### The Monty Hall 3-door Puzzle

You have a chance to win a large prize in a game show hosted by Monty. The prize is behind one of 3 closed doors, and the other two doors are losers.

- You are asked to select (but not to open yet) one of the doors.
- Then Monty, who knows what is behind each door, does the following: Whether or not you selected the winning door, he opens one of the other two doors that he knows for sure is a losing door (selecting random if both are losing doors).
- Then he asks you whether you would like to switch doors.

### Which strategy should you follow to have a better chance?

- Should you change doors?
- Or stick to your original selection?
- Or does it not matter?

# **Probability theory**

An <u>experiment</u> is a procedure that yields one of a given set of possible outcomes.

The **sample space** of the experiment is the set of possible outcomes:

$$S = \{s_1, s_2, \dots, s_n\}$$

A <u>probability distribution</u> in the sample space S is a function p assigning a number  $p(s_i)$  to each possible outcome  $s_i$  in S (the <u>probability</u> of  $s_i$ ) such the following two conditions hold:

- $0 \le p(s_i) \le 1$ , for every i = 1, 2, ..., n,
- $p(s_1) + p(s_2) + \cdots + p(s_n) = 1.$

FOR EXAMPLE: the **uniform distribution** on S

$$p(s_i) = \frac{1}{n}$$
, for every  $i = 1, 2, \dots, n$ .

If a fair dice is rolled, then there are six possible equally likely outcomes: 1, 2, 3, 4, 5, 6

So 
$$p(1) = p(2) = p(3) = p(4) = p(5) = p(6) = \frac{1}{6}$$

## Probability distribution: another example

If a *fair* coin is flipped, then there are two equally likely outcomes:

H (heads) and T (tails).

So 
$$p(H) = p(T) = \frac{1}{2}$$
.

But what probabilities should we assign to these outcomes when the coin is biased so that heads comes up twice as often as tails?

SOLUTION: We have p(H) = 2p(T). Since p(H) + p(T) = 1 should hold, it follows that 2p(T) + p(T) = 3p(T) = 1.

So 
$$p(T) = \frac{1}{3}$$
 and  $p(H) = \frac{2}{3}$ .

# **Events and their probabilities**

An **event** is a subset of the sample space, that is, a set of possible outcomes.

The **probability of an event** E is the sum of the probabilities

of the outcomes in E

FOR EXAMPLE: What is the probability that, when two fair dice are rolled, the sum of the two numbers on them is 7?

By the product rule, there are a total of  $6 \cdot 6 = 36$  possible outcomes, each of them equally likely with a probability of  $\frac{1}{36}$ .

There are 6 'successful' outcomes:

$$(1,6), (2,5), (3,4), (4,3), (5,2), (6,1).$$

Therefore, the probability of having 7 as the sum is  $6 \cdot \frac{1}{36} = \left| \frac{1}{6} \right|$ .

# Probability of events: another example

Suppose that a dice is loaded so that 3 appears twice as often as each of the other numbers, but the other five outcomes are equally likely.

What is the probability that an odd number appears when we roll this dice?

SOLUTION: We want to find the probability  $\ p(E)$  of the event  $\ E=\{1,3,5\}$  . We know that

- p(1) + p(2) + p(3) + p(4) + p(5) + p(6) = 1,
- p(1) = p(2) = p(4) = p(5) = p(6), and
- p(3) = 2p(1).

Therefore, 
$$p(1) = p(2) = p(4) = p(5) = p(6) = \frac{1}{7}$$
 and  $p(3) = \frac{2}{7}$ .

It follows that 
$$\boxed{p(E) \ = \ p(1) + p(3) + p(5) \ = \ \frac{4}{7}} \, .$$

### Probability of events: yet another example

In a certain kind of lottery, people can win the top prize if they correctly choose a set of 6 numbers out of the first 40 positive natural numbers.

What is the probability of winning the top prize?

SOLUTION: The total number of ways to choose 6 numbers out of 40 is

$$\binom{40}{6} = \frac{40!}{34! \cdot 6!} = 3838380.$$

Each of these outcomes is equally likely. The event W of winning the top prize contains just one of these. Consequently,

$$p(W) = \frac{1}{3838380} \approx 0.00000026.$$

## Probability of the complementary event

Let E be an event in a sample space S.

The probability of event  $\bar{E}$  , the **complement of** E **in** S ,

is the sum of the probabilities of the outcomes *not* in *E*:

$$p(\bar{E}) = 1 - p(E)$$

FOR EXAMPLE: We flip a fair coin 10 times.

What is the probability that heads comes up at least once?

Let E be the event that heads comes up at least once of the 10 flips.

Then  $\bar{E}$  is the event that each of the 10 times tails comes up. So

$$p(E) = 1 - p(\bar{E}) = 1 - \frac{1}{2^{10}} = 1 - \frac{1}{1024} = \frac{1023}{1024} \approx 0.999.$$

# Probability of the union of events

$$p(E_1 \cup E_2) = p(E_1) + p(E_2) - p(E_1 \cap E_2)$$

FOR EXAMPLE: What is the probability that a number randomly selected from the set  $X = \{n \in \mathbf{N}^+ \mid n \le 100\}$  is divisible by either 2 or 5 or by both (by 10)?

- ullet let  $E_1$  be the event that the selected number in X is divisible by 2
- let  $E_2$  be the event that the selected number in X is divisible by 5
- then  $E_1 \cap E_2$  is the event that it is divisible by both 2 and 5, that is, it is divisible by 10
- as |X| = 100,  $|E_1| = 50$ ,  $|E_2| = 20$ , and  $|E_1 \cap E_2| = 10$ , we obtain:

$$p(E_1 \cup E_2) = \frac{50}{100} + \frac{20}{100} - \frac{10}{100} = \frac{60}{100} = \left| \frac{3}{5} \right| = \boxed{0.6}$$

# **Conditional probability**

We flip a fair coin 3 times. We already know how to calculate the probability of the event E that at least 2 heads come up in a row:

The sample space consists of  $2^3=8$  outcomes, each with probability  $\frac{1}{8}$ .

3 of them is in 
$$E$$
:  $HHT$ ,  $THH$ ,  $HHH$ . So  $p(E) = \left| \frac{3}{8} \right|$ .

But what if the <u>first flip comes up heads</u> (event F occurs), and we are asked about the probability of E, knowing that F has already occurred?

Now the sample space is F: it consists of outcomes HHT, HTH, HTT, HHH For an outcome in E to occur, it must belong to  $E \cap F$ : either HHT or HHH

The conditional probability of E given F is

$$\boxed{p(E \mid F)} = \frac{|E \cap F|}{|F|} = \frac{\frac{|E \cap F|}{8}}{\frac{|F|}{8}} = \boxed{\frac{p(E \cap F)}{p(F)}}$$

IN THE EXAMPLE:

We have 
$$p(F) = \frac{4}{8} = \frac{1}{2}$$
 and  $p(E \cap F) = \frac{2}{8} = \frac{1}{4}$ . So  $p(E \mid F) = \frac{\frac{1}{4}}{\frac{1}{2}} = \boxed{\frac{1}{2}}$ .

### Conditional probability: another example

We throw a pair of fair dice. What is the probability that at least one dice is a 3, given that the sum of the two dice is 5?

#### SOLUTION:

- let A be the event that at least one of the dice is a 3
- let B the event that the sum of the two dice is 5, that is, B consists of the outcomes (1,4), (2,3), (3,2), (4,1).
- as the sample space consists of  $6^2=36$  equally probable outcomes,  $p(B)=\frac{4}{36}=\frac{1}{9}$
- the event  $A\cap B$  consists of two outcomes: (2,3) and (3,2) , and so  $p(A\cap B)=\frac{2}{36}=\frac{1}{18}$

Therefore,

$$p(A \mid B) = \frac{p(A \cap B)}{p(B)} = \frac{\frac{1}{18}}{\frac{1}{9}} = \frac{1}{2}$$

Observe that the unconditional probability of A is different:  $p(A) = \frac{11}{36}$ 

## Conditional probability: yet another example

We flip a fair coin 3 times. What is the probability that an odd number of tails appear, knowing that the first flip comes up tails?

#### SOLUTION:

- let X be the event that an odd number of tails appear
- let Y be the event that the first flip comes up tails
- as the sample space consists of  $2^3=8$  equally probable outcomes,  $p(Y)=\frac{4}{8}=\frac{1}{2}$
- the event  $X\cap Y$  consists of two outcomes: THH and TTT , and so  $p(X\cap Y)=\frac{2}{8}=\frac{1}{4}$

Therefore,

$$p(X \mid Y) = \frac{p(X \cap Y)}{p(Y)} = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{1}{2}$$

Observe that this time the unconditional probability of X is the same:

$$p(X) = \frac{4}{8} = \frac{1}{2}$$

# Independence

When two events E and F are <u>independent</u>, the occurrence of one of the events gives no information about the probability of the other event:

$$p(E \mid F) \ = \ p(E)$$
 or equivalently,  $p(F \mid E) \ = \ p(F)$ 

Using that  $p(E \mid F) = \frac{p(E \cap F)}{p(F)}$ , we obtain that E and F are independent events, whenever  $p(E \cap F) = p(E) \cdot p(F)$ 

#### FOR EXAMPLE:

- Events E and F on slide 104 are not independent.
- Events A and B on slide 105 are not independent.
- Events X and Y on slide 106 are independent.

# **Bayes' Theorem**

This might be very useful when E and F are events from the same sample space such that  $p(F) \neq 0$  and we can easily compute  $p(E \mid F)$ , but want to know  $p(F \mid E)$ :

$$p(F \mid E) = \frac{p(E \mid F)p(F)}{p(E \mid F)p(F) + p(E \mid \bar{F})p(\bar{F})}$$

FOR EXAMPLE: There are two boxes,  $Box_1$  and  $Box_2$ .

- $Box_1$  contains 2 green balls and 7 red balls
- $Box_2$  contains 4 green balls and 3 red balls

Bob is blindfolded, and

- selects a ball by randomly choosing one of the boxes first,
- then randomly choosing a ball from it.

If we know that Bob has selected a red ball, what is the probability that he selected a ball from  $Box_1$ ?

# Bayes' Theorem: Example (cont.)

If we know that Bob has selected a red ball, what is the probability that he selected a ball from  $Box_1$ ?

SOLUTION: Let  $\boxed{R}$  be the event that  $\begin{subarray}{l} Bob\ has\ chosen\ a\ ball\ from\ Box_1. \end{subarray}$ 

We want to find  $p(F \mid R)$ .

- $p(R \mid F) = \frac{7}{9}$ . as  $Box_1$  contains 2 green balls and 7 red balls
- $\bar{F}$ : is the event that <u>Bob has chosen a ball from Box\_2</u>. Then  $p(F)=p(\bar{F})=\frac{1}{2}$ .
- $p(R \mid \bar{F}) = \frac{3}{7}$ . as  $Box_2$  contains 4 green balls and 3 red balls

Then by Bayes' Theorem:

$$\underline{p(F \mid R)} = \frac{\frac{7}{9} \cdot \frac{1}{2}}{\frac{7}{9} \cdot \frac{1}{2} + \frac{3}{7} \cdot \frac{1}{2}} = \frac{\frac{7}{9}}{\frac{49+27}{63}} = \frac{\frac{7}{9}}{\frac{76}{63}} = \frac{49}{76} \approx \underline{0.645}$$

# Applying Bayes' Theorem: another example

Suppose that 1 person in 100000 has a particular rare disease for which there is a quite accurate diagnostic test:

- It is correct 99% of the time when given to someone with the disease.
- It is correct 99.5% of the time when given to someone who does not have the disease.

What is the probability that someone who tests positive for the disease actually has the disease?

SOLUTION: Let  $\boxed{D}$  be the event that a person has the disease, and  $\boxed{T}$  be the event that a person tests positive for the disease.

- $p(T \mid D) = 0.99$
- $p(\bar{T} \mid \bar{D}) = 0.995$ , so  $p(T \mid \bar{D}) \stackrel{Ex.5.5}{=} 1 p(\bar{T} \mid \bar{D}) = 1 0.995 = 0.005$
- $p(D) = \frac{1}{100000} = 0.00001$ , so  $p(\bar{D}) = 1 0.00001 = 0.99999$
- $p(D \mid T) = \frac{0.99 \cdot 0.00001}{0.99 \cdot 0.00001 + 0.005 \cdot 0.99999} \approx 0.002$

## Bayesian spam filters

A Bayesian spam filter uses information about previously seen e-mail messages to guess whether an incoming e-mail is spam. It looks for occurrences of particular words in messages. For a word w, the probability that w appears in a spam message is estimated by determining

- the number of times w appears in a message from a large set of messages known to be spam, and
- the number of times w appears in a large set of messages that are known not to be spam.

Unfortunately, spam filters sometimes fail to identify a spam message as spam: this is called a **false negative**. And they sometimes identify a message that is not spam as spam: this is called a **false positive**.

When testing for spam, it is important to minimise false positives, because filtering out wanted e-mail is much worse than letting some spam through.

## Bayesian spam filters: an example

Suppose we have found that the word "*Rolex"* occurs in 250 of 2000 messages known to be spam, and in 5 of 1000 messages known not to be spam.

Estimate the probability that an incoming message containing the word "*Rolex"* is spam, assuming that it is equally likely that an incoming message is spam or not spam. If our threshold for rejecting a message as spam is 0.9, will we reject this message?

SOLUTION: Let S be the event that an incoming message is spam, and R be the event that the message contains the word "Rolex". Then:

- $p(R \mid S) = \frac{250}{2000} = 0.125$
- $p(R \mid \bar{S}) = \frac{5}{1000} = 0.005$
- $p(S) = p(\bar{S}) = 0.5$
- $p(S \mid R) = \frac{0.125 \cdot 0.5}{0.125 \cdot 0.5 + 0.005 \cdot 0.5} = \frac{0.125}{0.13} \approx 0.962$

As 0.962 > 0.9, we would reject the message as spam.

## The 3-door puzzle revisited

Suppose the doors are numbered: door(1), door(2), door(3).

- You selected door(you) (where you can be either 1, 2, or 3).
- Then Monty opened door(Monty).
- The remaining 3rd door is door(other).

Now observe that you, Monty, and other are three distinct numbers.

Consider the following events:

$$W_1$$
: door(1) is the winning door

 $W_3$ : door(3) is the winning door

 $W_2$ : door(2) is the winning door

M: Monty opened door(Monty)

We are interested in  $p(W_{you} \mid M)$  :

- If  $p(W_{you} \mid M) > \frac{1}{2}$  then you should stick to door(you)
- if  $p(W_{\text{you}} \mid M) < \frac{1}{2}$  then you should switch to door(**other**)

### The 3-door puzzle revisited (cont.)

We want to use Bayes' theorem to compute  $p(W_{you} \mid M)$ :

$$p(W_{\mathsf{you}} \mid M) \ = \ \frac{p(M \mid W_{\mathsf{you}}) \cdot p(W_{\mathsf{you}})}{p(M \mid W_{\mathsf{you}}) \cdot p(W_{\mathsf{you}}) + p(M \mid \overline{W}_{\mathsf{you}}) \cdot p(\overline{W}_{\mathsf{you}})}$$

So we need to compute:

- ullet  $p(W_{\mathsf{you}})$
- $\bullet \quad p(M \mid W_{\mathsf{you}})$
- $\bullet \quad p(\overline{W}_{\mathsf{you}})$
- $p(M \mid \overline{W}_{you})$

# The 3-door puzzle revisited (cont.)

- $\underline{p(W_{you})} = p(W_{other}) = p(W_{Monty}) = \frac{1}{3}$ These are the 3 different outcomes and they are equally likely.
- $p(\overline{W}_{you}) = 1 p(W_{you}) = 1 \frac{1}{3} = \frac{2}{3}$
- $p(M \mid W_{you}) = \frac{1}{2}$

If your selected door(**you**) is the winning one, then Monty opens randomly one of the other two.

- $p(M \mid W_{\mathbf{other}}) = 1$ If door(**other**) is winning and you selected door(**you**), then Monty surely opens the 3rd door, door(**Monty**) (as he never opens the winning door).
- $p(M \cap W_{\mathbf{other}}) = p(M \mid W_{\mathbf{other}}) \cdot p(W_{\mathbf{other}}) = 1 \cdot \frac{1}{3} = \frac{1}{3}$
- $p(M \cap \overline{W}_{\mathbf{you}}) = p(M \cap (W_{\mathbf{Monty}} \cup W_{\mathbf{other}})) =$

$$p((M \cap W_{\mathbf{Monty}}) \cup (M \cap W_{\mathbf{other}})) = p(M \cap W_{\mathbf{other}}) = \frac{1}{3}$$

• 
$$p(M \mid \overline{W}_{\mathbf{you}}) = \frac{p(M \cap \overline{W}_{\mathbf{you}})}{p(\overline{W}_{\mathbf{you}})} = \frac{\frac{1}{3}}{\frac{2}{3}} = \frac{1}{2}$$

## The 3-door puzzle revisited: the conclusion

Now we can apply Bayes' theorem:

$$\frac{p(W_{\mathbf{you}} \mid M)}{p(M \mid W_{\mathbf{you}}) \cdot p(W_{\mathbf{you}}) + p(M \mid \overline{W}_{\mathbf{you}}) \cdot p(\overline{W}_{\mathbf{you}})}$$

$$= \frac{\frac{1}{2} \cdot \frac{1}{3}}{\frac{1}{2} \cdot \frac{1}{3} + \frac{1}{2} \cdot \frac{2}{3}} = \frac{\frac{1}{6}}{\frac{3}{6}} = \boxed{\frac{1}{3}}$$

In other words:

$$p(\overline{W}_{you} \mid M) = 1 - p(W_{you} \mid M) = 1 - \frac{1}{3} = \frac{2}{3} > \frac{1}{2}$$

So you'll have a better chance if you switch from door(you) to door(other).