

# STAT572 - Homework Assignment 9

Juan Carlos Apitz, ID 012523821

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# Bayesian Inference with MCMC vs. Random Walk Sampler

In this exercise we approximate the posterior distribution of the parameter  $\theta$ , which represents a hypothetical proportion of defective items in a sample of size  $n = 100$ , generated by a Bernoulli process with  $p = 0.20$ .

The left panel represents the distribution of the Markov Chain generated using Bayesian inference with MCMC. We use a  $\text{Unif}(0, 1)$  as the prior and to calculate the transition probability use the likelihood function:

$$L(\theta, X) = \theta^{\sum x_i} (1 - \theta)^{n - \sum x_i}$$

The right panel represents the Markov Chain generated using the Metropolis Random Walk sampler. In this case, the target distribution is proportional to the:

$$\text{Beta}(\sum x_i + 1, n - \sum x_i + 1)$$

We sample using the random walk:

$$Y = X_t + \frac{Z_t}{\sqrt{12}}$$

where

$$Z_t = a + (b - a)u(0, 1)$$

and  $a = -0.5$ ,  $b = 0.5$ . The result is that both processes generate random samples that follow a Beta distribution.

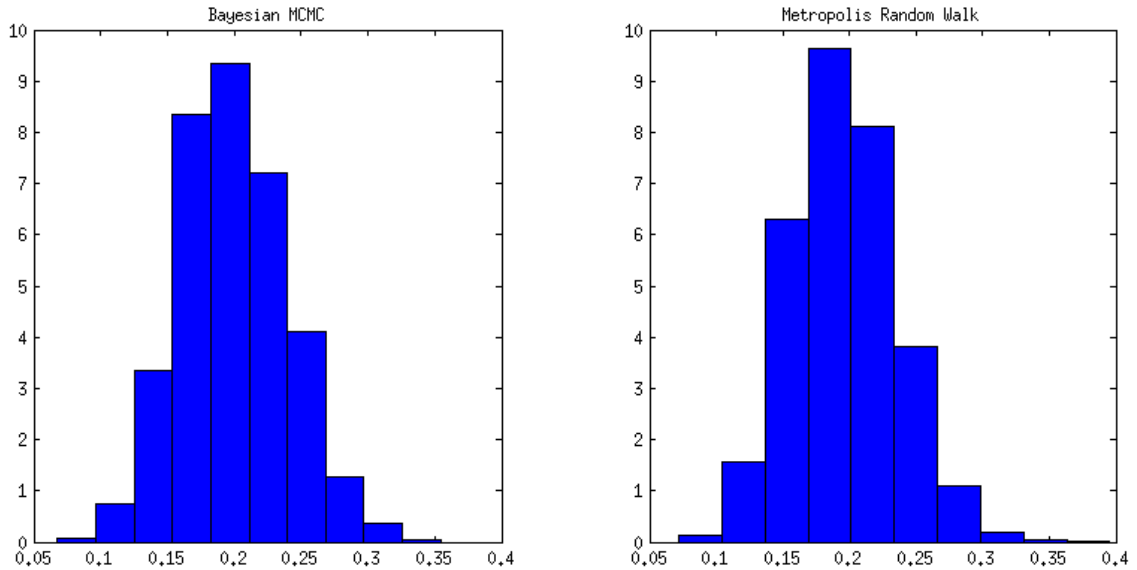


Figure 1: The Markov Chain generated using the Bayesian approach represented by the graph on the left panel, has a mean  $\bar{x} = 0.1993$  and a standard deviation  $s = 0.0402$ . The Markov Chain generated using the Metropolis Random Walk sampler represented by the graph on the right panel, has a mean  $\bar{x} = 0.1959$  and a standard deviation  $s = 0.0393$ . Note that the original data set used for this exercise was generated by a Bernoulli process with  $p = 0.20$ , thus the results are expected.

## Mean & Standard Deviation Results from MATLAB

The mean of the Bayesian MC is: 0.1993

The std deviation of the Bayesian MC is: 0.0402

The mean of the Random Walk MC is: 0.1959

The std deviation of the Random Walk MC is: 0.0393

### Code

```
% PART 0.
% generate hypothesized rs from bernoulli(0.2)
data = binornd(1,0.2,1,100); m = length(data);

% PART 1.
% use the M-H sampler to generate MC of size 2000 whose invariance
% (target) distribution is given by the posterior distribution of
%  $\theta|X$ 

% Set up function handle to evaluate the likelihood.
likelihood = @(theta,x,n) theta.^sum(x).*(1-theta).^(n-sum(x));

% Generate 10000 samples in the chain.
n = 20000; % random sample size
% initialize the chain
mc1 = zeros(1,n);
mc1(1) = rand(1); % generate the starting point
for i = 2:n
    % generate a candidate from the chosen prior unif(0,1)
    y = unifrnd(0,1);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, likelihood(y,data,m)/(likelihood(mc1(i-1),data,m))]);
    if u <= alphaf
        mc1(i) = y;
    else
        mc1(i) = mc1(i-1);
    end
end
% burn-in 5%
mc1 = mc1(0.05*n+1:n);

% PART 2.
% Given that we know the posterior  $\theta|x$  is dist Beta(sum(x)+1,n-sum(x)+1)
% generate a MC from this dist using the random walk sampler.

% Set up function handle to evaluate the Beta kernel.
betapdfker = @(x,a,b) (x.^(a-1)).*((1-x).^(b-1));
alpha = sum(data)+1; beta = m-sum(data)+1; % parameters for the beata
```

```

% generate the MC
mc2 = zeros(1,n);
mc2(1) = rand(1); % generate the starting point
for i = 2:n
    % generate a candidate from the chosen prior unif(0,1)
    a = -0.5; b = 0.5;
    z = a + (b-a)*unifrnd(0,1);
    y = mc2(i-1)+z/sqrt(12);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, betapdfker(y,alpha,beta)/(betapdfker(mc2(i-1),alpha,beta))]);
    if u <= alphaf
        mc2(i) = y;
    else
        mc2(i) = mc2(i-1);
    end
end
end
% burn-in 5%
mc2= mc2(0.05*n+1:n);

% Part 3.
% Calculate mean and sd
meanMC1 = mean(mc1);
stdMC1 = std(mc1);
meanMC2 = mean(mc2);
stdMC2 = std(mc2);
fprintf('\nThe mean of the Bayesian MC is: %2.4f\n',meanMC1)
fprintf('\nThe std deviation of the Bayesian MC is: %2.4f\n',stdMC1)
fprintf('\nThe mean of the Random Walk MC is: %2.4f\n',meanMC2)
fprintf('\nThe std deviation of the Random Walk MC is: %2.4f\n',stdMC2)

% Part 4.
% plot histograms and compare
figure(1)
subplot(1,2,1)
[fhath, bc] = hist(mc1);
fhath = fhath/((bc(2)-bc(1))*sum(fhath));
bar(bc,fhath,1,'b')
title('Bayesian MCMC')
subplot(1,2,2)
[fhath, bc] = hist(mc2);
fhath = fhath/((bc(2)-bc(1))*sum(fhath));
bar(bc,fhath,1,'b')
title('Metropolis Random Walk')

```

## Exercise 14.1

The general von Mises pdf has the form:

$$f(x) = \frac{1}{2\pi I_0(b)} e^{b \cos(x-\mu)}$$

The parameter  $\mu$  is the mean, so in this example we expect the mean of our sample to be close to zero. Additionally, our target distribution has the proportionality:

$$f(x|b) \propto e^{b \cos(x)}$$

Thus our algorithm uses the above relation to calculate  $\alpha$ , the transition probability, and  $Y_t = X_t + Z_t$ , where  $Z_t \sim u(-1, 1)$ , as our random walk sampler. The generated Markov Chain and its histogram is shown in figure 2 below.

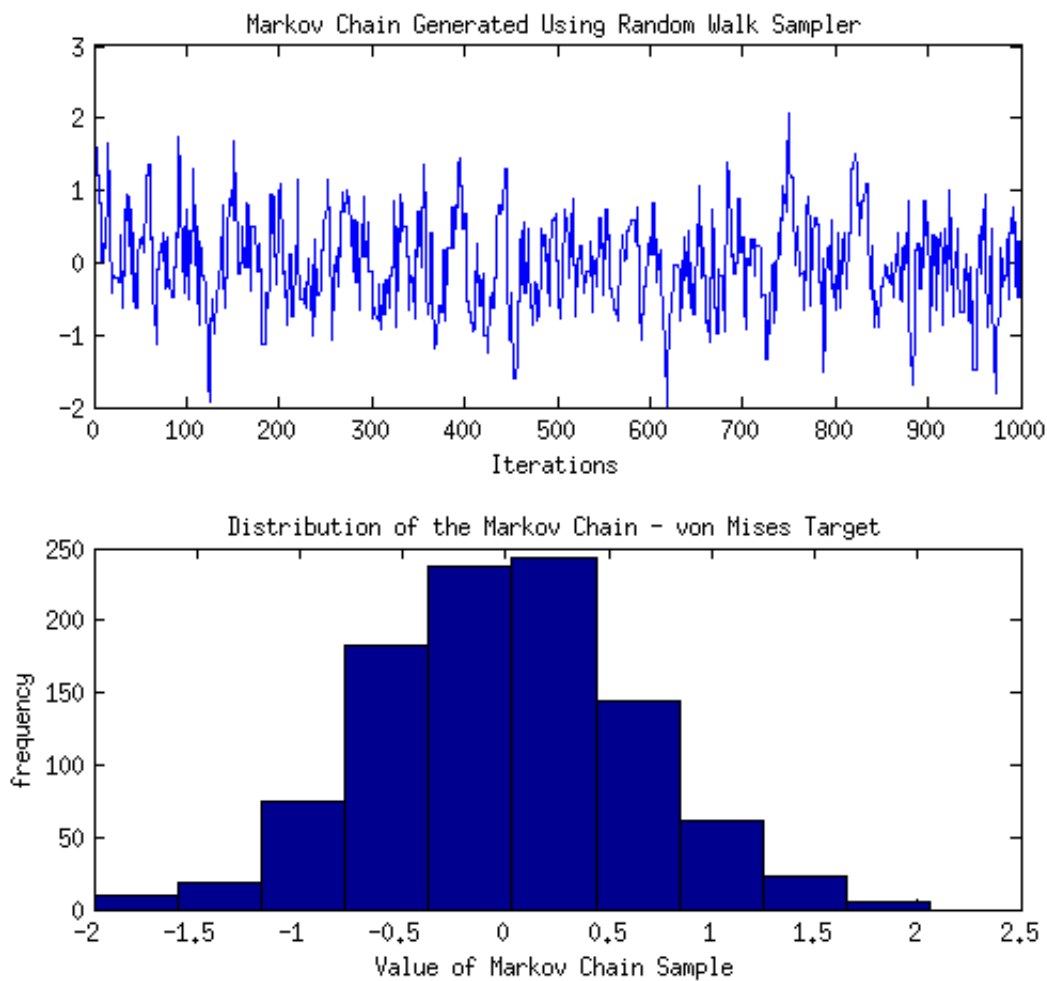


Figure 2: Markov Chain generated using the Metropolis Random Walk sampler (top panel) and histogram of the Markov Chain (bottom panel). Notice the random sample is centered at  $\mu = 0$ .

**Burn-In:** Since the starting point of the chain is 1, but our target distribution has a mean of zero, it would be advisable to "burn-in" the first few observations. That way, it is more likely that the chain obtained after the burn-in will be stationary and centered around the mean of the target distribution. Notice that the first few values start at about 1, and then converge towards zero.

## Code

```
% Set up function handle to evaluate the von Mises kernel.
vonmisespdfker = @(x,b) exp(b.*cos(x));
b = 3; % parameters for the von mises

% generate the MC
n = 1000;
mc = zeros(1,n);
mc(1) = 2; % generate the starting point
for i = 2:n
    % generate a candidate from the random walk
    y = mc(i-1)+unifrnd(-1,1);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, vonmisespdfker(y,b)/(vonmisespdfker(mc(i-1),b))]);
    if u <= alphaf
        mc(i) = y;
    else
        mc(i) = mc(i-1);
    end
end

% plots
figure(1)
subplot(2,1,1)
plot(mc) % entire MC
title('Markov Chain Generated Using Random Walk Sampler')
xlabel('Iterations')
subplot(2,1,2)
hist(mc)
title('Distribution of the Markov Chain - von Mises Target')
xlabel('Value of Markov Chain Sample')
ylabel('frequency')
```

## Exercise 14.2

In this exercise we generate a random sample from the  $\text{Beta}(0.5, 0.5)$ , the target distribution, using the Metropolis-Hastings algorithm. We sample candidates from the proposal distribution  $\text{Unif}(\max(0, X_t - 1), \min(1, X_t + 1))$ . The results in figure 3 show that the sample generated closely resembles the density of the target distribution.

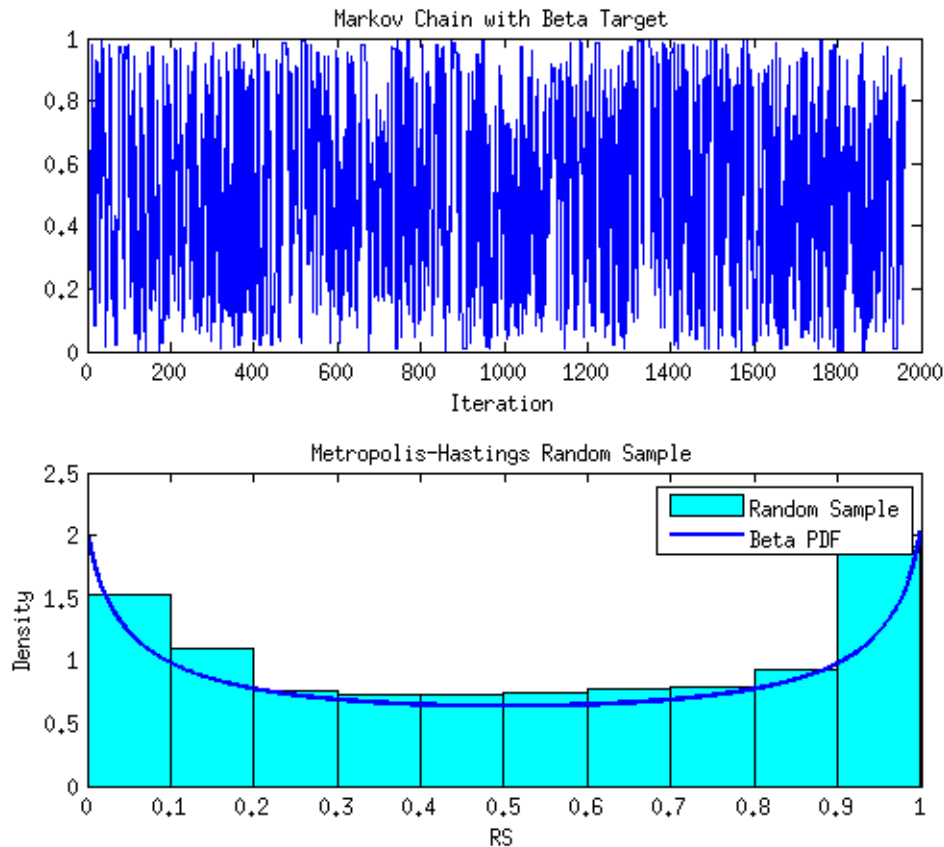


Figure 3: Random sample generated using Metropolis-Hastings. **Proposal:**  $\text{Unif}(\max(0, X_t - 1), \min(1, X_t + 1))$ ; **Target:**  $\text{Beta}(0.5, 0.5)$ .

### Code

```
% METROPOLIS-HASTINGS TO GENERATE BETA SAMPLES

% Set up function handle to evaluate the Beta.
betapdfker = @(x,a,b) (x.^(a-1)).*((1-x).^(b-1));
a = 0.5; b = 0.5; % parameters for the beta
% set up function handle to evaluate the proposal distribution
unipdf = @(theta1,theta2) 1./(theta2-theta1);

% 1.GENERATE RANDOM SAMPLE OF SIZE n
% Generate 10000 samples in the chain.
n = 2000; % random sample size
% initialize the chain
x = zeros(1,n);
x(1) = rand(1); % generate the starting point
```

```

d = 1;
for i = 2:n
    % generate a candidate from the proposal distribution. This will be a
    % proposal with parameters given by the previous value in the chain.
    theta1 = max(0,x(i-1)-d);
    theta2 = min(x(i-1)+d,1);
    y = unifrnd(theta1,theta2,1,1);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, betapdfker(y,a,b)*unipdf(y-d,y+d)/...
        (betapdfker(x(i-1),a,b)*unipdf(x(i-1)-d,x(i-1)+d))]);
    if u <= alphaf
        x(i) = y;
    else
        x(i) = x(i-1);
    end
end
% burn-in 2%
x= x(0.02*n+1:n);

% 2. PLOT THE RESULTS
subplot(211)
plot(x)
xlabel('Iteration')
title('Markov Chain with Beta Target', 'FontSize',14)
subplot(212)
[fhath, bc] = hist(x);
fhath = fhath/((bc(2)-bc(1))*sum(fhath));
bar(bc,fhath,1,'c')
hold on
linebetapdf = plot(linspace(0,1,5000),betapdf(linspace(0.025,0.975,5000),a,b),'b','LineW
xlabel('RS')
ylabel('Density')
title('Metropolis-Hastings Random Sample', 'FontSize',14)
legend('Random Sample','Beta PDF')
hold off

```



## Exercise 14.3

In this exercise we generate a random sample from the  $\text{Gamma}(2,3)$ , the target distribution, using the Metropolis-Hastings algorithm. We sample candidates from the proposal distribution  $\text{Unif}(\max(0, X_t - 5), X_t + 5)$ . The results in figure 4 show that the sample generated closely resembles the density of the target distribution.

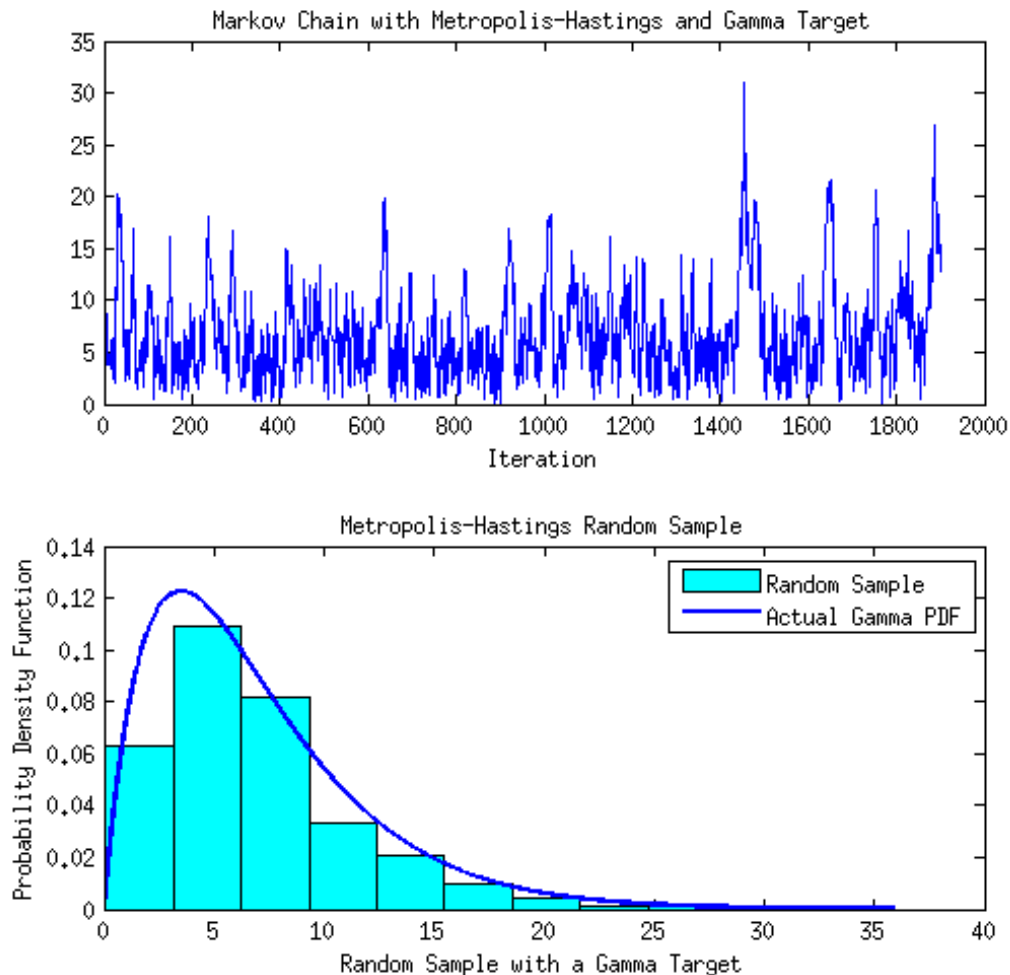


Figure 4: Random sample generated using Metropolis-Hastings. **Proposal:**  $\text{Unif}(\max(0, X_t - 5), X_t + 5)$ ; **Target:**  $\text{Gamma}(2,3)$ .

### Code

```
% METROPOLIS-HASTINGS

% Set up function handle to evaluate the gamma.
gammampdfker = @(x,a,b) (x.^(a-1)).*exp(-x./b);
a = 2; b = 3; % parameters for the gamma
% set up function handle to evaluate the proposal distribution
unipdf = @(theta1,theta2) 1./(theta2-theta1);

% 1.GENERATE RANDOM SAMPLE OF SIZE n
% Generate 20000 samples in the chain.
```

```

n = 2000; % random sample size
% initialize the chain
x = zeros(1,n);
x(1) = 1; %rand(1); % generate the starting point
d = 5;
for i = 2:n
    % generate a candidate from the proposal distribution. This will be a
    % proposal with parameters given by the previous value in the chain.
    theta1 = max(0,x(i-1)-d);
    theta2 = x(i-1)+d;
    y = unifrnd(theta1,theta2,1,1);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, gammapdfker(y,a,b)*unipdf(y-d,y+d)/...
        (gammapdfker(x(i-1),a,b)*unipdf(x(i-1)-d,x(i-1)+d))]);
    if u <= alphaf
        x(i) = y;
    else
        x(i) = x(i-1);
    end
end
% burn-in 5%
x= x(0.05*n+1:n);

% 2. PLOT THE RESULTS
subplot(211)
plot(x)
xlabel('Iteration')
title('Markov Chain with Metropolis-Hastings and Gamma Target','FontSize',14)
subplot(212)
[fhath, bc] = hist(x);
fhath = fhath/((bc(2)-bc(1))*sum(fhath));
bar(bc,fhath,1,'c')
hold on
plot(linspace(0,max(x)+5,5000),gampdf(linspace(0,max(x),5000),a,b),'b','LineWidth',2);
xlabel('Random Sample with a Gamma Target')
ylabel('Probability Density Function')
title('Metropolis-Hastings Random Sample', 'FontSize',14)
legend('Metropolis Hasting Random Sample','Actual Gamma PDF')
hold off

```

## Exercise 14.4

Figure 5 below shows the random sample generated from the standard normal via Metropolis-Hastings with a chain starting value of 20 (which is very far from the actual mean of zero) and a burn-in rate of 2%. Clearly, the sample is not quite adequate, having a denser upper tail, and not quite resembling the shape of the target distribution. We can also see, that it takes quite a few iterations for the chain to achieve stationarity around  $\mu = 0$ .

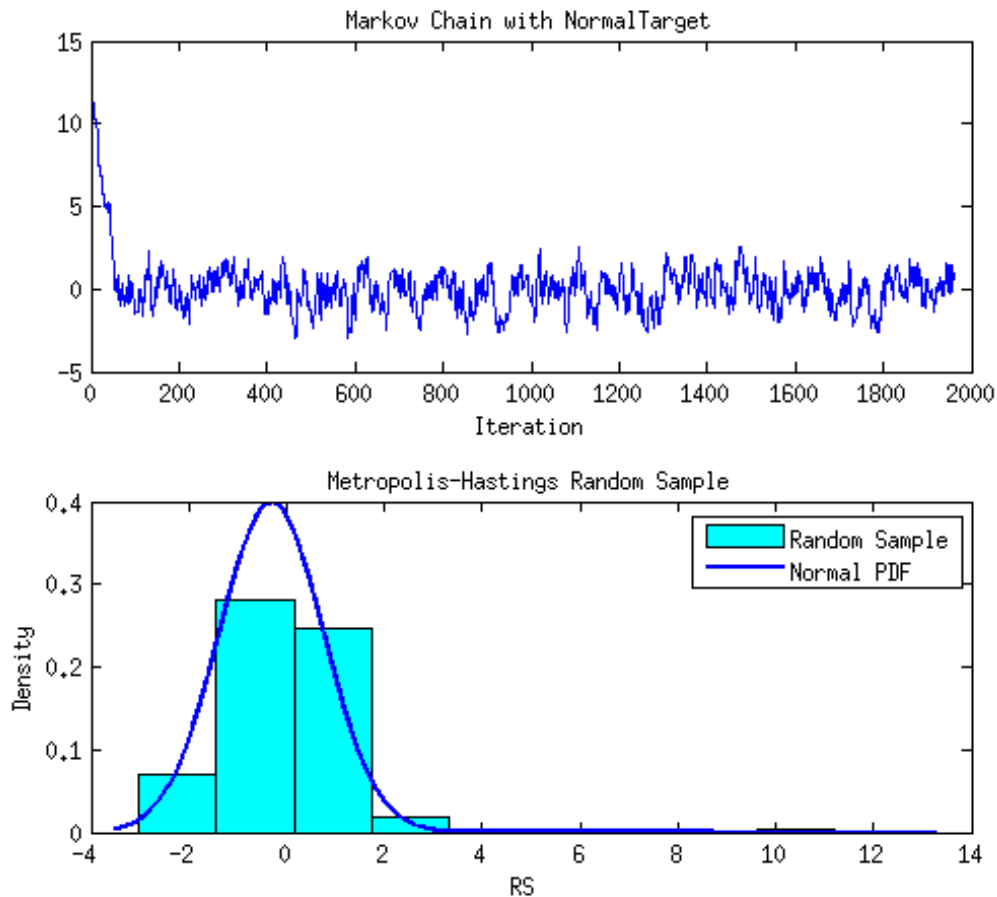


Figure 5: Markov Chain/Random Sample from the standard normal. **Starting value:** 20, **burn-in:** rate 2%. **Target:**  $\text{Normal}(0,1)$ ; **Proposal:**  $\text{Unif}(X_t - 1, X_t + 1)$ .

Figure 6 below shows a similarly generated random sample as in figure 5 using the same starting value of 20 but now the burn-in rate is 20%. Clearly, this sample is much more adequate, closely resembling the shape of the target distribution. We can also see, that the chain is more or less stationary around  $\mu = 0$ .

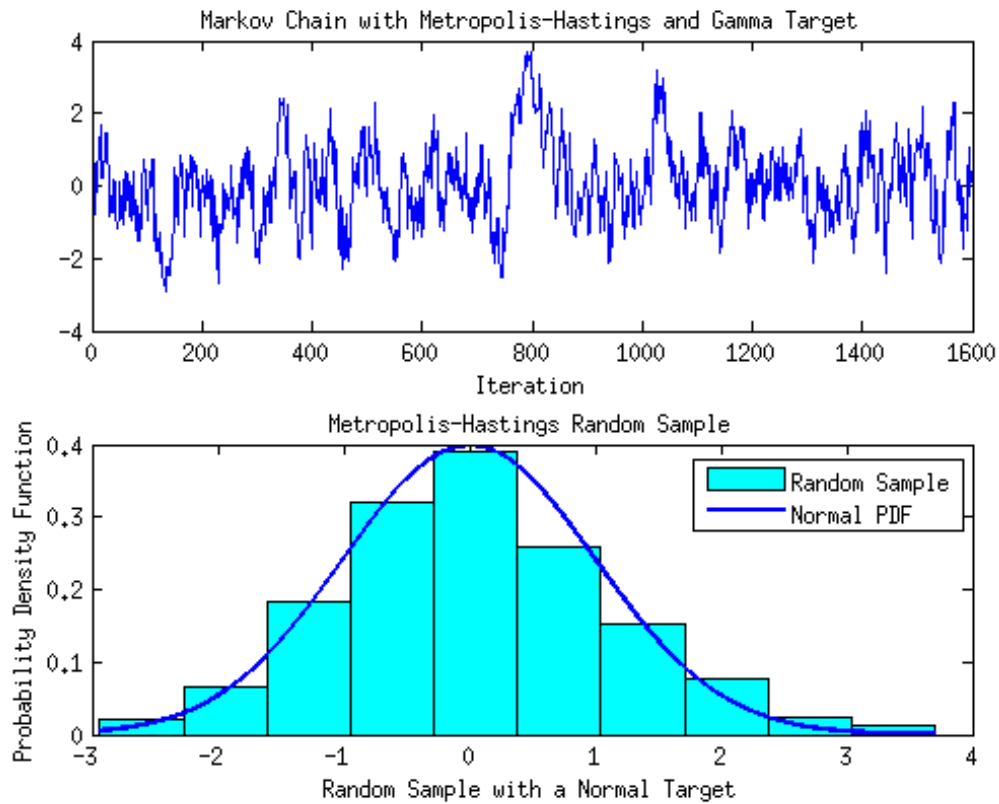


Figure 6: Markov Chain/Random Sample from the standard normal. **Starting value:** 20, **burn-in rate:** 20%. **Target:**  $\text{Normal}(0,1)$ ; **Proposal:**  $\text{Unif}(X_t - 1, X_t + 1)$ .

Figure 7 below shows a similarly generated random sample as in figure 5 but using a starting value of 0 and a burn-in rate of 2%. Clearly, this sample is adequate, closely resembling the shape of the target distribution. We can also see, that the chain is more or less stationary around  $\mu = 0$ . In the case, since the chain starts close to its mean, we don't expect the burn-in rate to make much of a difference on an already adequate sample.

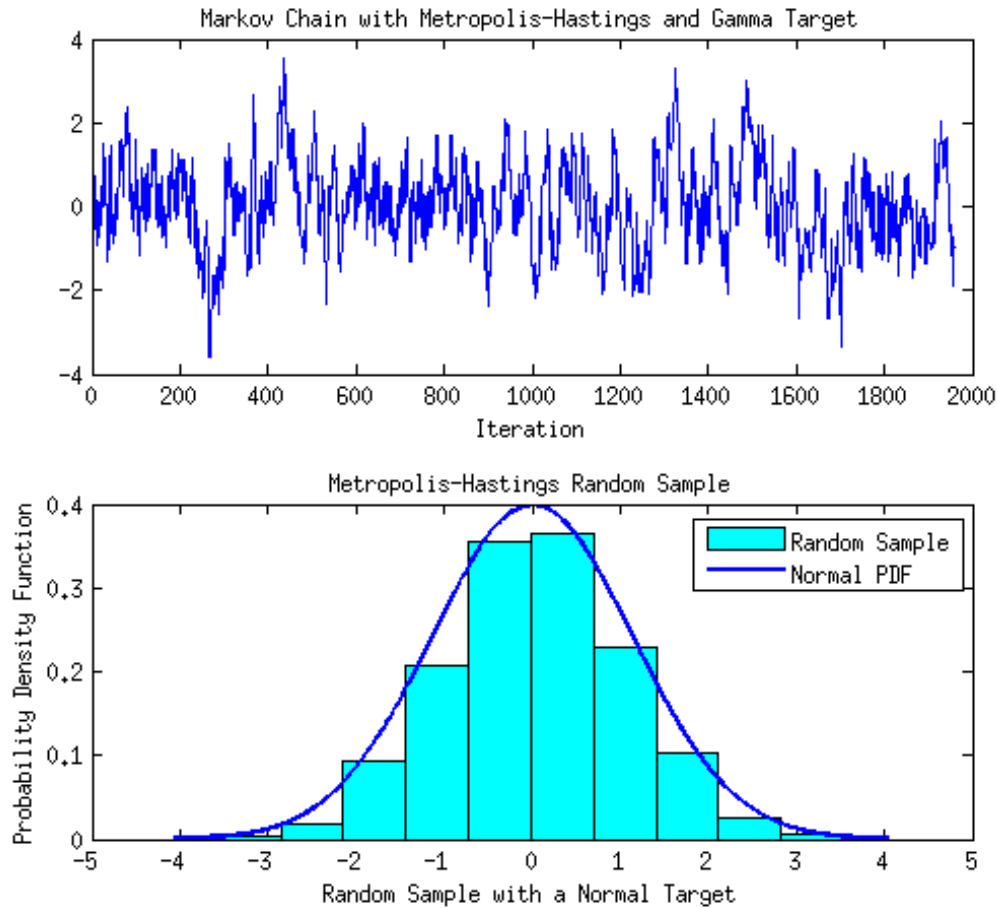


Figure 7: Markov Chain/Random Sample from the standard normal. **Starting value:** 0, **burn-in rate:** 2%. **Target:** Normal(0,1); **Proposal:**  $\text{Unif}(X_t - 1, X_t + 1)$ .

Figure 8 below shows a similarly generated random sample as in figure 5 but using a starting value of 0 and a burn-in rate of 20%. Clearly, this sample is adequate, closely resembling the shape of the target distribution. We can also see, that the chain is more or less stationary around  $\mu = 0$ . In the case, since the chain starts close to its mean, we confirm that the burn-in rate does not impact the adequacy of the sample generated given that the starting value of the chain is close to the mean of the target distribution.

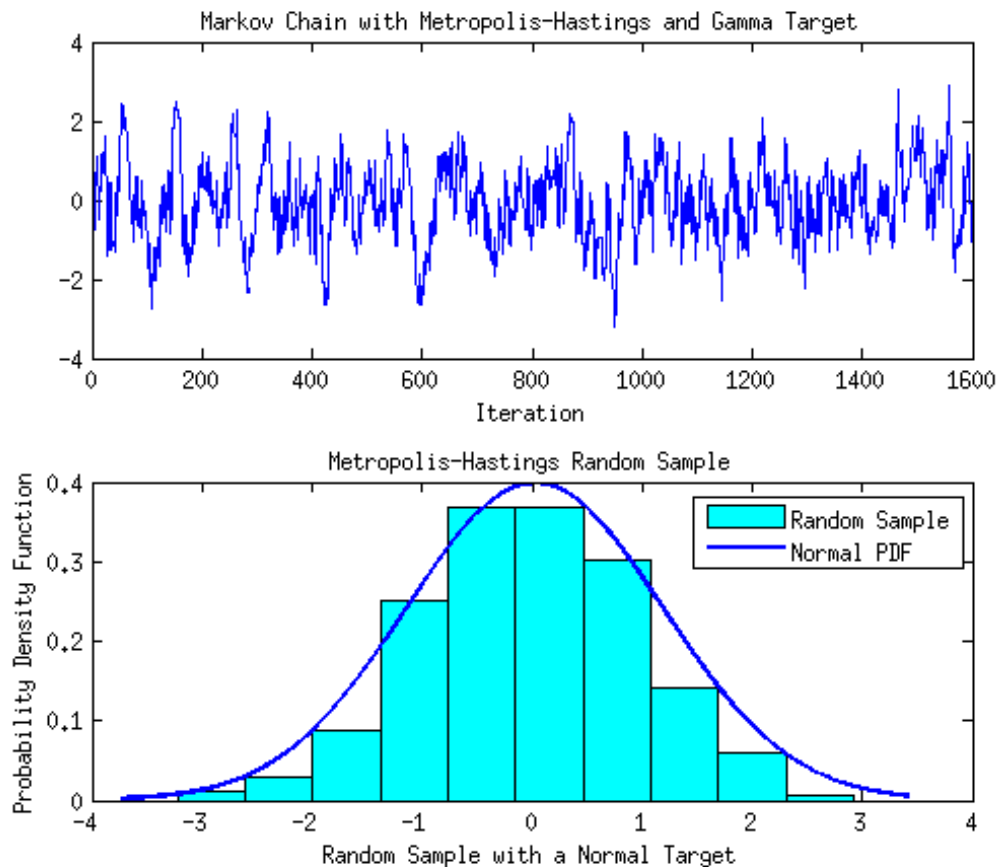


Figure 8: Markov Chain/Random Sample from the standard normal. **Starting value:** 0, **burn-in rate:** 20%. **Target:** Normal(0,1); **Proposal:** Unif( $X_t - 1, X_t + 1$ ).

### Code

```
% METROPOLIS-HASTINGS

% Set up function handle to evaluate the target distribution.
normpdfker = @(x,mu,sigma) exp(-0.5*((x-mu)/sig).^2);
mu = 0; sig = 1; % parameters for the normal
% set up function handle to evaluate the proposal distribution
unipdf = @(theta1,theta2) 1./(theta2-theta1);

% 1.GENERATE RANDOM SAMPLE OF SIZE n
% Generate 20000 samples in the chain.
n = 2000; % random sample size
% initialize the chain
x = zeros(1,n);
```

```

x(1) = 20; %rand(1); % generate the starting point
for i = 2:n
    % generate a candidate from the proposal distribution. This will be a
    % proposal with parameters given by the previous value in the chain.
    theta1 = x(i-1)-1; theta2 = x(i-1)+1;
    y = unifrnd(theta1,theta2);
    % generate a uniform for comparison
    u = rand(1);
    alphaf = min([1, normpdfker(y,mu,sig)*unipdf(y-1,y+1)/...
        (normpdfker(x(i-1),mu,sig)*unipdf(x(i-1)-1,x(i-1)+1))]);
    if u <= alphaf
        x(i) = y;
    else
        x(i) = x(i-1);
    end
end
% burn-in 5%
x= x(0.02*n+1:n);

% 2. PLOT THE RESULTS
subplot(211)
plot(x)
xlabel('Iteration')
title('Markov Chain with Metropolis-Hastings and Gamma Target','FontSize',14)
subplot(212)
[fhath, bc] = hist(x);
fhath = fhath/((bc(2)-bc(1))*sum(fhath));
bar(bc,fhath,1,'c')
hold on
plot(linspace(min(x)-0.5,max(x)+0.5,5000),normpdf(linspace(min(x),max(x),5000),mu,sig),',
xlabel('Random Sample with a Normal Target')
ylabel('Probability Density Function')
title('Metropolis-Hastings Random Sample', 'FontSize',14)
legend('Random Sample','Normal PDF')
hold off

```