

## AFRE 802

### Statistical Methods for Agricultural, Food, & Resource Economists



**Linear models & estimation by least squares – Part 1 of 3**  
**(WMS Ch. 11.1-11.3, Wooldridge pp. 22-37)**

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## GAME PLAN

- Collect Ch. 10 HW
- Hand out graded exercise (due Thursday)
- 
- Linear models & estimation by least squares  
 – Part 1 of 3
  - The simple linear regression model
    - Examples
    - Terminology
    - Assumptions about the error term (and concept of endogeneity)
    - Deriving estimates of the simple linear regression parameters:  
ordinary least squares (OLS)
    - Compute OLS estimates by hand & in Stata

# Simple Linear Regression

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## The simple linear regression model: motivation

- Suppose  $y$  and  $x$  are two variables that represent some population
- **How does  $y$  change when  $x$  changes? What is the causal effect (*ceteris paribus* effect) of  $x$  on  $y$ ?**
- *Examples?*

$y$	$x$
Corn yield	Fertilizer
Beef demand	Beef price
Wheat acreage	Wheat price
Hourly wage	Years of education
Community crime rate	# of police officers

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## The simple linear regression model

$$y = \beta_0 + \beta_1 x + u$$

$\beta_0$  (intercept) and  $\beta_1$  (slope) are the population parameters to be estimated

- $u$  is the **error term** or **disturbance**
  - $u$  for “**unobserved**”
  - Represents **all factors other than  $x$  that affect  $y$**
  - Some use  $\varepsilon$  instead of  $u$
- Terminology for  $y$  and  $x$ :

$y$	$x$
Dependent variable	Independent variable
Explained variable	Explanatory variable
Response variable	Control variable
Predicted variable	Predictor variable
Regressand	Regressor
	Covariate

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To use data to get unbiased estimates of  $\beta_0$  and  $\beta_1$ , need to restrict the relationship b/w  $x$  and  $u$

$$y = \beta_0 + \beta_1 x + u$$

- $E(u) = 0$  (not restrictive if have an intercept,  $\beta_0$ )
- \*\*\*  $E(u|x) = E(u)$  (i.e. the average value of  $u$  does not depend on the value of  $x$ )

#1 & #2  $\rightarrow E(u|x) = E(u) = 0$  (zero conditional mean)

- If this holds,  $x$  is “**exogenous**”; but **if  $x$  is correlated with  $u$ ,  $x$  is “endogenous”** (next class)

*What does this assumption imply below?*

- $yield = \beta_0 + \beta_1 fertilizer + u$ , where  $u$  is unobserved land quality (*inter alia*)
- $wage = \beta_0 + \beta_1 educ + u$ , where  $u$  is unobserved ability (*inter alia*)

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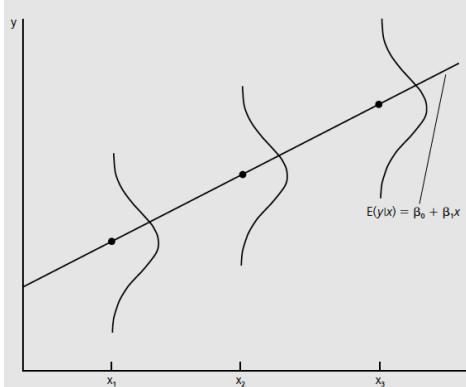
What is  $E(y|x)$  if we assume  $E(u|x)=0$ ?

Hint: Apply the rules for conditional expectations.

$$y = \beta_0 + \beta_1 x + u$$

$$E(y|x) = \beta_0 + \beta_1 x$$

$E(y|x)$  as a linear function of  $x$ .



Source: Wooldridge (2003)

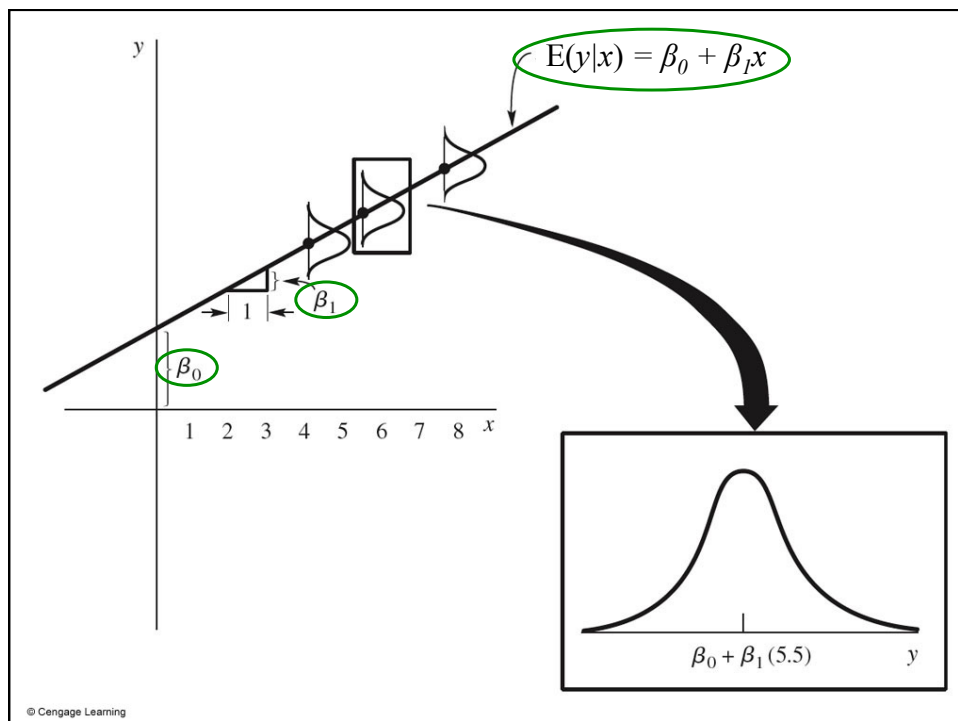
What is  $\frac{\partial E(y|x)}{\partial x}$  and how do we interpret this result?

$$\frac{\partial E(y|x)}{\partial x} = \beta_1$$

Interpretation:  $\beta_1$  is the expected change in  $y$  given a one unit increase in  $x$ , *ceteris paribus* (slope)

What is the interpretation of  $\beta_0$ ?

Interpretation:  $\beta_0$  is the expected value of  $y$  when  $x = 0$  (intercept)



### Why is it called linear regression?

$$y = \beta_0 + \beta_1 x + u$$

- **Linear in parameters**,  $\beta_0$  and  $\beta_1$
- Does NOT limit us to linear relationships between  $x$  and  $y$
- But rules out models that are **nonlinear in parameters**, e.g.:

$$y = \frac{1}{\beta_0 + \beta_1 x} + u$$

$$y = \Phi(\beta_0 + \beta_1 x) + u$$

$$y = \frac{\beta_0}{\beta_1} x + u$$

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### Estimating $\beta_0$ and $\beta_1$

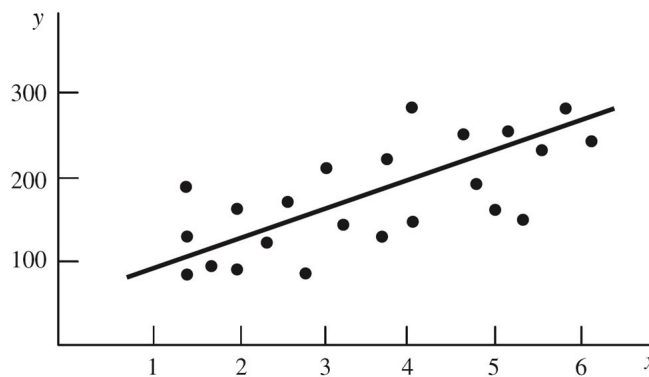
$$y = \beta_0 + \beta_1 x + u$$

- Suppose we have a random sample of size  $N$  from the population of interest. Then can write:

$$y_i = \beta_0 + \beta_1 x_i + u_i, \quad i = 1, 2, 3, \dots, N$$

We don't know  $\beta_0$  and  $\beta_1$  but want to estimate them.

How could we use the data in our sample to estimate  $\beta_0$  and  $\beta_1$ ?



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## Recall 3 common methods of estimation

1. Method of moments
  2. Maximum likelihood
  3. Least squares
- All 3 of these approaches lead to the same estimators for  $\beta_0$  and  $\beta_1$  (under certain assumptions)
  - We'll focus on the least squares approach
  - See Wooldridge (2003: 27-29) for method of moments discussion and his panel data book for MLE discussion

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## (Ordinary) least squares (OLS) approach

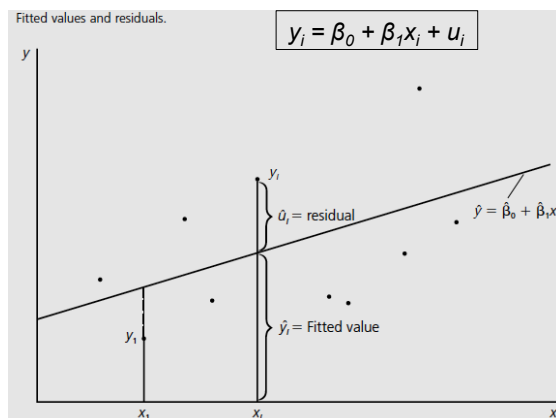
- *What was the gist of this approach?*
  - Choose estimator to minimize the sum of squared deviations b/w observed & estimated values
- “Fitted” values of y and residuals:

Fitted (estimated, predicted) values of y:  $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$

Residuals:  
 $\hat{u}_i = y_i - \hat{y}_i$   
 $= y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i$

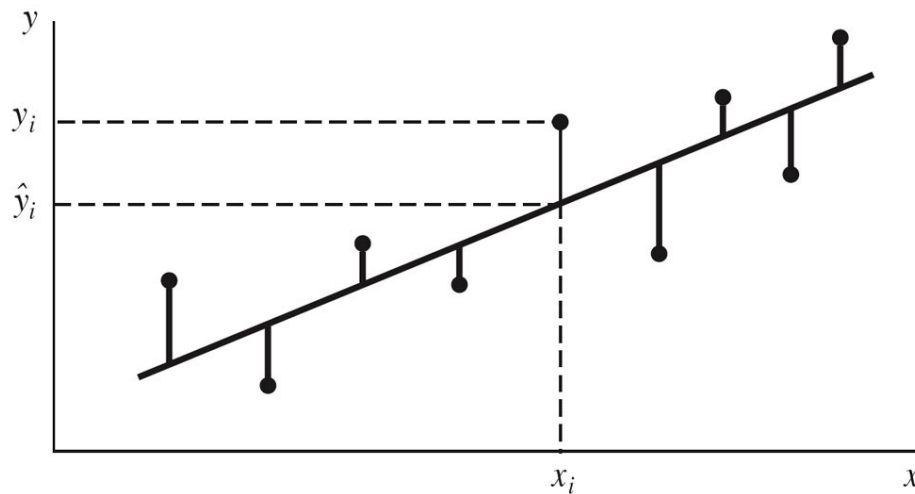
OLS:  
 Choose  $\hat{\beta}_0$  and  $\hat{\beta}_1$  to minimize:  

$$\sum_{i=1}^N \hat{u}_i^2 = \sum_{i=1}^N (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$



Source: Wooldridge (2003)

Where are the fitted values and residuals in this figure?



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The OLS estimators for  $\beta_0$  and  $\beta_1$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^N (x_i - \bar{x})^2}$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

Work through proof of formulas above. \*\*This is a proof you should know.\*\*

Another useful expression for  $\hat{\beta}_1$ :

$$\hat{\beta}_1 = \frac{\sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^N (x_i - \bar{x})^2} = \frac{\sum_{i=1}^N x_i y_i - \frac{1}{N} \sum_{i=1}^N x_i \sum_{i=1}^N y_i}{\sum_{i=1}^N x_i^2 - \frac{1}{N} \left( \sum_{i=1}^N x_i \right)^2}$$

Proof will be posted to D2L.

## Obtaining OLS estimates – example (by hand)

### EXAMPLE 11.1

Use the method of least squares to fit a straight line to the  $n = 5$  data points given in Table 11.1.

$$\hat{\beta}_1 = \frac{\sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^N (x_i - \bar{x})^2} = \frac{\sum_{i=1}^N x_i y_i - \frac{1}{N} \sum_{i=1}^N x_i \sum_{i=1}^N y_i}{\sum_{i=1}^N x_i^2 - \frac{1}{N} \left( \sum_{i=1}^N x_i \right)^2}$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

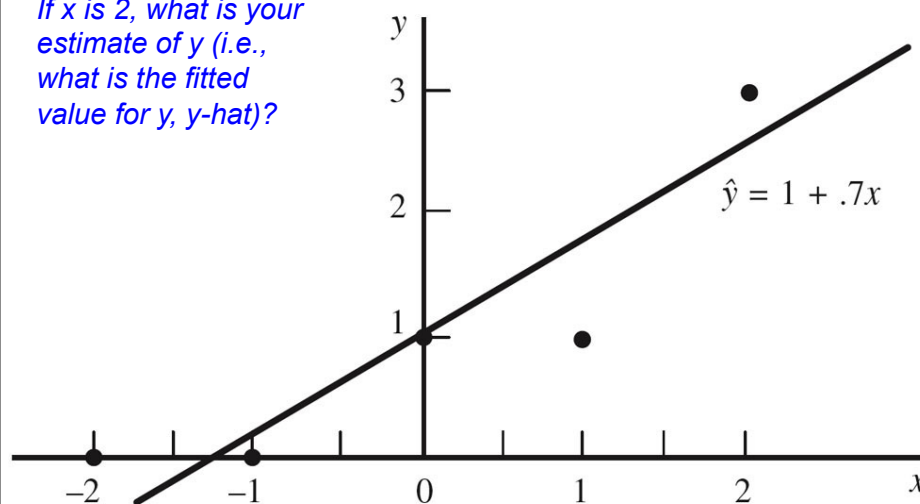
Table 11.1 Data for Example 11.1

$x$	$y$
-2	0
-1	0
0	1
1	1
2	3

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*If  $x$  is 2, what is your estimate of  $y$  (i.e., what is the fitted value for  $y$ ,  $\hat{y}$ )?*



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## Calculations using alternative formula

Table 11.2 Calculations for finding the coefficients

$x_i$	$y_i$	$x_i y_i$	$x_i^2$
-2	0	0	4
-1	0	0	1
0	1	0	0
1	1	1	1
2	3	6	4
$\sum_{i=1}^n x_i = 0$	$\sum_{i=1}^n y_i = 5$	$\sum_{i=1}^n x_i y_i = 7$	$\sum_{i=1}^n x_i^2 = 10$

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$$\hat{\beta}_1 = \frac{\sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^N (x_i - \bar{x})^2} = \frac{\sum_{i=1}^N x_i y_i - \frac{1}{N} \sum_{i=1}^N x_i \sum_{i=1}^N y_i}{\sum_{i=1}^N x_i^2 - \frac{1}{N} \left( \sum_{i=1}^N x_i \right)^2} = \frac{7 - \frac{1}{5}(0)(5)}{10 - \frac{1}{5}(0)^2} = 0.7$$

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## Basic Stata Commands

- **regress**  $y$   $x$       Linear regression of  $y$  on  $x$ 
  - EX) regress wage educ
- **predict newvar1, xb**      Compute fitted values
  - EX) predict wagehat, xb    (I just made up the name wagehat)
- **predict newvar2, resid**      Compute residuals
  - EX) predict uhat, resid    (I just made up the name uhat)

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## Obtaining OLS estimates – example (Stata)

### Wooldridge (2003) Example 2.4: Wage and education

Use Stata to run the simple linear regression of wage ( $y$ ) on educ ( $x$ ).

$$wage_i = \beta_0 + \beta_1 educ_i + u_i$$

Command: regress wage educ (or: reg wage educ)

*What are  $\beta_0$  and  $\beta_1$  below?*

```
reg wage educ
```

Source	SS	df	MS	
Model	1179.73204	1	1179.73204	
Residual	5980.68225	524	11.4135158	
Total	7160.41429	525	13.6388844	

Number of obs = 526  
 F( 1, 524) = 103.36  
 Prob > F = 0.0000  
 R-squared = 0.1648  
 Adj R-squared = 0.1632  
 Root MSE = 3.3784

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
wage	$\hat{\beta}_1$				
educ	.5413593	.053248	10.17	0.000	.4367534 .6459651
_cons	$\hat{\beta}_0$	.6849678	-1.32	0.187	-2.250472 .4407687

## Homework:

- WMS Ch. 11:
  - 11.1, 11.3 (do calculations by hand)
  - 11.4 and 11.5 (for these two problems, do calculations using Excel and the formulas on slide 17, and in Stata using the “regress” command; the data are on D2L in the Stata folder)
- Try to complete all Ch. 11 HW before last day of class (Dec. 8) so that we can go over it then. (You won’t turn in Ch. 11.)

## Next class:

- Linear regression part 2 of 3
  - Properties of OLS estimators

## Reading for next class:

- WMS Ch. 11: section 11.4
- Wooldridge *Introductory Econometrics* (2003): pp. 38-60, 101-102

Aside: NPR “Hidden Brain” example of a natural experiment, and when it might be reasonable to assume  $E(u|x)=E(u)$

- Listen for the following:
  - *What is the dependent variable?*
  - *What is the main explanatory variable of interest?*
  - *Why might it be reasonable to assume  $E(u|x)=E(u)$  here?*
  - *What is a natural experiment?*
- Dependent variable: cognitive function of elderly
- Main explanatory variable: wealth
- $E(u|x)=E(u)$  might be reasonable – Congress computational mistake – people in one cohort got higher benefits than next cohort (level of benefits shouldn’t be correlated with unobservables)



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## Aside: Natural experiments

A *natural experiment* occurs when **some exogenous event**—often a change in government policy—**changes the environment** in which individuals, families, firms, or cities operate. A natural experiment always has a **control group**, which is **not affected by the policy change**, and a **treatment group**, which is **thought to be affected by the policy change**. Unlike with a true experiment, where treatment and control groups are randomly and explicitly chosen, the control and treatment groups in natural experiments arise from the particular policy change. (Wooldridge, 2003: 417) <sup>21</sup>