AFRE 802 Statistical Methods for Agricultural, Food, & Resource Economists



Multivariate probability distributions – Part 3 of 3 (WMS Ch. 5.5-5.8, 5.11, 5.12)

October 12, 2017
Nicole Mason
Michigan State University
Fall 2017

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GAME PLAN

- Review
- Graded in-class exercise

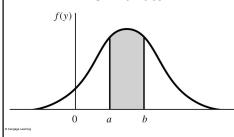
Multivariate probability distributions (Part 3 of 3)

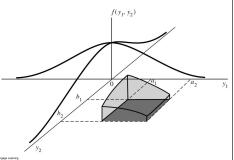
- 1. Finish coverage of conditional probability distributions
- 2. Independent random variables
- 3. Expected values of (general) functions of RVs
- 4. Covariances and correlation coefficients
- 5. Expected values, variances, covariances, and correlations of linear functions of RVs
- 6. Conditional expectations

Graphical representation of univariate vs. bivariate PDFs



Bivariate





Recall that in the univariate case, area under the PDF between a and $b = P(a \le Y \le b)$

$$P(a \le Y \le b)$$

$$= \int_{a}^{b} f(y) dy$$

Bivariate PDFs are 2-dimensional surfaces, so the volume under the surface corresponds to a probability – e.g., above:

$$P(a_1 \le Y_1 \le a_2, b_1 \le Y_2 \le b_2)$$

= $\int_{b_1}^{b_2} \int_{a_1}^{a_2} f(y_1, y_2) dy_1 dy_2$

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Marginal probability distributions

Discrete RVs:

$$p_1(y_1) = \sum_{\text{all } y_2} p(y_1, y_2) \text{ and } p_2(y_2) = \sum_{\text{all } y_1} p(y_1, y_2)$$

Table 5.2 Joint probability function for Y_1 and Y_2 , Example 5.5

		y_1		
<i>y</i> ₂	0	1	2	p ₂ (y ₂): Total
0	0	3/15	3/15	6/15
1	2/15	6/15	0	8/15
2	1/15	0	0	1/15
Total	3/15	9/15	3/15	1

Continuous RVs:

 $p_1(y_1)$:

$$f_1(y_1) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_2$$
 and $f_2(y_2) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_1$

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Conditional probability distributions

Recall that
$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Discrete RVs:

$$p(y_1 | y_2) = \frac{p(y_1, y_2)}{p_2(y_2)} \iff P(Y_1 = y_1 | Y_2 = y_2) = \frac{P(Y_1 = y_1, Y_2 = y_2)}{P(Y_2 = y_2)}$$

Continuous RVs:

$$f(y_1|y_2) = \frac{f(y_1, y_2)}{f_2(y_2)}$$
 and $f(y_2|y_1) = \frac{f(y_1, y_2)}{f_1(y_1)}$

Conditional CDF:

 $F(y_1 | y_2)$ means $P(Y_1 \le y_1 | Y_2 = y_2)$

Compute by integrating the conditional PDF over the relevant range.

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EX) Conditional probability distributions for discrete RVs

Continuing with our previous example about a 2-person committee drawn from Democrats, Republicans, and independents (with bivariate probability distribution below), find the conditional probability distribution of Y_1 given that Y_2 =1.

$$p(y_1 | y_2) = \frac{p(y_1, y_2)}{p_2(y_2)}$$

Answer

$$P(Y_1 = 0 | Y_2 = I) = \frac{p(0, I)}{p_2(I)} = \frac{2/15}{8/15} = \frac{I}{4}$$

$$P(Y_1 = I | Y_2 = I) = \frac{p(I, I)}{p_2(I)} = \frac{6/15}{8/15} = \frac{3}{4}$$

$$P(Y_1 = 2 | Y_2 = I) = \frac{p(2, I)}{p_2(I)} = \frac{0/15}{8/15} = 0$$

Table 5.2 Joint probability function for Y_1 and Y_2 , Example 5.5

	<i>y</i> 1			$p_2(y_2)$:
y_2	0	1	2	Total
0	0	3/15	3/15	6/15
1	2/15	6/15	0	8/15
2	1/15	0	0	1/15
Total	3/15	9/15	3/15	1
	0 1 2	0 0 1 2/15 2 1/15	$\begin{array}{c cccc} y_2 & 0 & 1 \\ 0 & 0 & 3/15 \\ 1 & 2/15 & 6/15 \\ 2 & 1/15 & 0 \end{array}$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

Graded in-class exercise

EX) Conditional probability distributions for continuous RVs

Find the conditional probability, $P(Y_1 \le 0.5 \mid Y_2 = 1.5)$, for the continuous bivariate PDF:

$$f(y_1, y_2) = \begin{cases} 0.5, \ 0 \le y_1 \le y_2 \le 2\\ 0, \text{ elsewhere} \end{cases} \qquad f(y_1 | y_2) = \frac{f(y_1, y_2)}{f_2(y_2)}$$

Conditional PDF:

$$f(y_1 | y_2) = \frac{f(y_1, y_2)}{f_2(y_2)}$$

Independent random variables

What is $P(A \cap B)$ if events A and B are **independent**?

$$P(A \cap B) = P(A)P(B)$$

Can use similar approach to see if two RVs, Y_1 and Y_2 , are independent:

CDF: $F(y_1, y_2) = F_1(y_1)F_2(y_2)$

Discrete probability distribution : $p(y_1, y_2) = p_1(y_1)p_2(y_2)$

Continuous probability density function: $f(y_1, y_2) = f_1(y_1) f_2(y_2)$

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EX) Checking for independence – discrete RVs

Continuing with our previous example about a 2-person committee drawn from Democrats, Republicans, and independents (with bivariate probability distribution below), is Y_1 independent of Y_2 ? Can check any point, but let's try (0,0).

Table 5.2 Joint probability function for Y_1 and Y_2 , Example 5.5

		y_1			$p_2(y_2)$:
	y_2	0	1	2	Total
	0	0	3/15	3/15	6/15
	1	2/15	6/15	0	8/15
	2	1/15	0	0	1/15
$p_I(y_I)$:	Total	3/15	9/15	3/15	1
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Discrete probability distribution if independent : $p(y_1, y_2) = p_1(y_1)p_2(y_2)$

EX) Checking for independence – continuous RVs

Suppose Y_1 and Y_2 have the continuous bivariate PDF below. Are these two RVs independent?

$$f(y_1, y_2) = \begin{cases} 6y_1 y_2^2, & 0 \le y_1 \le 1, \\ 0, & \text{elsewhere} \end{cases}$$

Continuous probability density function if independent: $f(y_1, y_2) = f_1(y_1) f_2(y_2)$

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Another way to check for independence

 Y_1 and Y_2 are independent RVs if:

$$f(y_1, y_2) = g(y_1)h(y_2)$$

where $g(y_I)$ is a nonnegative function of y_I alone and $h(y_2)$ is a nonnegative function of y_2 alone ***AND only if

$$f(y_1, y_2) > 0$$
 for $a \le y_1 \le b$ and $c \le y_2 \le d$ for constants $a, b, c, d ***$.

EX) Suppose Y_1 and Y_2 have the continuous bivariate PDF below. Are these two RVs independent?

$$f(y_1, y_2) = \begin{cases} 2y_1, & 0 \le y_1 \le 1, \\ 0, & \text{elsewhere} \end{cases}$$

Another example

$$f(y_1, y_2) = \begin{cases} 8y_1 y_2, & 0 \le y_1 \le 1, 0 \le y_2 \le y_1 \\ 0, & \text{elsewhere} \end{cases}$$

We cannot use the alternative method here because the range over which y_2 has positive probability is a function of y_1 .

If we found the marginal distributions for y_1 and y_2 , we would see that:

$$f_I(y_I) = 4y_I^3$$

$$f_2(y_2) = 4y_2(1-y_2^2)$$

So $f(y_1, y_2) \neq f_1(y_1) f_2(y_2)$, thus Y_1 and Y_2 not independent

The expected value of a function of RVs

Recall that in the *univariate* case,

Discrete RVs:
$$E[g(Y)] = \sum_{i} g(y_i)p(y_i)$$

Continuous RVs:
$$E[g(Y)] = \int_{-\infty}^{\infty} g(y) f(y) dy$$

For the *bivariate* case,

Discrete RVs:

Continuous RVs:

$$E[g(Y_1, Y_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(y_1, y_2) f(y_1, y_2) dy_1 dy_2$$

Expected value of a function of RVs – example

Let Y_1 and Y_2 have joint density given by

$$f(y_1, y_2) = \begin{cases} 2y_1, & 0 \le y_1 \le 1, 0 \le y_2 \le 1, \\ 0, & \text{elsewhere.} \end{cases}$$

Find $E(Y_1Y_2)$.

$$E[g(Y_1, Y_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(y_1, y_2) f(y_1, y_2) dy_1 dy_2$$

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Expected value of a function of RVs - Rules

1.
$$E[cg(Y_1, Y_2)] = cE[g(Y_1, Y_2)]$$
 for any constant c

2.
$$E[g_I(Y_1, Y_2) + g_2(Y_1, Y_2) + ... + g_k(Y_1, Y_2)]$$

= $E[g_I(Y_1, Y_2)] + E[g_2(Y_1, Y_2)] + ... + E[g_k(Y_1, Y_2)]$

3.
$$E(c_1Y_1 + c_2Y_2 + ... + c_kY_k) = c_1E(Y_1) + c_2E(Y_2) + ... + c_kE(Y_k)$$

4. If
$$Y_1$$
 and Y_2 are independent then
$$E[g(Y_1)h(Y_2)] = E[g(Y_1)]E[h(Y_2)]$$

Expected value of a function of RVs - Rules - examples

a. What is $E(Y_1 - Y_2)$ if $E(Y_1) = 0.75$ and $E(Y_2) = 0.375$

b. What is $E(3Y_1 - 7Y_2)$?

c. What is $E(Y_1Y_2)$ if Y_1 and Y_2 are independent?

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Covariance & correlation coefficient of 2 RVs

Recall the formula for the variance of an RV:

$$V(Y) = E\Big[(Y - \mu)^2\Big] = E\Big[(Y - \mu)(Y - \mu)\Big] = E(Y^2) - \Big[E(Y)\Big]^2 = E(Y^2) - \mu^2$$

The formula for the covariance of Y_1 and Y_2 is similar:

$$Cov(Y_1, Y_2) = E[(Y_1 - \mu_1)(Y_2 - \mu_2)] = E(Y_1 Y_2) - E(Y_1)E(Y_2) = E(Y_1 Y_2) - \mu_1 \mu_2$$

where $\mu_1 = E(Y_1)$ and $\mu_2 = E(Y_2)$.

It is difficult to interpret the magnitude of the covariance, but we can standardize it to get the correlation coefficient, ρ . Note that $-1 \le \rho \le 1$.

$$Corr(Y_1, Y_2) \equiv \rho = \frac{Cov(Y_1, Y_2)}{\sigma_1 \sigma_2}$$

where σ_1 and σ_2 are the standard deviations of Y_1 and Y_2 .

What is the relationship between the signs of Covand Corr?

Covariance & correlation coefficient of 2 RVs

- What does it mean to have a positive covariance (or correlation)? Negative? Zero?
- Correlation = +1? -1?

 Perfect Positive Correlation

 Perfect Negative Correlation

 Perfect Neg

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Covariance, correlation, & independence

What is $E(Y_1Y_2)$ if Y_1 and Y_2 are **independent**?

$$E(Y_1Y_2) = E(Y_1)E(Y_2)$$
 if Y_1 and Y_2 are **independent**

What is $Cov(Y_1Y_2)$ if Y_1 and Y_2 are **independent**?

$$Cov(Y_1, Y_2) = E(Y_1Y_2) - E(Y_1)E(Y_2) = E(Y_1)E(Y_2) - E(Y_1)E(Y_2) = 0$$

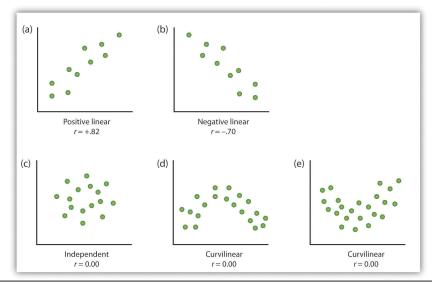
if Y_1 and Y_2 are **independent**

What is $Corr(Y_1, Y_2)$ if Y_1 and Y_2 are **independent**?

$$Corr(Y_1, Y_2) = \frac{Cov(Y_1, Y_2)}{\sigma_1 \sigma_2} = 0$$
 if Y_1 and Y_2 are **independent**

***Note: independence implies zero covariance (correlation)
BUT zero covariance (correlation) does NOT imply independence.
Why?

Because <u>covariance</u> & <u>correlation</u> are about <u>linear</u> dependence and it is possible for two variables to have a <u>non-linear relationship</u> but no linear relationship



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Calculating the covariance – example #1

$$f(y_1, y_2) = \begin{cases} 2y_1, & 0 \le y_1 \le 1, 0 \le y_2 \le 1, \\ 0, & \text{elsewhere.} \end{cases}$$

For this joint PDF, $E(Y_1Y_2)=1/3$, $E(Y_1)=2/3$, and $E(Y_2)=1/2$. Find $Cov(Y_1, Y_2)$. Does the answer surprise you? Why or why not?

$$Cov(Y_1, Y_2) = E(Y_1Y_2) - E(Y_1)E(Y_2)$$

Calculating the covariance – example #2

Show that Y_1 and Y_2 are dependent but have zero covariance.

		<i>y</i> ₁	
<i>y</i> ₂	-1	0	+1
-1	1/16	3/16	1/16
0	3/16	0	3/16
+1	1/16	3/16	1/16

$$E[g(Y_1, Y_2)] = \sum_{\text{all } y_2} \sum_{\text{all } y_1} g(y_1, y_2) p(y_1, y_2) \frac{1}{Cov(Y_1, Y_2) = E(Y_1 Y_2) - E(Y_1) E(Y_2)}$$

Rules for the expected value, variance, and covariance of linear functions of RVs:

The bivariate case (see WMS pp. 271-273 for proof & multivariate case)

Random variables Y_1 and Y_2 , and constants a_1, a_2, b_1 and b_2 :

1.
$$E(a_1Y_1 + b_1 + a_2Y_2 + b_2) = a_1E(Y_1) + b_1 + a_2E(Y_2) + b_2$$

EX) $E(3Y_1 - 2 - 8Y_2 + 5)$

2.
$$V(a_1Y_1 + b_1 + a_2Y_2 + b_2) = a_1^2V(Y_1) + a_2^2V(Y_2) + 2a_1a_2Cov(Y_1, Y_2)$$

EX)
$$V(Y_1 + Y_2)$$

EX)
$$V(Y_1 - Y_2)$$

EX)
$$V(3Y_1 - 2 - 8Y_2 + 5)$$

3.
$$Cov(a_1Y_1 + b_1, a_2Y_2 + b_2) = a_1a_2Cov(Y_1, Y_2)$$

EX)
$$Cov(Y_1, -Y_2)$$

EX)
$$Cov(3Y_1 - 2, -8Y_2 + 5)$$

Conditional expectations

Motivation

- Covariance and correlation measure the <u>linear</u> relationship (linear dependence) between two RVs and treat them symmetrically
- In applied economics, we often want to explain one RV (Y) in terms of another RV (X)
- Call Y the "explained" variable, X the "explanatory" variable
- Recall conditional probability distributions and PDFs: p(y|x) and f(y|x)
- We are often interested in the <u>conditional expectation</u>
 (a.k.a. the <u>conditional mean</u>):
 E(Y|X=x) or, for shorthand, E(Y|X) or sometimes E(Y|x)

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Conditional expectations & variances - formulas

Conditional expectation of Y given X

Discrete RVs:

$$E(Y | X = x) = E(Y | X) = \sum_{\text{all } y} y \ p(y | x)$$

Continuous RVs:

$$E(Y \mid X = x) = E(Y \mid X) = \int_{-\infty}^{\infty} y f(y \mid x) dy$$

How would you use the E[g(Y)|X] formula to find the conditional variance, V(Y|X)?

$$V(Y|X)$$

$$= E(Y^2|X) - \left[E(Y|X)\right]^2$$

Conditional expectation of g(Y) given X

Discrete RVs:

$$E[g(Y) | X = x] = E[g(Y) | X] = \sum_{\text{all } y} g(y) p(y | x)$$

Continuous RVs:

$$E[g(Y) | X = x] = E[g(Y) | X] = \int_{-\infty}^{\infty} g(y) f(y | x) dy$$

Conditional expectations & variances - rules Gist: treat the variable you are conditioning on as a constant

1.
$$E[g(Y)|Y] = g(Y)$$
 for any function $g(.)$

 $EX) E(Y^2 | Y)$

$$2. E[g(X)Y | X] = g(X)E(Y | X)$$

EX)
$$E(2X^2Y|X)$$

3. If *X* and *Y* are independent,

then
$$E(Y|X) = E(Y)$$
 and $V(Y|X) = V(Y)$

4. If
$$E(Y|X) = E(Y)$$
, then $Cov(X,Y) = 0$

5.
$$E[E(Y|X)] = E(Y)$$
 "the law of iterated expectations"

EX) If $E(WAGE \mid EDUC) = 4 + 0.6 \, EDUC$ and E(EDUC) = 11.5, find E(WAGE).

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Homework:

- WMS Ch. 5 (part 3 of 3)
 - Expected value of a function of RVs & special theorems: 5.72, 5.74
 - Covariance: 5.89, 5.91, 5.92 (Hint: $E(Y_1)=0.25$ and $E(Y_2)=0.5$)
 - Expected values, variances, covariances, and correlations of linear functions of RVs: 5.102, 5.103 (consult Theorem 5.12), 5.110
 - Conditional expectations: none but review & internalize the rules (and include them on your cheat sheet!)

Next class:

 Finish Ch. 5 (if need be) and answer any questions you have about the material for the midterm (Ch. 1-5 in WMS and integration)

In-class exercise #1: calculating the covariance

$$f(y_1, y_2) = \begin{cases} 3y_1, & 0 \le y_2 \le y_1 \le 1, \\ 0, & \text{elsewhere.} \end{cases}$$

 $E(Y_1)=0.75$, $E(Y_2)=0.375$. Find $Cov(Y_1, Y_2)$.

$$E[g(Y_1, Y_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(y_1, y_2) f(y_1, y_2) dy_1 dy_2$$

$$Cov(Y_1, Y_2) = E(Y_1Y_2) - E(Y_1)E(Y_2)$$

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In-class exercise #2: variance of a linear function of <u>3</u> random variables

Use Theorem 5.12 (copied on the next slide) and find the formula for:

$$V(a_1Y_1 + b_1 + a_2Y_2 + b_2 + a_3Y_3 + b_3)$$

Rules for the expected value, variance, and covariance of <u>linear</u> functions of RVs:

The general multivariate case Proof on pp. 272-3

Let Y_1, Y_2, \ldots, Y_n and X_1, X_2, \ldots, X_m be random variables with $E(Y_i) = \mu_i$ and $E(X_i) = \xi_i$. Define

$$U_1 = \sum_{i=1}^{n} a_i Y_i$$
 and $U_2 = \sum_{j=1}^{m} b_j X_j$

for constants a_1, a_2, \ldots, a_n and b_1, b_2, \ldots, b_m . Then the following hold:

- **a** $E(U_1) = \sum_{i=1}^{n} a_i \mu_i$.
- **b** $V(U_1) = \sum_{i=1}^n a_i^2 V(Y_i) + 2 \sum_{1 \le i < j \le n} a_i a_j \text{Cov}(Y_i, Y_j)$, where the double sum is over all pairs (i, j) with i < j.
- **c** $Cov(U_1, U_2) = \sum_{i=1}^{n} \sum_{j=1}^{m} a_i b_j Cov(Y_i, X_j).$