{XY}$  exists and is the unique solution of (4.1) in  $R(\Gamma_{XX})$ ;
- (b)  $\beta_0^*$  has the representation

$$\beta_0^*(s, t) = \sum_{i,j=1}^{\infty} \frac{E[\xi_i \zeta_i]}{\lambda_{Xi}} \theta_i(s) \phi_j(t);$$

- (c) *The set of the solutions of (4.1) is:*

$$\beta_0^* + \ker(\Gamma_{XX}) := \{\beta_0^* + h | h \in \ker(\Gamma_{XX})\},$$

*where  $\ker(\Gamma_{XX})$  is the kernel space of  $\Gamma_{XX}$ , i.e.,  $\ker(\Gamma_{XX}) = \{h \in L(T_1 \times T_2): \Gamma_{XX}h = 0\}$ .*

# Main result

**THEOREM 4.4.** *Assume condition 2.4 holds for  $X$  and  $Y$ . Then*

$$\arg \min_{\beta} E \|Y - \mathcal{L}_X \beta\|^2 = \beta_0^* + \ker(\Gamma_{XX}),$$

*and this coincides with the set of solutions for the functional normal equation (4.1).*

# Bibliography

- Conway – A Course in Functional Analysis, 2<sup>nd</sup> edition , 1990
  - Chapter 5\* - The Diagonalization of Compact Self-Adjoint Operators
- He, Mueller & Wang - Extending correlation and regression from multivariate to functional data , 2000.