

Gibbs Sampler

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The Gibbs Sampler

Specially adapted for multidimensional target distribution.

The goal is to construct MC whose stationary distribution equals to the target π .

Idea: Sequential sampling from marginal distribution.

Bivariate Case

$$\underline{X} = (X_1, X_2) \sim f(x_1, x_2) = \pi$$

We want to generate a MC from the target $f(x_1, x_2)$

$$\underline{X}_t = (X_{t,1}, X_{t,2})$$

Suppose that $X_1 | X_2 \sim f(x_1 | x_2)$ are easily sampled.
and $X_2 | X_1 \sim f(x_2 | x_1)$

Procedure

1. Set an initial point $\underline{X}_0 = (X_{0,1}, X_{0,2})$ at $t=0$
2. Generate $X_{t+1,1}$ from $f(X_{t,1} | X_{t,2} = x_{t,2})$
3. " $X_{t+1,2}$ from $f(X_{t,2} | X_{t+1,1} = x_{t+1,1})$

\swarrow prev. value for $X_{t,2}$
 \searrow from Step 2
4. Set $t = t+1$

Repeat step 2 ~ 4.

The Gibbs Sampler always accept the generated value.

Example 14.6 Bivariate (p594)

$$f(x, y) \propto \binom{n}{x} y^{x+p-1} (1-y)^{n-x+p-1} \quad x = 0, 1, \dots, n$$

Want to generate sample from $f(x)$ \rightarrow marginal $0 \leq y \leq 1$

Conditional distribution

$$f(x|y) \propto \binom{n}{x} y^x (1-y)^{n-x} \cdot \underbrace{y^{\alpha-1} (1-y)^{\beta-1}}_{\text{Constant given } y, \alpha, \beta}$$

$$\propto \binom{n}{x} y^x (1-y)^{n-x} \sim \text{Binomial}(n, y)$$

$$f(y|x) \propto \binom{n}{x} y^{x+\alpha-1} (1-y)^{n-x+\beta-1} \sim \text{Beta}(x+\alpha, n-x+\beta)$$

\rightarrow given

Procedure

1. Initial (x_0, y_0)
 x_0 r.s. from $\text{Binomial}(n, 0.5)$ \rightarrow guess
 y_0 r.s. from $\text{Beta}(x_0 + \alpha, n - x_0 + \beta)$
2. x_t r.s. from $\text{Binomial}(n, y_{t-1})$
 y_t r.s. from $\text{Beta}(x_t + \alpha, n - x_t + \beta)$
3. Continue.

Example 14.7 (p547)

Random sampling from Bivariate Normal distribution

Background

$$z_1, z_2 \stackrel{\text{iid}}{\sim} N(0, 1) \Rightarrow f(z_1, z_2) = \frac{1}{2\pi} e^{-\frac{1}{2}(z_1^2 + z_2^2)}$$

Let $x_1 = \sigma_1 z_1 + \mu_1$ and

$$x_2 = \sigma_2 [\rho z_1 + (1-\rho^2)^{1/2} z_2] + \mu_2$$

Then

$$f(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{(x_1-\mu_1)^2}{\sigma_1^2} - \frac{2\rho(x_1-\mu_1)(x_2-\mu_2)}{\sigma_1\sigma_2} + \frac{(x_2-\mu_2)^2}{\sigma_2^2}\right]\right)$$

$$f(x_1, x_2) = \frac{1}{2\pi \sqrt{1-\rho^2} \sigma_1 \sigma_2} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right] \right\}$$

Here $E(X_1) = \mu_1$, $\text{Var}(X_1) = \sigma_1^2$
 $E(X_2) = \mu_2$, $\text{Var}(X_2) = \sigma_2^2$
 correlation $\rho(X_1, X_2) = \rho$

Then X_1 and X_2 have a bivariate Normal distribution

$$N \left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{bmatrix} \right)$$

" \neq

Conditional distribution

$$X_2 | X_1 = x_1$$

$$Z_1 = \frac{X_1 - \mu_1}{\sigma_1} \quad \text{and} \quad Z_2 = \frac{X_2 - \mu_2}{\sigma_2}$$

$$X_2 = \sigma_2 \left[\rho \left(\frac{X_1 - \mu_1}{\sigma_1} \right) + \sqrt{1-\rho^2} \left(\frac{X_2 - \mu_2}{\sigma_2} \right) \right] + \mu_2$$

$$E[X_2 | X_1 = x_1] = \rho \cdot \sigma_2 \left(\frac{x_1 - \mu_1}{\sigma_1} \right) + \mu_2$$

$$\begin{aligned} \text{Var}[X_2 | X_1 = x_1] &= \sigma_2^2 (1-\rho^2) \cancel{V\left(\frac{X_2 - \mu_2}{\sigma_2}\right)} \\ &= \sigma_2^2 (1-\rho^2) \end{aligned}$$

$$X_2 | X_1 = x_1 \sim N \left(\mu_2 + \rho \cdot \frac{\sigma_2}{\sigma_1} (x_1 - \mu_1), \sigma_2^2 (1-\rho^2) \right)$$

Noting that $f(x_1, x_2)$ is symmetric in Z_1 and Z_2

$$X_1 | X_2 = x_2 \sim N \left(\mu_1 + \rho \frac{\sigma_1}{\sigma_2} (x_2 - \mu_2), \sigma_1^2 (1-\rho^2) \right)$$

Procedure

1. Take r.s. $z_{1,0} \sim N(0,1)$, $z_{2,0} \sim N(0,1)$

Then $x_{1,0} = \mu_1 + \sigma_1 z_{1,0}$

$$x_{2,0} = \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x_{1,0} - \mu_1) + \sigma_2 \sqrt{1-\rho^2} \cdot z_{2,0}$$

2. $x_{1,1} = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (x_{2,0} - \mu_2) + \sigma_1 \sqrt{1-\rho^2} \cdot z_{1,1}$

$$x_{2,1} = \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x_{1,1} - \mu_1) + \sigma_2 \sqrt{1-\rho^2} \cdot z_{2,1}$$

⋮

$$x_{1,t+1} = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (x_{2,t} - \mu_2) + \sigma_1 \sqrt{1-\rho^2} \cdot z_{1,t+1}$$

$$x_{2,t+1} = \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x_{1,t+1} - \mu_1) + \sigma_2 \sqrt{1-\rho^2} \cdot z_{2,t+1}$$

3. Repeat.

In-class of size 5000 (burn-in 20%)

Generate MC from Bivariate normal distribution

with mean $\underline{\mu} = \begin{bmatrix} 3 \\ 2 \end{bmatrix}$ and $\Sigma = \begin{bmatrix} 1.5 & 0.6 \\ 0.6 & 1 \end{bmatrix}$

Plot each of x_1 and x_2 MC and histograms

Plot scatterplot of x_1 vs x_2 MC.

$$\begin{aligned}
 \mu_1 &= 3 \\
 \mu_2 &= 2 \\
 \sigma_1 &= \sqrt{1.5} \\
 \sigma_2 &= 1 \\
 \rho &= \frac{0.6}{\sqrt{1.5} \cdot 1}
 \end{aligned}$$

2. ... Mixture

Bivariate Mixture

$$\begin{bmatrix} X \\ Y \end{bmatrix} \sim p \cdot N_2(\underline{\mu}, \underline{\Sigma}_1) + (1-p) \cdot N_2(\underline{\nu}, \underline{\Sigma}_2)$$

$$\underline{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \underline{\nu} = \begin{pmatrix} \nu_1 \\ \nu_2 \end{pmatrix}, \underline{\Sigma}_1 = \begin{bmatrix} a_1 & c_1 \\ c_1 & b_1 \end{bmatrix}, \underline{\Sigma}_2 = \begin{bmatrix} a_2 & c_2 \\ c_2 & b_2 \end{bmatrix}$$

It can be shown that

$$X|Y=y \sim w_y \cdot N\left(\mu_1 + \frac{(y-\mu_2) \cdot c_1}{b_1}, \frac{|\underline{\Sigma}_1|}{b_1}\right) + (1-w_y) \cdot N\left(\nu_1 + \frac{(y-\nu_2) \cdot c_2}{b_2}, \frac{|\underline{\Sigma}_2|}{b_2}\right)$$

determinant

$$Y|X=x \sim w_x \cdot N\left(\mu_2 + \frac{(x-\mu_1) \cdot c_1}{a_1}, \frac{|\underline{\Sigma}_1|}{a_1}\right) + (1-w_x) \cdot N\left(\nu_2 + \frac{(x-\nu_1) \cdot c_2}{a_2}, \frac{|\underline{\Sigma}_2|}{a_2}\right)$$

where

$$w_x = \frac{p \cdot a_1^{-1/2} \cdot \exp\left(-\frac{(x-\mu_1)^2}{2a_1}\right)}{p \cdot a_1^{-1/2} \cdot \exp\left(-\frac{(x-\mu_1)^2}{2a_1}\right) + (1-p) \cdot a_2^{-1/2} \cdot \exp\left(-\frac{(x-\nu_1)^2}{2a_2}\right)}$$

$$w_y = \frac{p \cdot b_1^{-1/2} \cdot \exp\left(-\frac{(y-\mu_2)^2}{2b_1}\right)}{p \cdot b_1^{-1/2} \cdot \exp\left(-\frac{(y-\mu_2)^2}{2b_1}\right) + (1-p) \cdot b_2^{-1/2} \cdot \exp\left(-\frac{(y-\nu_2)^2}{2b_2}\right)}$$

bivariate MC, $\begin{bmatrix} X \\ Y \end{bmatrix}$,

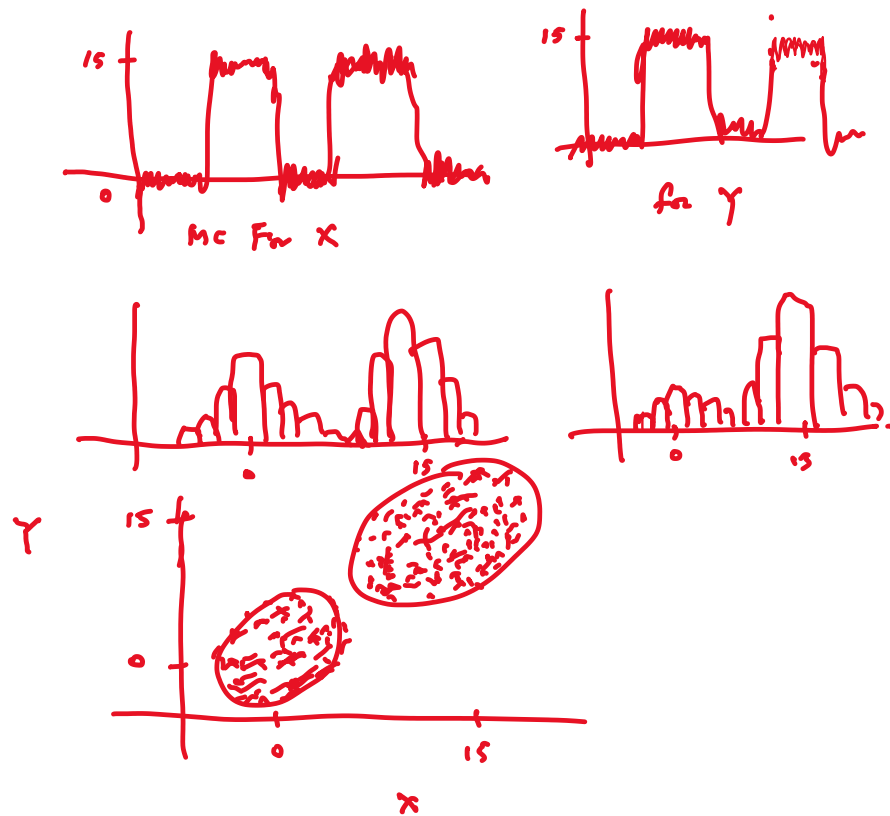
In-class

Use Gibbs sampler to generate \sim of size 5000 from bivariate mixture normal with

$$\underline{\mu} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \underline{\nu} = \begin{pmatrix} 15 \\ 15 \end{pmatrix}, p = 0.3, \underline{\Sigma}_1 = \underline{\Sigma}_2 = \begin{pmatrix} 3 & 1 \\ 1 & 3 \end{pmatrix}$$

After burn 10%, plot each MC and histogram.

Give the scatterplot of X vs Y MCs.



In-class 2 Given $X \sim \text{Exp}(\frac{1}{y})$

$$f(x|y) = y \cdot e^{-xy}, \quad 0 < x < \infty$$

$$f(y|x) = x \cdot e^{-xy}, \quad 0 < y < \infty$$

$\sim \text{Exp}(\frac{1}{x})$

Use Gibbs sampler to generate MC
whose invariant distribution is the marginal
distribution $f(x)$

- Procedure
1. Initial (x_0, y_0)
 x_0 r.s. from $\text{Exp}(1)$
 y_0 " " " "
 2. x_i r.s. from $\text{Exp}(\frac{1}{y_0})$
 y_i r.s. from $\text{Exp}(\frac{1}{x_i})$

$$\begin{aligned}
 & \text{3. } X_2, \text{ r.s. from } \text{Exp}\left(\frac{1}{y_1}\right) \\
 & \quad Y_2 \text{ " " } \text{Exp}\left(\frac{1}{X_2}\right) \\
 & \quad \vdots \\
 & \quad X_k \text{ r.s. from } \text{Exp}\left(\frac{1}{y_{k-1}}\right) \\
 & \quad Y_k \text{ " " } \text{Exp}\left(\frac{1}{X_k}\right)
 \end{aligned}$$

Continue

- Estimate $E(X)$, $V(X)$, $\text{skewness}(X)$, $\text{kurtosis}(X)$
- Histogram of X , plot of MC vs t
- Find $\hat{f}(x)$ at $x = 0.1, 1.8, 3, 9$