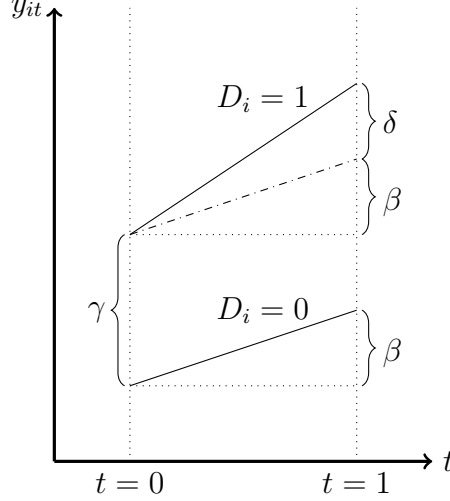


Problem Set 4: Solution**Part One: Hand-Written Exercise**

1. .



2. (a)

$$\begin{aligned}\tilde{\beta}_1 &= \frac{\sum (x_{i1} - \bar{x}_1)y_i}{\sum (x_{i1} - \bar{x}_1)^2} = \frac{\sum (x_{i1} - \bar{x}_1)(\beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + u_i)}{\sum (x_{i1} - \bar{x}_1)^2} \\ &= \beta_1 + \beta_2 \frac{\sum (x_{i1} - \bar{x}_1)x_{i2}}{\sum (x_{i1} - \bar{x}_1)^2} + \frac{\sum (x_{i1} - \bar{x}_1)u_i}{\sum (x_{i1} - \bar{x}_1)^2}.\end{aligned}$$

By the given assumptions, we have:

$$\begin{aligned}\frac{1}{n} \sum (x_{i1} - \bar{x}_1)u_i &\xrightarrow{p} \mathbb{E}(x_1 u) - \mu_{x_1} \mathbb{E}(u) = 0 \\ \frac{1}{n} \sum (x_{i1} - \bar{x}_1)^2 &\xrightarrow{p} \sigma_{x_1}^2 \\ \frac{1}{n} \sum (x_{i1} - \bar{x}_1)x_{i2} &= \frac{1}{n} \sum (x_{i1} - \bar{x}_1)(x_{i2} - \bar{x}_2) \xrightarrow{p} \sigma_{x_1 x_2}.\end{aligned}$$

Therefore, we have:

$$\tilde{\beta}_1 \xrightarrow{p} \beta_1 + \beta_2 \frac{\sigma_{x_1 x_2}}{\sigma_{x_1}^2}.$$

 $\tilde{\beta}_1$ is consistent for β_1 only when $\sigma_{x_1 x_2} = 0$, otherwise $\tilde{\beta}_1$ is not consistent.

(b) Since $\sigma_{x_1 x_2} > 0$, then if $\beta_2 > 0$, $\tilde{\beta}_1$ overestimates β_1 by $\beta_2 \frac{\sigma_{x_1 x_2}}{\sigma_{x_1}^2}$ as $n \rightarrow \infty$. On the other hand, if $\beta_2 < 0$, then $\tilde{\beta}_1$ underestimates β_1 by $-\beta_2 \frac{\sigma_{x_1 x_2}}{\sigma_{x_1}^2}$ as $n \rightarrow \infty$.

(c) $\sqrt{n}(\tilde{\beta}_1 - \beta_1)$ does not follow any distributions. Since $\sigma_{x_1x_2} > 0$, thus

$$\tilde{\beta} - \beta_1 \xrightarrow{p} \beta_2 \frac{\sigma_{x_1x_2}}{\sigma_{x_1}^2} \neq 0.$$

So $\sqrt{n}(\tilde{\beta}_1 - \beta_1)$ clearly diverges as $n \rightarrow \infty$.

3. (a) False. Recall that $\hat{\sigma}_{OLS}^2 = \frac{1}{n-k-1} \sum \hat{e}_i^2$ is an unbiased estimator of σ^2 . Hence, for $s \in \mathbb{R}$, $\hat{\sigma}^2(s) = \frac{1}{n-s} \sum \hat{e}_i^2$ is a biased estimator as long as $s \neq k+1$. However, $\forall s \in \mathbb{R}$, $\hat{\sigma}^2(s)$ is consistent for σ^2 .

(b) False. Consider a case where we have the data x_i , $i = 1, \dots, n$, and the true population mean $\mu_x = 0$. Our estimator is designed as:

$$\hat{\mu}_n = \begin{cases} -1 & \text{with } p = 1/2 \\ 1 & \text{with } p = 1/2 \end{cases}.$$

This estimator, although completely disregards the data x_i , is still unbiased. It is, however, not consistent ($\lim_{n \rightarrow \infty} \hat{\mu}_n \neq 0$).

(c) False. $\hat{\beta}$ could be consistent even when \mathbf{V} doesn't exist.