Landscape Analysis for Surrogate Models in the Evolutionary Black-Box Context (Supplementary Material)

Zbyněk Pitra

z.pitra@gmail.com

Faculty of Nuclear Sciences and Physical Engineering, Czech Technical University Břehová 7, 115 19 Prague, Czech Republic

Jan Koza koza@cs.cas.cz

Faculty of Information Technology, CTU in Prague, Thákurova 9, 160 00 Prague 6, Czech Republic

Jiří Tumpach tumpach@cs.cas.cz

Faculty of Mathematics and Physics, Charles University in Prague, Malostran. nám. 25, 11800 Prague, Czech Republic

Martin Holeňa martin@cs.cas.cz

Institute of Computer Science, Czech Academy of Sciences, Pod Vodárenskou věží 2, 182 07 Prague, Czech Republic

Abstract

Surrogate modeling has become a valuable technique for black-box optimization tasks with expensive evaluation of the objective function. In this supplementary material for the original article *Landscape Analysis for Surrogate Models in the Evolutionary Black-Box Context* (Pitra et al., 2022), we report additional information for investigation of the relationship between the predictive accuracy of surrogate models and features of the black-box function landscape. We also report additional information on study of properties of features for landscape analysis in the context of different transformations and ways of selecting the input data.

Keywords

Adaptive learning, Metalearning, Black-box optimization, Surrogate model, Landscape analysis

1 Introduction

In (Pitra et al., 2022) we have presented an investigation of relationships between the prediction error of the surrogate models, their settings, and features describing the fitness landscape during evolutionary optimization by the Covariance Matrix Adaptation Evolution Strategy (CMA-ES) (Hansen and Ostermeier, 1996; Hansen, 2006) state-of-the-art optimizer for expensive continuous black-box tasks. Its pseudocode is outlined in Algorithm 1. Four models in 39 different settings for the Doubly Trained Surrogate CMA-ES (DTS-CMA-ES) by Pitra et al. (2016), a surrogate-assisted version of the CMA-ES, were compared using three sets of 14 landscape features selected according to their properties, especially according to their robustness and similarity to other features. We analysed MSE and RDE dependence of various models and model settings

Algorithm 1 Pseudocode of the CMA-ES

Input: λ (population-size), \mathfrak{b} (original fitness function), $(w_i)_{i=1}^{\mu}$, μ_w (selection and recombination parameters), c_{σ} , d_{σ} (step-size control parameters), c_c , c_I , c_{μ} (covariance matrix adaptation parameters)

```
1: \operatorname{init} g = 0, \sigma^{(0)} > 0, \mathbf{m}^{(0)} \in \mathbb{R}^D, \mathbf{C}^{(0)} = \mathbf{I}, \mathbf{p}_{\sigma}^{(0)} = \mathbf{0}, \mathbf{p}_{c}^{(0)} = \mathbf{0}, \mu = \lfloor \lambda/2 \rfloor

2: \operatorname{repeat}

3: \mathbf{x}_k \leftarrow \mathbf{m}^{(g)} + \sigma^{(g)} \mathcal{N}(\mathbf{0}, \mathbf{C}^{(g)}) k = 1, \dots, \lambda

4: \mathfrak{w}_k \leftarrow \mathfrak{w}(\mathbf{x}_k) k = 1, \dots, \lambda

5: \mathbf{m}^{(g+1)} \leftarrow \sum_{i=1}^{\mu} w_i \mathbf{x}_{i:\lambda}, where \mathbf{x}_{i:\lambda} is the i-th best individual out of \mathbf{x}_1, \dots, \mathbf{x}_{\lambda}

6: \mathbf{p}_{\sigma}^{(g+1)} \leftarrow (1 - c_{\sigma}) \mathbf{p}_{\sigma}^{(g)} + \sqrt{c_{\sigma}(2 - c_{\sigma})\mu_{w}} \mathbf{C}^{(g)^{-1/2}} \frac{\mathbf{m}^{(g+1)} - \mathbf{m}^{(g)}}{\sigma^{(g)}}

7: h_{\sigma} \leftarrow \mathbb{I} \left( \| \mathbf{p}_{\sigma}^{(g+1)} \| < \sqrt{1 - (1 - c_{\sigma})^{2(g+1)}} \left( 1.4 + \frac{2}{D+1} \right) E \| \mathcal{N}(0, 1) \| \right)

8: \mathbf{p}_{c}^{(g+1)} \leftarrow (1 - c_{c}) \mathbf{p}_{c}^{(g)} + h_{\sigma} \sqrt{c_{c}(2 - c_{c})\mu_{w}} \frac{\mathbf{m}^{(g+1)} - \mathbf{m}^{(g)}}{\sigma^{(g)}}

9: \mathbf{C}_{\mu} \leftarrow \sum_{i=1}^{\mu} w_{i} \sigma^{(g)^{-2}} \left( \mathbf{x}_{i:\lambda}^{(g+1)} - \mathbf{m}^{(g)} \right) \left( \mathbf{x}_{i:\lambda}^{(g+1)} - \mathbf{m}^{(g)} \right)^{\top}

10: \mathbf{C}^{(g+1)} \leftarrow (1 - c_{1} - c_{\mu}) \mathbf{C}^{(g)} + c_{1} \mathbf{p}_{c}^{(g+1)} \mathbf{p}_{c}^{(g+1)} + c_{\mu} \mathbf{C}_{\mu}

11: \sigma^{(g+1)} \leftarrow \sigma^{(g)} \exp \left( \frac{c_{\sigma}}{d_{\sigma}} \left( \frac{\| \mathbf{p}_{\sigma}^{(g+1)} \|}{E \| \mathcal{N}(\mathbf{0}, \mathbf{I}) \|} - 1 \right) \right)

12: g \leftarrow g + 1

13: \mathbf{until} stopping criterion is \mathbf{met}

Output: \mathbf{x}_{\text{res}} (resulting optimum)
```

on the features calculated using three methods selecting sample sets extracted from DTS-CMA-ES runs on noiseless benchmarks from the Comparing Continuous Optimisers (COCO) benchmark function testbed (Hansen et al., 2016). Up to our knowledge, this analysis of landscape feature properties and their connection to surrogate models in evolutionary optimization is more profound than any other published so far.

The employed training set selection (TSS) methods were:

- \diamond *TSS full*, taking all the already evaluated points, i. e., $\mathcal{T} = \mathcal{A}$,
- ⋄ TSS knn, selecting the union of the sets of k-nearest neighbors of all points for which the fitness should be predicted, where k is user defined (e.g., in Kern et al. (2006)),
- \diamond *TSS nearest*, selecting the union of the sets of k-nearest neighbors of all points for which the fitness should be predicted, where k is maximal such that the total number of selected points does not exceed a given maximum number of points N_{max} and no point is further from current mean $\mathbf{m}^{(g)}$ than a given maximal distance r_{max} (e. g., in Bajer et al. (2019)).

We have performed all the feature investigations for each TSS method separately. Employed extracted *basic sample sets* were:

- ♦ *archive A* containing all points evaluated so far by the original fitness;
- \diamond training set \mathcal{T} selected from \mathcal{A} used for model training.

By combining the two basic sample sets with a population \mathcal{P} consisting of the points without a known value of the original fitness to be evaluated by the surrogate model, we have obtained two new sets $\mathcal{A}_{\mathcal{P}} = \mathcal{A} \cup \mathcal{P}$ and $\mathcal{T}_{\mathcal{P}} = \mathcal{T} \cup \mathcal{P}$. We have utilized either

transformed and non-transformed sets for feature calculations, resulting in 8 different sample sets (4 in case of TSS full due to $\mathcal{T}=\mathcal{A}$): $\mathcal{A}, \mathcal{A}^{\scriptscriptstyle T}, \mathcal{A}_{\mathcal{P}}, \mathcal{A}^{\scriptscriptstyle T}_{\mathcal{P}}, \mathcal{T}, \mathcal{T}^{\scriptscriptstyle T}, \mathcal{T}_{\mathcal{P}}, \mathcal{T}^{\scriptscriptstyle T}_{\mathcal{P}}$, where $^{\scriptscriptstyle T}$ denotes transformation into the $\sigma^2\mathbf{C}$ basis.

The supplementary material is structured as follows. Section 2 describes landscape features utilized in (Pitra et al., 2022) in more detail. Sections 3 and 4 supply the testing and analysis, respectively, in the original paper by more detailed results.

2 Feature Definitions

Let us consider an input set X of N points

$$\mathbf{X} = \left\{ \mathbf{x}_i \middle| \mathbf{x}_i \in \sum_{j=1}^{D} [l_j, u_j], l_j, u_j \in \mathbb{R}, l_j < u_j \right\}_{i=1}^{N}.$$

Then we can define a *sample set* S of N pairs of observations

$$S = \{(\mathbf{x}_i, y_i) \mid \mathbf{x}_i \in \mathbf{X}, y_i \in \mathbb{R} \cup \{\circ\}, i = 1, \dots, N\},$$

where \circ denotes missing y_i value (e. g., \mathbf{x}_i was not evaluated yet). Then the sample set can be utilized to describe landscape properties using a *landscape feature* φ

$$\varphi: \bigcup_{N \in \mathbb{N}} \mathbb{R}^{N,D} \times (\mathbb{R} \cup \{\circ\})^{N,1} \mapsto \mathbb{R} \cup \{\pm \infty, \bullet\},$$

where \bullet denotes impossibility of feature computation. The y_i values can be gathered to a set $\mathbf{y} = \{y_i\}_{i=1}^N$.

The following features were reimplemented in Matlab according to the R-package flacco by Kerschke and Dagefoerde (2017). In the words of Kerschke (2017a), "It is important to notice that, independent of the research domain, most of these features do not provide intuitively understandable numbers. Therefore, we strongly recommend not to interpret them on their own. Moreover, some of them are stochastic and hence should be evaluated multiple times on an instance and afterwards be aggregated in a reasonable manner. Nevertheless, they definitely provide information that can be of great importance to scientific models, such as machine learning algorithms in general or algorithm selectors in particular." Following the remark about stochasticity, we have fixed random seeds which are not fixed in the original implementation to always return the same values for identical input.

2.1 Basic Features Φ_{Basic}

Kerschke (2017b) summarized the following features providing obvious information about the input space:

- \diamond **Dimension** of the input space $\varphi_{\text{dim}} = D$.
- $\diamond \ \ \mathbf{Number \ of \ observations} \ \varphi_{\mathrm{obs}} = N.$
- \diamond Minimum and maximum of lower bounds $\varphi_{\text{lower_min}} = \min_{i \in D} l_i$, $\varphi_{\text{lower_max}} = \max_{i \in D} l_i$.
- \diamond Minimum and maximum of upper bounds $\varphi_{\text{upper_min}} = \min_{i \in D} u_i$, $\varphi_{\text{upper_max}} = \max_{i \in D} u_i$.
- \diamond Minimum and maximum of y values $\varphi_{\texttt{objective_min}} = \min_{i \in N} y_i$, $\varphi_{\texttt{objective_max}} = \max_{i \in N} y_i$.
- \diamond Minimum and maximum of cell blocks per dimension $\varphi_{\texttt{blocks_min}}$, $\varphi_{\texttt{blocks_max}}$.
- \diamond Total number of cells $\varphi_{\tt cells_total}.$
- $\diamond \ \ \mathbf{Number \ of \ filled \ cells} \ \varphi_{\mathtt{cells_filled}}.$
- \diamond Binary flag stating whether the **objective function** should be **minimized** $\varphi_{\min \text{minimize},\text{fun}}$.

In this research, we have used only φ_{dim} and φ_{obs} because the remaining ones were constant on the whole \mathcal{D}_{100} dataset. φ_{dim} is identical regardless the sample set and φ_{obs} was not computed in the $\sigma^2\mathbf{C}$ basis because it does not influence the resulting feature values.

2.2 CMA Features Φ_{CMA}

In each CMA-ES generation g during the fitness evaluation step (Algorithm 1, step 4), we have defined the following features in Pitra et al. (2019) for the set of points X:

- \diamond Generation number $\varphi_{\tt generation}^{\tt CMA} = g$ indicates the phase of the optimization process.
- \diamond **Step-size** $\varphi^{\text{CMA}}_{\text{step-size}} = \sigma^{(g)}$ provides an information about the extent of the approximated region.
- \diamond Number of restarts $\varphi^{\rm CMA}_{\rm restart}=n^{(g)}_{\rm r}$ performed till generation g may indicate land-scape difficulty.
- \diamond **Mahalanobis mean distance** of the CMA-ES mean $\mathbf{m}^{(g)}$ to the sample mean $\mu_{\mathbf{X}}$ of \mathbf{X}

$$\varphi_{\text{mean_dist}}^{\text{CMA}}(\mathbf{X}) = \sqrt{(\mathbf{m}^{(g)} - \mu_{\mathbf{X}})^{\top} \mathbf{C}_{\mathbf{X}}^{-1} (\mathbf{m}^{(g)} - \mu_{\mathbf{X}})},$$
(1)

where C_X is the sample covariance of X. This feature indicates suitability of X for model training from the point of view of the current state of the CMA-ES algorithm.

- \diamond Square of the \mathbf{p}_c evolution path length $\varphi_{\mathtt{evopath_c_norm}}^{\mathtt{CMA}} = \|\mathbf{p}_c^{(g)}\|^2$ is the only possible non-zero eigenvalue of *rank-one update* covariance matrix $\mathbf{p}_c^{(g+1)}\mathbf{p}_c^{(g+1)}^{\mathsf{T}}$ (see step 10 in Algorithm 1). That feature providing information about the correlations between consecutive CMA-ES steps indicates a similarity of function landscapes among subsequent generations.
- \diamond \mathbf{p}_{σ} **evolution path** ratio, i. e., the ratio between the evolution path length $\|\mathbf{p}_{\sigma}^{(g)}\|$ and the expected length of a random evolution path used to update step-size. It provides a useful information about distribution changes:

$$\varphi_{\text{evopath_s_norm}}^{\text{CMA}} = \frac{\|\mathbf{p}_{\sigma}^{(g)}\|}{E \|\mathcal{N}(\mathbf{0}, \mathbf{I})\|} = \frac{\|\mathbf{p}_{\sigma}^{(g)}\| \Gamma(\frac{D}{2})}{\sqrt{2} \Gamma(\frac{D+1}{2})}.$$
 (2)

♦ CMA similarity likelihood. The log-likelihood of the set of points X with respect to the CMA-ES distribution may also serve as a measure of its suitability for training

$$\varphi_{\text{cma-lik}}^{\text{CMA}}(\mathbf{X}) = -\frac{N}{2} \left(D \log 2\pi \sigma^{(g)}^2 + \log \det \mathbf{C}^{(g)} \right) - \frac{1}{2} \sum_{\mathbf{x} \in \mathbf{X}} \left(\frac{\mathbf{x} - \mathbf{m}^{(g)}}{\sigma^{(g)}} \right)^{-1} \mathbf{C}^{(g)} - 1 \left(\frac{\mathbf{x} - \mathbf{m}^{(g)}}{\sigma^{(g)}} \right).$$
(3)

2.3 Dispersion Features Φ_{Dis}

The dispersion features by Lunacek and Whitley (2006) compare the dispersion among observations from S and among a subset of these points. The subsets are created based on tresholds using the quantiles 0.02, 0.05, 0.1, and 0.25 of the objective values y.

The set of all distances within the set S:

$$D_{\text{all}} = \left\{ \text{dist}(\mathbf{x}_i, \mathbf{x}_j) \mid i, j = 1, \dots, N \right\}. \tag{4}$$

The quantile subset of all distances:

$$D_{\text{quantile}} = \left\{ \text{dist}(\mathbf{x}_i, \mathbf{x}_j) \mid y_i, y_j \le Q_{\text{quantile}}(\mathbf{y}), \right\}. \tag{5}$$

⋄ Ratio of the quantile subset and all points median distances

$$\varphi_{\text{ratio_median_[quantile]}}^{\text{Dis}}(\mathcal{S}) = \frac{\text{median } D_{\text{quantile}}}{\text{median } D_{\text{all}}}.$$
 (6)

 $\diamond~$ Ratio of the quantile subset and all points mean distances

$$\varphi_{\text{ratio_mean_[quantile]}}^{\text{Dis}}(\mathcal{S}) = \frac{\text{mean } D_{\text{quantile}}}{\text{mean } D_{\text{all}}}.$$
(7)

⋄ **Difference between** the quantile subset and all points **median distances**

$$\varphi_{\text{diff}_median_[quantile]}^{\text{Dis}}(\mathcal{S}) = \text{median } D_{\text{quantile}} - \text{median } D_{\text{all}}.$$
(8)

♦ **Difference between** the quantile subset and all points **mean distances**

$$\varphi_{\text{diff_mean_[quantile]}}^{\text{Dis}}(\mathcal{S}) = \operatorname{mean} D_{\text{quantile}} - \operatorname{mean} D_{\text{all}}.$$
(9)

2.4 Information Content Features Φ_{Inf}

The Information Content of Fitness Sequences by Muñoz et al. (2015) approach is based on a symbol sequence $\Psi = \{\psi_1, \dots, \psi_{N-1}\}$, where

$$\psi_{i} = \begin{cases} \bar{1}, & \text{if} \quad \frac{y_{i+1} - y_{i}}{\|\mathbf{x}_{i+1} - \mathbf{x}_{i}\|} < -\varepsilon, \\ 0, & \text{if} \quad \left| \frac{y_{i+1} - y_{i}}{\|\mathbf{x}_{i+1} - \mathbf{x}_{i}\|} \right| \leq \varepsilon, \\ 1, & \text{if} \quad \frac{y_{i+1} - y_{i}}{\|\mathbf{x}_{i+1} - \mathbf{x}_{i}\|} > \varepsilon. \end{cases}$$

$$(10)$$

The sequence considers observations from a sample set S as a random walk accross the objective landscape and depends on the information sensitivity parameter $\varepsilon > 0$.

The symbol sequence Ψ is aggregated by the information content $H(\varepsilon) = -\sum_{i \neq j} p_{ij} \log_6 p_{ij}$, where p_{ij} is the probability of having the block $\psi_i \psi_j$, where $\psi_i, \psi_j \in \{\bar{1}, 0, 1\}$, within the sequence. Note that the base of the logarithm was set to six as this equals the number of possible blocks $\psi_i \psi_j$ for which $\psi_i \neq \psi_j$, i.e., $\psi_i \psi_j \in \{\bar{1}0, 0\bar{1}, \bar{1}1, 1\bar{1}, 01, 10\}$.

Another aggregation of the information is the so-called partial information content $M(\varepsilon) = |\Psi'|/(N-1)$, where Ψ' is the symbol sequence of alternating $\bar{1}$ s and 1s, which is derived from Ψ by removing all 0 and repeated symbols.

When we do consider \circ as a valid state of y, the sequence Ψ consists of $\psi_i \in \{\bar{1},0,1,\bar{N}\}$, where $\psi_i = \bar{N}$, if $y_{i+1} = \circ$ or $y_i = \circ$. Thus, $H(\varepsilon) = -\sum_{i\neq j} p_{ij} \log_{12} p_{ij}$ due to the increased number of possible $\psi_i\psi_j$ blocks, i. e., $\psi_i\psi_j \in \{\bar{1}0,0\bar{1},\bar{1}1,1\bar{1},01,10,\bar{N}\bar{1},\bar{1}\bar{N},\bar{N}0,0\bar{N},\bar{N}1,1\bar{N}\}$. Therefore, features based on Ψ' can be utilized only when \circ state is not present in \mathbf{y} .

Based on sequences Ψ and Ψ' the following features can be defined according to Muñoz et al. (2015):

- \diamond Maximum information content $\varphi_{h,\max}^{Inf}(\mathcal{S}) = \max_{\varepsilon} H(\varepsilon)$.
- \diamond Settling sensitivity $\varphi_{\text{eps.s}}^{\text{Inf}}(\mathcal{S}) = \log_{10} \min(\varepsilon | H(\varepsilon) < s)$, where default s = 0.05 (see Muñoz et al. (2015)).
- $\diamond \ \, \mathbf{Maximum \ sensitivity} \ \varphi^{\mathrm{Inf}}_{\mathrm{eps_max}}(\mathcal{S}) = \mathrm{arg} \ \mathrm{max}_{\varepsilon} \, H(\varepsilon).$
- \diamond Initial partial information $\varphi_{m0}^{\text{Inf}}(\mathcal{S}) = M(\varepsilon = 0)$.
- \diamond Ratio of partial information sensitivity $\varphi^{\inf}_{\text{eps.ratio}}(\mathcal{S}) = \log_{10} \max(\varepsilon|M(\varepsilon)) > r\varphi^{\inf}_{m0}(\mathcal{S}))$, where default r=0.5 (see also Muñoz et al. (2015)).

2.5 Levelset Features Φ_{Lvl}

In Levelset features by Mersmann et al. (2011), the sample set \mathcal{S} is split into two classes by a specific treshold calculated using the quantiles 0.1, 0.25, and 0.5. Linear, quadratic, and mixture discriminant analysis (lda, qda, and mda) are used to predict whether the objective values y fall below or exceed the calculated threshold. The extracted features are based on the distribution of the resulting cross-validated mean misclassification errors of each classifier.

- $\Leftrightarrow \begin{tabular}{ll} \bf Mean \ misclassification \ error \ of \ appropriate \ discriminant \ analysis \ method \ using \ defined \ quantile \ \varphi^{Lvl}_{\tt mmce-[method]-[quantile]}(\mathcal{S}). \end{tabular}$
- \diamond Ratio between mean misclassification errors of two discriminant analysis methods using a given quantile $\varphi^{\mathrm{Lvl}}_{[\mathtt{method1}]_[\mathtt{method2}]_[\mathtt{quantile}]}(\mathcal{S}).$

2.6 Metamodel Features Φ_{MM}

To calculate metamodel features by Mersmann et al. (2011), linear and quadratic regression models (lin and quad) with or without interactions are fitted to a sample set S. The adjusted coefficient of determination R^2 and features reflecting the size relations of the model coefficients are extracted:

- $\diamond \ \, \textbf{Adjusted} \,\, R^2 \,\, \text{of a simple model} \,\, \varphi^{\text{MM}}_{\text{[model]_simple_adj_r2}}(\mathcal{S}).$
- \diamond Adjusted R^2 of a model with interactions $\varphi^{ ext{MM}}_{[ext{model}]_ ext{w.interact_adj_r2}}(\mathcal{S}).$
- $\diamond \ \ \textbf{Intercept} \ \text{of a simple linear model} \ \varphi^{\text{MM}}_{\texttt{lin.simple.intercept}}(\mathcal{S}).$
- $\diamond \ \ \textbf{Minimal} \ \ \text{absolute value of linear model coefficients} \ \ \varphi^{\text{MM}}_{\texttt{lin_simple_coef_min}}(\mathcal{S}).$
- $\diamond \ \, \textbf{Maximal} \ \, \text{absolute value of linear model coefficients} \ \, \varphi^{\text{MM}}_{\texttt{lin_simple_coef_max}}(\mathcal{S}).$
- \diamond Ratio of maximal and minimal abs. value of linear model coefficients $\varphi_{\text{lin.simple.coef.max.by.min}}^{\text{MM}}(\mathcal{S}).$
- \diamond Ratio of maximal and minimal abs. value of quadratic model coefficients $\varphi_{{\rm quad_simple_cond}}^{\rm MM}(\mathcal{S}).$

The feature $\varphi_{\text{lin.simple.intercept}}^{\text{MM}}$ was excluded because it is useless if fitness normalization is performed. The remaining features were not computed on sample sets with $\mathcal P$ because it does not influence the resulting feature values.

2.7 Nearest Better Clustering Features Φ_{NBC}

Nearest better clustering features by Kerschke et al. (2015) extract information based on the comparison of the sets of distances from all observations towards their nearest neighbors (D_{nn}) and their nearest better neighbors (D_{nb}).

The distance to the nearest neighbor of a search point \mathbf{x}_i , i = 1, ..., N from a set of points \mathbf{X} :

$$d_{\text{nn}}(\mathbf{x}_i, \mathbf{X}) = \min(\operatorname{dist}(\mathbf{x}_i, \mathbf{x}_j) | \mathbf{x}_j \in \mathbf{X}, i \neq j).$$
(11)

The distance to the nearest better neighbor:

$$d_{\text{nb}}(\mathbf{x}_i, \mathbf{X}) = \min(\operatorname{dist}(\mathbf{x}_i, \mathbf{x}_j) | y_j < y_i, \mathbf{x}_j \in \mathbf{X})).$$
(12)

The set of all nearest neighbor distances within the set X:

$$D_{\rm nn} = \{d_{\rm nn}(\mathbf{x}_i, \mathbf{X}) | i = 1, \dots, N\}.$$
 (13)

The set of all nearest better distances:

$$D_{\text{nb}} = \left\{ d_{\text{nb}}(\mathbf{x}_i, \mathbf{X}) | i = 1, \dots, N \right\}. \tag{14}$$

The features are defined as follows:

- \diamond Ratio of the standard deviations between the $D_{\rm nn}$ and $D_{\rm nb}$ $\varphi^{\rm NBC}_{\rm nb_std_ratio}(\mathcal{S}) = \frac{{\rm std}\,D_{\rm nn}}{{\rm std}\,D_{\rm n}}$.
- \diamond Ratio of the means between the $D_{\rm nn}$ and $D_{\rm nb}$ $\varphi^{\rm NBC}_{\rm nb_mean_ratio}(\mathcal{S}) = \frac{{\rm mean}\,D_{\rm nn}}{{\rm mean}\,D_{\rm nb}}$.
- \diamond Correlation between the distances of the nearest neighbors and nearest better neighbors $\varphi_{\text{nb_cor}}^{\text{NBC}}(\mathcal{S}) = \text{corr}(D_{\text{nn}}, D_{\text{nb}}).$
- $\diamond \ \, \textbf{Coefficient of } \ \, \textbf{variation of the distance ratios} \ \, \varphi^{\text{NBC}}_{\texttt{dist_ratio}}(\mathcal{S}) = \frac{\text{std}\,\mathcal{W}}{\text{mean}\,\mathcal{W}} \text{, where } \mathcal{W} = \left\{ \frac{d_{\texttt{nn}}(\mathbf{x},\mathbf{X})}{d_{\texttt{nb}}(\mathbf{x},\mathbf{X})} \middle| \mathbf{x} \in \mathbf{X} \right\}.$
- \diamond Correlation between the fitness value, and the count of observations to whom the current observation is the nearest better neighbor $\varphi^{\mathrm{NBC}}_{\mathrm{nb-fitness.cor}}(\mathcal{S}) = -\operatorname{corr}(\{\deg^-(\mathbf{x}_i), y_i | i=1,\dots,N\})$, where $\deg^-(\mathbf{x}_i)$ is the indegree of \mathbf{x}_i in the nearest better graph (the number of points for which a certain point is the nearest better point).

The features from Φ_{NBC} were not computed on sample sets with \mathcal{P} because it does not influence the resulting feature values.

2.8 y-Distribution Features Φ_{y-D}

The y-distribution features from Mersmann et al. (2011) compute the basic statistics of the fitness values y.

- $\diamond \ \ \textbf{Skewness} \ \ \text{and} \ \ \textbf{kurtosis} \ \ \text{of} \ \ \textbf{y} \ \ \text{values} \ \ \varphi_{\texttt{skewness}}^{\texttt{y-D}}(\textbf{y}) \ \ \text{and} \ \ \varphi_{\texttt{kurtosis}}^{\texttt{y-D}}(\textbf{y}).$
- \diamond **Number of peaks** of the kernel-based estimation of the density of **y**-distribution $\varphi_{\text{number_of_peaks}}^{\text{y-D}}(\mathbf{y})$, where the peak is defined according to Kerschke and Dagefoerde (2017).

All three features were computed only on \mathcal{A} and \mathcal{T} because neither \mathcal{P} nor the $\sigma^2 \mathbf{C}$ basis influence the resulting feature values.

3 Landscape Feature Investigation Results

First, we have investigated the impossibility of feature calculation (i.e., the feature value \bullet) for each feature. Considering that large amount of \bullet values on the tested dataset suggests low usability of the respective feature, we have excluded features resulting in \bullet in more than 25% of all measured values. Many features are difficult to calculate using low numbers of points. Therefore, for each feature we have measured the minimal number of points N_{\bullet} in a particular combination of feature and sample set, for which the calculation resulted in \bullet in at most 1% of cases. All measured values can be found in Tables 1–6.

We have tested the dependency of individual features on the dimension using feature medians from 100 samples for each distribution from the dataset. The Friedman's test rejected the hypothesis that the feature medians are independent of the dimension for all features at the family-wise significance level 0.05 using the Bonferroni-Holm correction. Moreover, for most of the features, the subsequently performed pairwise tests rejected the hypothesis of equality of feature medians for all pairs od dimensions. There were only several features for which the hypothesis was not rejected for some pairs of dimensions (see Tables 1–6).

Any analysis of the influence of multiple landscape features on the predictive error of surrogate models requires high robustness of features against random sampling of points. Therefore, we define feature *robustness* as a proportion of cases for which the difference between the 1st and 100th percentile calculated after standardization on samples from the same CMA-ES distribution is ≤ 0.05 . The robustness calculated for individual features is listed in Tables 1–6.

Table 1: TSS full features ($\mathcal{A}=\mathcal{T}$ for TSS full) from feature sets Φ_{Basic} , Φ_{CMA} , Φ_{Dis} , and $\Phi_{\text{y-D}}$. Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		\mathcal{A}			$\mathcal{A}^{ op}$			$\mathcal{A}_{\mathcal{P}}$			$\mathcal{A}_{\mathcal{P}}^{ op}$	
	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)
$arphi_{ ext{dim}}$	0	100.00		0	100.00		0	100.00		0	100.00	
$arphi_{ m obs}$	1	100.00		1	100.00		13	100.00		13	100.00	
CMA queneration CMA Step.size CMA restart CMA cma_mean.dist	0	100.00		0	100.00		0	100.00		0	100.00	
φĈMA step_size	0	100.00		0	100.00		0	100.00		0	100.00	
φCMA γrestart	0	100.00		0	100.00		0	100.00		0	100.00	
φCMA φcma mean dist	6	70.11		6	56.83		13	63.61		13	54.12	
CMA evopath_c_norm CMA evopath_s_norm	0	100.00		0	100.00		0	100.00		0	100.00	
CMA Pevopath s norm	0	100.00		0	100.00		0	100.00		0	100.00	
φ ^{CMA} _{cma_lik}	1	99.75		1	99.96		13	99.75		13	99.96	
$arphi_{ m ratio_mean_02}^{ m Dis}$	71	57.69	(2,5), (3,10), (3,20)	71	62.33	(2,5), (5,20)	86	57.36	(2,5), (3,20)	86	62.16	
φDis	71	65.95	(3,5)	71	69.28		86	65.41		86	69.60	
φDis γatio_median_02 φDis diff_mean_02 φDis diff_median_02	71	47.75		71	99.52		86	48.77		86	99.52	
Pdiff median 02	71	50.21		71	99.44		86	51.44		86	99.45	
φ ^{Dis} ratio_mean_05	27	55.22	(5, 20)	27	60.59		42	54.76	(5, 20)	42	60.32	
φ ^{Dis} ratio_median_05	27	59.80		27	64.43		42	59.66	(5, 20)	42	64.57	
$\varphi_{\text{diff_mean_05}}^{\text{Dis}}$	27	47.99		27	99.49		42	49.16		42	99.49	
obs diff.median.02 obs ratio.mean.05 obs diff.median.05 obs diff.median.05	27	51.16		27	99.46		42	52.88		42	99.48	
Dis diff.median.05 Dis ratio.mean.10 Dis ratio.median.10 Dis diff.mean.10 Dis ratio.median.25	12	54.63	(5, 20)	12	59.33		25	54.06	(5, 20)	25	58.71	
φ ^{Dis} ratio_median_10	12	58.29		12	62.38	(2,3)	25	57.80		25	61.82	
φdiff_mean_10	12	49.70		12	99.52		25	50.88		25	99.52	
φdiff_median_10	12	53.51		12	99.48		25	55.53		25	99.49	
φ _{ratio_mean_25}	6	54.80		6	58.86		15	53.87		15	58.06	
φratio_median_25	6	56.21		6	59.74		15	54.91		15	58.26	
φdiff_mean_25	6	53.78		6	99.55		15	55.29		15	99.55	
φ Dis Dis Platio_median_25 Dis Patio_median_25 G diff_mean_25 G diff_median_25	6	57.40		6	99.49		15	60.24		15	99.50	
$\varphi_{ ext{skewness}}^{ ext{y-D}}$	6	36.71		_	_	_	_	_	_	_	_	_
$\varphi_{\text{kurtosis}}^{\text{y-D}}$	6	63.06		_	_	_	_	_	_	_	_	_
$\varphi_{\text{skewness}}^{\text{y-D}}$ $\varphi_{\text{surtosis}}^{\text{y-D}}$ $\varphi_{\text{number_of_peaks}}^{\text{number_of_peaks}}$	6	32.34		_	-	_	_	-	_	_	_	_

Table 2: TSS full features ($\mathcal{A}=\mathcal{T}$ for TSS full) from feature sets Φ_{Lvl} , Φ_{MM} , Φ_{Inf} , and Φ_{NBC} . Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		\mathcal{A}			$\mathcal{A}^{ op}$			$\mathcal{A}_{\mathcal{P}}$			$\mathcal{A}_{\mathcal{P}}^{\top}$	
	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)	N_{\bullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)
φLvl mmce_lda_10	6	39.47		6	39.47		13	53.01		13	53.01	
φLvI	6	69.30		6	69.88		13	65.30		13	65.91	
ν Lvl φ Lvl	6	28.07		6	28.14		13	27.38		13	27.64	
Pinter da 10 Pul Pul 1da qda 10	6	80.09		6	80.12		18	92.37		18	92.38	
	6	44.32		6	42.71		13	46.42		13	48.00	
θlda_mda_10 LvI φgda_mda_10	6	81.14		6	80.51		13	78.27		13	77.33	
(2LVI	6	32.64		6	32.64		13	54.19		13	54.19	
φ ^{LVI}	6	62.01		6	62.41		13	56.49		13	56.69	
ν Lvl ν Lvl	6	22.86		6	22.63		13	27.03		13	26.99	
mmce.qda.25 Lvl mmce.mda.25 Lvl flda.qda.25 Lvl lda.qda.25	6	76.25		6	76.29	(3,5)	15	94.13		15	94.12	
Lvl	6	32.07	(3, 20)	6	27.68	(3, 10),	13	22.86		13	20.39	
			(, ,			(10, 20)						
$\varphi^{\mathrm{Lvl}}_{\mathrm{qda_mda_25}}$	6	77.24		6	77.68		13	63.08		13	63.11	
Ľvl	6	22.40		6	22.39		13	18.46		13	18.46	
Ψ ^{LvI} Ψmmce ada 50	6	48.24		6	49.46		13	32.88		13	33.15	
φ ^{LvI} φ _{mmce mda} 50	6	20.13		6	19.60		13	38.52		13	37.58	
(OLVI	6	69.22	(10, 20)	6	69.24	(10, 20)	15	84.45		15	84.41	
LX/	6	36.05		6	30.90		13	6.80	(5,20)	13	4.51	
φ1da_mda_50 LvI φqda_mda_50	6	42.65	(3, 20)	6	40.68		13	21.96	(5, 20)	13	22.47	(3,5), (5,20)
$\varphi_{\text{lin_simple_adj_r2}}^{\text{MM}}$	6	19.64		6	19.62		_	_	_	_	_	_
ω ^{MM}	6	51.95		6	85.41		_	_	_	_	_	_
(OMM	6	29.73		6	77.68		_	_	_	_	_	_
Ψlin_simple_coef_max ΨMM lin_simple_coef_max by	6	27.78		6	69.78		_	_	_	_	_	_
Ψlin_simple_coef_max_by_ ωMM	100	44.77		100	43.89		_	_	_	_	_	_
φ MM lin.simple.coef.max.by- φ MM γ lin.w.interact.adj.r2 φ MM guad.simple.adj.r2	6	50.23		6	52.16		_	_	_	_	_	_
MM (OMM	6	46.45		6	95.21		_	_	_	_	_	_
$\begin{array}{l} \varphi_{\mathrm{quad_simple_adj_r2}}^{\mathrm{NIM}} \\ \varphi_{\mathrm{quad_simple_cond}}^{\mathrm{MM}} \\ \varphi_{\mathrm{quad_w_interact_adj_r2}}^{\mathrm{MM}} \end{array}$	629	82.50		531	77.81		_	_	_	_	_	_
			(2.5)		22.02		10	25.74		10	25.42	
φInf h_max yInf φeps_s	6	14.46	(3,5)	6	33.83		13	35.74		13	35.42	
φ _{eps_s} Inf	6	6.59		6	34.65		3056	26.05		3196	54.56	
Vinf φeps_max .Inf	6	64.93		6	84.99		13	65.17		13	84.18	
φ ^{Inf} _{m0}	6 6	1.30 27.67		6	3.88 44.91		_	_	_	_	_	_
$\varphi_{ ext{eps_ratio}}^{ ext{Inf}}$	ь	27.67		6	44.91					_		
φNBC nb.std.ratio γNBC γNBC	6	31.11		6	18.27		_	_	_	_	_	_
	6	38.98		6	32.96		_	-	_	_	_	_
φNBC nb-cor φNBC dist_ratio	6	45.94		6	32.39		_	_	_	_	_	_
NBC	6	55.23		6	50.78		_	-	_	_	_	_
φ ^{NBC} φ ^{NBC} nb_fitness_cor	6	77.17		6	77.13		_	_	_	_	_	_

Table 3: TSS nearest features from feature sets $\Phi_{\rm Basic}$, $\Phi_{\rm CMA}$, $\Phi_{\rm Dis}$, and $\Phi_{\rm y-D}$ (only \mathcal{T} -based and sample set indepent, \mathcal{A} -based are identical to TSS full in Table 1). Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		$\mathcal T$			$\mathcal{T}^{ op}$			$\mathcal{T}_{\mathcal{P}}$			$\mathcal{T}_{\mathcal{P}}^{ op}$	
	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)
$arphi_{ ext{dim}}$	0	100.00		0	100.00		0	100.00		0	100.00	
$arphi_{ ext{obs}}$	1	100.00		1	100.00		13	100.00		13	100.00	
$\varphi_{\text{generation}}^{\text{CMA}}$	0	100.00		0	100.00		0	100.00		0	100.00	
φCMA Step size	0	100.00		0	100.00		0	100.00		0	100.00	
φCMA restart	0	100.00		0	100.00		0	100.00		0	100.00	
$\varphi_{\text{CMA}}^{\text{CMA}}$ $\varphi_{\text{restart}}^{\text{CMA}}$ $\varphi_{\text{CMA}}^{\text{CMA}}$ $\varphi_{\text{CMA}}^{\text{CMA}}$	6	70.11		6	56.83		13	63.61		13	54.12	
φCMA evopath c norm	0	100.00		0	100.00		0	100.00		0	100.00	
φCMA evopath s norm	0	100.00		0	100.00		0	100.00		0	100.00	
CMA evopath_c_norm CMA evopath_c_norm CMA evopath_s_norm CMA cma_lik	1	99.75		1	99.96		13	99.75		13	99.96	
$arphi_{ m ratio_mean_02}^{ m Dis}$	71	57.69	(2,5), (3,10), (3,20)	71	62.33	(2,5), (5,20)	86	57.36	(2,5), (3,20)	86	62.16	
$arphi^{ ext{Dis}}_{ ext{ratio_median_02}} \ arphi^{ ext{Dis}}_{ ext{diff_mean_02}}$	71	65.95	(3,5)	71	69.28		86	65.41		86	69.60	
Pratio_median_02 Dis diff_mean_02 Pois pois gatio_mean_02	71	47.75	` '	71	99.52		86	48.77		86	99.52	
φDis φdiff median 02	71	50.21		71	99.44		86	51.44		86	99.45	
Pratio mean 05	27	55.22	(5, 20)	27	60.59		42	54.76	(5, 20)	42	60.32	
Post ratio_mean_05 Post ratio_median_05 Post ratio_median_05 Post ratio_median_05 Post ratio_median_05 Post ratio_median_05	27	59.80		27	64.43		42	59.66	(5, 20)	42	64.57	
$\varphi_{ t diff_mean_05}^{ t Dis}$	27	47.99		27	99.49		42	49.16		42	99.49	
φDis diff_median_05	27	51.16		27	99.46		42	52.88		42	99.48	
(ODIS	12	54.63	(5, 20)	12	59.33		25	54.06	(5, 20)	25	58.71	
Pratio_mean_10 Dis ratio_median_10 Dis diff_mean_10	12	58.29		12	62.38	(2,3)	25	57.80		25	61.82	
Ψratio_median_10 ψ Dis ψ diff_mean_10 Ψ Dis ψ diff_median_10 ψ Dis ψ Dis ξ atio_mean_25	12	49.70		12	99.52		25	50.88		25	99.52	
φdiff_median_10	12	53.51		12	99.48		25	55.53		25	99.49	
Pdiff_median_10 PDis γ ratio_mean_25 γ ratio_median_25	6	54.80		6	58.86		15	53.87		15	58.06	
Pratio_mean_25 Plis Pratio_median_25 Plis Pdiff_mean_25	6	56.21		6	59.74		15	54.91		15	58.26	
φdiff_mean_25	6	53.78		6	99.55		15	55.29		15	99.55	
φ _{diff_mean_25} φ _{Dis} φ _{diff_median_25}	6	57.40		6	99.49		15	60.24		15	99.50	
y-D \(\text{\$\psi_{skewness}} \) \(\text{\$\psi_{v}\$-D} \) \(\text{\$\psi_{kurtosis}} \) \(\text{\$\psi_{v}\$-D} \) \(\text{\$\psi_{number_of_peaks}} \)	6	36.71		_	-	_	_	-	-	_	_	-
φ _{kurtosis}	6	63.06		_	_	_	_	_	_	_	_	_
φy-D γnumber of peaks	6	32.34		-	_	_	_	_	-	_	_	_

Table 4: TSS nearest features from feature sets $\Phi_{\rm Lvl}$, $\Phi_{\rm MM}$, $\Phi_{\rm Inf}$, and $\Phi_{\rm NBC}$ (only \mathcal{T} -based and sample set indepent, \mathcal{A} -based are identical to TSS full in Table 2). Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		$\mathcal T$			$\mathcal{T}^{ op}$			$\mathcal{T}_{\mathcal{P}}$			$\mathcal{T}_{\mathcal{P}}^{ op}$	
	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)) N _•	rob.(%)	(D_i, D_j)	N_{ullet}		(D_i, D_j)
φ ^{Lvl} mmce_lda_10	6	39.47		6	39.47		13	53.01		13	53.01	
ωLvI	6	69.30		6	69.88		13	65.30		13	65.91	
Lvl	6	28.07		6	28.14		13	27.38		13	27.64	
Finded dall Value of the second of the secon	6	80.09		6	80.12		18	92.37		18	92.38	
Plda_qda_10 PLvl Plda_mda_10	6	44.32		6	42.71		13	46.42		13	48.00	
Lvl	6	81.14		6	80.51		13	78.27		13	77.33	
_Lvi	6	32.64		6	32.64		13	54.19		13	54.19	
Pmmce_1da_25 PLvl Pmmce_qda_25	6	62.01		6	62.41		13	56.49		13	56.69	
r mmce.qda.25 φLvl pmce.mda.25	6	22.86		6	22.63		13	27.03		13	26.99	
Lvl Lvl	6	76.25		6	76.29	(3,5)	15	94.13		15	94.12	
$arphi_{ ext{Lvl}}^{ ext{Lul}}$ $arphi_{ ext{1da-qda-25}}^{ ext{Lvl}}$ $arphi_{ ext{1da-mda-25}}^{ ext{Lul}}$	6	32.07	(3, 20)	6	27.68	(3, 10),	13	22.86		13	20.39	
	O		(0,20)	Ü		(10, 20)						
$arphi_{ ext{gda_mda_25}}^{ ext{Lvl}}$	6	77.24		6	77.68		13	63.08		13	63.11	
L XV	6	22.40		6	22.39		13	18.46		13	18.46	
Pmmce_lda_50 Lvl Pmmce_qda_50	6	48.24		6	49.46		13	32.88		13	33.15	
Ψmmce mda 50	6	20.13		6	19.60		13	38.52		13	37.58	
(OLVI	6	69.22	(10, 20)	6	69.24	(10, 20)	15	84.45		15	84.41	
L XV	6	36.05		6	30.90		13	6.80	(5, 20)	13	4.51	
φ _{1da_mda_50} LvI φ _{qda_mda_50}	6	42.65	(3, 20)	6	40.68		13	21.96	(5, 20)	13	22.47	(3,5), (5,20)
$\varphi_{\texttt{lin_simple_adj_r2}}^{\texttt{MM}}$	6	19.64		6	19.62		_	_	_	_	_	_
(OMM	6	51.95		6	85.41		_	_	_	_	_	_
γ lin_simple_coef_min γ MM lin_simple_coef_max	6	29.73		6	77.68		_	_	_	_	_	_
	min 6	27.78		6	69.78		_	_	_	_	_	_
MM	100	44.77		100	43.89		_	_	_	_	_	_
$\varphi_{\text{lin_simple_coef_max_by.}}^{\text{MM}}$ $\varphi_{\text{lin_w.interact_adj_r2}}^{\text{mu}}$ $\varphi_{\text{guad.simple_adj_r2}}^{\text{MM}}$	6	50.23		6	52.16		_	_	_	_	_	_
Ψquad_simple_adj_r2 (MM)		46.45		6	95.21		_	_	_	_	_	_
$\varphi_{\text{quad_simple_cond}}^{\text{MM}}$ $\varphi_{\text{quad_w_interact_adj_r2}}^{\text{MM}}$	629	82.50		531	77.81		_	_	_	_	_	_
Inf	6	14.46	(3,5)	6	33.83		13	35.74		13	35.42	
θh.max Inf φeps_s	6	6.59	(3,3)	6	34.65		3056	26.05		3196	54.56	
Ψeps_s ,Inf	6	64.93		6	34.63 84.99		13	65.17		13	84.18	
φInf eps_max Inf	6	1.30		6	3.88		13	05.17			04.10	
$arphi_{ ext{m0}}^{ ext{Inf}}$ $arphi_{ ext{Inf}}^{ ext{eps_ratio}}$	6	27.67		6	3.00 44.91		_	_	_	_	_	_
φNBC nb.std.ratio γNBC nb.mean_ratio	6	31.11		6	18.27		_	_	_	_	_	_
Phb.std_ratio Phb.std_ratio Phb.std_ratio Phb.scor	6	38.98		6	32.96		_	_	_	-	_	_
$\varphi_{ ext{nb_cor}}^{ ext{NBC}}$ $\varphi_{ ext{dist_ratio}}^{ ext{NBC}}$	6	45.94		6	32.39		_	-	-	-	_	-
φ ^{INDC} dist_ratio	6	55.23		6	50.78		_	-	-	-	_	_
$\varphi_{ ext{nist_ratio}}^{ ext{dist_ratio}}$ $\varphi_{ ext{nb_fitness_cor}}^{ ext{NBC}}$	6	77.17		6	77.13		_	_	_	-	_	_

Table 5: TSS knn features from feature sets Φ_{Basic} , Φ_{CMA} , Φ_{Dis} , and $\Phi_{\text{y-D}}$ (only \mathcal{T} -based and sample set indepent, \mathcal{A} -based are identical to TSS full in Table 1). Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		\mathcal{T}			$\mathcal{T}^{ op}$			$\mathcal{T}_{\mathcal{P}}$			$\mathcal{T}_{\mathcal{P}}^{ op}$	
	N_{ullet}	rob.(%)	(D_i, D_j)) N _•	rob.(%)	(D_i, D_j)) N _•	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)
$arphi_{ exttt{dim}}$	0	100.00		0	100.00		0	100.00		0	100.00	
$arphi_{ ext{obs}}$	1	92.16		1	92.16		13	92.26		13	92.26	
$\varphi_{ ext{qeneration}}^{ ext{CMA}}$	0	100.00		0	100.00		0	100.00		0	100.00	
CMA CMA restart CMA CMA CMA CMA	0	100.00		0	100.00		0	100.00		0	100.00	
Prestart	0	100.00		0	100.00		0	100.00		0	100.00	
CMA Cma maan di at	6	78.64		6	98.40		13	93.96		13	98.71	
CMA Peyopath c norm	0	100.00		0	100.00		0	100.00		0	100.00	
CMA	0	100.00		0	100.00		0	100.00		0	100.00	
φ cMA φ cMA φ cMA φ cvopath_c_norm φ cMA φ copath_s_norm φ cMA φ cma_lik	1	98.81		1	99.95		13	98.80		13	99.96	
Dis	74	30.53		74	23.95		109	29.18		109	23.33	
Pratio_mean_02 Pis Pratio_median_02 Pis Pdiff_mean_02	74	34.49	(5, 10)	74	25.97		109	32.08		109	24.94	
Dis	74	50.17	` , ,	74	96.75		109	50.99		109	96.76	
_Θ Dis	74	49.55		74	95.54		109	50.37		109	95.47	
Dis	30	28.43		30	21.63		53	26.56		53	21.19	
₍₂ Dis	30	35.85		30	24.73	(3, 10)	53	31.45	(2,20)	53	23.58	
Yratio_median_05 YDis diff_mean_05 YDis diff_median_05	30	55.94		30	96.80		53	56.94	` '	53	96.85	
Dis	30	55.11		30	95.80		53	56.23		53	95.82	
$arphi^{ extsf{Dis}}_{ ext{ratio_mean_10}}$	15	25.57	(2, 3), (2, 10), (3, 10), (5, 10)	15	19.13		28	24.33	(5, 10)	28	18.97	
$\varphi_{\text{ratio_median_10}}^{\text{Dis}}$	15	32.86		15	21.96		28	29.68		28	21.49	
ρDis Odiff mean 10	15	58.33		15	96.96		28	59.42		28	97.02	
φDis diff median 10	15	57.49		15	96.09		28	58.85		28	96.16	
Pratio mean 25	6	19.39		6	14.96		15	19.01		15	14.16	
Plis ratio_median_10 Plis diff_mean_10 Plis diff_median_10 Plis ratio_mean_25 Plis ratio_mean_25 Plis ratio_median_25	6	23.38		6	15.76	(3,5)	15	24.08		15	16.25	
	6	62.31		6	97.06		15	62.72		15	97.12	
$\varphi_{ ext{diff_mean_25}}^{ ext{diff_mean_25}} otag diff_median_25$	6	61.67		6	96.27		15	62.60		15	96.36	
$\varphi_{ ext{skewness}}^{ ext{y-D}}$	6	30.49		_	_	_	_	_	_	_	_	
φy-D	6	72.61		_	_	_	_	_	_	_	_	_
Y-D Ykurtosis y-D number_of_peaks	6	26.63		-	-	_	-	-	_	_	_	-

Table 6: TSS knn features from feature sets Φ_{Lvl} , Φ_{MM} , Φ_{Inf} , and Φ_{NBC} (only \mathcal{T} -based, \mathcal{A} -based are identical to TSS full in Table 2). Features are grouped according to their feature sets (separated by horizontal lines). Features with less than 25% of values equal to \bullet and robustness greater than 0.9, are in gray. N_{\bullet} denotes the lowest measured number of points from which at most 1% of feature calculations resulted in \bullet ($N_{\bullet}=0$ for sample set independent φ). The (D_i,D_j) column shows the pairs of feature dimensions for which the two-sided Wilcoxon signed rank test with the Bonferroni-Holm correction does not reject the hypothesis of equality of median feature values, at the family-wise level 0.05 for each individual feature.

		$\mathcal T$			$\mathcal{T}^{ op}$			$\mathcal{T}_{\mathcal{P}}$			$\mathcal{T}^ op_\mathcal{P}$	
	N_{ullet}	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)) N _•	rob.(%)	(D_i, D_j)	N_{ullet}	rob.(%)	(D_i, D_j)
$arphi_{ t mmce_lda_10}^{ t Lvl}$	6	11.39		6	11.39		13	9.63		13	9.93	
ω ^{LVI}	6	67.83		6	69.40		13	69.58		13	70.79	(3,5)
Lvl *	6	21.15		6	20.32		13	19.93		13	19.25	
φ^{LvI}	23	67.38		23	67.38		37	70.31		37	70.32	
Pida_qda_10 PLvl plda_mda_10 PLvl φda_mda_10	6	15.74		6	13.69		13	22.17		13	20.80	
(aLvl	6	57.54		6	56.51		13	55.10		13	55.06	
Lvl	6	9.73		6	9.63		13	15.42	(3,20)	13	15.65	(3,20)
ω ^{Lvl}	6	37.27		6	38.00		13	36.70		13	37.25	
I.VI	6	17.79		6	16.60		13	15.60		13	14.90	
Ψmmce_mda_25 Lvl Ψlda_qda_25	6	73.93	(5, 10)	6	73.87	(5, 10)	18	89.97		18	89.95	
ν Lvl ν Lda mda 25	6	8.97		6	6.33	, , ,	13	4.46		13	2.95	
PLvI 1da_qda_25 LvI Plda_mda_25 PLvI qda_mda_25	6	47.47		6	49.72		13	35.90		13	39.50	(2, 10), (3, 5)
$\varphi^{\mathrm{Lvl}}_{\mathtt{mmce_lda_50}}$	6	17.46		6	16.62		13	7.86		13	7.88	
ωLvI ωmmaa ada 50	6	34.21		6	34.32		13	23.54		13	27.64	
Ψ Lvl 1da qda 50	6	24.08		6	20.56		13	20.74		13	19.07	(2,3)
φLvl φlda gda 50	6	41.38		6	41.37		18	73.19		18	73.01	,
	6	26.79		6	32.00		13	2.70	(5, 10)	13	2.30	
Ψlda_mda_50 Ψqda_mda_50	6	17.30		6	24.22		13	5.95		13	7.69	(5, 20)
,oMM	6	15.67		6	15.63		-	_	_	-	_	_
Ψlin simple coef min	6	97.40		6	63.93		_	_	_	-	_	_
φMM	6	98.12	(2,3)	6	87.04		_	-	_	_	-	-
φMM Vlin_simple_coef_max_by	6	34.41		6	45.65		_	_	_	_	_	_
	100	37.50		100	30.25		_	_	_	_	_	_
Ψlin_w_interact_adj_r2 ΨMM quad_simple_adj_r2	6	41.50		6	37.91		_	_	_	_	_	_
φMM quad_simple_cond	6	92.55		6	87.21		_	_	_	_	_	_
$arphi_{ ext{quad.w.interact.adj.r2}}^{ ext{MM}}$	688	69.48		632	60.02		_	_	_	_	_	_
$arphi_{ ext{h_max}}^{ ext{Inf}}$ $arphi_{ ext{lnf}}^{ ext{lnf}}$ $arphi_{ ext{eps-s}}^{ ext{cps-s}}$	6	1.94		6	1.57	(3,5)	13	31.85		13	38.67	
φInf eps-s	6	38.77		6	3.58		2424	66.99		2485	70.18	(5, 10)
φ ^{inr}	6	97.84		6	59.04		13	97.70		13	58.95	
φ_{m0}^{Inf}	6	0.84		6	0.44		_	-	_	_	_	_
φInf m0 γInf φInf eps_ratio	6	17.19		6	5.59		_	_	_	-	_	_
φNBC	6	20.10		6	14.20		_	_	_	_	_	_
NBC γ nb mean ratio γ NBC γ nb mean ratio	6	36.38		6	30.42		_	_	_	_	_	-
φNBC nb.cor γNBC dist_ratio	6	28.49		6	26.50		_	_	_	_	_	-
. NBC	6	41.24		6	26.53		-	_	_	-	_	_
φ dist_ratio NBC γnb_fitness_cor	6	45.12		6	45.14		_	_	-	-	-	_

4 Connection of Landscape Features and Model Error Results

We have tested the statistical significance of the MSE and RDE differences for 19 surrogate model settings using TSS full and TSS nearest methods and also the lmm surrogate model utilizing TSS knn, i. e., 39 different combinations of model settings ψ and TSS methods (ψ , TSS), on all available sample sets using the non-parametric two-sided Wilcoxon signed rank test with the Holm correction for the family-wise error. To better illustrate the differences between individual settings, we also count the percentage of cases at which one setting had the error lower than the other. The pairwise score and the statistical significance of the pairwise differences are summarized in Tables 7 and 8.

To compare the convenience of individual features as descriptors of areas where the surrogate model $\mathcal M$ with a particular (ψ, TSS) combination has the best performance, we use the Kolmogorov-Smirnov test (KS test) testing the equality of the distribution of values of individual features calculated on the whole testing dataset and on only those sample sets from the testing dataset for which the (ψ, TSS) combination leads to the lowest error (MSE or RDE) among all tested combinations. The hypothesis of distribution equality is tested at the family-wise significance level $\alpha=0.05$ after the Holm correction. The resulting p-values are summarized in Tables 9–13.

A pairwise comparison of the model settings MSE in different TSS. The percentage of wins of i-th model setting against j-th model setting over all available data is given in the i-th row and j-th column. The numbers in bold mark the row model setting being significantly better than the column model setting according to the two-sided Wilcoxon signed rank test with the Holm correction at = 0.05. family-wise significance level α Table 7:

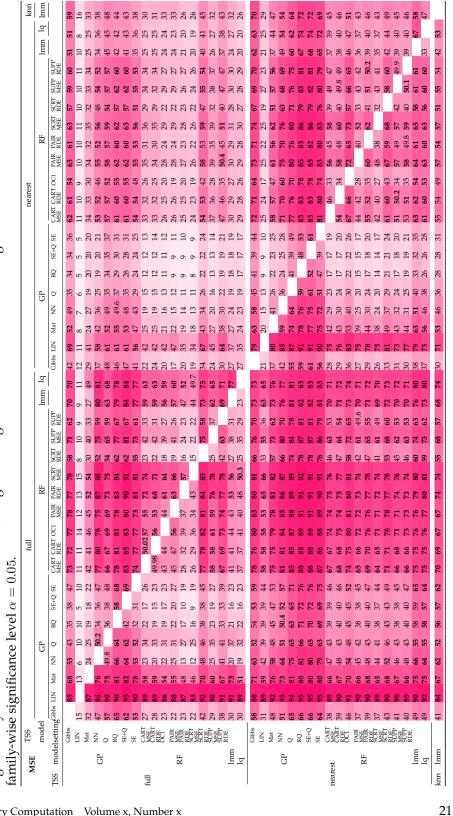


Table 8: A pairwise comparison of the model settings RDE in different TSS. The percentage of wins of i-th model setting against j-th model setting over all available data is given in the i-th row and j-th column. The numbers in bold mark the row model setting being significantly better than the column model setting according to the two-sided Wilcoxon signed rank test with the Holm correction at family-wise significance level $\alpha = 0.05$.

knn	nearest	full	TSS
lmm	GP est RF	GP RF	RDE
	Gibs	SE S	DE TSS model modelsetting Gibbs LIN
49.9	58 58 61 45 58 58 58 58 58 58 58 58 58 58 58 58 58	24 37 37 37 38 33 33 33 38 38	ingGibt
9 83	83 59 81 81 81 82 89 89 75 77 77 77 78 83	76 64 64 64 64 64 64 64 64 64 64 64 64 64	S LIN
67	68 38 60 60 63 77 63 77 77 67 67 67 68	63 64 66 66 66 66 66 67 68 68 68 68 68 68 68 68 68 68 68 68 68	Mat
66	65 58 67 67 67 67 67 67 67 67 67 67 67 67 67	55 29 29 40 40 40 40 40 40 40 40 40 40 40 40 40	N
56	56 28 42 49 65 66 66 66 64 42 42 42 43 43 43 43 43 44 44 44 44 44 45 46 46 46 46 46 46 46 46 46 46 46 46 46	26 26 33 33 33 35 36 44 44 44 44 44 44 44 44 44 44 44 44 44	ρ Θ
59	56 26 28 38 38 47 70 69 69 69 39 39 39 39 39 39 39 39 39 43 38	35 35 36 36 37 37 37 37 37 37 37 37 37 37 37 37 37	RQ
60	57 228 39 39 57 70 60 60	226 226 32 32 32 32 33 34 46 46 46 46 47 37	SE+Q
63	60 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	8 4 5 2 5 6 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8	SE
70	70 54 54 55 56 56 56 56 56 57 57 57 57 57 57 57 57 57 57 57 57 57	6	CART MSE
68	770 37 57 77 77 77 56 66 56 56 57 77 57 57 57 57 57 57 57 57 57 57 57	60 60 60 60 60 60 60 60 60 60 60 60 60 6	full CART OCI
70	73 40 60 69 80 80 80 80 63 63 63 63 63 63 63 63 63 63 63 63 69 69 69 69 69 69 69 69 69 69 69 69 69	67 552 665 665 67 57 57 57 57 57 57	
70	773 337 466 660 660 661 811 811 811 811 661 661 661 661 662 663 663 663 663 664 664 665 665 665 665 665 665 665 665	66 66 66 66 66 66 66 66 66 66 66 66 66	PAIR I
74	774 444 444 447 448 484 884 884 884 884	53 53 54 67 67 67 67 67	RF PAIR S
75 6	774 6 54 54 54 54 54 54 54 55 55 55 55 55 55 5	668 668 668 668 668 668	SCRT S
51 6	65 66 66 772 773 771 771 771 771 771 771 771 771 771	58 58 60 60 60 60 60 60 60 60 60 60	SCRT S RDE N
68 6	669 669 665 57 4469.7 44	662 5 661 5 662 5 662 5 662 5 662 5 663 5 644 5 644 5 644 5 64 6 64 6 64 6 64 6	SUPP S MSE R
60 7	557 666 772 7771 7771 7771 7771 7771 7774 756 66 66 66 66 66 66 66 66 66 66 66 66 6	558 6655 6655 6655 6655 6655 6655 6655	SUPP RDE
71 72	66 68 68 68 68 68 69 60 60 60 60 60 60 60 60 60 60 60 60 60	2 62 62 63 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6	mm
2 40	8 20 20 20 20 20 20 20 20 20 20 20 20 20	5 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	[] [4]
) 79	86 83 85 77 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7	7 411 7 7 411 7 7 7 411 7 7 7 411 7 7 7 616 6 8 6 8 6 8 6 8 6 8 6 8 7 7 7 6 6 6 6	Gibbs LIN
63	58 567 764 444 44 44 45 45 45 45 45 45 45 45 45 4	255 440 450 37 57 444 454 454 454 454 454 454 454 454	N Mat
56	62 21 62 74 42 73 74 44 40 40 40 41 41 41 41 41 41 41 41 41 41 41 41 41	55 19 52 53 54 54 54 54 36 37 37 37 37 37 37	Z Z
47	28 28 38 36 60 60 60 32 32 32 32 32 33 33 33 33 33 33 33 33	16 16 27 33 33 34 44 44 43 29 29 29 20 20 20 21 21 21 21 21 21 21 21 21 21 21 21 21	0 Q
37	12 224 40 25 25 25 25 25 26 26 26 27 27 27 27 27 27 27 27 27 27 27 27 27	39 224 23 30 30 30 30 30 30 20 20 20 20 20 20 20 20 20 20 20 20 20	RQ
36	12 24 24 27 26 38 38 25 25 25 25 25 26 26 26 26 26 26 26 26 26 26 26 26 26	39 24 24 23 31 31 31 31 29 29 20 20 21 21 21 21 21 21 21 21 21 21 21 21 21	SE+Q
38	49 29 29 40 40 40 40 27 27 27 27 27 27 27 28 29 29 29 29 29 29 29 29 29 29 29 29 29	41 11 12 25 25 25 25 25 25 27 27 27 27 27 27 27 27 27 27 27 27 27	Q SE
62	66 65 65 65 65 65 65 65 65 65 65 65 65 6	8 23 23 23 24 24 25 25 25 25 25 25 25 25 25 25 25 25 25 	CART
62	68 68 68 68 68 68 68 68 68 68 68 68 68 6	8 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	near T CAR RDE
58	55 55 55 55 55 55 55 55 55 55 55 55 55	3 3 4 4 4 5 3 3 3 3 8 8 8 9 9 1 4 5 5 5 8 8	est T OC1
65	71 32 32 78 78 78 78 78 78 78 78 78 78 78 78 78	63 28 47 61 64 63 60 48 49 49 49 49 49 49 49 49 49 49 49 49 49	PAIR MSE
62	67 51 51 66 68 68 68 68 49 49 49 49 49 49 66 60 60 60 60 60 60 60 60 60 60 60 60	60 60 60 61 61 60 60 60 60 60 60 60 60 60 60 60 60 60	RF PAIR RDE
65	70 32 32 72 72 72 78 78 78 78 78 77 77 77 77 77 77 77 77	63 228 447 661 661 661 661 661 661 661 661 661 66	SCRI
61	67 27 50.2 59 67 74 72 48 49 55 42 49 41	59 42 42 57 60 59 56 42 42 43 33 33 34 41 41	SCRT
60		57 22 22 21 41 41 51 51 55 58 58 58 58 58 58 58 58 58 58 58 58	
62	53 53 61 61 69 75 75 75 75 74 47 53 53 53 53	61 45 54 54 62 61 61 61 61 61 61 61 61 61 61 61 61 61	
49.9	59 20 20 34 43 43 53 53 66 66 61 61 61 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 41 31 31 31 31 31 31 31 31 31 31 31 31 31	50.03 117 31 34 42 40 40 30 30 30 30 30 30 30 30 30 30 30 30 30	lmm
55	667 667 667 667 667 667 667 667 667 667	51 118 118 33 33 33 33 34 44 45 44 45 32 32 32 33 33 33 33 33 33 33 33 33 33	р М
	60 21 21 22 23 24 24 24 25 26 26 26 26 27 27 27 27 28 28 28 28 28 28 28 28 28 28 28 28 28	50.1 50.1 50.1 50.1 50.1 50.1 50.1 50.1	lmm

Table 9: The p-values of the Kolmogorov-Smirnov (KS) test comparing the equality of probability distributions of individual TSS full feature representatives on all data and on those data on which a particular model setting scored best in MSE. The p-values are after the Holm correction and they are shown only if the KS test rejects the equality of both distributions at the family-wise significance level $\alpha=0.05$, non-rejecting the equality hypothesis is indicated with —. Zeros indicate p-values below the smallest double precision number.

$\overline{\mathcal{M}}$				G	P								RF]	lmm	lq
settings	Gibbs	LIN	Mat	NN	Q	RQ	SE+Q		CART MSE		OC1					SUPP MSE		_	
$arphi_{ ext{dim}}$	2e- 101	_	3.2e-2	1.3e-3	2.9e-5	1.7e- 23	2.8e- 13	_	1.3e- 59	1.9e- 16	_	_	-2.4e-2	3.9e-2	1.5e-3	1.0e- 16	7.4e-3	4e- 115	7.8e- 46
$arphi_{ exttt{obs}}(\mathcal{A})$	1.8e- 40	_	1.0e-7	1.7e-7	7.0e- 29	5.9e- 36	1.6e- 21	_	2.6e- 43	1.7e-7	8.4e- 11	_	2.3e- 10	1.1e-6	6.0e- 15	1.4e- 10	_	3.2e- 14	3e- 259
$arphi_{ ext{evopath_c_norm}}^{ ext{CMA}}$	3.5e-7	3.5e-8	7.9e- 37	_	2.6e- 69	7.6e- 62	9.1e- 30	2.1e- 10	9.9e- 20	3.7e- 18	4.5e-4	_	1.2e-8	9.3e-8	7.0e- 18	2.2e- 10	4.3e- 24	1e- 216	5e- 307
$arphi_{ ext{evopath_s_norm}}^{ ext{CMA}}$	1.3e- 88	2.1e-3	9.7e- 10	1.5e- 25	3.3e- 48	6.0e- 33	9.2e- 17	7.4e- 15	5.1e- 18	7.0e- 13	6.9e- 14	2.7e-6	2.6e-5	3.9e-5	1.3e- 10	2.1e- 14	6.3e- 11	1.3e- 16	1.1e- 13
$arphi_{ ext{restart}}^{ ext{CMA}}$	7e- 105	1.5e-7	9.7e- 20	2.3e- 61	7e- 273	5e- 193	3e- 137	3.5e- 46	3.3e- 21	8.2e- 18	1.7e- 60	_	2.0e-4	4.1e-9	8.7e- 98	5.5e- 14	5.1e- 37	1.9e- 12	2.2e- 15
$arphi_{ exttt{step_size}}^{ exttt{CMA}}$	0	5.2e- 19	1.5e- 62	2e- 123				5.4e-	5.0e-	1.5e-	1.1e-	2.0e-	3.4e-	2.2e-	2.5e-	5.6e-			9e- 194
$arphi_{ exttt{cma_lik}}^{ exttt{CMA}}(\mathcal{A}_{\mathcal{P}}^{ op})$	1.5e- 37	1.8e-2	8.2e-7	2.7e- 14	2.9e- 30	2.8e- 33	3.8e-8	6.6e- 14	1.7e-9	7.8e-9	4.3e-8	2.0e-3	2.2e-6	1.6e-5	9.3e- 14	9.7e-8	4.5e- 12	6.0e- 34	1.5e- 22
$arphi_{ ext{diff_median_02}}^{ ext{Dis}}(\mathcal{A}^{ op})$) 0	2.7e- 30	3e- 103	3e- 159	0	0	4e- 160	2e- 112	3.9e- 94	7e- 110	1.1e- 91	6.1e-7	, 8.3e- 50	2.3e- 38	3.8e-	5.9e- 62	6e- 107	3e- 103	6e- 250
$arphi_{ ext{diff_median_10}}^{ ext{Dis}}(\mathcal{A}^{ ext{T}}$) 0	1.9e- 35		9e- 173	0	0	3e- 166	2e- 130	2.8e- 89		5.0e- 93		1.8e- 62				4e- 125	2.5e- 82	0
$\varphi_{\text{diff_mean_05}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	0	1.6e- 35	5e- 132	7e- 188	0	0	1e- 185	7e- 138	6e- 115		1e- 104	3.1e-9	3.0e- 68	5.8e- 55	7e-	6.5e- 81	5e- 129	2.0e- 81	0
$arphi_{ ext{diff_mean_25}}^{ ext{Dis}}(\mathcal{A}_{\mathcal{P}}^{ op})$	0	6.2e- 40	3e- 157	6e- 186	0	0	2e- 181		3e- 102				4.9e-				1e- 128	1e- 103	0
$arphi_{ ext{lda_qda_l0}}^{ ext{Lvl}}(\mathcal{A}_{\mathcal{P}})$	9.7e- 29	1.7e-5	8.2e- 18	3.4e- 19	1.6e- 22	4.4e-5	7.8e- 25	1.5e-4	5.4e-4	5.4e- 12	1.7e- 10	_	1.0e- 11	1.1e-7	, 1.7e- 35	2.5e-3	1.7e-9	2.0e- 56	0
$arphi_{ exttt{1da_qda_25}}^{ exttt{Lvl}}(\mathcal{A}_{\mathcal{P}})$			6.2e- 31																3e- 181
$arphi_{ ext{quad_simple_cond}}^{ ext{MM}}(arphi_{ ext{quad_simple_cond}}(arphi_{ ext{quad_simple}}))$																			1e- 127

Table 10: The p-values of the Kolmogorov-Smirnov (KS) test comparing the equality of probability distributions of individual TSS full feature representatives on all data and on those data on which a particular model setting scored best in RDE. The p-values are after the Holm correction and they are shown only if the KS test rejects the equality of both distributions at the family-wise significance level $\alpha=0.05$, non-rejecting the equality hypothesis is indicated with —. Zeros indicate p-values below the smallest double precision number.

$\overline{\mathcal{M}}$				G	Р								RF					lmm	lq
settings	Gibbs	LIN	Mat	NN	Q	RQ:	SE+Q		CARTO MSE I		OC1				SCRT RDE				
$arphi_{ ext{dim}}$	6e- 112	3e- 182	4.8e- 30	4.8e- 18	3.0e- 14	1.6e- 25	2.5e- 14	2.2e-8	_	2.8e- 12	1.2e- 10	1.1e- 17	1.1e- 26	1.9e- 15	1.5e- 22	6.6e-5	1.3e- 25	1.5e- 17	4.1e- 72
$arphi_{ exttt{obs}}(\mathcal{A})$	2e- 286	2e- 152	6e- 308	5e- 103	3e-	4e- 111			1.5e- 17										4e- 280
$\varphi_{\text{evopath_c_norm}}^{\text{CMA}}$	2.6e- 56	0	0	3e- 131	6e- 295	7e- 152	1e- 170	7e- 140	6.9e-4	9.5e-6	2.2e-3	1.2e-4	_	_	3.9e-6	1.4e-7	1.1e-3	1e- 172	8e- 309
$\varphi_{\rm evopath_s_norm}^{\rm CMA}$	5.6e- 69	8.8e- 27	5.5e-4	3.1e- 62	9.8e- 13	6.1e- 25	3.7e- 16	9.0e- 20	4.3e- 11	7.3e-4	3.9e-2	_	1.4e- 11	1.7e-5	1.3e-8	9.6e-4	3.3e-7	4.9e- 27	1.7e- 21
$arphi_{ ext{restart}}^{ ext{CMA}}$	4.3e-2	2.2e- 77	1.7e- 59	3.5e-6	3.1e-9	1.5e-7	2.2e- 25	6.6e- 37	5.1e- 31	1.1e-4	3.8e-2	_	2.0e-2	_	3.3e-8	1.0e-4	_	1.7e- 54	7e- 138
$\varphi_{\text{step_size}}^{\text{CMA}}$	0	3e- 234	1e- 227	0	0	0	0	0	5.3e- 75	2.3e- 23	1.8e-4	5.5e-6	1.3e- 68	2.7e- 48	2.4e- 23	5.0e- 40	3.3e- 17	8.7e- 36	4.7e- 87
$\varphi_{\text{cma_lik}}^{\text{CMA}}(\mathcal{A}_{\mathcal{P}}^{\top})$	6.3e- 65	9.4e- 19	1.3e- 53	1.4e- 36	2.1e- 74	1.3e- 48	1.3e- 46	5.4e- 38	_	_	_	_	7.8e-4	5.3e-3	4.2e-4	4.5e-2	5.8e-9	2.1e- 46	5.1e- 42
$arphi_{ ext{diff_median_02}}^{ ext{Dis}}(\mathcal{A}^{ op}$) 0	9e- 234	0	0	0	0	0	0	1.2e- 88	1.7e- 23	1.0e-7	_	2.7e- 72	7.0e- 57	9.6e- 15	1.8e- 37	3.3e- 15	2e- 104	5e- 110
$arphi_{ ext{diff_median_10}}^{ ext{Dis}}(\mathcal{A}^{ op}$) 0	0	0	0	0	0	0	0	4.8e- 88	3.6e- 21	1.3e-9	7.6e-6	1.0e- 93	3.3e- 79	1.8e- 13	1.1e- 43	1.6e- 12	3.8e- 76	1e- 130
$\varphi_{\text{diff_mean_05}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	0	0	0	0	0	0	0	0	5.2e- 98	4.2e- 22	1.2e-9	8.6e-4	1.9e- 93	4.7e- 71	1.7e-8	3.6e- 47	1.1e-8	2.9e- 66	1e- 128
$\varphi_{\text{diff_mean_25}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	0	0	0	0	0	0	0	0	1.3e- 93	1.3e- 20	5.5e- 11	6.1e-4	2e- 103	2.8e- 86	5.5e-9	2.0e- 44	3.6e-8	1e- 106	5e- 201
$\varphi^{\operatorname{Lvl}}_{\operatorname{lda_qda_l0}}(\mathcal{A}_{\mathcal{P}})$	6e- 221	1e- 268	0	1e- 105	0	4e- 167	2e- 241		6.8e-6										0
$arphi_{ ext{lda_qda_25}}^{ ext{Lvl}}(\mathcal{A}_{\mathcal{P}})$	1.6e- 54	2.0e- 40	3.4e- 90	7.7e- 24	1.3e- 74	1.4e- 35	6.3e- 45	1.0e- 51	5.2e-5	1.2e-3	5.2e- 10	1.8e- 18	3.0e- 36	5.1e- 33	3.3e-3	7.5e-5	1.1e-7	5.5e- 25	4.8e- 53
$\varphi_{\rm quad_simple_cond}^{\rm MM}(\mathcal{A}$	1 ⊤) ^{1e-} ₁₉₀	8.5e- 68	6e- 220	2.1e- 64	0	2e- 177	3e- 180	4e- 148	7.1e-5	1.6e-3	_	8.8e-3	4.3e-9	6.0e-6	2.6e-4	2.0e-3	3.5e-2	0	0

Table 11: The p-values of the Kolmogorov-Smirnov (KS) test comparing the equality of probability distributions of individual TSS nearest feature representatives on all data and on those data on which a particular model setting scored best in MSE. The p-values are after the Holm correction and they are shown only if the KS test rejects the equality of both distributions at the family-wise significance level $\alpha=0.05$, non-rejecting the equality hypothesis is indicated with —. Zeros indicate p-values below the smallest double precision number.

$\overline{\mathcal{M}}$				G	P								RF					lmm	lq
settings	Gibbs	LIN	Mat	NN	Q	RQ	SE+Q		CART MSE		OC1				SCRT RDE				
$arphi_{ ext{dim}}$	0	_	1.9e-7	8.2e- 19	3.8e-3	1.8e-9	5.5e- 81	7.9e-3	9.7e- 30	5.7e- 14	2.9e-3	7.2e-3	2.9e- 10	4.4e-7	2.5e- 13	_	1.5e-2	6.7e- 36	1.5e- 77
$arphi_{ t obs}(\mathcal{T})$	0.5	5.0e-3	6.1e-6	6.0e- 20	6.2e-6	2.2e-9	3.7e- 50	1.5e-6	4.5e- 35	4.7e- 22	2.1e-9	2.0e-5	8.2e- 13	2.2e-6	2.3e- 24	9.3e-3	3.4e- 11	2e- 124	2.7e- 12
$arphi_{ ext{evopath_c_norm}}^{ ext{CMA}}$	5e- 131	6.1e-5	1.9e- 46	1.1e-7	7.0e- 32	5.2e- 18	1.6e- 33	1.5e- 80	3.5e- 26	7.9e- 19	3.8e- 23	5.5e- 12	2.7e- 19	2.2e- 17	7.4e- 19	4.7e- 30	9.2e- 13	2.6e- 61	1e- 183
$\varphi_{\text{evopath_s_norm}}^{\text{CMA}}$															1.6e-6				
$arphi_{ ext{restart}}^{ ext{CMA}}$	2.4e- ,	2.5e-4	_	1.7e-2	1.2e-3	2.3e- 75	3.8e- 42	1.0e- 45	2.6e-2	_	3.2e-2	_	3.3e-2	_	_	7.9e-4	2.5e-4	1.1e- 42	7.0e- 64
$arphi_{ t step_size}^{ ext{CMA}}$	8e- , 311 °	2.5e-5	2.4e- 21	1.3e-4	1.6e- 27	0	0	6e- 178	6.3e- 13	1.6e- 14	6.6e- 29	3.1e-8	2.1e- 16	1.9e-9	9.7e- 25	1.7e- 34	6.3e- 29	2e- 136	2.0e- 78
$arphi_{ exttt{cma_lik}}^{ exttt{CMA}}(\mathcal{A}_{\mathcal{P}}^{ op})$	8.2e- 60	_	6.1e-9	2.0e-4	1.1e- 28	3e- 182	1e- 202	6e- 180	1.0e-6	1.2e-4	1.2e- 15	2.6e-7	1.9e-5	1.1e-6	1.4e- 11	1.6e- 10	2.8e- 10	5.1e- 83	3.7e- 22
$\varphi_{\text{diff_median_10}}^{\text{Dis}}(\mathcal{A}^\top$) 3e- 146	1.9e- 10	4.9e- 37	7.4e- 16	5.2e- 57	0	0		23			10	20	23		00	40		2.7e- 14
$\varphi_{\text{diff_mean_05}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	5e- 192	2.4e- 10	4.3e- 38	4.6e- 15	1e- 100	0	0	0	6.3e- 34	1.6e- 25	9.3e- 52	9.9e- 27	5.7e- 31	5.8e- 35	5.2e- 43	1.7e- 73	2.2e- 51	0	2.3e- 19
$\varphi_{\text{diff_mean_25}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	4e- 197	3.7e- 11	6.3e- 47	5.8e- 18	8e- 101	0	0	0	5.8e- 31	2.3e- 25	5.2e- 50	9.3e- 26	2.0e- 30	5.4e- 35	1.3e- 40	1.0e- 79	8.4e- 50	0	9.6e- 43
$\varphi_{\text{diff_median_02}}^{\text{Dis}}(\mathcal{T}^{\top}$) 1.2e- 26	1.9e- 14	5.6e-7	2.0e- 33	4.8e- 21	2e- 208	7e- 228	3e- 106	_	_	6.3e-5	2.5e-3	_	_	1.6e-4	_	5.0e-9	2.7e- 81	2.1e- 28
$arphi_{ ext{lda_qda_10}}^{ ext{Lvl}}(\mathcal{A}_{\mathcal{P}})$							8e- 143	6.8e- 66	1.8e-4	1.3e-5	3.9e-5	2.0e-7	3.4e-5	2.5e-6	1.2e-2	2.1e- 14	6.2e-4	1.3e- 21	4e- 150
$arphi_{ ext{lda_qda_25}}^{ ext{Lvl}}(\mathcal{T}_{\mathcal{P}})$	9.5e- 15	3.2e- 11	_	8.0e- 14	2.9e-5	1.4e- 15	5.7e- 23	2.7e- 22	1.1e-4	2.6e-5	6.1e- 15	8.9e- 15	1.6e-4	2.4e- 12	9.7e- 13	4.5e- 16	3.7e- 12	9e- 123	9.8e- 57
$arphi_{ ext{quad_simple_cond}}^{ ext{MM}}(\mathcal{T}$	⊤ §.8e- 96	_	6.8e-6	1.9e- 10	3.3e- 38	4e- 111	2e- 118	3.0e- 90	2.2e- 11	1.7e-8	5.2e-8	5.6e-4	2.7e-9	7.4e-3	9.1e-8	1.3e-4	4.4e-5	2.5e- 20	3.6e-3

Table 12: The p-values of the Kolmogorov-Smirnov (KS) test comparing the equality of probability distributions of individual TSS nearest feature representatives on all data and on those data on which a particular model setting scored best in RDE. The p-values are after the Holm correction and they are shown only if the KS test rejects the equality of both distributions at the family-wise significance level $\alpha=0.05$, non-rejecting the equality hypothesis is indicated with —. Zeros indicate p-values below the smallest double precision number.

\mathcal{M}				G	P							RF					lmm	lq
settings	Gibbs	LIN	Mat	NN	Q	RQ:	SE+Q	SECARTO MSE		OC1			SCRT MSE					
$arphi_{ ext{dim}}$	0	2e- 189	1.2e- 31	1.5e- 28	9.2e- 32	2.6e- 15	7.0e- 10	5.6e- 44 9.5e-7	2.3e- 10	2.2e- 34	2.4e- 26	8.3e- 11	1.5e- 24	1.4e-8	4.3e- 59	2.0e- 29	7.8e- 54	2.7e-7
$arphi_{ exttt{obs}}(\mathcal{T})$	0	9.3e- 50	1.2e- 11	4.7e- 16	7.0e- 54	1.7e- 15	1.5e- 12	3.1e- 48 2.1e-5	8.1e-5	7.5e- 20	4.4e- 16	7.3e-5	5.3e- 16	9.8e-5	6.6e- 37	1.0e- 15	4.8e- 58	1.2e- 28
$arphi_{ ext{evopath_c_norm}}^{ ext{CMA}}$	2e- 171	0	0	1e- 132	6.0e- 83	5e- 159	4e- 202	2e- 284 3.3e-2	3.3e-2	2.2e-5	_	3.4e-2	1.5e-2	1.9e-3	4.2e- 12	4.9e-5	3e- 202	4.5e- 13
$\varphi_{\text{evopath_s_norm}}^{\text{CMA}}$	1.0e- 67	6.7e- 34	1.0e- 17	2.6e- 45	4.1e- 12	1.1e-9	5.8e- 17	7.7e- 10 5.3e-3	2.7e-3	1.4e-3	1.3e-3	2.1e-2	6.0e-6	_	7.3e-5	1.5e-4	2.2e- 71	1.5e- 59
$arphi_{ ext{restart}}^{ ext{CMA}}$	4.0e- 93				1e- 146	2e- 280	0	2e- 322 1.1e-2										3e- 270
$arphi_{ ext{step_size}}^{ ext{CMA}}$	0	1e- 236	0	2e- 285	7e- 234	8.6e- 80	3e- 139	2.8e- 1.0e- 72 22	5.3e- 22		3.5e- 20		9.8e- 27	2.5e- 16	3.9e- 22		2e- 191	9e- 138
$arphi_{ exttt{cma_lik}}^{ exttt{CMA}}(\mathcal{A}_{\mathcal{P}}^{ op})$	6.9e- 64	2.1e- 14	1.4e- 50	2.5e- 31	2.5e- 86	6.3e- 71	5.2e- 95	1.3e- 86 4.3e-8	1.9e-8	9.0e- 11	6.4e- 11	2.1e-8	7.3e- 13	1.0e-9	2.0e- 11	3.4e-7	5.2e- 72	3.4e- 12
$arphi_{ ext{diff_median_10}}^{ ext{Dis}}(\mathcal{A}^{ op}$	1e- 214	0	0	0	3e- 307	0	0	0 1.2e- 23		1.5e- 15			5.9e- 28	3.8e- 16	1.5e- 15		()	5e- 225
$\varphi_{\text{diff_mean_05}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	1e- 191	0	0	0	0	0	0	0 2.6e- 22	8.3e- 21	9.9e- 20	1.6e- 29	3.8e- 15	1.9e- 29	3.8e- 19	9.5e- 15	4.9e- 14	()	0
$\varphi_{\text{diff_mean_25}}^{\text{Dis}}(\mathcal{A}_{\mathcal{P}}^{\top})$	6e- 197	0	0	0	0	0	0	0 2.6e- 22	2.7e- 21	1.7e- 20	1.3e- 30		9.5e- 31	3.6e- 19	2.7e- 15	5.0e- 14	0	2e- 317
$arphi_{ ext{diff_median_02}}^{ ext{Dis}}(\mathcal{T}^{ op}$) 1.0e- 22	8e- 132	9.6e- 62	2.9e- 84	1e- 107	5e- 262	0	4e- 256 —	1.3e-3	_	1.2e-3	1.1e-3	5.3e-3	1.7e-2	3.4e-5	9.7e-6	6e- 200	1e- 215
$\varphi^{\operatorname{Lvl}}_{\operatorname{lda_qda_l0}}(\mathcal{A}_{\mathcal{P}})$	8.2e- 52	3e- 266	0	2e- 116	9e- 130	2.1e- 26	1.4e- 19	6.8e- 31 7.1e-3	2.4e-5	3.5e-3	1.7e-5	8.8e-6	2.8e-4	_	1.7e-3	1.3e-2	1.8e- 49	5e- 107
$arphi_{ t 1da_qda_25}^{ t Lvl}(\mathcal{T}_{\mathcal{P}})$	6.8e- 88	8.2e- 12	1.0e- 22	2.3e- 17	1.2e- 42	2e- 268	9e- 231	1e- 186 4.0e-6	1.4e-3	5.1e-7	1.3e- 10	6.1e-5	7.5e- 10	2.7e-9	2.2e- 11	2.5e-8	7e- 146	8.2e- 85
$\varphi_{\text{quad_simple_cond}}^{\text{MM}}(\mathcal{T}$	-⊤³3.4e- 30	3.9e- 71	4.2e- 27	1.7e- 22	3.3e- 58	1.1e- 46	9.8e- 55	4.2e- 55 2.4e-4	6.2e-4	_	1.2e-3	3.5e-3	6.8e-4	_	_	_	1.1e- 21	6.0e- 16

Table 13: The p-values of the Kolmogorov-Smirnov (KS) test comparing the equality of probability distributions of individual TSS knn feature representatives on all data and on those data on which the lmm model setting scored best in MSE and RDE. The p-values are after the Holm correction and they are shown only if the KS test rejects the equality of both distributions at the family-wise significance level $\alpha=0.05$. Zeros indicate p-values below the smallest double precision number.

	$arphi_{ ext{dim}}$	$arphi_{ m obs}(\mathcal{A}_{\mathcal{P}})$	$ ho_{ extsf{evopath.s.norm}}^{ extsf{CMA}}$	$arphi_{ m restart}$	$arphi_{ ext{cma.lik}}^{ ext{CMA}}(\mathcal{A}_{\mathcal{P}}^{ op})$	$arphi_{ ext{cma-lik}}^{ ext{CMA}}(\mathcal{T}_{\mathcal{P}})$	$arphi_{ ext{diff-mean.10}}(\mathcal{A}^{ extsf{T}})$	$arphi_{ ext{diff_median_02}}^{ ext{Dis}}(\mathcal{A}^{ extsf{T}})$	$arphi_{ ext{diff_mean_10}}(\mathcal{T}_{\mathcal{P}}^{T})$	$arphi_{ ext{diff_mean_05}}(\mathcal{T}^{ extsf{T}})$	$arphi_{ extsf{eps_max}}^{ ext{Inf}}(\mathcal{T})$	$arphi_{ ext{Lda-qda-10}}^{ ext{Lvl}}(\mathcal{A}_{\mathcal{P}})$	$arphi_{ ext{Lda-qda-25}}^{ ext{Lvl}}(\mathcal{A}_{\mathcal{P}})$	$ ho_{ ext{quad-simple_cond}}^{ ext{MM}}(\mathcal{T})$
MSE	5.2e-32	9.2e-08	1.6e-63	1.9e-75	1.2e-78	3e-267	0	0	8.3e-126	7.4e-161	0	1.1e-23	2.8e-60	2e-25
RDE	9.5e-45	1.2e-28	2.8e-67	1.4e-196	1.6e-46	0	0	0	2.8e-188	4.9e-209	0	9.2e-14	7.4e-119	6.6e-319

References

- Bajer, L., Pitra, Z., Repický, J., and Holeňa, M. (2019). Gaussian process surrogate models for the CMA Evolution Strategy. Evolutionary Computation, 27(4):665–697.
- Hansen, N. (2006). The CMA evolution strategy: A comparing review. In *Towards a New Evolutionary Computation*, Studies in Fuzziness and Soft Comp., pages 75–102. Springer.
- Hansen, N., Auger, A., Mersmann, O., Tusar, T., and Brockhoff, D. (2016). COCO: A platform for comparing continuous optimizers in a black-box setting. arXiv:1603.08785.
- Hansen, N. and Ostermeier, A. (1996). Adapting arbitrary normal mutation distributions in evolution strategies: The covariance matrix adaptation. CEC '96., pages 312–317. IEEE.
- Kern, S., Hansen, N., and Koumoutsakos, P. (2006). Local Meta-models for Optimization Using Evolution Strategies. PPSN '06, pages 939–948. Springer.
- Kerschke, P. (2017a). Automated and Feature-Based Problem Characterization and Algorithm Selection Through Machine Learning. Dissertation, University of Münster. Publication status: Published.
- Kerschke, P. (2017b). Comprehensive feature-based landscape analysis of continuous and constrained optimization problems using the R-package flacco. *ArXiv e-prints*.
- Kerschke, P. and Dagefoerde, J. (2017). flacco: Feature-Based Landscape Analysis of Continuous and Constraint Optimization Problems. R-package v. 1.7.
- Kerschke, P., Preuss, M., Wessing, S., and Trautmann, H. (2015). Detecting funnel structures by means of exploratory landscape analysis. GECCO '15, pages 265–272.
- Lunacek, M. and Whitley, D. (2006). The dispersion metric and the CMA evolution strategy. GECCO '06, pages 477–484.
- Mersmann, O., Bischl, B., Trautmann, H., Preuss, M., Weihs, C., and Rudolph, G. (2011). Exploratory landscape analysis. GECCO '11, pages 829–836.
- Muñoz, M. A., Kirley, M., and Halgamuge, S. K. (2015). Exploratory landscape analysis of continuous space optimization problems using information content. *IEEE Transactions on Evolutionary Computation*, 19(1):74–87.
- Pitra, Z., Bajer, L., and Holeňa, M. (2016). Doubly trained evolution control for the Surrogate CMA-ES. PPSN '16, pages 59–68. Springer.
- Pitra, Z., Koza, J., Tumpach, J., and Holeňa, M. (2022). Landscape analysis for surrogate models in the evolutionary black-box context. *arXiv*, 2203.11315:1–34. Under review in journal.
- Pitra, Z., Repický, J., and Holeňa, M. (2019). Landscape analysis of Gaussian process surrogates for the covariance matrix adaptation evolution strategy. GECCO '19, pages 691–699.