

The multilevel random coefficient model

Lecture 19

Multivariate Statistics

Psychology 613 – Spring 2022

The logic: Moving from *sets* to *distributions* of coefficients

Our groups → A random sample of possible groups

Our lines → A random sample of possible lines (i.e., slopes and intercepts)

So, rather than characterize the particular lines, our goal is to characterize the *distribution* of lines

- “Average” intercept and some measure of the variability around that intercept
- “Average” slope and the variability around that slope

Random coefficient model: Multiple equation form

Within-Group (L1) Model: $Y_{ij} = \beta_{0j} + \beta_{1j} X_{ij} + e_{ij}$

Coefficients of the within-group model then serve as criteria in between-group (L2) models:

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

Each β is a function of an “average” coefficient and random error (variation)

(Notation)

$$Y_{ij} = \beta_{0j} + \beta_{1j} X_{ij} + e_{ij}$$

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

← “0”: equation for β_0

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

← “1”: equation for β_1

- ij subscripts indicate individuals within groups
- Double subscripts on gamma coefficients are positional: the 1st indicates equation, the 2nd indicates the position within the equation

Random coefficient model:

Single equation form

Start with multiple equation form:

$$\text{L1: } Y_{ij} = \beta_{0j} + \beta_{1j} X_{ij} + e_{ij}$$

$$\text{L2: } \beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

$$\begin{aligned} \text{Combine: } Y_{ij} &= \gamma_{00} + u_{0j} + (\gamma_{10} + u_{1j}) X_{ij} + e_{ij} \\ Y_{ij} &= \gamma_{00} + u_{0j} + \gamma_{10} X_{ij} + u_{1j} X_{ij} + e_{ij} \end{aligned}$$

Rearrange: $Y_{ij} = \gamma_{00} + \gamma_{10} X_{ij} + (u_{0j} + u_{1j} X_{ij} + e_{ij})$
→ *Looks like a standard regression with a complex error term*

Estimates

$$Y_{ij} = \gamma_{00} + \gamma_{10} X_{ij} + (u_{0j} + u_{1j} X_{ij} + e_{ij})$$

- (1) “average” intercept - γ_{00}
- (2) “average” slope - γ_{10}
- (3) variability of intercepts - $\text{Var}(u_{0j})$
- (4) variability of slopes - $\text{Var}(u_{1j})$
- (5) within-group variability - $\text{Var}(e_{ij})$

Types of parameters / estimates

Fixed effects: γ_s (population parameters)
 g_s (sample estimates)

Similar to unstandardized regression params

Variance components: co/vars of error terms
L1: $\text{var}(e_{ij}) = \sigma^2$ (population parameter)
 s^2 (sample estimate)

Similar to regression SS-error

Level 2 variance components

$$\begin{bmatrix} \text{Var}(u_{0j}) & \text{Cov}(u_{1j}, u_{0j}) \\ \text{Cov}(u_{0j}, u_{1j}) & \text{Var}(u_{1j}) \end{bmatrix} = \begin{bmatrix} \tau_{00} & \tau_{01} \\ \tau_{10} & \tau_{11} \end{bmatrix} = \begin{bmatrix} t_{00} & t_{10} \\ t_{01} & t_{11} \end{bmatrix}$$

“True” population values
(unknowable) Sample estimates

These represent the variance of the slopes and intercepts around the gammas

NOTE: $t_{10} = t_{01}$ and $\tau_{10} = \tau_{01}$, so there are only **3** variance components to estimate at the second level

β s

$$Y_{ij} = \gamma_{00} + \gamma_{10} X_{ij} + (u_{0j} + u_{1j} X_{ij} + e_{ij})$$

Notice that single equation form contains no β 's!

Need not calculate β s in order to estimate fixed effects and variance components!

Allows estimation in presence of rank deficiency

If desired, can obtain post hoc estimates of β 's

Example from NELS-88 data

$$\text{L1: } \text{MathAch}_{ij} = \beta_{0j} + \beta_{1j} \text{Homework}_{ij} + e_{ij}$$

$$\text{L2: } \beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

The so-called *random effects regression* model:

- Normal intercept (b_0) and slope (b_1) but those terms vary randomly across groups

Example from NELS-88 data

Predicted Math = 44.77^{***} + 2.04 Homework

$$\sigma^2 = 43.07 (3.93)^{***}$$

$$t_{00} = 69.24 (34.97)^*$$

$$t_{01} = -31.72 (18.14)^+ \quad t_{11} = 22.43 (11.49)^+$$

→ *Significant variability to be explained in the intercepts across schools*

Example from NELS-88 data

```
Formula: mathscore ~ timeonmath + (timeonmath | Schoolid)
```

```
REML criterion at convergence: 1764
```

```
Scaled residuals:
```

Min	1Q	Median	3Q	Max
-2.5110	-0.5357	0.0175	0.6121	2.5708

```
Random effects:
```

Groups	Name	Variance	Std.Dev.	Corr
Schoolid	(Intercept)	69.30	8.325	
	timeonmath	22.45	4.738	-0.81
	Residual	43.07	6.563	

Number of obs: 260, groups: Schoolid, 10

Annotations: t_{00} points to 69.30, t_{11} points to 22.45, t_{01} points to -0.81, s^2 points to 43.07, g_{00} points to 69.30.

```
Fixed effects:
```

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	44.771	2.744	9.227	16.318	4.06e-08 ***
timeonmath	2.040	1.554	8.785	1.313	0.222

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
Correlation of Fixed Effects:
```

(Intr)
timeonmath -0.804

```
> |
```

g_{10}

Example from NELS-88 data

Predicted Math = $44.77^{***} + 2.04$ Homework

$$\sigma^2 = 43.07 (3.93)^{***}$$

$$t_{00} = 69.24 (34.97)^*$$

$$t_{01} = -31.72 (18.14)^+ \quad t_{11} = 22.43 (11.49)^+$$

$$\begin{aligned} ICC &= \textit{between} / (\textit{between} + \textit{within}) \\ &= 69.24 / (69.24 + 43.07) = 61.6\% \end{aligned}$$

Adding L2 predictors

Multiple equation form:

$$\text{L1: } Y_{ij} = \beta_{0j} + \beta_{1j} X_{ij} + e_{ij}$$

$$\text{L2: } \beta_{0j} = \gamma_{00} + \gamma_{01} W_j + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11} W_j + u_{1j}$$

In the combined (single) equation, this is:

$$Y_{ij} = \gamma_{00} + \gamma_{01} W_j + \gamma_{10} X_{ij} + \gamma_{11} W_j X_{ij} \dots \\ + (e_{ij} + u_{0j} + u_{1j} X_{ij})$$

Fixed coefficients (w/ L2 predictors)

γ_{00} – an overall intercept

γ_{01} – the main effect of the level 2 (W_j) variable

γ_{10} – the main effect of the level 1 (X_{ij}) variable

γ_{11} – the *cross-level interaction* (the effect of the W_j variable on the relationship between X_{ij} and Y_{ij})

Variance components (with L2 predictors)

With a level 2 (W_j) predictor now in the model, our distribution of intercepts/slopes is conditional, i.e., expected value depends on the value of W_j for a particular group

L2 error terms (u 's) now represent *residuals*, after controlling for W_j (no longer total variability in intercepts/slopes, but variability remaining after adjusting for W_j)

Example from NELS-88

Predicted Math = $59.21^{***} + 1.09 \text{ Homework ...}$
 $-15.97^+ \text{ public} + .95 \text{ HW}^* \text{Public}$

$$\sigma^2 = 42.96 (3.91)^{***}$$

$$t_{00} = 51.81 (28.64)^+$$

$$t_{01} = -36.70 (20.07)^+ \quad t_{11} = 27.26 (14.59)^+$$

→ *There is no longer significant L2 variability in the intercepts to be explained!*

Example from NELS-88

g_{00}

g_{01}

Estimates of Fixed Effects^a

Parameter	Estimate	Std. Error	df	t	Sig.	95% Confidence Interval	
						Lower Bound	Upper Bound
Intercept	59.210218	7.407272	7.323	7.994	.000	41.850398	76.570038
[schooltype=1]	-15.965674	7.829124	7.402	-2.039	.079	-34.276521	2.345172
[schooltype=4]	0 ^b	0
timeonmath	1.094640	5.243213	7.093	.209	.840	-11.270805	13.460085
[schooltype=1] * timeonmath	.951160	5.542392	7.172	.172	.868	-12.090995	13.993315
[schooltype=4] * timeonmath	0 ^b	0

a. Dependent Variable: Math score.

b. This parameter is set to zero because it is redundant.

g_{10}

g_{11}

Example from NELS-88 (R)

g_{00}

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)	
(Intercept)	43.2445	2.5353	8.1316	17.057	1.18e-07	***
timeonmath	2.0458	1.7963	7.9063	1.139	0.2881	
schooltype4	15.9657	7.8291	7.4024	2.039	0.0786	.
timeonmath:schooltype4	-0.9512	5.5424	7.1722	-0.172	0.8685	

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

g_{10}

g_{11}

g_{01}

Example from NELS-88

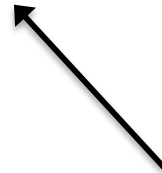
	Private schools (Public = 0)	Public schools (Public = 1)
E(Math) for Homework = 0	59.21	$59.21 - 15.97 = 43.26$
Homework – Math slope	1.09	$1.09 + 0.95 = 2.04$

Example from NELS-88

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Schoolid	(Intercept)	51.84	7.200	
	timeonmath	27.27	5.222	-0.98
Residual		42.96	6.554	

Number of obs: 260, groups: Schoolid, 10



Check the variance of the intercepts (t_{00}) and the slopes (t_{11})...

Assumptions of the multilevel random coefficient models

$$\text{Cov}(X_{ij}, e_{ij}) = 0$$

$$\text{Cov}(W_j, e_{ij}) = 0$$

$$\text{Cov}(W_j, u_j) = 0$$

$$\text{Cov}(X_{ij}, u_j) = 0$$

$$\text{Cov}(e_{ij}, u_j) = 0$$

$e_{ij} \sim \text{iid } N(0, \sigma^2)$ within each group j

$$\begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} \sim \text{iid}, N \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} t_{00} & t_{10} \\ t_{01} & t_{11} \end{bmatrix} \right)$$

Full multilevel random coefficient model: Recap

Multi-equation form:

$$\text{L1: } Y_{ij} = \beta_{0j} + \beta_{1j} X_{ij} + e_{ij}$$

$$\text{L2: } \beta_{0j} = \gamma_{00} + \gamma_{01} W_j + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11} W_j + u_{1j}$$

Single equation form:

$$Y_{ij} = \gamma_{00} + \gamma_{01} W_j + \gamma_{10} X_{ij} + \gamma_{11} W_j X_{ij} + (e_{ij} + u_{0j} + u_{1j} X_{ij})$$

- Predictors at both levels
- All L1 coefficients (intercept and slope(s)) treated as random
- **Simplifying the models gives several useful submodels**

Random effects ANOVA model

Null model / empty model <- ***USE THIS TO GET AN ICC***

No predictors, only a random intercept

$$\text{L1: } Y_{ij} = b_{0j} + e_{ij}$$

$$\text{L2: } b_{0j} = g_{00} + u_{0j}$$

$$\text{Single equation: } Y_{ij} = g_{00} + (u_{0j} + e_{ij})$$

g_{00} represents the overall level or “average” of the dependent variable

Variance components partition variance into within groups (σ^2) and between groups (τ_{00}) portions

μ_{00}

$$Y_{ij} = 48.86^{***} + (u_{0j} + e_{ij})$$

Estimates of Fixed Effects^a

Parameter	Estimate	Std. Error	df	t	Sig.	95% Confidence Interval	
						Lower Bound	Upper Bound
Intercept	48.861323	1.927392	9.631	25.351	.000	44.544428	53.178217

a. Dependent Variable: Math score.

Covariance Parameters

Estimates of Covariance Parameters^a

Parameter	Estimate	Std. Error	Wald Z	Sig.	95% Confidence Interval	
					Lower Bound	Upper Bound
Residual	72.255843	6.455032	11.194	.000	60.649918	86.082671
Intercept [subject = Schoolid] Variance	34.011005	16.920147	2.010	.044	12.827880	90.174558

a. Dependent Variable: Math score.

$\text{var}(u_{0j})$

$\text{Var}(e_{ij})$

Calculating the ICC

$$ICC = \tau_{00} / (\tau_{00} + \sigma^2)$$

NELS-88 Example: DV = math achievement

Covariance Parameter Estimates

Cov Parm	Subject	Estimate	SE	Z-score	P(z)
Intercept	school	34.0116.92	2.01	0.0222	
Residual		72.266.46	11.19	<.0001	

$$\rightarrow ICC = 34.01 / (34.01 + 72.26) = 0.32$$

One-way ANCOVA with random effects

Add non-random L1 variable X_{ij}

Allows intercepts to differ randomly by assumes a common slope

L1: $Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij}$ Multi-equation form

L2: $b_{0j} = g_{00} + u_{0j}$

$b_{1j} = g_{10}$ (No random error term)

$$Y_{ij} = g_{00} + g_{10}X_{ij} + (u_{0j} + e_{ij})$$

Single equation form

$$Y_{ij} = 44.98 + 2.21 \cdot \text{timeonmath} + (u_{0j} + e_{ij})$$

Formula: mathscore ~ timeonmath + (1 | Schoolid)

REML criterion at convergence: 1839.9

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.6060	-0.6872	-0.0244	0.5983	3.3770

Random effects:

Groups	Name	Variance	Std.Dev.
Schoolid	(Intercept)	25.22	5.022
	Residual	64.52	8.033

Number of obs: 260, groups: Schoolid, 10

$\text{var}(u_{0j})$

$\text{var}(e_{ij})$

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	44.982	1.803	12.670	24.949	3.76e-12 ***
timeonmath	2.207	0.379	257.150	5.823	1.72e-08 ***

g_{10}

g_{00}

Random coefficients regression (RCR)

Add random error to X_{ij} slope

L1: $Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij}$ Multi-equation form

L2: $b_{0j} = g_{00} + u_{0j}$

$b_{1j} = g_{10} + u_{1j}$ (random error term)

$Y_{ij} = g_{00} + g_{10}X_{ij} + (u_{0j} + u_{1j}X_{ij} + e_{ij})$ Single equation form

Allows both intercept and slope to vary randomly

Does not yet predict L2 variation in intercept or slope

$$Y_{ij} = 44.77 + 2.04 * \text{math} + (u_{0j} + u_{1j} * \text{math} + e_{ij})$$

Formula: `mathscore ~ timeonmath + (timeonmath | Schoolid)`

REML criterion at convergence: 1764

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.5110	-0.5357	0.0175	0.6121	2.5708

Random effects:

Groups	Name	Variance	Std.Dev	Corr
Schoolid	(Intercept)	69.30	8.325	
	timeonmath	22.45	4.738	-0.81
	Residual	43.07	6.563	

Number of obs: 260, groups: Schoolid, 10

$\text{var}(u_{0j})$

$\text{var}(u_{1j})$

$\text{var}(e_{ij})$

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	44.771	2.744	9.227	16.318	4.06e-08 ***
timeonmath	2.040	1.554	8.785	1.313	0.222

g_{10}

g_{00}

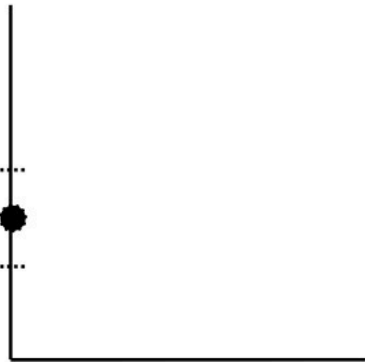
Testing for significant variance

- The significance tests of the variance components are suspect because variances are not normally distributed
- BUT, can create nested models that differ by only one variance term

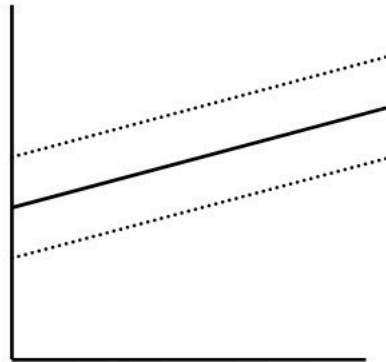
```
> anova(model1,model2)
refitting model(s) with ML (instead of REML)
Data: NULL
Models:
object: mathscore ~ timeonmath + (1 | Schoolid)
..1: mathscore ~ timeonmath + (timeonmath | Schoolid)
      Df    AIC    BIC logLik deviance Chisq Chi Df Pr(>Chisq)
object  4 1850.7 1864.9 -921.33   1842.7
..1      6 1781.4 1802.8 -884.69   1769.4 73.272      2 < 2.2e-16 ***
---
```

Graphs

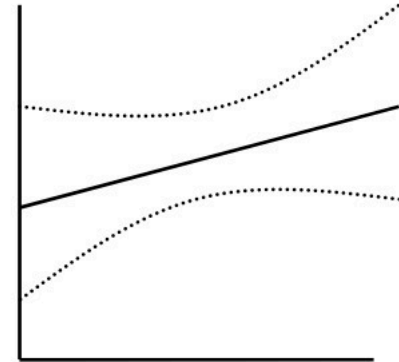
Random Coefficient Models:



ANOVA



ANCOVA



RCR

Add main effect of L2 variable

Add W_j to intercept equation (no cross-level interaction)

$$\begin{array}{ll} \text{L1:} & Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij} \\ \text{L2:} & b_{0j} = g_{00} + g_{01}W_j + u_{0j} \\ & b_{1j} = g_{10} + u_{1j} \end{array} \quad \begin{array}{l} \text{Multi-equation form} \\ \\ \text{(random error term)} \end{array}$$

$$Y_{ij} = g_{00} + g_{10}X_{ij} + g_{01}W_j + (u_{0j} + u_{1j}X_{ij} + e_{ij})$$

Single equation form

“Main effects” of L1 and L2 variables (g_{10} and g_{01} , respectively)

$$Y_{ij} = 38.5 + 1.9 \cdot \text{math} + 4.9 \cdot \text{type} + (u_{0j} + u_{1j}X_{ij} + e_{ij})$$

Formula: `mathscore ~ timeonmath + schooltype + (timeonmath | Schoolid)`

REML criterion at convergence: 1746.2

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.63744	-0.56207	-0.05469	0.64216	2.58580

Random effects:

Groups	Name	Variance	Std. Dev.	Corr
Schoolid	(Intercept)	45.84	6.771	
	timeonmath	23.96	4.895	-0.97
	Residual	42.96	6.554	

Number of obs: 260, groups: Schoolid, 10

$\text{var}(u_{0j})$

$\text{var}(u_{1j})$

$\text{var}(e_{ij})$

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	38.4930	2.4558	11.7740	15.675	3.01e-09 ***
timeonmath	1.9524	1.5990	8.7760	1.221	0.253869
schooltype	4.8869	0.7025	5.9730	6.957	0.000447 ***

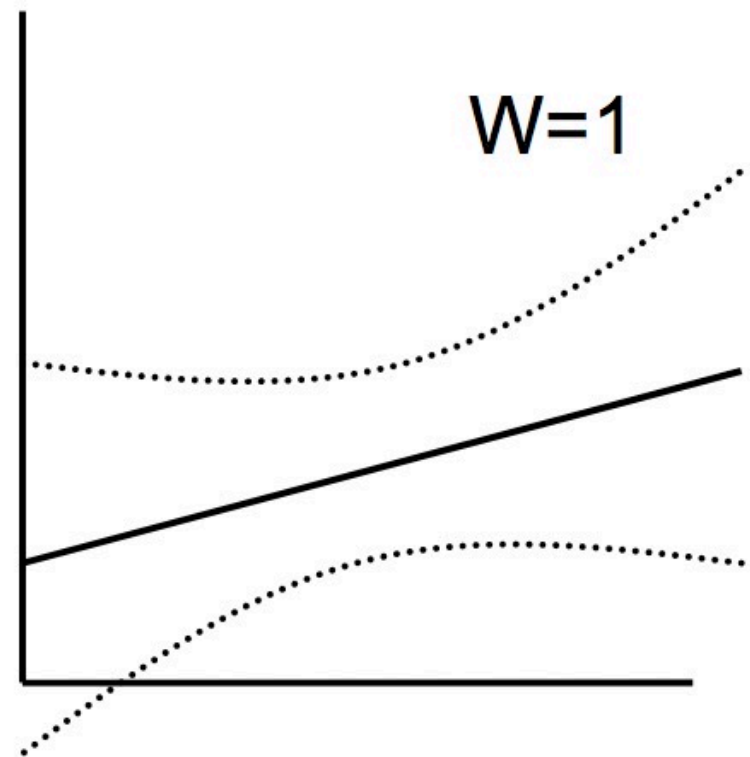
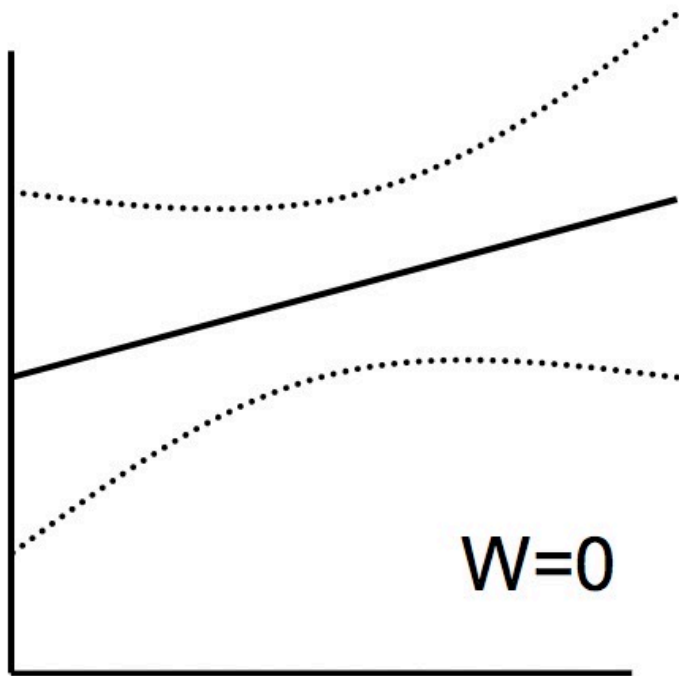
g_{10}

g_{01}

g_{00}

Graphs

Random coefficient models with dichotomous L2 predictor W_j (e.g., public=0, private=1).



Intercepts can vary within- and between-groups
Slopes only vary within-groups

Add cross-level interaction

Add W_j to slope equation

$$\begin{array}{ll} \text{L1:} & Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij} \\ \text{L2:} & b_{0j} = g_{00} + g_{01}W_j + u_{0j} \\ & b_{1j} = g_{10} + g_{11}W_j + u_{1j} \end{array} \quad \text{Multi-equation form}$$

$$Y_{ij} = g_{00} + g_{10}X_{ij} + g_{01}W_j + g_{11}X_{ij}W_j + (u_{0j} + u_{1j}X_{ij} + e_{ij})$$

Single equation form

g_{11} is the effect of W_j on the slope between X_{ij} and Y_{ij}

→ *This is the full multilevel model*

$$Y_{ij} = 38 + 2.4 \cdot \text{math} + 5.3 \cdot \text{type} - 0.3 \cdot \text{math} \cdot \text{type} + (u_{0j} + u_{1j}X_{ij} + e_{ij})$$

Formula: mathscore ~ timeonmath + schooltype + timeonmath * schooltype +
(timeonmath | Schoolid)

REML criterion at convergence: 1743.1

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.64023	-0.56436	-0.05373	0.65209	2.58438

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Schoolid	(Intercept)	51.84	7.200	
	timeonmath	27.27	5.222	-0.98
	Residual	42.96	6.554	

Number of obs: 260, groups: Schoolid, 10

$\text{var}(u_{0j})$

$\text{var}(u_{1j})$

$\text{var}(e_{ij})$

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	37.9227	4.1861	7.8360	9.059	2.02e-05 ***
timeonmath	2.3629	2.9650	7.6090	0.797	0.4496
schooltype	5.3219	2.6097	7.4020	2.039	0.0786 .
timeonmath:schooltype	-0.3171	1.8475	7.1720	-0.172	0.8685

g_{10}

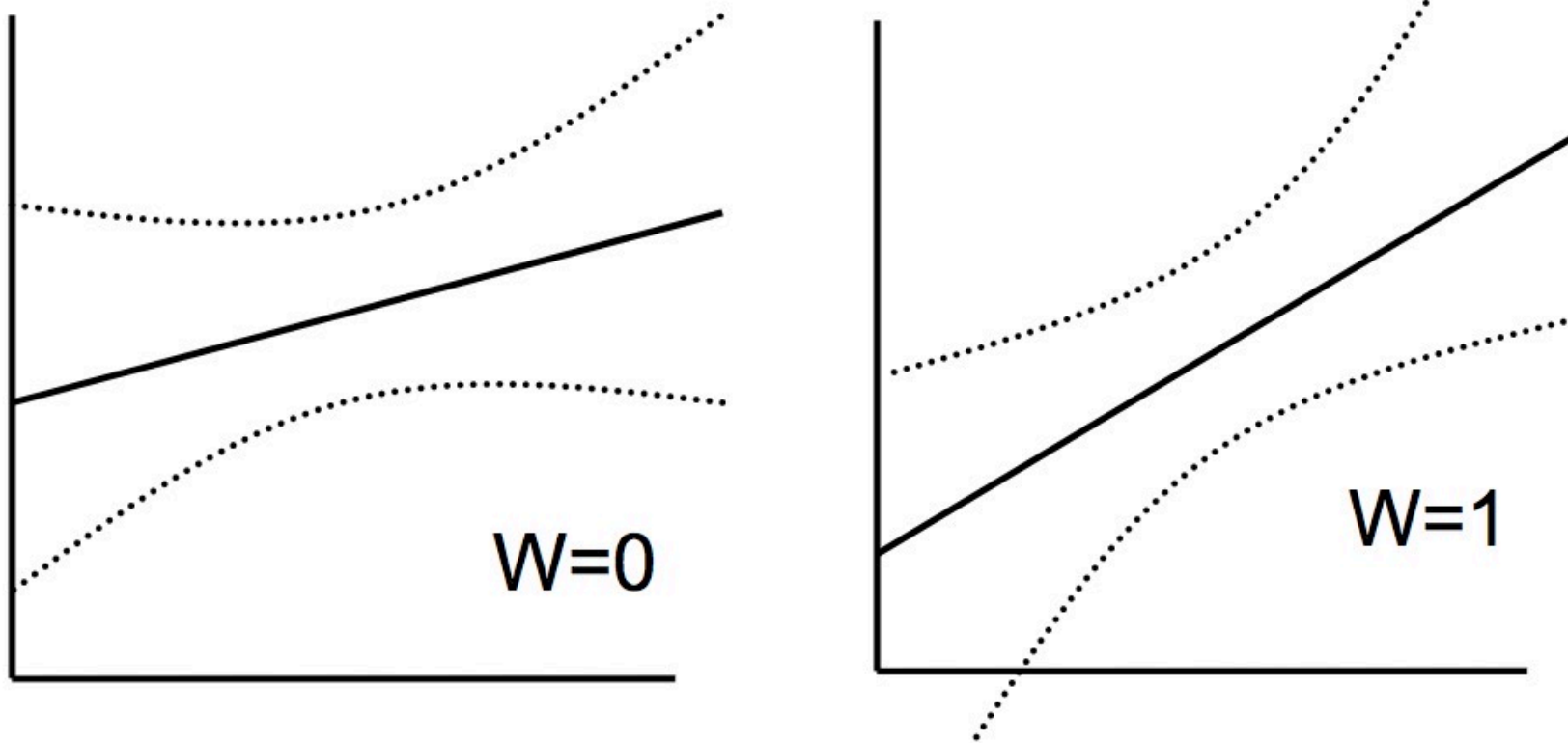
g_{11}

g_{01}

g_{00}

Graphs

Random coefficient models with dichotomous L2 predictor W_j (e.g., public=0, private=1)



Intercepts can vary within- and between-groups
Slopes can vary within- and between groups

Non-randomly varying slopes

No residual variability in slope after accounting for W_j

L1: $Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij}$ Multi-equation form

L2: $b_{0j} = g_{00} + g_{01}W_j + u_{0j}$
 $b_{1j} = g_{10} + g_{11}W_j$ (No random error)

$$Y_{ij} = g_{00} + g_{10}X_{ij} + g_{01}W_j + g_{11}X_{ij}W_j + (u_{0j} + e_{ij})$$

Single equation form

Slope of predictor (X) is entirely determined by overall mean and group effect

$$Y_{ij} = 37.2 + 3.4 \cdot \text{math} + 5.5 \cdot \text{type} - 0.6 \cdot \text{math} \cdot \text{type} + (u_{0j} + e_{ij})$$

Formula: mathscore ~ timeonmath + schooltype + timeonmath * schooltype +
(1 | Schoolid)

REML criterion at convergence: 1826.1

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.9172	-0.6792	-0.0361	0.6133	3.5261

Random effects:

Groups	Name	Variance	Std. Dev.
Schoolid	(Intercept)	12.13	3.483
	Residual	63.30	7.956

Number of obs: 260, groups: Schoolid, 10

$\text{var}(u_{0j})$

$\text{var}(e_{ij})$

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)	
(Intercept)	37.1984	2.4397	12.7000	15.247	1.54e-09	***
timeonmath	3.4401	0.6971	255.8200	4.935	1.45e-06	***
schooltype	5.5030	1.4491	10.0200	3.797	0.00349	**
timeonmath:schooltype	-0.5864	0.2530	253.0900	-2.318	0.02127	*

g_{10}

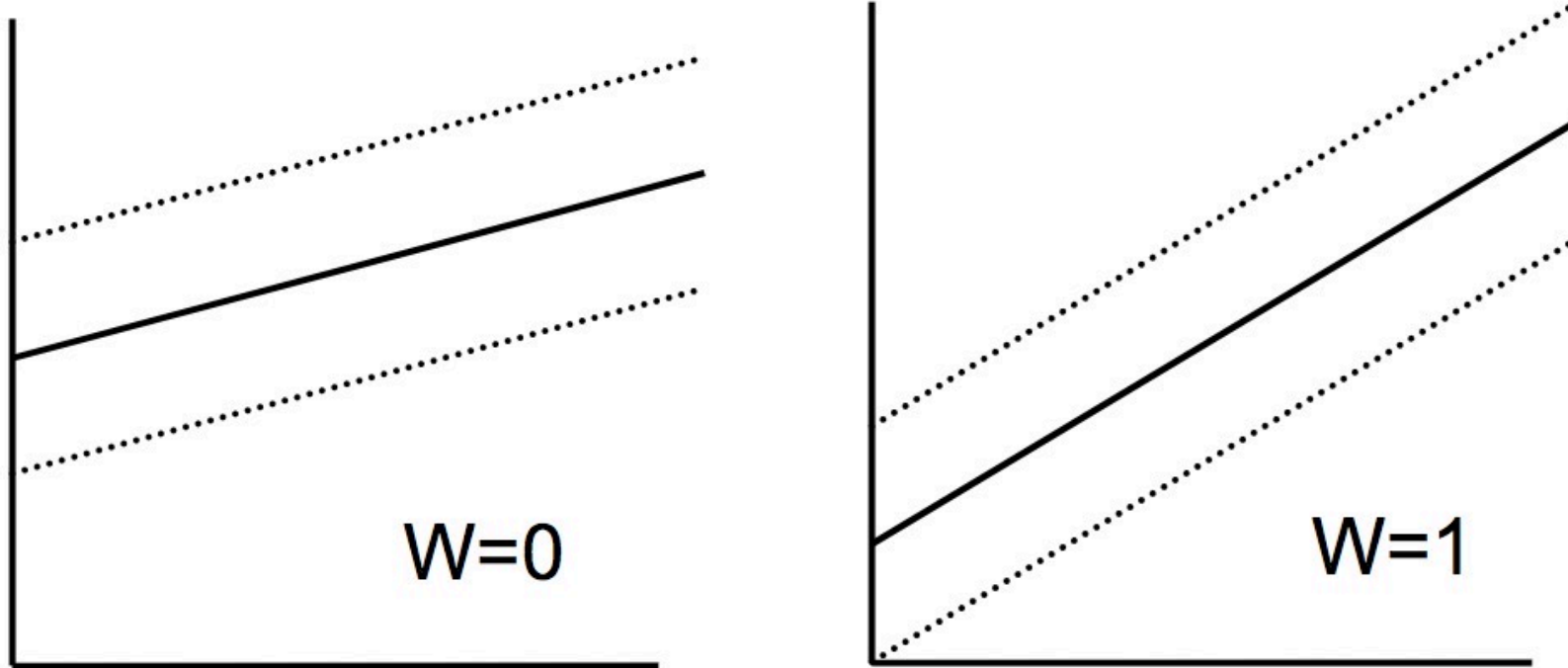
g_{11}

g_{01}

g_{00}

Graphs

Random coefficient models with dichotomous L2 predictor W_j (e.g., public=0, private=1).



Intercepts can vary within- and between-groups
Slopes only vary between groups (and not within groups)

Non-randomly varying slopes and intercepts

In the extreme, no residual variability in intercepts or slopes

L1:	$Y_{ij} = b_{0j} + b_{1j}X_{ij} + e_{ij}$	Multi-equation form
L2:	$b_{0j} = g_{00} + g_{01}W_j$	(No random error)
	$b_{1j} = g_{10} + g_{11}W_j$	(No random error)

$$Y_{ij} = g_{00} + g_{10}X_{ij} + g_{01}W_j + g_{11}X_{ij}W_j + (e_{ij})$$

Single equation form

This is equivalent to a single-level disaggregation model, and could be estimated with standard OLS regression