

Assignment 3

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Problem 1

a.

$$\begin{aligned}f_Y(y) &= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi} dx \\&= \left[\frac{1}{\pi} x \right]_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \\&= \frac{2}{\pi} \sqrt{1-y^2} \\f_{Y|X}(y|x) &= \frac{f_{X,Y}(x,y)}{f_X(x)} \\&= \frac{\frac{1}{\pi}}{\frac{2}{\pi} \sqrt{1-x^2}} \\&= \frac{1}{2\sqrt{1-x^2}}\end{aligned}$$

b.

$$\begin{aligned}\mathbb{E}[Y|X=x] &= \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{y}{2\sqrt{1-x^2}} dy \\&= 0\end{aligned}$$

Since

$$f(x,y) = \frac{y}{2\sqrt{1-x^2}}$$

is an odd function of y

Problem 2

a.

No it's not.

A valid CDF should have following property

$$\lim_{x,y \rightarrow 1} F(x,y) = 1 \quad (1)$$

$$\lim_{x,y \rightarrow 0} F(x,y) = 0 \quad (2)$$

And by definition of marginal CDF

$$\lim_{x \rightarrow 0, y \rightarrow 1} F(x,y) = 0 \quad (3)$$

$$\lim_{x \rightarrow 1, y \rightarrow 0} F(x,y) = 0 \quad (4)$$

And $F(x,y)$ does not satisfy property (3) and (4)

b.

By definition, we can derive marginal CDF by

$$F(1,y) = F_Y(y) = \frac{1+y}{2}, y \in [0,1]$$

$$F(x,1) = F_X(x) = \frac{1+x}{2}, x \in [0,1]$$

By definition of marginal pdf

$$\frac{\partial F(x,y)}{\partial x} = f_X(x) = \frac{1}{2}, x \in [0,1]$$

$$\frac{\partial F(x,y)}{\partial y} = f_Y(y) = \frac{1}{2}, y \in [0,1]$$

c.

No, since

$$\frac{\partial^2 F(x,y)}{\partial y \partial x} = 0$$

which is not a valid joint PDF

Problem 3

a.

$$\begin{aligned}P(X > Y) &= P(X = 2, Y = 1) + P(X = 3, Y \leq 2) + P(X = 4, Y \leq 3) + \dots + P(X = 12, Y \leq 6) \\&= \frac{1 \times 1 + 2 \times 2 + 3 \times 3 + 4 \times 4 + 5 \times 5 + 6 \times 6 + (6 + 5 + 4 + 3 + 2 + 1) \times 6}{6 \times 6 \times 6} \\&= \frac{181}{216}\end{aligned}$$

b.

Let the sum of Ray's daughter throwing $2i$ dices be X_i and him throwing i dices Y_i

$$\text{Let } \bar{W}_i = \bar{X}_i - \bar{Y}_i = \frac{X_i - Y_i}{i}$$

Then we know that

$$\begin{aligned}\mathbb{E}[\bar{W}_i] &= \mathbb{E}[\bar{X}_i] - \mathbb{E}[\bar{Y}_i] \\&= 7 - 3.5 \\&= 3.5 \\Var[\bar{W}_i] &= Var[\bar{X}_i] + Var[\bar{Y}_i] = \frac{35}{12} + \frac{70}{12} = \frac{35}{4}\end{aligned}$$

By Central limit theorem,

$$\begin{aligned}&\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \sigma^2) \\&\Rightarrow \sqrt{n}(\bar{W}_n - 3.5) \xrightarrow{d} N(0, \frac{35}{4}) \\&\Rightarrow \frac{\sqrt{n}(\bar{W}_n - 3.5)}{\sqrt{\frac{35}{4}}} \xrightarrow{d} N(0, 1) \\&\Rightarrow P(\bar{W}_n > 0) = \Phi\left(\frac{-3.5\sqrt{n}}{\sqrt{\frac{35}{4}}}\right)\end{aligned}$$

if n is large enough

Problem 4

a.

$$\begin{aligned}\mathbb{E}[X] &= \frac{1}{2} \times 3 + \frac{1}{2} \times \frac{1}{3} = \frac{5}{3} \\ \mathbb{E}[\ln(X)] &= 1/2 \times \ln(3) + 1/2 \times \ln\left(\frac{1}{3}\right) = 0\end{aligned}$$

b.

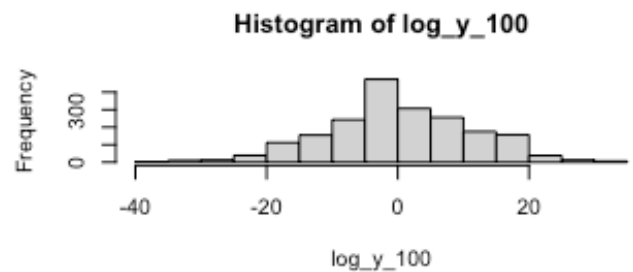
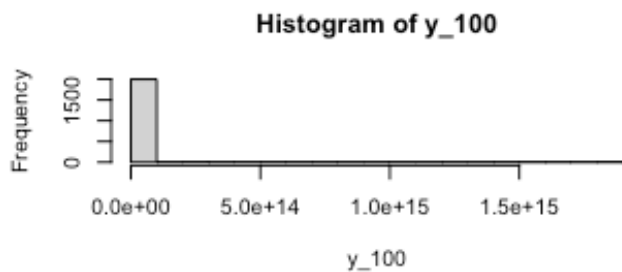
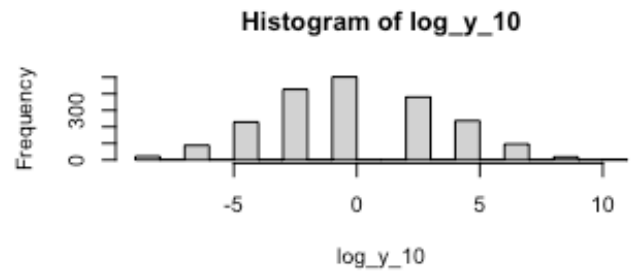
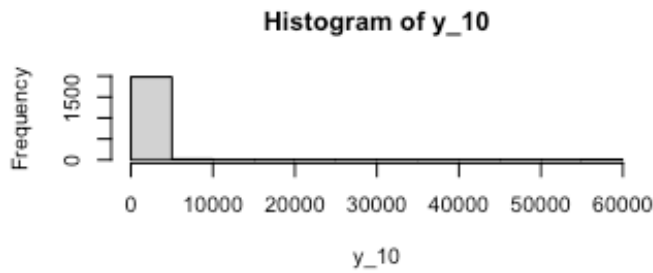
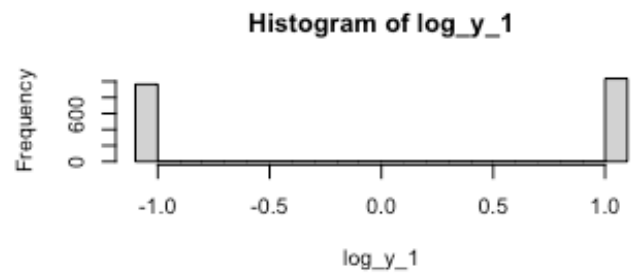
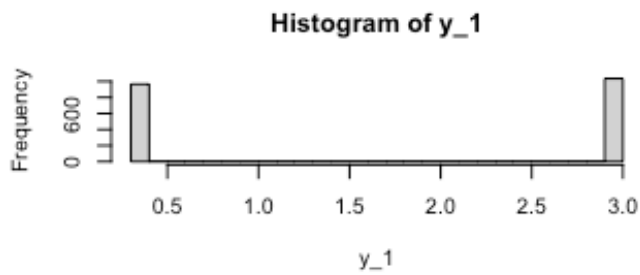
Since $\{X_i\}_{i=1}^n$ are i.i.d random variables

$$\mathbb{E}[Y_n] = \mathbb{E}\left[\prod_{i=1}^n X_i\right] = \prod_{i=1}^n \mathbb{E}[X_i] = \left(\frac{5}{3}\right)^n$$

$$\mathbb{E}[\ln(Y_n)] = \mathbb{E}\left[\sum_{i=1}^n \ln X_i\right] = \sum_{i=1}^n \mathbb{E}[\ln X_i] = 0$$

c.

```
par(mfcol=c(3,2))
y_n <- function(n){
  # do n trials(Xn) with flipping only once and count success time (0 or 1)
  return(replicate(2000,prod(8/3*rbinom(n,1,0.5)+1/3), simplify = "array"))
}
log_y_n <- function(n){
  return(replicate(2000,sum(log(8/3*rbinom(n,1,0.5)+1/3)), simplify = "array"))
}
y_1 <- y_n(1)
y_10 <- y_n(10)
y_100 <- y_n(100)
log_y_1 <- log_y_n(1)
log_y_10 <- log_y_n(10)
log_y_100 <- log_y_n(100)
hist(y_1,breaks = 20)
hist(y_10,breaks= 20)
hist(y_100,breaks= 20)
hist(log_y_1,breaks = 20)
hist(log_y_10,breaks = 20)
hist(log_y_100,breaks = 20)
```



Problem 5

a.

Let $\mathbb{E}[X] = \mu$ and $\mathbb{E}[Y] = \nu$

$$\begin{aligned} \mathbb{E} \left[\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}) \right] &= \mathbb{E} \left[\sum_{i=1}^n \left[(X_i - \mu) + (\mu - \bar{X}) \right] \left[(Y_i - \nu) + (\nu - \bar{Y}) \right] \right] \\ &= \mathbb{E} \left[\sum_{i=1}^n \left[(X_i - \mu)(Y_i - \mu) + (X_i - \mu)(\nu - \bar{Y}) + (\mu - \bar{X})(Y_i - \nu) + (\mu - \bar{X})(\nu - \bar{Y}) \right] \right] \\ &= ncov(X, Y) - cov(X, Y) - cov(X, Y) + cov(X, Y) \\ &= (n - 1)cov(X, Y) \end{aligned}$$

Therefore $\hat{\sigma}_{XY} = \frac{1}{n} \mathbb{E} \left[\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}) \right] = \frac{n-1}{n} cov(X, Y)$

which is a downward biased estimator

b.

$$\begin{aligned}\hat{\sigma}_{XY} &= \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}) \\ &= \frac{1}{n} \sum_{i=1}^n X_i Y_i - \bar{X}\bar{Y}\end{aligned}$$

By WLLN

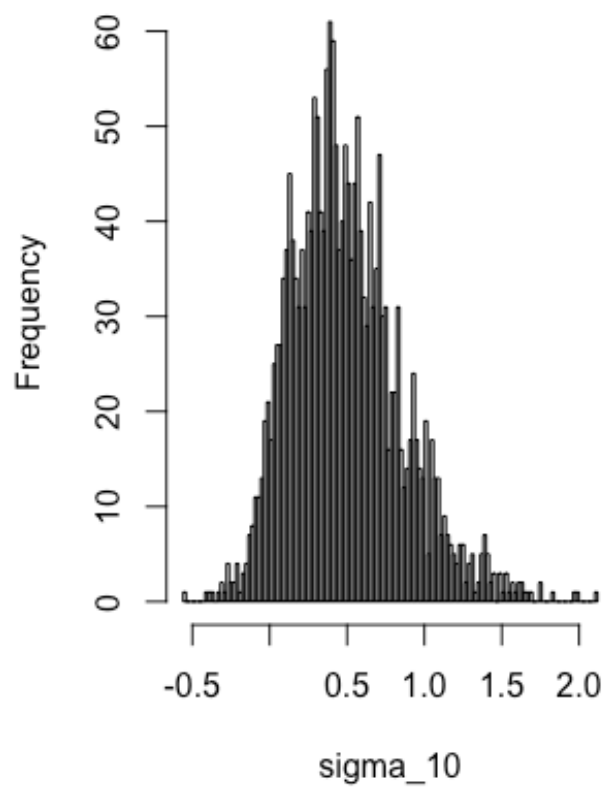
$$\frac{1}{n} \sum_{i=1}^n X_i Y_i - \bar{X}\bar{Y} \xrightarrow[W.L.L.N]{p} \frac{1}{n} \sum_{i=1}^n X_i Y_i - \mu\nu = \text{cov}(X, Y)$$

Therefore, $\sigma_{\hat{XY}}$ is a consistent estimator

c.

```
library("MASS")
cov_mat <- matrix(c(1,0.5,0.5,1),2,2)
mu <- c(0,0)
sigma_n <- function(n){
  #cov return a covariance matrix
  return(replicate(2000,cov(mvrnorm(n,mu,cov_mat))[1,2],simplify = "array"))
}
sigma_10 <- sigma_n(10)
sigma_100 <- sigma_n(100)
par(mfrow=c(1,2))
hist(sigma_10,breaks=100)
hist(sigma_100,breaks=100)
```

Histogram of sigma_10



Histogram of sigma_100

