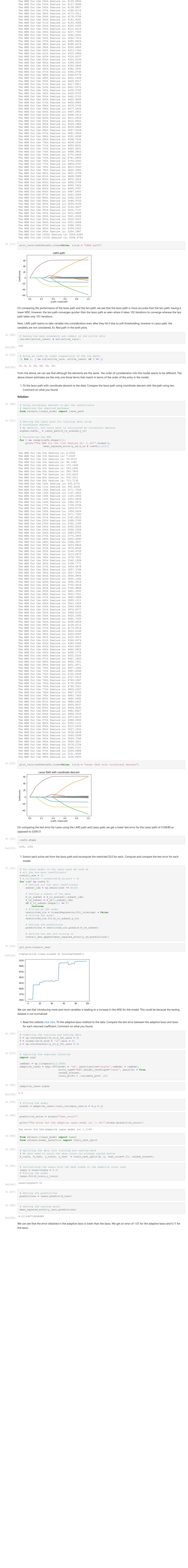
STA 208: Homework 1 (Do not distribute) Due 04/24/2022 midnight (11:59pm) **Instructions:** 1. Submit your homework using one file name "LastName FirstName hw1.html" on canvas. 2. The written portions can be either done in markdown and TeX in new cells or written by hand and scanned. Using TeX is strongly preferred. However, if you have scanned solutions for handwriting, you can submit a zip file. Please make sure your handwriting is clear and readable and your scanned files are displayed properly in your jupyter notebook. 3. Your code should be readable; writing a piece of code should be compared to writing a page of a book. Adopt the one-statement-perline rule. Consider splitting a lengthy statement into multiple lines to improve readability. (You will lose one point for each line that does not follow the one-statementper-line rule) 4. To help understand and maintain code, you should always add comments to explain your code. (homework with no comments will receive 0 points). For a very long comment, please break it into multiple lines. 5. In your Jupyter Notebook, put your answers in new cells after each exercise. You can make as many new cells as you like. Use code cells for code and Markdown cells for text. 6. Please make sure to print out the necessary results to avoid losing points. We should not run your code to figure out your answers. 7. However, also make sure we are able to open this notebook and run everything here by running the cells in sequence; in case that the TA wants to check the details. 8. You will be graded on correctness of your math, code efficiency and succinctness, and conclusions and modelling decisions Exercise 1 (Empirical risk minimization) (20 pts, 5 pts each) Consider Poisson model with rate parameter  $\lambda$  which has PMF,  $p(y|\lambda) = \frac{\lambda^y}{v!}e^{-\lambda},$ where y = 0, 1, ... is some count variable. In Poison regression, we model  $\lambda = e^{\beta^T x}$  to obtain  $p(y|x,\beta)$ . 1. Let the loss function for Poisson regression be  $\ell_i(\beta) \propto -\log p(y_i|x_i,\beta)$  for a dataset consisting of predictor variables and count values  $\{x_i, y_i\}_{i=1}^n$ . Here  $\propto$  means that we disregard any additive terms that are not dependent on  $\beta$ . Write an expression for  $\ell_i$  and derive its gradient. **Solution:** Based on the above expression, we can find the log density of the poisson distribution after substituting  $\lambda = e^{\beta^T x}$  back into the probability function. On doing so, we get the density function as:  $p(y|x,\beta) = \frac{(e^{\beta^T x})^y}{y!} e^{-e^{\beta^T x}}$ On taking the log of this, we get:  $\log p(y|x,\beta) = y\beta^{T}x - \log(y!) - e^{\beta^{T}x}$ Based on the definition of the  $\alpha$  proportionality, we just ignore the additive terms not dependent on  $\beta$  to get the loss function which is given as:  $l_i(\beta) = -\log p(y_i | x_i, \beta) = -(y_i \beta^T x_i - e^{\beta^T x_i}) = e^{\beta^T x_i} - y_i \beta^T x_i$ The gradient of this loss function is given as:  $\frac{\partial l_i(\beta)}{\partial \beta} = x_i (e^{\beta^T x_i} - y_i)$ 1. Show that the empirical risk  $R_n(\beta)$  is a convex function of  $\beta$ . **Solution:** The risk function  $R_n(\beta)$  is defined as summation of loss functions for each data point. It is formally defined as:  $R_n(\beta) = \frac{1}{n} \sum_{i=1}^{n} l_i(\beta)$ In order to show that the risk function is convex, we can use the fact that if loss function is convex, then the risk function is convex as risk function is a linear combination or sum of the loss functions at every 'i'. Now in order to show the loss function is convex, we have to show that the double derivative of the loss function is  $\geq 0$ . From part 1, we have the gradient as:  $\frac{\partial l_i(\beta)}{\partial \beta} = x_i (e^{\beta^T x_i} - y_i)$ Taking the double derivate of the above expression as:  $\frac{\partial^2 l_i(\beta)}{\partial R^2} = x_i^T x_i (e^{\beta^T x_i}) \ge 0$ Since for any values of  $x_i$ , the double derivative would be greater than 0, we can conclude that the loss function is convex. Since the loss function is convex and the risk function is the sum of loss functions, we can conclude that the risk function is convex as well. 1. Consider the mapping  $F_{\eta}(\beta) = \beta - \eta \nabla R_{\eta}(\beta)$  which is the iteration of gradient descent ( $\eta > 0$  is called the learning parameter). Show that at the minimizer of  $R_{n'}$ ,  $\hat{\beta}$ , we have that  $F(\hat{\beta}) = \hat{\beta}$ . **Solution:** In order to show that at the minimizer of  $R_{n'}\hat{\beta}$ , we have that  $F(\hat{\beta}) = \hat{\beta}$ , we have to show that the gradient of the risk function is zero at the minimizer of  $R_n$ .  $\frac{\partial R_n(\beta)}{\partial \beta} = \frac{\partial}{\partial \beta} \left( \frac{1}{n} \sum_{i=1}^n l_i(\beta) \right) = \frac{1}{n} \sum_{i=1}^n \frac{\partial}{\partial \beta} l_i(\beta)$  $= \frac{1}{n} \sum_{i=1}^{n} x_i (e^{\beta^T x_i} - y_i)$ To minimize this, we need to set the above equation to 0 in order to get the estimate  $\hat{\beta}$  ie.  $\frac{1}{n} \sum_{i=1}^{n} x_i (e^{\beta^T x_i} - y_i) = 0$  $e^{\hat{\beta}^T x_i} - y_i = 0$ Taking the log on both the sides we get  $\hat{\beta}^T x_i - \log y_i = 0$ Therefore, we have  $\hat{\beta}^T = \log y_i x_i^{-1}$ We know as a fact that since the  $R_n(\beta)$  is convex, the gradient of the risk function is zero at the minimizer of  $R_n$ . Therefore, we can conclude that the estimate  $\hat{\beta}$  is the minimizer of  $R_n$ . At the minimizer of  $R_{n'}$  we have that  $F(\hat{\beta}) = \hat{\beta}$ , as  $\nabla = 0 \mid_{\beta = \hat{\beta}}$  as we can see from the above calculations. 1. I have a script to simulate from this model below. Implement the gradient descent algorithm above and show that with enough data (n large enough) the estimated  $\hat{\beta}$  approaches the true  $\beta$  (you can look at the sum of square error between these two vectors). Solution: In [1]: # Importing the required libraries for question 1 import numpy as np import matplotlib.pyplot as plt In [2]: ## Simulate from the Poisson regression model (use y, X) np.random.seed(2022) n, p = 1000, 20X = np.random.normal(0,1,size = (n,p))beta = np.random.normal(0,.2,size = (p))lamb = np.exp(X @ beta) y = np.random.poisson(lamb) In [3]: # Defining gradient descent function def gradient descent(X, y, eta, maxiter = 10000): Summary: The gradient descent function takes in the X, y, eta and maxiter as inputs and returns the optimal beta. It initially starts from a random beta and tries to approximate F(beta) = beta hatArgs: X (np.array): n x p matrix of covariates y (np.array): n x 1 vector of response eta (np.float64): Learning rate maxiter (int, optional): Iterations for which the approximation is done. Defaults to 1000. # Defining the initial beta random beta = np.random.normal(0,0.2,size = (p))# Empty array to store the beta values betas = np.zeros((maxiter,20)) for i in range(maxiter): # Getting the predictions predictions = np.exp(X @ random\_beta) # Computing the gradient descent beta for the ith iteration random\_beta = random\_beta - eta \* 1/(X.shape[0]) \* X.T @ (predictions - y) # Storing the beta values betas[i,:] = random\_beta # returning the all obtained betas return betas In [4]: # Checking for various learning rates that might be suitable for the eta = [10 \*\* -x for x in range(1,6)]# Getting the mse for all etas overall mse = [] for e in eta: # Getting the estimates estimated beta = gradient descent(X, y, e) # Storing the mses at ebery iteration mse = []for estimate in estimated beta: mse.append(((estimate - beta) \*\* 2).mean()) # Appending the mse for the particular eta" overall mse.append(mse) In [5]: # Getting a plot for every eta value and seeing how quickly does # the beta converge for mse in overall mse: # Plotting the mse vs iteration plt.plot(range(10000), mse) plt.xlabel('Iterations') plt.ylabel('MSE') plt.legend(eta) <matplotlib.legend.Legend at 0x1e7ea3ee760> Out[5]: 0.10 0.1 0.01 0.001 0.08 0.0001 1e-05 0.06 MSE 0.04 0.02 0.00 2000 4000 6000 8000 10000 Iterations From the above, we can conclude that the gradient descent algorithm converges to the true  $\beta$  as the learning rate  $\eta$  is small enough. In this case, the best seems to be  $\eta = 0.001$  where it converges very smoothly between 0 and 2000 iterations. Exercise 2 (Regression and OLS) (35 pts, 5 pts each) Consider the regression setting in which  $x_i \in \mathbb{R}^p$  and  $y_i \in \mathbb{R}$ , for i = 1, ..., n and p < n. 1. For a given regressor, let  $\hat{y}_i$  be prediction given  $x_i$  and  $\hat{y}$  be the vector form. Show that linear regression can be written in the form  $\hat{y} = Hy$ , where H is dependent on X (the matrix of where each row is  $x_i$ ), assuming that p < n and X is full rank. Give an expression for H or an algorithm for computing H. **Solution:** Given,  $\hat{y} = Hy$ , in general for any p < n, we can write the linear regression as: The predictions  $\hat{y}$  are given by  $X\hat{\beta}$  where X is the matrix of where each row is  $x_i$  and  $\hat{\beta}$  is the vector of estimated coefficients.  $\hat{v} = X\hat{\beta}$ By the OLS solution we know that  $\hat{\beta} = (X^T X)^{-1} X^T y$ . Substituting the above expression for  $\hat{\beta}$  we get:  $\hat{y} = X\hat{\beta} = X(X^TX)^{-1}X^Ty$ To find the  $\hat{\beta}$ , we can use the OLS solution, where we try to minimize the following:  $argmin \|y - X\beta\|_2^2$  $= argmin(y - X\beta)^{T}(y - X\beta)$ =  $argmin(y^Ty - 2\beta^TX^Ty + \beta^TX^TX\beta)$ On taking the derivative of the above expression, we get:  $\hat{\beta} = (X^T X)^{-1} X^T y$ Correlating with the given expression  $\hat{y} = Hy$ , we get:  $H = X(X^TX)^{-1}X^T$ 1. Assuming p < n and X is full rank, let  $X = UDV^T$  be the thin singular value decomposition where U is  $n \times p$ , and V, D is  $p \times p$  (D is diagonal). • a) Derive an expression for the OLS coefficients  $\beta = Ab$  such that A is  $p \times p$  and depends on V and D, and b is a p vector and does not depend on D. **Solution:** We know that the OLS coefficients are given by:  $X\beta = y$  $UDV^T\beta = y$ Multiplying by U^T on both the sides and by property of SVD we know that  $UU^T = I$ , we get  $U^{T}(UDV^{T})\beta = (UU^{T})DV^{T}\beta = DV^{T}\beta = U^{T}v$ Since D is a singular matrix, we can take the inverse of D and multiply it by V to get the OLS coefficients.  $V^T \beta = D^{-1} U^T v$  $VV^T\beta = VD^{-1}U^Tv$  $\beta = VD^{-1}U^{T}v$ On comparing the obtained expression with the given expression  $\beta = Ab$ , we can say that,  $A_{n\times n} = VD^{-1} \setminus \& b = U^T y$ • b) Describe a fit method that precomputes these quantities separately **Solution:** Scipy's svd function can be used to compute the SVD of a matrix. In order to compute the above A and B we can write the following In [6]: # Defining the librarry for question 2 from scipy.linalg import svd In [7]: # Defining a function to calculate the A and B values def custom svd(X,y): # Getting the U,D,V values from the X matrix U,D,V t = svd(X,full matrices = False) # Getting inverse of D D inv = np.diag(1/D)# Getting the A and B values A = V t.T @ D invb = U.T @ y return A, b • c) Use the simulated data y and X in below to find  $\hat{\beta}$  using SVD. In [8]: ## Simulate from the linear regression model (use y,X) np.random.seed(2022) n, p = 100,20X = np.random.normal(0,1,size = (n,p))beta = np.random.normal(0,.2,size = (p))sigma = 1y = np.random.normal(X @ beta, sigma\*\*2) # Calculating the estimated beta using the above function A,b = custom svd(X,y)# beta is given by A @ b beta hat = A @ b # The estimated beta is beta\_hat array([-0.17529727, 0.2409376, -0.44251575, -0.03407534, 0.23607846, Out[9]: -0.2524819, 0.50945566, 0.03948573, 0.21520295, -0.03151863, 0.02374895, 0.13474846, -0.37652418, -0.01026382, -0.33508302, -0.04318508, 0.2321108, -0.20681122, -0.05872932, -0.24552763]) • d) Call a new data  $\tilde{X} \in \mathbb{R}^{m \times p}$ , derive an expression for the predicted y with  $\tilde{X}$  using SVD. **Solution:** We know that the predicted y is given by:  $\hat{y} = X\hat{\beta}$  where  $\hat{\beta} = (X^T X)^{-1} X^T y$ On substituting the SVD decomposition of  $\hat{\beta}$  we get:  $\hat{\mathbf{v}} = \tilde{\mathbf{X}}\tilde{\mathbf{V}}\tilde{\mathbf{D}}^{-1}\tilde{\mathbf{U}}^T\mathbf{v}$  $= \tilde{U}\tilde{D}\tilde{V}^T\tilde{V}\tilde{D}^{-1}\tilde{U}^Tv$ On simplifying the above expression we get,  $\hat{\mathbf{y}} = \tilde{\mathbf{U}}\tilde{\mathbf{U}}^T\mathbf{y}$ 1. Consider a regressor that performs OLS using the SVD above, but every instance of D will only use the largest r values on the diagonal, the remainder will be set to 0. Call this new  $p \times p$  matrix  $D_r$  (r < p). Then the new coefficient vector is the OLS computed as if the design matrix is modified by  $X \leftarrow UD_rV^T$ . • a) Given that you have computed b already, how could you make a method change\_rank that recomputes A with  $D_r$  instead of D? **Solution:** Using the same implementation as before, we can modify the singular value decomposition to only use the largest r values on the diagonal. The function change\_rank will be used to recompute the OLS coefficients with a lower rank using the same computation as estimate\_using\_svd function written in the previous part. In [10]: # Defining the change rank function def change rank(X,y,r = np.linalg.matrix rank(X)): # Getting SVD decompostion U,D,V t = svd(X,full matrices = False) # svd() returms the singular values as an array so we convert it to a diagonal matrix D inv = 1/Dif r != np.linalg.matrix rank(X): # getting r values and setting rest to 0 D inv = np.append(D inv[:r], np.zeros(len(D) - r)) D inv = np.diag(D inv) # Getting the A and B values A = V t.T @ D invb = U.T @ y # retunrning A and b return A, b • b) Choose r=10, recompute  $\hat{\beta}$  (call it  $\hat{\beta}_{\text{LowRank}}$ ) in Question 2-c. Solution: In [11]: # Getting A and b based on r = 10 $A,b = change_rank(X,y,r = 10)$ # Estimated beta is given by A @ b estimated beta = A @ b estimated beta array([ 0.03112502, -0.0430854 , -0.16394756, 0.03728632, 0.03755968, Out[11]: -0.04653204, -0.04193316, 0.13440484, -0.02375678, -0.02505815, 0.26185551, 0.02535527, -0.04744292, -0.01710232, -0.03823046, -0.01175641, 0.06146285, -0.16952398, -0.01660738, -0.10822824]) Exercise 3 (Subset selection) (15 pts) Recall the subset selection problem with tuning parameter k,  $\min_{\beta: \|\beta\|_0 \le k} \|y - X\beta\|_2^2,$ where  $\|\beta\|_0 = \#\{j = 1 ..., p : \beta_i \neq 0\}.$ Notice that we can write this as  $\min_{\beta: |\operatorname{supp}(\beta)| \le k} |y - X\beta|_2^2,$ where supp( $\beta$ ) = { $j = 1..., p : \beta_i \neq 0$ } (supp( $\beta$ ) is the support of  $\beta$ ). 1. (5 points) Write the subset selection problem in the following form  $\min_{S \subseteq \{1, \dots, p\}, |S| \le k} y^{\mathsf{T}} P_{S} y,$ where  $P_S$  is a projection. **Solution:** We know that the predictions are given by:  $\hat{y} = X\hat{\beta} = Hy$  where  $H = X(X^TX)^{-1}X^T$ In order to minimize the above problem, we need to find a  $\beta$  that would minimize the difference between y and  $x\beta$ , which happens to be  $\hat{\beta}$ . If we resubtitute  $\beta = \hat{\beta}$ , then we get the expression mentioned above hence the problem further converts into:  $\min_{S\subseteq\left\{1,\,\ldots,p\right\},\,\left|S\right|\leq k}\left\|y-\hat{y}\right\|_{2}^{2}=\min_{\left|S\subseteq\left\{1,\,\ldots,p\right\},\,\left|S\right|\leq k}\left\|y-H_{s}y\right\|_{2}^{2}$  $= \min_{S\subseteq \left\{1,\,\ldots,p\right\},\,\left|S\right|\leq k} \left\|y(I-H_s)\right\|_2^2$ Note that in the above expression, instead of the normal H which would consider all the predictors,  $H_s$  would be selecting all the predictors that are in the support of  $\beta$  which are in the set S. On expanding the expression, we get:  $\min_{S\subseteq \{1,\ldots,p\},\,|S|\leq k}y^T(I-H_s)^T(I-H_s)y$ Since  $H_s$  is a projection, we can also say that  $(I - H_s)$  is symmetric, hence we can write the above expression as:  $\min_{S\subseteq \{1, \dots, p\}, |S| \le k} y^T (I - H_s) y$ We can substitute the expression for  $I - H_s$  as  $P_s$  and we get:  $S \subseteq \left\{\,1\,,\, \ldots\,,p\,\right\}\,,\, |S| \leq k$ 1. (10 points) Suppose that we have a nested sequence of models  $S_1 \subset S_2 \subset ... \subset S_p$  such that  $|S_k| = k$  ( $|S_k|$  is the cardinality of  $S_{k'}$ meaning that it contains k variables). Prove that  $y^{\mathsf{T}} P_{S_{t}} y \ge y^{\mathsf{T}} P_{S_{t+1}} y$ for k = 1, ..., p - 1. What does this tell us about the solution to the subset selection problem and the constraint  $|S| \le k$ ? (Hint: using the fact that  $X^TX$  is positive definite, write  $X^TX = VDV^T$ ) Solution: Using the given inequality, we have:  $y^{T}P_{k}y \ge y^{T}P_{k+1}y \Longrightarrow y^{T}(I - H_{k})y \ge y^{T}(I - H_{k+1})y$  $y^{T}y - y^{T}H_{k}y \ge y^{T}y - H_{k+1}y$  $y^T H_k y \leq y^T H_{k+1} y$ Substituting the expression for  $H_k$  in its SVD form we get the following:  $y^{T}UDV^{T}VD_{k}^{-1}U^{T}y \leq y^{T}UDV^{T}VD_{k+1}^{-1}U^{T}y$  $y^{T}UDD_{k}^{-1}U^{T}y \le y^{T}UDD_{k+1}^{-1}U^{T}y$ Noticing that  $DD_k^{-1}$  is 'k' 1's on the diagonal and rest 0's, we notice that the expression is in the form of  $x^TAx$  which can be simplified into:  $\sum_{i=1}^{k} (\sum_{j=1}^{n} U_{ji} y_{j})^{2} \leq \sum_{i=1}^{k+1} (\sum_{j=1}^{n} U_{ji} y_{j})^{2}$ We can see the RHS is the summation of 1 extra term than on the LHS. Hence the inequality is satisfied. Exercise 4 (Ridge, lasso and adaptive lasso) (40 pts, 5 pts each) For this exercise, it may be helpful to use the sklearn.linear\_model module. I have also included a plotting tool for making the lasso path in ESL. 1. Load the training and test data using the script below. Fit OLS on the training dataset and compute the test error. Throughout you do not need to compute an intercept but you should normalize the X (divide by the column norms). **Solution:** In [12]: # Importing all the required packages for question 4 from sklearn.linear model import LinearRegression from sklearn.metrics import mean squared error from sklearn.linear model import lars path from sklearn.preprocessing import normalize In [13]: # Loading our data import pickle with open('hw1.data','rb') as f: y\_tr, X\_tr, y\_te, X\_te = pickle.load(f) In [14]: # Normalizing the data by dividing the column by its norm X tr scaled = normalize(X tr,axis = 0) In [15]: # Using the Linear Regression model from sci-kit learn ols = LinearRegression(fit intercept = False) # Fitting the model ols.fit(X\_tr\_scaled,y\_tr) LinearRegression(fit intercept=False) Out[15]: In [16]: # Getting the testing accuracy predictions = ols.predict(X te) # Getting the MSE ols sse = mean squared error(y te,predictions) print("The MSE for the OLS model is: {:.4f}".format(ols sse)) The MSE for the OLS model is: 5209.3163 1. Ridge regression: • a) Train and tune ridge regression using a validation set (choose LOOCV) and compute the test error (square error loss). **Solution:** In [17]: # Importing the required packages from sklearn.linear model import RidgeCV from sklearn.model selection import KFold In [18]: # RidgeCV by default uses the LOO cross validation # By giving it no alphas, it assumes default to be [0.1,1,10] loocv\_ridge = RidgeCV(store\_cv\_values = True, fit\_intercept=False) # Fitting the model loocv ridge.fit(X tr scaled,y tr) RidgeCV(alphas=array([ 0.1, 1. , 10. ]), fit\_intercept=False, Out[18]: store\_cv\_values=True) In [19]: # Getting the predictions loocv ridge predictions = loocv ridge.predict(X te) # Getting the SSE for the RidgeCV model loocv\_ridge\_sse = mean\_squared\_error(y\_te,loocv\_ridge\_predictions) print("The MSE for the RidgeCV model is: {:.4f}".format(loocv ridge sse)) The MSE for the RidgeCV model is: 3651.7683 • b) Repeat a) but using K-fold (you can choose K = 5 or 10) cross validation, compute the test error. Compare the result to a). Comment on what you found. **Solution:** In [20]: # RidgeCV by default uses the LOO cross validation but by # setting cv = 10, we get a 10-fold cross validation # By giving it no alphas, it assumes default to be [0.1,1,10]kfold ridge = RidgeCV(cv = 10, store cv values = False, fit\_intercept = False) # Fitting the model kfold\_ridge.fit(X\_tr\_scaled,y\_tr) RidgeCV(alphas=array([ 0.1, 1. , 10. ]), cv=10, fit\_intercept=False) Out[20]: In [21]: # Getting the predictions kfold predictions = kfold ridge.predict(X te) # Getting the SSE for the RidgeCV model kfold\_ridge\_sse = mean\_squared\_error(y\_te,kfold\_predictions) print("The MSE for the RidgeCV model is: {:.4f}".format(kfold ridge sse)) The MSE for the RidgeCV model is: 3651.7683 On noticing the MSE values we get, we see basically no improvement in the K-fold cross validation and the LOOCV. Unlike LOOCV, the kfold CV does not explore all of the possible splits of the original data: it randomly splits the data set into k equally sized partitions, keeping one as validation set and k - 1 as training set. However LOOCV is more suitable when the dataset is smaller 1. Fit the lasso path with lars to the data and compute the test error for each returned lasso coefficient. **Solution:** In [22]: # Default function for getting the plots def plot lars(coefs, lines=False, title="Lars Path"): Plot the lasso path where coefs is a matrix - the columns are beta vectors xx = np.sum(np.abs(coefs.T), axis=1)xx /= xx[-1]plt.plot(xx, coefs.T) ymin, ymax = plt.ylim() if lines: plt.vlines(xx, ymin, ymax, linestyle='dashed') plt.xlabel('|coef| / max|coef|') plt.ylabel('Coefficients') plt.title(title) plt.axis('tight') In [23]: # Getting the lasso path for training data using the 'lars path' # function by setting the method as 'lasso' # The lars function returns the alphas , the coefficients and the # active set alphas,active lasso,coefs = lars path(X tr scaled, y tr, method = 'lasso') In [24]: # Calculating the test error for the lasso path for i in range(coefs.shape[1]): mse = mean squared error(y te, X te @ coefs[:,i]) print("The MSE for the {}th feature is: {:.4f}".format(i, mse)) The MSE for the 0th feature is: 6.0565 The MSE for the 1th feature is: 828.3130 The MSE for the 2th feature is: 2103.2714 The MSE for the 3th feature is: 2733.1611 The MSE for the 4th feature is: 3641.9393 The MSE for the 5th feature is: 3675.6728 The MSE for the 6th feature is: 3835.6989 The MSE for the 7th feature is: 3870.5148 The MSE for the 8th feature is: 3961.7864 The MSE for the 9th feature is: 3970.7311 The MSE for the 10th feature is: 3972.2181 The MSE for the 11th feature is: 3973.3568 The MSE for the 12th feature is: 3975.9204 The MSE for the 13th feature is: 3976.9031 The MSE for the 14th feature is: 3991.6420 The MSE for the 15th feature is: 4006.7771 The MSE for the 16th feature is: 4016.6011 The MSE for the 17th feature is: 4028.9334 The MSE for the 18th feature is: 4072.1554 The MSE for the 19th feature is: 4091.7907 The MSE for the 20th feature is: 4107.4824 The MSE for the 21th feature is: 4110.1380 The MSE for the 22th feature is: 4112.6543 The MSE for the 23th feature is: 4117.3531 The MSE for the 24th feature is: 4135.6905 The MSE for the 25th feature is: 4137.9284 The MSE for the 26th feature is: 4138.9837 The MSE for the 27th feature is: 4166.8116 The MSE for the 28th feature is: 4173.9311 The MSE for the 29th feature is: 4178.9758 The MSE for the 30th feature is: 4181.9081 The MSE for the 31th feature is: 4193.3584 The MSE for the 32th feature is: 4205.5185 The MSE for the 33th feature is: 4210.4216 The MSE for the 34th feature is: 4237.7393 The MSE for the 35th feature is: 4262.0260 The MSE for the 36th feature is: 4264.7013 The MSE for the 37th feature is: 4285.8409 The MSE for the 38th feature is: 4288.4476 The MSE for the 39th feature is: 4300.8689 The MSE for the 40th feature is: 4313.1343 The MSE for the 41th feature is: 4315.4986 The MSE for the 42th feature is: 4316.6237 The MSE for the 43th feature is: 4322.9194 The MSE for the 44th feature is: 4348.0643 The MSE for the 45th feature is: 4350.0849 The MSE for the 46th feature is: 4382.3931 The MSE for the 47th feature is: 4392.5744 The MSE for the 48th feature is: 4399.8778 The MSE for the 49th feature is: 4402.1606 The MSE for the 50th feature is: 4405.8317 The MSE for the 51th feature is: 4417.8611 The MSE for the 52th feature is: 4421.5575 The MSE for the 53th feature is: 4439.5782 The MSE for the 54th feature is: 4439.8755 The MSE for the 55th feature is: 4442.2722 The MSE for the 56th feature is: 4455.9866 The MSE for the 57th feature is: 4456.8983 The MSE for the 58th feature is: 4476.3749 The MSE for the 59th feature is: 4477.0502 The MSE for the 60th feature is: 4497.0902 The MSE for the 61th feature is: 4498.0514 The MSE for the 62th feature is: 4512.2826 The MSE for the 63th feature is: 4512.8922 The MSE for the 64th feature is: 4529.1880 The MSE for the 65th feature is: 4539.0866 The MSE for the 66th feature is: 4557.6328 The MSE for the 67th feature is: 4581.3009 The MSE for the 68th feature is: 4591.6098 The MSE for the 69th feature is: 4598.7044 The MSE for the 70th feature is: 4628.6873 The MSE for the 71th feature is: 4639.0437 The MSE for the 72th feature is: 4650.1517 The MSE for the 73th feature is: 4650.4156 The MSE for the 74th feature is: 4718.3010 The MSE for the 75th feature is: 4723.5013 The MSE for the 76th feature is: 4738.8062 The MSE for the 77th feature is: 4752.7452 The MSE for the 78th feature is: 4754.3405 The MSE for the 79th feature is: 4757.4373 The MSE for the 80th feature is: 4765.6359 The MSE for the 81th feature is: 4770.6977 The MSE for the 82th feature is: 4779.1586 The MSE for the 83th feature is: 4798.8972 The MSE for the 84th feature is: 4803.6949 The MSE for the 85th feature is: 4811.2433 The MSE for the 86th feature is: 4837.1730 The MSE for the 87th feature is: 4844.6195 The MSE for the 88th feature is: 4859.6660 The MSE for the 89th feature is: 4937.7018 The MSE for the 90th feature is: 4946.9575 The MSE for the 91th feature is: 4952.4210 The MSE for the 92th feature is: 4959.1634 The MSE for the 93th feature is: 5023.5971 The MSE for the 94th feature is: 5042.9633 The MSE for the 95th feature is: 5100.8378 The MSE for the 96th feature is: 5113.7773 The MSE for the 97th feature is: 5120.2302 The MSE for the 98th feature is: 5132.3888 The MSE for the 99th feature is: 5174.5057 The MSE for the 100th feature is: 5183.8481 The MSE for the 101th feature is: 5193.0188 The MSE for the 102th feature is: 5209.3163 One of the interesting thing to note is that the MSE eventually converges to the MSE obtained in the normal OLS case In [25]: # plotting for Lasso path obtained using LARS plot\_lars(coefs=coefs, lines=False, title = "Lasso Path with LARS") Lasso Path with LARS 60 40 20 Coefficients -20-400.0 0.2 0.8 0.6 |coef| / max|coef| 1. Perform 3 without the lasso modification generating the lars path. Compare and contrast the lars path to the lasso path, what is the key difference. Tell me when the active sets differ and how, if they do at all. **Solution:** In [26]: # Getting the lars path for training data alphas,active lars,coefs = lars path(X tr scaled,y tr, method = 'lars') # Calculating the MSE for i in range(coefs.shape[1]): print("The MSE for the {}th feature is: {:.4f}".format(i, mean squared error(y te, X te @ coefs[:,i])))



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