

Problem Set 6: Solution**Part One: Hand-Written Exercise**

1. For the probit model, its NLS estimator is obtained by solving the F.O.C.:

$$\frac{\partial}{\partial \boldsymbol{\theta}} \sum_{i=1}^n \left[y_i - \Phi(\mathbf{x}_i' \boldsymbol{\theta}) \right]^2 = -2 \sum_{i=1}^n \left[y_i - \Phi(\mathbf{x}_i' \boldsymbol{\theta}) \right] \phi(\mathbf{x}_i' \boldsymbol{\theta}) \mathbf{x}_i = \mathbf{0},$$

which is not the same as the F.O.C of the ML method:

$$\sum_{i=1}^n \frac{y_i - \Phi(x_i' \boldsymbol{\theta})}{\Phi(x_i' \boldsymbol{\theta}) [1 - \Phi(x_i' \boldsymbol{\theta})]} \phi(\mathbf{x}_i' \boldsymbol{\theta}) \mathbf{x}_i = \mathbf{0}.$$

For the logit model, its NLS estimator is obtained by solving the F.O.C.:

$$\frac{\partial}{\partial \boldsymbol{\theta}} \sum_{i=1}^n \left[y_i - G(\mathbf{x}_i' \boldsymbol{\theta}) \right]^2 = -2 \sum_{i=1}^n \left[y_i - G(\mathbf{x}_i' \boldsymbol{\theta}) \right] G'(\mathbf{x}_i' \boldsymbol{\theta}) \mathbf{x}_i = \mathbf{0},$$

which is not the same as the F.O.C of the ML method:

$$\sum_{i=1}^n \left[y_i - G(\mathbf{x}_i' \boldsymbol{\theta}) \right] \mathbf{x}_i = \mathbf{0}. \blacksquare$$

2. First, we have

$$\begin{aligned} \mathbf{H}(\boldsymbol{\theta}_0) &= \mathbb{E}[\nabla^2 \ln \ell(\boldsymbol{\theta}_0)] \\ &= \mathbb{E} \left(\nabla \left[\frac{[y_i - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)] \phi(\mathbf{x}_i' \boldsymbol{\theta}_0)}{\Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)]} \mathbf{x}_i \right] \right) \\ &= \mathbb{E} \left(\frac{\phi'(\mathbf{x}_i' \boldsymbol{\theta}_0) \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) y_i - \phi'(\mathbf{x}_i' \boldsymbol{\theta}_0) \Phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0) - \phi'(\mathbf{x}_i' \boldsymbol{\theta}_0) \Phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0) y_i + \phi'(\mathbf{x}_i' \boldsymbol{\theta}_0) \Phi^3(\mathbf{x}_i' \boldsymbol{\theta}_0)}{(\Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)])^2} \mathbf{x}_i \mathbf{x}_i' \right) \\ &\quad - \mathbb{E} \left(\frac{[y_i - 2y_i \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) + \Phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0)] \phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0)}{(\Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)])^2} \mathbf{x}_i \mathbf{x}_i' \right) \\ &= 0 - \mathbb{E} \left(\frac{[y_i - 2y_i \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) + \Phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0)] \phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0)}{(\Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)])^2} \mathbf{x}_i \mathbf{x}_i' \right) \\ &= -\mathbb{E} \left(\frac{\phi^2(\mathbf{x}_i' \boldsymbol{\theta}_0)}{\Phi(\mathbf{x}_i' \boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i' \boldsymbol{\theta}_0)]} \mathbf{x}_i \mathbf{x}_i' \right), \end{aligned}$$

where that last two equations are by the law of iterated expectation and the fact that $\mathbb{E}(y_i|\mathbf{x}_i) = \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)$.

Now, for the information matrix $\mathbf{B}(\boldsymbol{\theta}_0)$, we have:

$$\begin{aligned}\mathbf{B}(\boldsymbol{\theta}_0) &= \mathbb{E} \left(\frac{1}{N} (\nabla L_n(\boldsymbol{\theta}_0)) (\nabla L_n(\boldsymbol{\theta}_0))' \right) \\ &= \mathbb{E} \left(\frac{[y_i - \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)]^2 \phi^2(\mathbf{x}_i'\boldsymbol{\theta}_0)}{(\Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)])^2} \mathbf{x}_i \mathbf{x}_i' \right) \quad \text{when } (y_i, \mathbf{x}_i)' \text{ are iid data.} \\ &= \mathbb{E} \left(\frac{[y_i^2 - 2y_i \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) + \Phi^2(\mathbf{x}_i'\boldsymbol{\theta}_0)] \phi^2(\mathbf{x}_i'\boldsymbol{\theta}_0)}{(\Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)])^2} \mathbf{x}_i \mathbf{x}_i' \right) \\ &= \mathbb{E} \left(\frac{\phi^2(\mathbf{x}_i'\boldsymbol{\theta}_0)}{\Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) [1 - \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)]} \mathbf{x}_i \mathbf{x}_i' \right),\end{aligned}$$

where the last equation is due to the fact that:

$$\text{Var}(y_i|\mathbf{x}_i) = \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) (1 - \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0)),$$

and

$$\begin{aligned}\text{Var}(y_i|\mathbf{x}_i) &= \mathbb{E}(y_i^2|\mathbf{x}_i) - \mathbb{E}^2(y_i|\mathbf{x}_i) \\ &= \mathbb{E}(y_i^2|\mathbf{x}_i) - \Phi^2(\mathbf{x}_i'\boldsymbol{\theta}_0),\end{aligned}$$

which leads to the fact

$$\mathbb{E}(y_i^2|\mathbf{x}_i) = \Phi(\mathbf{x}_i'\boldsymbol{\theta}_0) = \mathbb{E}(y_i|\mathbf{x}_i).$$

It follows that the information equality holds: $\mathbf{H}(\boldsymbol{\theta}_0) + \mathbf{B}(\boldsymbol{\theta}_0) = 0$. ■

3. (a)

$$\begin{aligned}\therefore \frac{(n-k-1)\hat{\sigma}^2}{\sigma_o^2} | \mathbf{X} &\sim \chi^2(n-k-1) \\ \therefore \text{var}\left(\frac{(n-k-1)\hat{\sigma}^2}{\sigma_o^2} | \mathbf{X}\right) &= 2(n-k-1) \\ \Rightarrow \text{var}(\hat{\sigma}^2 | \mathbf{X}) &= \left(\frac{1}{n-k-1}\right) 2\sigma_o^4 \\ \Rightarrow \text{var}(\sqrt{n}\hat{\sigma}^2 | \mathbf{X}) &= \left(\frac{n}{n-k-1}\right) 2\sigma_o^4 > 2\sigma_o^4.\end{aligned}$$

(b)

$$\begin{aligned}\because \tilde{\sigma}^2 &= \frac{n-k-1}{n} \hat{\sigma}_o^2 \\ \therefore \text{var}(\tilde{\sigma}^2 | \mathbf{X}) &= \left(\frac{n-k-1}{n} \right) \frac{2\sigma_o^4}{n} \\ \Rightarrow \text{var}(\sqrt{n}\tilde{\sigma}^2 | \mathbf{X}) &= \left(\frac{n-k-1}{n} \right) 2\sigma_o^4 < 2\sigma_o^4.\end{aligned}$$

(c) Yes. In general, $\text{var}(\sqrt{n}\tilde{\sigma}^2 | \mathbf{X}) < 2\sigma_o^4 < \text{var}(\sqrt{n}\hat{\sigma}^2 | \mathbf{X})$.

(d) No. Since $\tilde{\sigma}^2$ is biased, the Cramér-Rao lower bound doesn't apply to $\tilde{\sigma}^2$. Also, note that $\hat{\sigma}^2$ doesn't reach the Cramér-Rao lower bound.