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GATE-EC2023

Jay Vikrant EE22BTECH11025

Q65ST.2023:Let $\{X_n\}_{n\geq 1}$ be a sequence of independent and identically distributed random variables with mean 0 and variance 1, all of them defined on the same probability space. For n=1,2,3,..., let

$$Y_n = \frac{1}{n}(X_1X_2 + X_3X_4 + \dots + X_{2n-1}X_{2n})$$
 (1)

Then which of the following statements is/are true?

- (A) $\{\sqrt{n}Y_n\}_{n\geq 1}$ converges in distribution to a standard normal random variable.
- (B) $\{Y_n\}_{n\geq 1}$ converges in 2nd mean to 0.
- (C) $\{Y_n + \frac{1}{n}\}_{n \ge 1}$ converges in probability to 0.
- (D) $\{X_n\}_{n\geq 1}$ converges almost surely to 0.

Solution:

1) X_n converges in distribution to a standard normal random variable X, $X_n \stackrel{d}{\rightarrow} X$, then for all X,

$$\lim_{n \to \infty} F_{X_n}(x) = F_X(x) \tag{2}$$

where, $X \sim \mathcal{N}(0, 1)$

2) X_n converges in p^{th} mean to X, then we have

$$\lim_{n \to \infty} E(|X_n - X|^p) = 0 \tag{3}$$

3) X_n converges in probability to X, $X_n \stackrel{p}{\to} X$, then for all $\epsilon > 0$,

$$\lim_{n \to \infty} \Pr(|X_n - X| \ge \epsilon) = 0 \tag{4}$$

4) X_n converges almost surely to X, $X_n \xrightarrow{a.s} X$, then for all $\epsilon > 0$,

$$\lim_{n \to \infty} \Pr(|X_n - X| < \epsilon) = 1 \tag{5}$$

$$E(X) = \int_{-\infty}^{\infty} x F_X x dx \tag{6}$$

$$\therefore X \text{ is positive-valued}$$
 (7)

$$= \int_0^\infty x F_X x dx \tag{8}$$

$$\geq \int_{\epsilon^2}^{\infty} x F_X x dx \tag{9}$$

for some
$$\epsilon > 0$$
 (10)

$$\geq \int_{c^2}^{\infty} k F_X x dx \tag{11}$$

$$= \epsilon^2 \int_{\epsilon^2}^{\infty} F_X x dx \tag{12}$$

$$E(X) = \epsilon^2 \Pr(X \ge \epsilon^2)$$
 (13)

This inequality is called Markov inequality Now, we take

$$\sigma^2 = E(X^2) - [E(X)]^2 \tag{14}$$

$$= E(X^2) - \mu^2 \tag{15}$$

$$= E(X^2) - E(\mu^2)$$
 (16)

$$\sigma^2 = \mathcal{E}(X^2 - \mu^2) \tag{17}$$

using (13),

$$\sigma^2 = \epsilon^2 \Pr\left(X^2 - \mu^2 \ge \epsilon^2\right) \tag{18}$$

$$\frac{\sigma^2}{\epsilon^2} = \Pr\left(X^2 - \mu^2 \ge \epsilon^2\right) \tag{19}$$

$$\Pr(|X - \mu| \ge \epsilon) = \frac{\sigma^2}{\epsilon^2}$$
 (20)

This above inequality is called Chebyshev's inequality

Now, as X_i is a sequence of independent and identically distributed random variables,

Let $Z_i = X_{2i-1}X_{2i}$.

$$p_{Z_i Z_i}(x) = p_{X_{2i-1} X_{2i} X_{2i-1} X_{2i}}(x)$$
 (21)

$$= p_{X_{2i-1}}(x)p_{X_{2i}}(x)p_{X_{2i-1}}(x)p_{X_{2i}}(x) \qquad (22)$$

$$= p_{X_{2i-1}X_{2i}}(x)p_{X_{2i-1}X_{2i}}(x)$$
 (23)

$$= p_{Z_i}(x)p_{Z_i}(x) \tag{24}$$

 Z_i is a independent R.V. Similarly,

$$F_{X_{2i-1}}(x) = F_{X_{2i}}(x) = F_{X_{2i-1}}(x) = F_{X_{2i}}(x)$$
 (25)

$$F_{X_{2i-1}}(x)F_{X_{2i}}(x) = F_{X_{2i-1}}(x)F_{X_{2i}}(x)$$
 (26)

$$F_{Z_i}(x) = F_{Z_i}(x) \qquad (27)$$

Thus, Z_i is an identical distributed R.V. Hence,

$$Z_n = \sum_{i=1}^n Z_i \tag{28}$$

is an i.i.d

$$E(Z_i) = E(X_{2i-1}X_{2i})$$
 (29)

$$= E(X_{2i-1}) E(X_{2i}) = 0$$
 (30)

$$Var(Z_i) = E(Z_i^2) - [E(Z_i)]^2$$
 (31)

$$= E(Z_i^2) - 0 (32)$$

$$= E(X_{2i-1}^2 X_{2i}^2) \tag{33}$$

$$= E(X_{2i-1}^2)E(X_{2i}^2) = 1$$
 (34)

Now,

$$Y_n = \frac{1}{n} \left(\sum_{i=1}^n Z_i \right) \tag{35}$$

$$E(Y_n) = E\left(\frac{1}{n}\sum_{i=1}^n Z_i\right)$$
 (36)

$$= \frac{1}{n} \sum_{i=1}^{n} E(Z_i) = 0$$
 (37)

As all the variables are i.i.d's and are thus uncorrelated,

$$\operatorname{cov}(X_{i}, X_{j}) = \begin{cases} 0 & \text{if } i \neq j \\ \operatorname{Var}(X_{i}) & \text{if } i = j \end{cases}$$
 (38)

$$Var(Y_n) = \frac{1}{n^2} \left(\sum_{i=1}^n \text{cov}(Z_i, Z_i) \right)$$
 (39)

$$= \frac{1}{n^2} \left(\sum_{i=1}^n \operatorname{var}(Z_i) \right) \tag{40}$$

$$= \frac{1}{n^2} \left(\sum_{i=1}^n 1 \right) \tag{41}$$

$$=\frac{1}{n}\tag{42}$$

(A) by central limit theorem,

$$\frac{Y_n - \mu}{\sigma} \sim \mathcal{N}(0, 1) \tag{43}$$

$$= \frac{Y_n - 0}{\sqrt{\frac{1}{n}}} \sim \mathcal{N}(0, 1) \tag{44}$$

$$= \sqrt{n}Y_n \sim \mathcal{N}(0,1) \tag{45}$$

Thus,

$$\lim_{n \to \infty} F_{\sqrt{n}Y_n} = F_P(x) \tag{46}$$

where, $P \sim \mathcal{N}(0, 1)$

So, option (A) is correct

(B) For 2nd mean to be converging to 0,

$$\lim_{n \to \infty} E(|Y_n - 0|^2) = \lim_{n \to \infty} E(Y_n^2)$$
 (47)

$$\therefore \frac{1}{n} = E(Y_n^2) - [E(Y_n)]^2$$
 (48)

$$= \lim_{n \to \infty} \frac{1}{n} + [E(Y_n)]^2 \tag{49}$$

$$: [E(Y_n)]^2 = 0 (50)$$

$$=\lim_{n\to\infty}\frac{1}{n}=0\tag{51}$$

Thus, $\{Y_n\}_{n\geq 1}$ converges in 2nd mean to 0 Hence, option (B) is correct.

(C) For $\{Y_n + \frac{1}{n}\}_{n \ge 1}$ to be converging $Y_n + \frac{1}{n} \xrightarrow{p} 0$

$$\lim_{n \to \infty} \Pr\left(\left|Y_n + \frac{1}{n} - Y\right| \ge \epsilon\right) = \lim_{n \to \infty} \Pr\left(\left|Y_n + \frac{1}{n}\right| \ge \epsilon\right)$$
(52)

$$\leq \lim_{n \to \infty} \frac{\operatorname{Var}(Y_n + \frac{1}{n})}{\epsilon^2}$$
(53)

$$\forall \epsilon > 0 \tag{54}$$

$$= \lim_{n \to \infty} \frac{\operatorname{Var}(Y_n) + \operatorname{Var}(\frac{1}{n})}{\epsilon^2}$$
(55)

$$=\lim_{n\to\infty}\frac{\frac{1}{n}+0}{\epsilon^2}\tag{56}$$

$$=0 (57)$$

hence, option (C) is correct

(D) As all the variables are i.i.d's and are thus uncorrelated,

$$\operatorname{Var}(X_n) = \sum_{i=1}^n \operatorname{cov}(X_i, X_i)$$
 (58)

$$= \sum_{i=1}^{n} \operatorname{var}(X_i) \tag{59}$$

$$=\sum_{i=1}^{n} 1 \tag{60}$$

$$= n \tag{61}$$

$$\lim_{n \to \infty} \Pr(|X_n - X| < \epsilon) = 1 - \lim_{n \to \infty} \Pr(|X_n - 0| \ge \epsilon)$$
(62)

$$\leq 1 - \lim_{n \to \infty} \frac{Var(X_n)}{\epsilon^2}$$
 (63)
$$\forall \epsilon > 0$$
 (64)

$$\forall \epsilon > 0 \tag{64}$$

$$= \lim_{n \to \infty} 1 - \frac{Var(X_n)}{\epsilon^2}$$
 (65)
$$= \lim_{n \to \infty} 1 - \frac{n}{\epsilon^2} = -\infty$$
 (66)

$$= \lim_{n \to \infty} 1 - \frac{n}{\epsilon^2} = -\infty \tag{66}$$

Hence, option (D) is incorrect