

Definition 1. Data that is counted is **discrete** while data that is measured is **continuous** (Gates et al., 2018b).

Definition 2. A **combination** is an arrangement of items without regard to the order. A **permutation** is an arrangement of items in a specific order. (Gates et al., 2018a).

Definition 3. An **experiment** is a procedure that yields one of a given set of possible outcomes (Rosen, 2002).

Definition 4. The **sample space** of the experiment is the set of possible outcomes (Rosen, 2002).

Definition 5. An **event** is a subset of the sample space (Rosen, 2002).

Definition 6. Probability is the measure of the likelihood that an event will occur. Probability is quantified as a number between 0 and 1, where 0 indicates impossibility and 1 certainty.

Definition 7. Finite probability. If S is a finite nonempty sample space of equally likely outcomes, and E is an event, that is, a subset of S , then the probability of E is $p(E) = \frac{|E|}{|S|}$ (Rosen, 2002).

Definition 8. A **random variable**¹ is a function from the sample space of an experiment to the set of real numbers.

$$X : S \rightarrow \mathbb{R}$$

A random variable $X(t)$ maps each outcome t to a real number (Rosen, 2002).

Note. It is worth clarifying that “random variable” is a misnomer. A random variable is a function. It is not a variable, and it is not random! It is a representation of an underlying random system. The name *random variable* (the translation of *variabile casuale*) was introduced by the Italian mathematician F. P. Cantelli in 1916 (Rosen, 2002).

1 Discrete Probability Distributions

Definition 9. The **distribution** of a random variable X on a sample space S is the set of pairs $(r, p(X = r))$ for all $r \in X(S)$, where $p(X = r)$ is the probability that X takes the value r . The set of pairs in this distribution is determined by the probabilities $p(X = r)$ for $r \in X(S)$ (Rosen, 2002).

1.1 Bernoulli Distribution

Definition 10. Discrete probability distribution of a random variable which takes the value 1 with probability p and the value 0 with probability $q = 1 - p$.

¹ Random variables translate outcomes in the sample space of an experiment to real numbers so that we can reason about them with mathematics.

Random variable

X = The outcome of the **Bernoulli experiment** is “success”.

Probability mass function

$$f(k; p) = \Pr(k; p) = \Pr(X = k) = p^k(1 - p)^{1-k} \quad \text{for } k \in \{0, 1\}$$

Parameter

p = Probability of the outcome of a **Bernoulli experiment** being a success.

1.2 Binomial Distribution

Definition 11. Discrete probability distribution of the number of successes in a sequence of n independent Bernoulli experiments. For a single trial, that is, $n = 1$, the binomial distribution is a Bernoulli distribution.

Note. The binomial distribution is frequently used to model the number of successes in a sample of size n drawn with replacement from a population of size N . If the sampling is carried out without replacement, the draws are not independent and so the resulting distribution is a **hypergeometric** distribution, not a binomial one. However, for N much larger than n , the binomial distribution remains a good approximation, and is widely used.

Random variable

X = The number of successes in a sequence of n independent **Bernoulli experiments** is k .

Probability mass function of getting k successes in n trials.

$$f(k; n, p) = \Pr(k; n, p) = \Pr(X = k) = \binom{n}{k} p^k (1 - p)^{n-k} \quad \text{for } k \in \{0, 1, \dots, n\}$$

Parameters

p = Probability of the outcome of a **Bernoulli experiment** being a success.

n = Number of trials (experiemnts).

Definition 12. The binomial coefficient, $\binom{n}{k} = \frac{n!}{k!(n-k)!}$, is the coefficient of the x^k term in the polynomial expansion of the binomial power $(1 + x)^n = \binom{n}{0}x^0 + \binom{n}{1}x^1 + \dots + \binom{n}{n}x^n$, that is, the **binomial theorem**.

Explanation k successes occur with probability p^k and $n - k$ failures occur with probability $(1 - p)^{n-k}$. The k successes can occur anywhere among the n trials, and there are $\binom{n}{k}$ different ways of distributing k successes in a sequence of n trials.

1.3 Poisson Distribution

Definition 13. Discrete probability distribution of a given number of events occurring in a fixed space interval, usually time, and these events occur with a known constant rate and independently of the time since the last event. The Poisson distribution is derived from the **binomial distribution** by splitting up the interval into n subintervals, each of which is so small that at most one event could occur in it with non-zero probability (Wackerly et al., 2008).

Random variable

X = The number of events in an interval is k .

Probability mass function of getting k events in an interval.

$$f(k; \lambda) = \Pr(k; \lambda) = \Pr(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

Parameter

λ = event rate, the average number of events in an interval.

1.4 Geometric Distribution

Geometric random variable has two alternate formulations:

X_1 = The number of **Bernoulli experiments** needed to get the first success.

X_2 = The number of failed **Bernoulli experiments** before the first success.

Probability mass function for each corresponding random variable.

$$\begin{aligned} f_1(k; p) &= \Pr_1(k; p) = \Pr(X_1 = k) = (1 - p)^{k-1}p && \text{for } k = 1, 2, \dots, 0 \leq p \leq 1 \\ f_2(k; p) &= \Pr_2(k; p) = \Pr(X_2 = k) = (1 - p)^k p && \text{for } k = 0, 1, 2, \dots, 0 \leq p < 1 \end{aligned}$$

Parameter

p = Probability of the outcome of a Bernoulli experiment being success.

Example Probability that the first coin flip turns up heads:

$$p(X_1 = 1) = (1 - p)^{1-1}p = (1 - p)^0 p = p(X_2 = 0)$$

2 Continuous Probability Distributions

2.1 Exponential Distribution

Definition 14. The exponential distribution (also known as negative exponential distribution) is the probability distribution that describes the time between events in a **Poisson point process**.

Note. It is the continuous analogue of the **geometric distribution**, and it has the key property of being memoryless. In addition to being used for the analysis of Poisson point processes it is found in various other contexts.

Exponential random variable

X = The amount of time that passes before the next event.

Probability density function

$$f(x; \lambda) = \lambda e^{-\lambda x} \quad \text{for } x \geq 0$$

Parameter

λ = event rate, the average number of events in an interval.

Derivation from the **geometric distribution** (MathHolt, 2018).

1. Split up each interval into n sub-intervals such that the probability of an event occurring in a certain sub-interval is $p = \frac{\lambda}{n}$
2. For a **geometric random variable** Y = The number of trials needed to get the first success, $p(X \leq b) \approx P(Y \leq b \cdot n)$

$$\begin{aligned} P(Y \leq b \cdot n) &= \sum_{k=1}^{b \cdot n} P(Y = k) \\ &= \sum_{k=1}^{b \cdot n} (1-p)^{k-1} p = \sum_{k=1}^{b \cdot n} \left(1 - \frac{\lambda}{n}\right)^{k-1} \cdot \frac{\lambda}{n} \\ &= \frac{\lambda}{n} \cdot \sum_{k=1}^{b \cdot n} \left(1 - \frac{\lambda}{n}\right)^{k-1} \\ &= \frac{\lambda}{n} \cdot \sum_{k=0}^{b \cdot n - 1} \left(1 - \frac{\lambda}{n}\right)^k \end{aligned}$$

$$\sum_{k=0}^m a^k = \frac{1 - a^{m+1}}{1 - a} \quad \text{finite geometric sum}$$

$$m = b \cdot n - 1 \text{ and } a = 1 - \frac{\lambda}{n}$$

$$\begin{aligned}
P(Y \leq b \cdot n) &= \frac{\lambda}{n} \cdot \frac{1 - \left(1 - \frac{\lambda}{n}\right)^{b \cdot n}}{1 - \left(1 - \frac{\lambda}{n}\right)} \\
&= 1 - \left(1 - \frac{\lambda}{n}\right)^{b \cdot n} \\
&= 1 - \left(\left(1 - \frac{\lambda}{n}\right)^n\right)^b
\end{aligned}$$

$$\begin{aligned}
p(X \leq b) &= \lim_{n \rightarrow \infty} P(Y \leq b \cdot n) \\
&= \lim_{n \rightarrow \infty} \left(1 - \left(\left(1 - \frac{\lambda}{n}\right)^n\right)^b\right) \\
&= 1 - \left(\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n\right)^b \\
&= 1 - (e^{-\lambda})^b \\
&= 1 - e^{-\lambda \cdot b}
\end{aligned}$$

$$\begin{aligned}
p(X \leq b) &= \int_{-\infty}^b \text{pdf}(x) dx \\
p(X \leq b) &= \int_0^b \text{pdf}(x) dx \\
\frac{d}{db} P(X \leq b) &= \frac{d}{db} \int_0^b \text{pdf}(x) dx \\
&= \text{pdf}(b)
\end{aligned}$$

$$\begin{aligned}
\text{pdf}(b) &= \frac{d}{db} P(X \leq b) \\
&= \frac{d}{db} (1 - e^{-\lambda \cdot b}) \\
&= \lambda e^{-\lambda b}
\end{aligned}$$

Yue (2018) present an alternate derivation from a **Markov process** representation of the exponential distribution.

3 Stochastic (Random) Process

Definition 15. A **stochastic process** is a random phenomenon that arises through a process which is developing in time in a manner controlled by probabilistic laws (Parzen, 1999). From a point of view of the mathematical theory of probability a stochastic process is best defined as a

collection of random variables defined on a common probability space (Ω, \mathcal{F}, P) and indexed by points in some space.

$$\{X(t), t \in T\}$$

A more appropriate name in mathematics is **random field**. The set used to index the random variables is called the **index set**. Historically, the index set was some subset of the real line, such as the natural numbers, giving the index set the interpretation of time. Each random variable in the collection takes values from the same mathematical space known as the **state space**. An **increment** is the amount that a stochastic process changes between two index values, often interpreted as two points in time.

Note. Even though the index set can be any set in any space, generally more results and theorems are possible for stochastic processes when the index set is ordered. Most commonly, random variables in a stochastic process are indexed by the positive numbers along the real number line, interpreted as time. Among these, the **Brownian motion process** and the **Poisson process** are considered the most important and central in the theory of stochastic processes. Other stochastic processes include **random walk** and **Markov Chain** and **martingales**. A martingale models a fair game. A simple random walk is a Markov chain and also a martingale.

4 Point Process

Definition 16. Point processes are not defined like stochastic processes, but are used to describe data that are localized in space or time. Point processes on the real line form an important special case that is particularly amenable to study because the points are ordered in a natural way, and the whole point process can be described completely by the (random) intervals between the points or by the event times or by the event counts within each non-overlapping intervals. A **temporal** point process is a random process whose realization is a collection of points in time.

4.1 Renewal (Point) Process

Definition 17. Historically the first point processes that were studied had the real half line $\mathbb{R}_+ = [0, \infty)$ as their state space, which in this context is usually interpreted as time. These studies were motivated by the wish to model telecommunication systems, in which the points represented events in time, such as calls to a telephone exchange. Point processes on \mathbb{R}_+ are typically described by giving the sequence of their (random) inter-event times (T_1, T_2, \dots) , from which the actual sequence (X_1, X_2, \dots) of event times can be obtained as

$$X_k = \sum_{j=1}^k T_j \quad \text{for } k \geq 1.$$

If inter-event times are independent and identically distributed, the point process is called a **renewal process**.

Note. The **intensity** $\lambda(t|H_t)$ of a point process on the real half-line with respect to a filtration H_t is defined as

$$\lambda(t|H_t) = \lim_{\Delta t \rightarrow 0} \frac{1}{\Delta t} \Pr(\text{One event occurs in time interval } [t, t + \Delta t] | H_t)$$

H_t can denote the history of event-point times preceding time t but can also correspond to other *filtrations* (for example in the case of a Cox process). The *compensator* of a point process, also known as the dual-predictable projection, is the integrated conditional intensity function defined by

$$\Lambda(s, u) = \int_s^u \lambda(t|H_t) dt$$

4.2 Poisson (Point)² Process

Definition 18. The number of points in a region of finite size within a Poisson process is a random variable with a **Poisson distribution**. The Poisson point process is often defined on the real line, where it can be considered as a **stochastic process**. The Poisson process is a point process with convenient mathematical properties. The Poisson point process has the property that each point is stochastically independent to all the other points in the process. The Poisson process has been used to build other point processes where the points are not independent of each other.

Definition 19. Poisson Counting Process on the positive half-line that represents the total number of occurrences or events that have happened up to and including time t .

$$\{N(t), t \geq 0\}$$

The probability of random variable $N(t)$ being equal to n :

$$P\{N(t) = n\} = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$$

Definition 20. Poisson Point Process on the positive half-line that represents the number of occurrences in the interval $(a, b]$.

$$P\{N(a_i, b_i) = n_i, i = 1, \dots, k\} = \prod_{i=1}^k \frac{[\lambda(b_i - a_i)]^{n_i}}{n_i!} e^{-\lambda(b_i - a_i)}$$

Definition 21. Poisson Point Process³ on the positive half-line that represents the inter-event times.

$$P\{X_i, i = 1, \dots, k\} = \prod_{i=1}^k \lambda e^{-\lambda x_i}$$

² The word point is often omitted, but there are other Poisson processes of objects, which, instead of points, consist of more complicated mathematical objects such as lines and polygons, and such processes can be based on the Poisson point process.

³This is the interpretation that we will use when describing PCIM.

4.3 Marked Point Process

Definition 22. Let \mathcal{X} be the space of random variables of an ordinary point process and \mathcal{L} be the space of random variables of labels. A marked point process can be regarded as either a point process in the product space $\mathcal{X} \times \mathcal{L}$, subject to the finiteness constraint of the ground process (that is, the point process of event times), or as an ordinary point process in \mathcal{X} with an associated sequence random variables taking their values in \mathcal{L} . A realization of a marked point process is a sequence $\{(x_i, l_i)\}_{i=1}^n$.

4.4 PCIM

The **Piecewise-Constant Conditional Intensity Model** (PCIM) is a class of **marked point processes** that can model the types and timing of events. This model captures the dependencies of each type of event on events in the past through a set of piecewise-constant conditional intensity functions. Decision trees represent these dependencies. The decision trees are the piecewise-constant conditional intensity functions (Gunawardana et al., 2011)

A conjugate prior for this model allows for closed-form computation of the marginal likelihood and parameter posteriors. Model selection then becomes a problem of choosing a decision tree. Decision tree induction can be done efficiently because of the closed form for the marginal likelihood.

Poisson Networks (Rajaram et al., 2005) (also piece-wise constant conditional intensity models) are closely related to PCIMs.

Definition 23. Event stream

$$y = \{(t_i, l_i)\}_{i=1}^n$$

with $0 < t_1 < \dots < t_n$, where $t_i \in [0, \infty)$ is the time of the i th event and l_i is its label, drawn from a finite label set \mathcal{L} .

Definition 24. History at time t of event sequence y is the sub-sequence

$$h(t, y) = \{(t_i, l_i) | (t_i, l_i) \in y, t_i \leq t\}$$

- $t_0 = 0$
- $h_i = h(t_{i-1}, y)$ is the history at time t_i
- $t(y) = \max(\{t : (t, l) \in y\})$ is the end time of event sequence y such that $t(h_i) = t_{i-1}$

Definition 25. Conditional Intensity Model (Didelez, 2008; Daley and Vere-Jones, 2003):

$$p(x|\theta) = \prod_{l \in \mathcal{L}} \prod_{i=1}^n \lambda_l(t_i | h_i, \theta)^{\mathbf{1}_l(l_i)} e^{-\Lambda_l(t_i | h_i; \theta)}$$

where $\Lambda_l(t|x; \theta) = \int_{-\infty}^t \lambda_l(\tau|x; \theta) d\tau$ for each event sequence x and the indicator function $\mathbf{1}_l(l')$ is one if $l' = l$ and zero otherwise. The rate parameter changes as a function of time and history and is computed by the conditional intensity function $\lambda_l(t_i | h_i, \theta)$. “The conditional intensities are assumed to satisfy $\lambda_l(t|x; \theta) = 0$ for $t \leq t(x)$ to ensure that $t_i > t_{i-1} = t(h_i)$.”⁴

⁴This sentence means that the conditional intensity functions will only evaluate to valid values for times into the future. “ $\lambda_l(t|y)$ is piecewise constant in t for all $t > t(y)$ ” (Parikh et al., 2012).

Note. These models offer a powerful approach for decomposing the dependencies of different event types on the past. In particular, this per-label conditional factorization allows one to model detailed label-specific dependence on past events (Gunawardana et al., 2011).

Note. The *Conditional Intensity Model* is analogous⁵ to the likelihood of a sequence of samples (**inter-event times**) from the exponential distribution where the rate parameter changes over time. Let X be exponential distributed with rate parameter λ and x_1, \dots, x_n are n independent samples from X . The likelihood function is given by:

$$L(\lambda; x) = p(x; \lambda) = \prod_{i=1}^n \lambda e^{-\lambda x_i}$$

When there is only one label, the conditional intensity model becomes

$$p(x|\theta) = \prod_{i=1}^n \lambda(t_i|h_i, \theta) e^{-\Lambda(t_i|h_i; \theta)}$$

and when the rate parameter is fixed $\lambda(t_i|h_i, \theta) = \lambda$ it becomes $L(\lambda; x)$.

Definition 26. Piecewise-Constant Conditional Intensity Models (PCIM) are a factorization of Conditional Intensity Models where the conditional intensity functions are assumed to be piecewise-constant. This assumption allows efficient learning and inference.

$$p(x|S, \Theta) = \prod_{l \in \mathcal{L}} \prod_{s \in \Sigma_l} \lambda_{ls}^{c_{ls}(x)} e^{-\lambda_{ls} d_{ls}(x)}$$

- $S = \{S_l\}_{l \in \mathcal{L}}$, where $S_l = (\Sigma_l, \sigma_l(t, x))$: Set of local structures, one per event type (label l)
- Σ_l is a set of discrete **states**
- $s = \sigma_l(t, x)$ is the state given by the **piecewise-constant state function** (a **decision tree**)
- $\Theta = \{\theta_l\}_{l \in \mathcal{L}}$, where $\theta_l = \{\lambda_{ls}\}_{s \in \Sigma_l}$: Set of local parameters
- $\lambda_l(t|x) = \lambda_{ls}$: The **piecewise-constant conditional intensity functions**
- $c_{ls}(x)$: The number of times label l occurs in state s
- $d_{ls}(x)$: The total time during which $\sigma_l(t|x)$ maps to state s

Note. “The state s summarizes all the information about t and y necessary for computing $\lambda_l(t|y)$ in that given s , $\lambda_l(t|y)$ can be computed without further information about t and y . However, unlike in Markov models such as CTBNs, the state does not contain all the information about t and y for predicting future states.” (Parikh et al., 2012)

⁵ Despite the similarity to the likelihood of a non-homogeneous Poisson process, this likelihood does not in general define a Poisson process as the conditioning on history can cause the independent increments property of Poisson processes to not hold.

Note. Gunawardana et al. (2011) show that given the structure S , a product of Gamma distributions is a conjugate prior for Θ , and that under this prior, the marginal likelihood of the data can be given in closed form. Thus, parameter estimation can be done in closed form given a structure, and imposing a structural prior allows a closed form Bayesian score to be computed for a structure.

Note. Gunawardana et al. (2011) demonstrate their method on two problems. The forecast the user's web search queries and failure events based on system logs on a machine.

4.5 C-PCIM

Definition 27. Conjoint Piecewise-Constant Conditional Intensity Models generalize PCIMs by allowing the sharing of parameters across event types. This parameter sharing allows to better model event streams of fine-grained events, which are likely to be rare events.

Note. In contrast, PCIMs learn dependencies of each event type separately, using independent sub-models for each event type. If some event types are rare, there would not be sufficient data to learn an independent sub-model for such types.

Note. During structure learning, the C-PCIM learns what event types in what historical contexts can be modeled by shared parameters, thereby allowing more efficient use of data during parameter estimation. In cases where events are structured (encoded via the basis test functions?), with the different event types having known attributes, we show how structure learning can take advantage of these attributes to distinguish between different event types when their dependencies differ, while sharing parameters when they do not.

Definition 28. Conjoint Piecewise-Constant Conditional Intensity Models (C-PCIM) are PCIMs where the parameters are shared across labels by a single state function $\sigma(l, t, y)$ that now also depends on the label l .

$$p(x|S, \Theta) = \prod_{s \in \Sigma} \lambda_s^{c_s(x)} e^{-\lambda_s d_s(x)}$$

- $S = (\Sigma, \sigma(\cdot, \cdot, \cdot))$: Structure
- Σ : Set of discrete **states**
- $s = \sigma(l, t, x)$ is the state given by the **piecewise-constant state function** (a **decision tree**)
- $\Theta = \{\lambda_s\}_{s \in \Sigma}$: Set of parameters shared across labels $l \in \mathcal{L}$
- $\lambda_l(t|x) = \lambda_s$: The **piecewise-constant conditional intensity functions**
- $c_s(x) = \sum_i \mathbf{1}_s(\sigma(l_i, t_i, h_i))$: The number of times label l occurs in x when the state function maps to state s for label l
- $d_s(x) = \sum_i \mathbf{1}_s(\sigma(l_{i-1}, t_{i-1}, h_{i-1}))(t_i - t_{i-1})$: The total time spent in state s

Parameter Learning Using a product of Gamma priors on λ_s as conjugate for Θ , the point estimate $\hat{\Theta} = \mathbf{E}[\Theta|S, x]$, given a structure and the data, is given by

$$\hat{\lambda}_s = \frac{\alpha + c_s(x)}{\beta + d_s(x)}$$

Structure Learning Using a factored prior on the structure, $p(S) \propto \kappa^{|\Sigma|}$, learn a decision tree (Chickering et al., 1997) by selecting the next node on the tree to split based on the gain:

$$\text{Gain}(S \rightarrow S') = \frac{\prod_{s' \in \Sigma_f} \kappa \gamma_{s \odot s'}(x)}{\kappa \gamma_s(x)}$$

where

$$\gamma_s(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} \frac{\Gamma(\alpha + c_s(x))}{(\beta + d_s(x))^{\alpha + c_s(x)}}$$

where

$$\Gamma(n) = (n - 1)!$$

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