

GATE-EC2023

Jay Vikrant EE22BTECH11025

Q65ST.2023: Let $\{X_n\}_{n \geq 1}$ be a sequence of independent and identically distributed random variables with mean 0 and variance 1, all of them defined on the same probability space. For $n=1,2,3,\dots$, let

$$Y_n = \frac{1}{n}(X_1X_2 + X_3X_4 + \dots + X_{2n-1}X_{2n}) \quad (1)$$

Then which of the following statements is/are true?

- (A) $\{\sqrt{n}Y_n\}_{n \geq 1}$ converges in distribution to a standard normal random variable.
- (B) $\{Y_n\}_{n \geq 1}$ converges in 2nd mean to 0.
- (C) $\{Y_n + \frac{1}{n}\}_{n \geq 1}$ converges in probability to 0.
- (D) $\{X_n\}_{n \geq 1}$ converges almost surely to 0.

Solution: As X_i is a sequence of independent and identically distributed random variables,

Since X_1, X_2, X_3, X_4 are independent R.V ,

$$P_{X_1X_2X_3X_4}(x) = p_{X_1}(x)p_{X_2}(x)p_{X_3}(x)p_{X_4}(x) \quad (2)$$

$$= p_{X_1X_2}(x)p_{X_3X_4}(x) \quad (3)$$

$$= p_{(X_1X_2)(X_3X_4)}(x) \quad (4)$$

X_iX_{i+1} is a independent R.V

Also, X_1, X_2, X_3, X_4 are identical distributed R.V

Then, X_1X_2 and X_3X_4 are also identical distributed R.V

Thus, X_iX_{i+1} is a identical distributed R.V.

Hence $\sum_{i=1}^{2n-1} X_iX_{i+1}$ is iid.

Now, let $Z_i = X_iX_{i+1}$ then,

$$E(Z_i) = E(X_i)E(X_{i+1}) = 0 \quad (5)$$

$$Var(Z_i) = E(Z_i^2) - [E(Z_i)]^2 = 1 \quad (6)$$

$$\implies Z \sim \mathcal{N}(0, 1) \quad (7)$$

$$Y_n = \frac{1}{n} \left(\sum_{i=1}^{2n-1} Z_i \right) \quad (8)$$

$$\implies E(Y_n) = 0 \quad (9)$$

$$\implies Var(Y_n) = \frac{1}{n} \quad (10)$$

$$Y \sim \mathcal{N}\left(0, \frac{1}{\sqrt{n}}\right) \quad (11)$$

(A) Using (11),

$$\sqrt{n}Y \sim \mathcal{N}(0, 1) \quad (12)$$

For some $\{P_n\}_{n \geq 1}$ converges in distribution to P , $P_n \xrightarrow{d} P$, then for all p,

$$\lim_{n \rightarrow \infty} F_{P_n}(x) = F_P(x) \quad (13)$$

where, $P \sim \mathcal{N}(0, 1)$

$$\lim_{n \rightarrow \infty} F_{\sqrt{n}Y_n} = \lim_{n \rightarrow \infty} \int_{-\infty}^x \frac{1}{\sqrt{2\pi}(1)} e^{-\frac{(x-0)^2}{2(1)^2}} dx \quad (14)$$

$$= \lim_{n \rightarrow \infty} \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{(x)^2}{2}} dx \quad (15)$$

$$= F_P(x) \quad (16)$$

where, $P \sim \mathcal{N}(0, 1)$

So, option (A) is correct

(B) For 2nd mean to be converging to 0,

$$\lim_{n \rightarrow \infty} E(|Y_n - 0|^2) = \lim_{n \rightarrow \infty} E(Y_n^2) \quad (17)$$

$$\because \frac{1}{n} = E(Y_n^2) - [E(Y_n)]^2 \quad (18)$$

$$= \lim_{n \rightarrow \infty} \frac{1}{n} + [E(Y_n)]^2 \quad (19)$$

$$\because [E(Y_n)]^2 = 0 \quad (20)$$

$$= \lim_{n \rightarrow \infty} \frac{1}{n} = 0 \quad (21)$$

Thus, $\{Y_n\}_{n \geq 1}$ converges in 2nd mean to 0

Hence, option (B) is correct.

(C) For $\{Y_n + \frac{1}{n}\}_{n \geq 1}$ to be converging $Y_n + \frac{1}{n} \xrightarrow{p} 0$ Using (11) ,

$$Y + \frac{1}{n} \sim \mathcal{N}\left(\frac{1}{n}, \frac{1}{\sqrt{n}}\right) \quad (22)$$

$$\lim_{n \rightarrow \infty} F_{Y_n + \frac{1}{n}} = \lim_{n \rightarrow \infty} \int_{-\infty}^x \frac{1}{\sqrt{2\pi}\left(\frac{1}{n}\right)} e^{-\frac{(x-\frac{1}{n})^2}{2\left(\frac{1}{n}\right)^2}} dx \quad (23)$$

$$= \lim_{n \rightarrow \infty} \int_{-\infty}^x \frac{n}{\sqrt{2\pi}} e^{-\frac{(n^2x-n)^2}{2}} dx \quad (24)$$

$$= 0 \quad (25)$$

Thus, $Y_n + \frac{1}{n} \xrightarrow{d} 0 \implies Y_n + \frac{1}{n} \xrightarrow{p} 0$

Note: This condition is only true for converges at real constants.

Hence, option (C) is correct.

(D) let $\{X_n\}_{n \geq 1}$ almost surely converges to 0. Then,
 $X_n \xrightarrow{p} 1$.

$$\lim_{n \rightarrow \infty} F_{X_n} = \lim_{n \rightarrow \infty} \int_{-\infty}^x \frac{1}{\sqrt{2\pi}(1)} e^{\frac{-(x-0)^2}{2(1)^2}} dx \quad (26)$$

$$= F_X(x) \quad (27)$$

$$X_n \xrightarrow{d} X \quad (28)$$

$$\implies X_n \xrightarrow{p} X \quad (29)$$

Hence, option (D) is incorrect