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GATE-EC2023

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Q65ST.2023:Let $\{X_n\}_{n\geq 1}$ be a sequence of independent and identically distributed random variables with mean 0 and variance 1, all of them defined on the same probability space. For n=1,2,3,..., let

$$Y_n = \frac{1}{n}(X_1X_2 + X_3X_4 + \dots + X_{2n-1}X_{2n})$$
 (1)

Then which of the following statements is/are true?

- (A) $\{\sqrt{n}Y_n\}_{n\geq 1}$ converges in distribution to a standard normal random variable.
- (B) $\{Y_n\}_{n\geq 1}$ converges in 2nd mean to 0.
- (C) $\{Y_n + \frac{1}{n}\}_{n \ge 1}$ converges in probability to 0.
- (D) $\{X_n\}_{n\geq 1}$ converges almost surely to 0.

Solution: As X_i is a sequence of independent and identically distributed random variables,

Since X_1, X_2, X_3, X_4 are independent R.V,

$$Pr(X_1X_2X_3X_4) = Pr(X_1) Pr(X_2) Pr(X_3) Pr(X_4)$$
 (2)

$$= \Pr(X_1 X_2) \Pr(X_3 X_4) \tag{3}$$

$$= \Pr((X_1 X_2) (X_3 X_4)) \tag{4}$$

 $X_i X_{i+1}$ is a independent R.V

Also, X_1, X_2, X_3, X_4 are identical distributed R.V Then, X_1X_2 and X_3X_4 are also identical distributed R.V

Thus, $X_i X_{i+1}$ is a identical distributed R.V. Hence $\sum_{i=1}^{2n-1} X_i X_{i+1}$ is iid.

$$E(X_i X_{i+1}) = E(X_i) E(X_{i+1}) = 0$$
 (5)

$$\implies E(Y_n) = 0$$
 (6)

$$Var(X_{i}X_{i+1}) = E(X_{i}^{2}X_{i+1}^{2}) - [E(X_{i}X_{i+1})]^{2} = 1$$
 (7)

$$\implies Var(Y_n) = \frac{1}{n}$$
 (8)

(9)

(A) Using (6) and (8),

$$E(\sqrt{n}Y_n) = 0 \tag{10}$$

$$Var(\sqrt{n}Y_n) = \frac{1}{\sqrt{n}}$$
 (11)

For some $\{Z_n\}_{n\geq 1}$ converges in distribution to $Z, Z_n \xrightarrow{d} Z$, then for all z,

$$\lim_{n\to\infty} F_{Z_n}(x) = F_Z(x) \tag{12}$$

where, $Z \sim \mathcal{N}(0, 1)$

$$\lim_{n \to \infty} F_{\sqrt{n}Y_n} = \lim_{n \to \infty} \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi \left(\frac{1}{\sqrt{n}}\right)}} e^{-\frac{(x-0)^2}{2\left(\frac{1}{\sqrt{n}}\right)}} dx$$

$$= \lim_{n \to \infty} \int_{-\infty}^{x} \frac{\sqrt[4]{n}}{\sqrt{2\pi}} e^{-\frac{\left(\sqrt[4]{n}x\right)^2}{2}} dx$$

$$= \lim_{n \to \infty} \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} e^{-\frac{(x)^2}{2}} dx$$

$$= F_{Z}(x)$$
(15)
$$= F_{Z}(x)$$
(16)

where, $Z \sim \mathcal{N}(0, 1)$

So, option (A) is correct

(B) For 2nd mean to be converging to 0,

$$\lim_{n\to\infty} E(|Y_n - 0|^2) = \lim_{n\to\infty} E(Y_n^2)$$
 (17)

$$\therefore \frac{1}{n} = E(Y_n^2) - [E(Y_n)]^2$$
 (18)

$$= \lim_{n\to\infty} \frac{1}{n} + [E(Y_n)]^2$$
 (19)

$$\therefore [E(Y_n)]^2 = 0$$
 (20)

$$= \lim_{n\to\infty} \frac{1}{n} = 0$$
 (21)

Thus, $\{Y_n\}_{n\geq 1}$ converges in 2nd mean to 0 Hence, option (B) is correct.

(C) For $\{Y_n + \frac{1}{n}\}_{n \ge 1}$ to be converging $Y_n + \frac{1}{n} \xrightarrow{p} 0$

$$E(Y_n + \frac{1}{n}) = \frac{1}{n}$$
 (22)

$$Var(Y_n + \frac{1}{n}) = \frac{1}{n} \tag{23}$$

$$\lim_{n \to \infty} F_{Y_{n} + \frac{1}{n}} = \lim_{n \to \infty} \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi \left(\frac{1}{n}\right)}} e^{-\frac{\left(x - \frac{1}{n}\right)^{2}}{2\left(\frac{1}{n}\right)}} dx$$

$$= \lim_{n \to \infty} \int_{-\infty}^{x} \frac{n}{\sqrt{2\pi}} e^{-\frac{\left(n^{2}x - n\right)^{2}}{2}} dx$$

$$= 0$$
(24)
$$= 0$$
(25)

Thus,
$$Y_n + \frac{1}{n} \xrightarrow{d} 0 \implies Y_n + \frac{1}{n} \xrightarrow{p} 0$$

Note: This condition is only true for converges

at real constants.

Hence, option (C) is correct.

(D) let $\{X_n\}_{n\geq 1}$ almost surely converges to 0. Then, $X_n \xrightarrow{p} 1$.

$$\lim_{n\to\infty} F_{X_n} = \lim_{n\to\infty} \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi(1)}} e^{\frac{-(x-0)^2}{2(1)^2}} dx$$
(27)

$$=F_X(x) \tag{28}$$

$$X_n \xrightarrow{d} X$$
 (29)

$$\implies X_n \stackrel{p}{\to} X \tag{30}$$

Hence, option (D) is incorrect