$$X \sim Exp(\lambda_1), Y \sim Exp(\lambda_2)$$

$$P(min(X,Y) > z) = P(X > z, Y > z) = P(X > z)P(Y > z) = e^{-z\lambda_1}e^{-z\lambda_2} = e^{-z(\lambda_1 + \lambda_2)}$$

(Note that X and Y are independent random variables.)

min(X,Y) is also exponentially distributed with parameter $\lambda = \lambda_1 + \lambda_2$

Let
$$Z = max(X, Y), Z \ge 0$$

$$F_Z(z) = P(max(X,Y) \le z) = P(X \le z, Y \le z) = (1 - e^{-z\lambda_1})(1 - e^{-z\lambda_2})$$

$$f_z(z) = \frac{dF_Z(z)}{dz} = \lambda_2 e^{-z\lambda_2} (1 - e^{-z\lambda_1}) + \lambda_1 e^{-z\lambda_1} (1 - e^{-z\lambda_2})$$
$$= \lambda_1 e^{-z\lambda_1} + \lambda_2 e^{-z\lambda_2} - (\lambda_1 + \lambda_2) e^{-z(\lambda_1 + \lambda_2)}$$

Exercise 2

Let
$$P(X = x | Y = 3) = P(Z = x)$$
. Then, $Z \sim Bin\left(3, \frac{1}{3}\right)$

Explanation. For the remaining 3 non-blue selections, we can choose either a white ball or a red ball and the probability of getting a white ball is half the probability of red ball.

Thus,
$$E[Z] = E[X|Y = 3] = 3 \cdot \frac{1}{3} = 1$$
 (if $A \sim Bin(n, p)$, then, $E[A] = np$.)

Exercise 3

 X_1 and X_2 are independent binomial random variables with respective parameters (n_1, p) and (n_2, p) .

$$\begin{split} P(X_1 = k | X_1 + X_2 = m) &= \frac{P(X_1 = k, X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{P(X_1 = k)P(X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{\binom{n_1}{k}p^k(1 - p)^{n_1 - k}\binom{n_2}{m - k}p^{m - k}(1 - p)^{n_2 - m + k}}{\binom{n_1 + n_2}{m}p^m(1 - p)^{n_1 + n_2 - m}} \quad [X_1 + X_2 \sim Bin(n_1 + n_2, p)] \\ &= \frac{\binom{n_1}{k}\binom{n_2}{m - k}p^m(1 - p)^{n_1 + n_2 - m}}{\binom{n_1 + n_2}{m}p^m(1 - p)^{n_1 + n_2 - m}} \\ &= \frac{\binom{n_1}{k}\binom{n_2}{m - k}}{\binom{n_1 + n_2}{m}} \end{split}$$

Exercise 4

Let $X \sim Unif(-1,1)$ and $Y = X^2$; Clearly dependent. Also, E[X] = 0.

Correlation.
$$Cov(X,Y) = E[XY] - E[X]E[Y] = E[X^3] = \int_{-1}^{1} 0.5x^3 = 0$$
 (uncorrelated)

 $X \sim Poi(\lambda), \ \lambda \sim Exp(1)$ (The mean of an exponential random variable is the inverse of its parameter.)

$$\begin{split} P(X=n) &= E[P(X=n|\lambda)] = \int_0^\infty e^{-\lambda} \frac{\lambda^n}{n!} f(\lambda) d\lambda \\ &= \int_0^\infty e^{-\lambda} \frac{\lambda^n}{n!} e^{-\lambda} d\lambda \\ &= \int_0^\infty e^{-2\lambda} \frac{\lambda^n}{n!} d\lambda \\ &= \frac{1}{2^{n+1}} \int_0^\infty e^{-x} \frac{x^n}{n!} dx \quad \text{[Change of Variable: } x = 2\lambda \text{]} \\ &= \frac{1}{2^{n+1}} \frac{\Gamma(n+1)}{n!} \\ &= \frac{1}{2^{n+1}} \quad \text{since, } n \text{ is a natural number.} \end{split}$$

Exercise 6

$$f_{X,Y}(x,y) = \begin{cases} c(1+xy) & 2 \le x \le 3\\ 0 & otherwise \end{cases}$$

1. For valid density function, $\iint_{\mathbb{R}^2} f_{X,Y}(x,y) \, dx \, dy = 1$

$$\implies \int_{1}^{2} \int_{2}^{3} c(1+xy) \, dx \, dy = 1$$

$$\implies c \int_{1}^{2} x + y \frac{x^{2}}{2} \Big|_{2}^{3} \, dy = 1$$

$$\implies c \int_{1}^{2} 1 + \frac{5y}{2} \, dy = 1$$

$$\implies c \Big[y + \frac{5y^{2}}{4} \Big]_{1}^{2} = 1$$

$$\implies \frac{19c}{4} = 1$$

$$\implies c = \frac{4}{19}$$

2.

$$f_X(x) = \int f_{XY}(x, y) dy = \int_1^2 \frac{4}{19} (1 + xy) dy$$
$$= \frac{4}{19} \left[y + \frac{xy^2}{2} \right]_1^2$$
$$= \frac{4}{19} \left(1 + \frac{3}{2} x \right)$$

$$f_X(x) = \int f_{XY}(x, y) dx = \int_2^3 \frac{4}{19} (1 + xy) dx$$
$$= \frac{4}{19} \left[x + \frac{yx^2}{2} \right]_2^3$$
$$= \frac{4}{19} \left(1 + \frac{5}{2} x \right)$$

Let X = no. of accidents a policyholder have in a year. Given. $X \sim Pois(\lambda)$, $g(\lambda) = \lambda e^{-\lambda}$

$$\begin{split} P(X=n) &= E[P(X=n|\lambda)] = \int_0^\infty e^{-\lambda} \frac{\lambda^n}{n!} g(\lambda) d\lambda \\ &= \int_0^\infty e^{-\lambda} \frac{\lambda^n}{n!} \lambda e^{-\lambda} d\lambda \\ &= \int_0^\infty e^{-2\lambda} \frac{\lambda^{n+1}}{n!} d\lambda \\ &= \frac{1}{2^{n+2}} \int_0^\infty e^{-x} \frac{x^{n+1}}{n!} dx \quad \text{[Change of Variable: } x = 2\lambda \text{]} \\ &= \frac{1}{2^{n+2}} \frac{\Gamma(n+2)}{n!} \\ &= \frac{n+1}{2^{n+2}} \quad \text{since, } n \text{ is a natural number.} \end{split}$$

Exercise 8

Let X ad Y denotes the number of females and males respectively who visit the yoga studio. Given. $X + Y \sim Pois(\lambda)$

$$P(X = n, Y = m) = P(X = n, Y = m | X + Y = m + n).P(X + Y = m + n)$$
$$= {\binom{m+n}{n}} p^n (1-p)^m e^{-\lambda} \frac{\lambda^{m+n}}{(m+n)!}$$

(Each person who visits is, independently, female with probability p or male with probability (1-p) thus, X or Y follow binomial distribution for a fixed number of visitors)

Exercise 9

We will make use of the linearity properties of the E[X] function i.e.,

$$E[X + Y] = E[X] + E[Y]$$
$$E[kX] = kE[X]$$

$$\begin{split} Cov(aX_1+b,cX_2+b) &= E[(aX_1+b)(cX_2+b)] - E[aX_1+b]E[cX_2+b] \\ &= E[acX_1X_2+abX_1+bcX_2+b^2] - (aE[X_1]+b)(cE[X_2]+b) \\ &= acE[X_1X_2]+abE[X_1]+bcE[X_2]+b^2 - acE[X_1]E[X_2] \\ &- abE[X_1] - bcE[X_2] - b^2 \\ &= ac(E[X_1X_2]-E[X_1]E[X_2]) \\ &= acCov(X_1,X_2) \end{split}$$

$$Cov(X_1 + X_2, X_3) = E[(X_1 + X_2)X_3] - E[X_1 + X_2]E[X_3]$$

$$= E[X_1X_3 + X_2X_3] - (E[X_1] + E[X_2])E_(X_3)$$

$$= E[X_1X_3] + E[X_2X_3] - E[X_1]E[X_3] - E[X_2]E[X_3]$$

$$= E[X_1X_3] - E[X_1]E[X_3] + E[X_2X_3] - E[X_2]E[X_3]$$

$$= Cov(X_1, X_3) + Cov(X_2 + X_3)$$

Given. n = 100, $\delta = 0.95$, $\hat{\mu} = 0.45$ Let the true mean by μ .

$$P(|\hat{\mu} - \mu| > \epsilon) \le 2e^{n\epsilon^2} = 0.05$$

$$\implies 2e^{100\epsilon^2} = 0.05$$

$$\implies \epsilon = 0.192$$

So, our confidence inteval is (0.45 - 0.192, 0.45 + 0.192) = (0.258, 0.642)

$$\epsilon_{new} = \epsilon/2 = 0.096$$

Now,

$$n_{new} = \frac{1}{\epsilon_{new}^2} log\left(\frac{2}{0.05}\right)$$
$$= \frac{1}{(0.096)^2} log\left(\frac{2}{0.05}\right)$$
$$= 400$$

So, we need 300 more sample points.