

Neural mechanisms underlying verification and falsification in human reasoning

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ABSTRACT

Humans can learn about the world by obtaining evidence that either verifies (confirms) or falsifies (disconfirms) current beliefs. Where multiple theories compete to explain data, such as during scientific reasoning, falsification allows a hypothesis to be definitively ruled out, whereas verification only provides incremental evidence for one view over another. Nevertheless, using a rule-learning task, we show that human participants learn faster from evidence that confirms (rather than disconfirms) a hypothesis, even though computational simulations showed that performance is maximised by learning faster from disconfirmation. Correspondingly, functional neuroimaging revealed greater BOLD responses to confirmatory relative to disconfirmatory feedback in the striatum and prefrontal cortex. These behavioural and neural biases were more pronounced when the number of possible hypotheses was greater. A bias to learn faster from confirmation may underpin suboptimal human reasoning, contributing to flawed human scientific inference.

INTRODUCTION

The capacity to create new knowledge from experience lies at the heart of human scientific endeavour and technological advancement. Primates can select actions conditioned on conjoint information about a stimulus and its context, a capacity that underpins flexible, rule-based learning and cognitive control^{1,2}. Rule-based learning occurs when clusters of sensorimotor contingencies repeatedly co-occur in a given context^{3,4}. For example, after multiple visits to London, one might learn to respect the following rule: *in London, bring an umbrella*. However, humans can also use inductive reasoning to derive abstract propositional knowledge about the world (theories), which can in turn be used to make new predictions⁵⁻⁷. For example, a tourist who has never visited Oxford can draw upon their experience with the weather in London, Bristol and Edinburgh, to derive the theory *in the UK, it usually rains*, which in turn prompts the rule *in Oxford, bring an umbrella*.

According to one philosophical tradition⁸, science advances as general theories are derived from observation, and evidence for a theory mounts as new phenomena are reported that verify its predictions. For example, our belief in the Standard Model of particle physics was increased by the discovery of the Higgs Boson particle⁹. Accordingly, psychologists have studied the mechanisms by which both adults and infants acquire new knowledge through induction¹⁰, or derived metrics for imputing cause to effect on the basis of observed contingency¹¹⁻¹³. However, an alternative philosophical position is that scientific knowledge is acquired not by verification but by falsification, because a limitless number of confirmatory experiments are needed to verify a theory exhaustively¹⁴. To give one oft-cited illustration, a definitive test of the theory *all swans are white* would require the colour of all swans past, present and future to be measured, whereas this theory could be disproved by the discovery of a single black swan. This position is enshrined in conventions for the reporting of frequentist statistics, where the null hypothesis is assumed to be true unless disproved¹⁵.

Philosophical considerations aside, how science advances will depend on the nature of the cognitive architecture that is employed by researchers as they design experiments and scrutinize data. Here, thus, we sought to understand how humans learn to reason from observation, and how verification and falsification of hypotheses are implemented in the neural circuitry underpinning voluntary choices. It has been shown that even when conclusive information is obtained only from disconfirmation, humans seek out confirmatory (i.e. verifiatory) evidence by preference^{16,17} and may overvalue confirmatory evidence when it arises^{6,18,19}. Confirmatory biases may allow humans to generalise old observations to new situations, but may also hinder scientific reasoning, prompting researchers to overlook findings that are inconvenient for a favoured theory, or engage in flawed statistical methods such as circular analysis²⁰. Here, we asked human participants to perform an experience-based rule learning task, allowing us to probe how humans learned from feedback that verified (confirmed) or falsified (disconfirmed) a current contention. Simultaneously acquiring functional magnetic resonance imaging (fMRI) allowed us to measure behavioural and neural concomitants of learning from confirmatory and disconfirmatory feedback, offering insights into the neural locus of the decision biases that limit human inductive inference.

Confirmatory biases may be particularly prevalent in situations that require multiple possible hypotheses to be entertained, because capacity limits preclude the online maintenance of information about which contentions remain valid and which have been disproved. Hand-in-hand

with the capacity to reason, however, humans have evolved systems of symbolic representation and communication that allow rules to be conveyed and communicated without costly trial-and-error search over possible hypotheses. A related goal of our experiment was to understand how human reasoning changes as the information burden is alleviated by symbolic information. We thus included two conditions: in the *familiar cues* condition, participants viewed symbolic cues that offered partial information about the rule that was valid in that block, whereas in the *novel cues* condition, no such information was available. Comparing behaviour and brain activity in these conditions allowed us to assess how two landmark human cognitive advances – reasoning and symbolic communication – interact during the learning of task rules.

RESULTS

Task and design. Eighteen healthy human participants learned to classify stimuli (each composed of a pair of coloured shapes occurring left and right of fixation) as ‘target’ or ‘nontarget’ on the basis of their shape (square, circle, triangle) and/or colour (red, green, blue), responding with a button press, and receiving fully informative trial-and-error feedback after each response (Fig. 1a). Each decision rule, which remained constant over a block of 16 trials, defined targets (50% of trials) as a disjunctive combination of one feature on the left and another on the right. For example, in one block stimuli were targets if the stimulus on the left was red OR the stimulus on the right was a triangle (i.e. rule={red-left, triangle-right}). The use of a disjunctive (OR rule) ensured that an effect (target) could have multiple possible causes (i.e. potential decision rules). In the previous example, positive feedback could be received for the response ‘target’ to the stimulus {red-left, circle-right}, even though circle-right is not part of the rule. This feature of the task allowed us to tap into a fundamental aspect of scientific reasoning, whereby multiple competing hypotheses vie to explain an empirically observed phenomenon, and inferences should be based on the hypothesis providing the strongest positive evidence.

The disjunctive task also ensured that targets and nontargets provided asymmetric information about the decision rule. By analogy with the ‘black swan’ example introduced above, feedback that a stimulus was a nontarget allowed participants to eliminate hypotheses about the decision rule component {e.g. red-left \notin target}, whereas information that the stimulus was a target provided only incremental evidence for {red-left \in target} but precludes definitive inclusion or exclusion of candidate decision rules. This design thus allowed us to pursue our main question of interest: how participants learn differently from verification, i.e. ‘confirmatory’ evidence (indicating that a stimulus was a target) and falsification, i.e. ‘disconfirmatory’ evidence (indicating that it was a nontarget).

In the *familiar cues* condition, symbolic cues, whose meaning had been learned in a previous training session, disclosed the *class* of rule that applied in that block. The rule ‘class’ referred to the dimension (e.g. {colour-left, shape-right}) but not the feature (e.g. {red-left, triangle-right}) that was relevant for decisions. There were four rule classes: {colour-left, colour-right}, {colour-left, shape-right}, {shape-left, colour-right}, and {shape-left, shape-right}. In the *novel cues* condition, a distinct set of cues were paired randomly with blocks, so that they offered no information about the relevant rule. An illustration of the task is provided in Fig. 1a, and more detailed information is available in the Online Methods.

Computational model. Our first concern was to understand the computational mechanisms by which humans performed the task. A large family of models have been proposed to describe human inductive learning^{11-13,21-26}, but when causal relationships are disclosed gradually by experience, a class of model in which evidence is integrated over time performs particularly well²¹. We thus assumed that participants entertained a number of parallel competing hypotheses about the possible decision rule (potential causes), and responded according to whether there was criterial evidence that either the left or right stimulus was a target, i.e. accumulated confirmatory vs. disconfirmatory evidence²¹. At the beginning of blocks with novel cues, there were 12 potential causes: 4 side-dimension pairs (colour-right, colour-left, shape-right and shape-left) x 3 possible features (red, green, blue or circle, square, triangle); in blocks with familiar cues this space was reduced by half, as the irrelevant dimensions on each side were eliminated according to the information afforded by the cue. Each potential cause was associated with an estimate of the value of responding “target” (positive values) vs. “nontarget” (negative values), that we refer to as hypothesis value (or H_{ij} where i and j are the indices over all the possible side-dimensions and features respectively).

For each trial, we calculated the target value (TV) of each stimulus by combining H_{ij} for the two (familiar cues) or four (novel cues) relevant features according to the following equation:

$$TV = \tau \cdot \min_{(i,j) \in S(t)} (H_{(i,j)}) + (1-\tau) \cdot \max_{(i,j) \in S(t)} (H_{(i,j)}) \quad (1)$$

where $S(t)$ is the subset of two (*familiar cues* condition) or four (*novel cues* condition) pairs (i,j) representing each feature presented in trial t .

The model responds “target” when $TV > 0$ and “nontarget” otherwise. The free parameter τ thus encodes the extent to which decisions are based on the maximum hypothesis value across all potentially relevant features (i.e. the hypothesis with strongest support; $\tau = 0$), the minimum hypothesis value (the hypothesis with weakest support; $\tau = 1$) or on a mixture information in the observed features ($0 < \tau < 1$).

Following each decision, the model updates estimates of the relevant hypotheses according to a delta rule:

$$dH_{(i,j) \in S(t)} = \frac{1}{2} \cdot \cos(\frac{1}{2} \pi \alpha) \cdot ((+1) - H_{(i,j) \in S(t)}) \quad \text{if target} \quad (2.1)$$

$$dH_{(i,j) \in S(t)} = \frac{1}{2} \cdot \sin(\frac{1}{2} \pi \alpha) \cdot ((-1) - H_{(i,j) \in S(t)}) \quad \text{if nontarget} \quad (2.2)$$

$$H_{(i,j) \in S(t)} = H_{(i,j) \in S(t)} + dH_{(i,j) \in S(t)} \quad (3)$$

The free parameter α thus controls the relative rate at which participants learned from targets ($\alpha = 0$), from nontargets ($\alpha = 1$) or a mixture of the two strategies ($1 > \alpha > 0$). Our model is thus closely allied to standard reinforcement learning accounts of choice behaviour²⁷ and bears a relation to connectionist models of induction in which category belongingness is updated with a delta rule²⁵. Critically however, we propose separate parameters controlling the relative rate at which participants update their beliefs from confirmatory or disconfirmatory feedback (α), and whether their decisions are determined by the strongest evidence, or by some mixture of the evidence for multiple probable candidate causes (τ).

Model performance. Using the sequences of stimulation and feedback viewed by human participants as inputs to the model, we began by calculating how the model performed under different parameterisations of τ and α (Fig. 2a). Model performance peaked close to $\alpha = 1$ (maximal learning from disconfirmation) and $\tau = 0$ (decisions based on the strongest evidence), for blocks with both novel and familiar cues (Fig. 2a). Intuitively, this follows from the disjunctive rule, which ensures that disconfirmatory evidence is more informative. Consider a stimulus composed of a red element on the left and a blue element on the right. Feedback that the stimulus is a ‘nontarget’ allows the candidate rules {red-left} and {blue-right} to both be eliminated. Feedback that the stimulus is a ‘target’ implies that one component of the rule is either {red-left} or {blue-right}, but does not indicate which is correct. The best learning strategy is thus to weight disconfirmatory more than confirmatory evidence ($\alpha = 1$). At the time of choice, however, if strong evidence exists that {red-left} is a part of the rule, under a disjunctive rule the stimulus is likely to be a target irrespective of the hypothesis value of {blue-right}. The best policy is thus to base decisions on the maximum hypothesis value ($\tau = 0$) – not on a mixture of the available feature values.

Model fitting to human choice and accuracy data. Next, we adjusted these parameters to minimize the mean square error (MSE) between model and human behaviour. Average human accuracy (% correct responses) and choices (% target responses) across the 16 trials that constituted each block are shown in Fig. 1b. Accuracy increased across the block overall ($F_{(15,255)} = 29.9$, $p < 1 \times 10^{-9}$), but did so faster in blocks with familiar cues (trial \times cue interaction, $F_{(1,17)} = 12.89$, $p < 0.008$). Participants began with a bias to respond ‘target’ that abated across the block ($F_{(15,255)} = 46.1$, $p < 1 \times 10^{-9}$) in roughly equal measure for the two conditions ($F < 1$). The choice and accuracy of the best-fitting model parameterisation (lines) is rendered onto equivalent human data (dots) in Fig. 1b-c. As can be seen, the model was able to capture the pattern of behaviour accurately across the block in each condition. For completeness, we also compared human performance to that of a hierarchical Bayesian model that estimated the conditional probability that each stimulus was a target. This model outperformed participants by a wide margin, and provided poorer fits to human data even when the number of parameters was taken into account (supplementary section 1).

Interestingly, however, the parameters that allowed these fits to human performance to be obtained took on values that diverged from those predicting maximal model performance (Fig. 2b). Specifically, mean values of α were 0.64 ± 0.28 and 0.38 ± 0.21 for blocks with familiar and novel cues respectively, diverging reliably from the respective parameters ($\alpha = 0.89$ [familiar cues] and $\alpha = 0.95$ [novel cues]) that yielded maximal performance under this model in both cases (both $t_{(17)} > 3.8$, $p < 0.001$). Similarly, mean values for τ were 0.4 ± 0.31 (familiar cues) and 0.64 ± 0.21 (novel cues), both showing a divergence from the performance-maximising parameters ($\tau = 0.05$ in both cases) that was statistically reliable (both $t_{(17)} > 4.3$, $p < 0.001$). In other words, participants eschewed the normative strategy that would yield maximal performance, i.e. to eliminate disconfirmed hypotheses and base choices on the strongest evidence offered by the available stimulus features. Rather, descriptively, they learned more from confirmatory feedback, and based their choices on a mixture of information from both stimuli. However, values of τ were smaller ($t_{(17)} = 2.8$, $p < 0.006$), and values of α were larger ($t_{(17)} = 3.3$, $p < 0.004$), in the familiar relative to novel cues condition, indicating that participants were closer to the performance-maximising strategy when given advance information that reduced the space of possible hypotheses.

Best-fitting parameters predict subject performance. As shown above, model parameterisations for which α was high (learning from disconfirmation) and τ was low (decisions guided by the strongest evidence) achieved the highest levels of performance. If our model is faithfully describing human behaviour, one might expect that the performance shown by each human subject would depend on his or her individual parameterisation, as estimated by the model fit. We thus calculated correlations across the cohort between mean performance levels (% correct) and the estimated values for α and τ for each participant. Consistent with this view, α positively predicted performance ($r = 0.64$, $p < 0.004$) and τ negatively predicted performance ($r = -0.67$, $p < 0.002$). In other words, to the extent that participants did learn from disconfirmation, and chose according to the strongest evidence, they were more successful on the task. Scatter plots of parameters and participant performance are shown in Fig. 2c.

Influence of feedback history. Next, we used the best-fitting parameterisation of the model to predict additional aspects of human performance, with a particular focus on how learning from past veridical and falsidical feedback might influence behaviour. We plotted how the human probability of responding “target” [$p(\text{target})$] varied as a function of the history of feedback indicating that each of its features was part of a target or nontarget. Specifically, for each trial t ($t > 1$), $p(\text{target})$ was computed as a function of the number of previous times that the relevant feature (e.g. in a colour-colour block, red-left or green-right) and irrelevant feature (e.g. circle-left, square-right) had received confirmatory feedback (i.e. feedback indicating that the stimulus contained a target) and disconfirmatory feedback (i.e. indicating that it did not). We collapsed over instances where there were 4 or more confirmations or disconfirmations, and subjected the data for both relevant and irrelevant dimensions to an analysis of variance crossing the factors cues (novel, familiar) \times feedback (confirmatory, disconfirmatory) \times bin (1-4). A three-way interaction between these factors indicated that the slope of the line relating feedback history to choice was steeper for targets than nontargets, but that this difference was particularly acute in the novel cues condition ($F_{(2,31)} = 73$, $p < 0.001$). In other words, the receipt of confirmatory feedback seemed to have a more potent effect on future behaviour than disconfirmatory feedback, particularly when a large number of hypotheses were possible. A similar pattern was observed for the model (using the previously-estimated best-fitting parameterisation) where it was driven by the relative differential learning rate for targets and nontargets. Human (dots) and model (lines) data are shown for comparison in Fig. 2d.

Functional neuroimaging. Together, these behavioural and modelling analyses suggest that during inductive reasoning, humans pursue a suboptimal strategy in which learning is over-reliant on confirmatory feedback, and that this tendency is particularly pronounced in the novel cues condition, when the number of possible hypotheses about the rule exceeds the likely capacity of online maintenance processes. Next, thus, we sought to characterise the neural mechanisms which might give rise to this bias, by examining blood-oxygen (BOLD) data from functional magnetic resonance imaging (fMRI) whilst participants performed the task.

Model-based fMRI analysis. We began by using the imaging data to validate our computational model. On each trial the model provides an estimate expected value (EV), which reflects the distance of the target value from zero (unsigned TV, or $|TV|$) and thus the probability of a correct response. This quantity was calculated under the best-fitting parameterisation and the performance-maximising parameterisation, and regressed against the BOLD response at each

voxel in the brain (Fig. 3a). For the best-fitting parameterisation of the model, we observed a strong correlation between EV and BOLD signal in the medial orbitofrontal cortex, (or ventromedial prefrontal cortex; VMPFC) (peak: -6, 52, -2, $t_{(17)} = 6.19$, $p < 1 \times 10^{-5}$), a brain region where BOLD signals²⁸ and single-unit activity usually correlate with expected value²⁹. However, no reliable correlations with VMPFC BOLD were observed when the expected value was calculated under the performance-maximising parameterisation (Fig. 4a, right panel). This finding provides further support for our computational model. We also included in the same design matrix regressors that encoded the maximum and minimum hypothesis value on that trial, independent of target value (i.e. the maximum belief in any hypothesis, independent of the stimulus presented). Interestingly, we positive correlations between the maximum and minimum hypothesis values were observed in the medial and lateral orbitofrontal cortices respectively (Fig. 4b), regions that have previously been implicated in respectively encoding the value of a chosen option and of the next best alternative³⁰.

Next, we examined BOLD signals in the VMPFC as a function of the stimulus (target vs. nontarget), and cue (novel vs. familiar) on each trial. To provide a neural index of the learning occurring following feedback, we contrasted the BOLD signals elicited by correct and error feedback on these trials. Although there was a strong effect of feedback ($p < 0.0001$), the divergence between correct and error signals occurred in roughly equal measure for targets in novel and familiar cues conditions (time x cues interaction, $p = 0.38$) and for nontargets in novel and familiar cues conditions ($p = 0.33$). However, a reliable stimulus x time interaction showed that overall, there was more robust neural correlates of learning in the VMPFC from targets (i.e. confirmatory feedback) than from nontargets (disconfirmatory feedback), irrespective of the cues (stimulus x time interaction, $F_{(5,85)} = 5.2$, $p < 0.02$). This suggests the existence of a distinct mechanism that favours learning from targets over nontargets. These data are shown in Fig. 3c (left panels). When we explored other brain regions that responded robustly to correct > error trials, such as the putamen (Fig. 3c, right panel and Fig. S2), we observed no differences between the signals elicited by target/nontarget stimuli or familiar/novel cues (all p -values > 0.4).

fMRI analyses. The BOLD responses observed in the VMPFC thus provide neural evidence that participants learned faster from confirmatory feedback (indicating the stimulus was a target) than from disconfirmatory feedback (indicating it was a nontarget) irrespective of whether the cues were novel or familiar. However, behavioural data above indicate that disconfirmatory learning is particularly weak when the cues are novel and the number of possible hypotheses is large. To pinpoint the neural locus of this bias, we identified voxels that responded to the three-way interactions between stimulus, cue and feedback. Significant three-way interactions among were observed prominently in dorsal stream cortical areas implicated in online maintenance of task-relevant information and model-based learning (Fig. 4a), including caudal portions of the dorsolateral prefrontal cortex (DLPFC; right peak: 38, 8, 58; $t_{(17)} = 10.03$, $p < 1 \times 10^{-8}$), the rostromedial prefrontal cortex (RLPFC; left peak: -38, 60, 2; $t_{(17)} = 5.97$, $p < 1 \times 10^{-5}$; right peak: 34, 60, 2; $t_{(17)} = 5.92$, $p < 1 \times 10^{-5}$), and the inferior parietal lobule (IPL) (right peak: 42, -44, 42; $t_{(17)} = 7.61$, $p < 1 \times 10^{-6}$). All activations reported here survive whole-brain correction for multiple comparisons using the voxelwise false discovery rate approach³¹. A full description of the regions activated is reported in tables S1 and S2.

To understand better how these regions were responding during the task, we then re-estimated brain activity using a finite impulse response (FIR) filter, allowing us to plot peristimulus BOLD

signals for each condition of the 2 x 2 x 2 design (Fig. 4b). In each of the regions, for the familiar cues condition BOLD signals initially encoded the choice made by participants (choose target > choose nontarget) and then later encoded the feedback received (error > correct). These observations were qualified by reliable stimulus x feedback interactions (i.e. main effect of choice, target vs. nontarget) in an early time window at ~6-8s post-stimulus (IPL: $F_{(1,17)} = 34.6$, $p < 0.0001$; DLPFC: $F_{(1,17)} = 11.7$, $p < 0.004$; RLPFC: $F_{(1,17)} = 11$, $p < 0.005$), accompanied by a nonsignificant effect of feedback (all p-values > 0.1), alongside a significant main effect of feedback in a later window some ~8-10s post-feedback (IPL: $F_{(1,17)} = 41.2$, $p < 1 \times 10^{-5}$; DLPFC: $F_{(1,17)} = 43$, $p < 1 \times 10^{-5}$; RLPFC: $F_{(1,17)} = 41.4$, $p < 1 \times 10^{-5}$) with no reliable stimulus x feedback interaction (all p-values > 0.1). Although robust BOLD responses were observed in the novel cues condition, they did not differ as a function of choice or feedback in this way (all p-values > 0.1).

Comparing error and correct responses. According to reinforcement learning models, the relative brain response evoked by positive and negative feedback provides an index the effectiveness of learning in a given condition. Next, thus, we explored the three-way interaction identified above in a complementary fashion, replottting the data from Fig. 4b to show the differential BOLD response for error and correct trials (Fig. 5a) for each combination of stimulus (target, nontarget) and cue (familiar, novel) and performing appropriate post-hoc statistical tests. In the DLPFC and IPL, for *targets* the error–correct signal diverged strongly from zero, but did not differ according to whether the cues were familiar (full lines) or novel (dashed lines) (time x cues interaction for targets: DLPFC, $p = 0.38$; IPL $p = 0.13$), although this effect did just reach significance in the RLPFC ($F_{(5,83)} = 2.5$, $p < 0.04$). However, for *non-targets*, the error-correct signal diverged strongly from zero for familiar but not for novel cues (time x cues interaction for nontargets: DLPFC, $F_{(4,75)} = 6.8$, $p < 0.0001$; IPL, $F_{(4,74)} = 5.2$, $p < 0.001$; RLPFC, $F_{(4,74)} = 5.8$, $p < 0.0001$). Of note, maximal differences were observed in the later of the two time-windows, consistent with an effect driven by the feedback. In other words, for each of these regions, neural signals accompanying learning from targets were equally strong for familiar and novel cues, but neural signals indexing learning from nontargets were stronger for familiar than novel cues. This is consistent with the finding from behaviour that learning from nontargets is specifically dampened when the cues are novel and capacity is stretched to the limit.

Correlations between brain activity and model parameters. If brain activity in these structures underlies the inductive reasoning occurring during the task, then one might expect that across the cohort, those participants whose strategy more closely approximated that of the performance-maximising model – i.e. those who were more prone to learn from disconfirmation – would show greater BOLD signals in these regions. We thus correlated individual subjects' values of α with the strength of the peak BOLD signal (error-correct, in the late window) obtained in each of the conditions. In the DLPFC and RLPFC, individual values of α predicted the BOLD signal in all conditions except for nontargets in the novel cues condition (Fig. 5b-c). No such correlation was observed in the IPL.

DISCUSSION

Humans can generalise rule-based knowledge to make inferences in novel situations. For example, knowledge of the features that denote membership of a given category allows previously unseen exemplars to be accurately classified ('feature generalisation'). This allows high-level categories and concepts to be formed ('that is a dog') and contributes to linguistic development during

infancy⁵. However, even in maturity humans are subject to stereotypical inductive biases, betraying a tendency to learn preferentially from information that verifies, rather than falsifies, a currently-entertained hypothesis. Failures of human induction can stymie scientific reasoning, provoking researchers to discount evidence that contradicts an established or currently-favoured hypothesis. Excessively confirmatory scientific reasoning may increase the volatility of scientific progress, as evidence accretes strongly around a single theory until it is dramatically swept away in a ‘paradigm shift’³². In other domains, confirmatory biases may be yet more pernicious, leading for example to stereotyping and prejudice on the basis of race, gender, or social background¹⁷.

Here, we explored the neural and computational basis for this confirmatory bias in inductive reasoning, creating a trial-and-error rule-learning task in which falsification allowed definitive inferences about a rule, whereas verification only allowed beliefs to be updated incrementally. Our task simulates the circumstances that scientists face when interpreting data, whereby the observation of an empirical effect often provides evidence for multiple competing hypotheses, whereas the disconfirmatory evidence can allow a theory to be definitively ruled out. We modelled the task with a learning model in which evidence that each stimulus produced an effect was updated following trial-and-error feedback. Many theories compete to describe how humans impute cause to effect, including models that rely on associative mechanisms³³ and those that assume probabilistic inference over the structure of the world^{11,34}. Our experiment was not designed to speak directly to this controversy, but rather to investigate how humans weight confirmatory and disconfirmatory evidence when reasoning about causes.

Humans performed suboptimally on the task, in two distinct ways. Firstly, humans failed to base decisions on the most relevant of the two potential sources of information – that which provided the strongest evidence. In other words, even where an effect can be provoked by any one of several competing causes, humans diagnoses tend to be based on a mixture of information provided by all causes, rather than that furnished by the most probable cause. A similar bias has been shown in a learning task in which participants were required to assign credit to multiple possible reward-predicting features³⁵.

Secondly, and consistent with an established tradition in cognitive psychology¹⁷, humans exhibited a bias towards confirmatory learning on the task, even though model simulations showed that disconfirmatory learning would have led to improved performance. Model fits demonstrated that suboptimal performance was driven by steeper learning from feedback that verified (rather than falsified) a contention. Accordingly, when we measured brain activity in the VMPFC, a brain region characterised by its robust response to informative feedback and reinforcement, the BOLD signal evoked by errors and correct trials diverged more sharply for targets (providing confirmatory information) than for nontargets (permitting definitive disconfirmation).

Behaviourally, the confirmation bias was less pronounced when an advance cue provided partial information about the rule (familiar cues condition), reducing the number of competing alternative hypotheses by a factor of two (relative to the novel cues condition). Neurally, this interaction between the target and the cue was reflected in BOLD activity recorded from dorsal stream structures previously implicated in learning and acting upon theories about the world (‘model-based’ rather than ‘model-free’ learning)³⁶⁻³⁹. In these regions, neural signatures of learning were particularly dampened following disconfirmatory feedback in the novel cues condition. In our model, efficient performance required the parallel maintenance and updating of

multiple competing hypotheses about the rule, a process that is likely to draw heavily upon control structures in the parietal and prefrontal cortices. The reduction in learning-related activity in these regions following nontargets may reflect a failure of this updating process, consistent with the PFC as a locus for prediction errors on causal learning tasks^{40,41}. The link between neural activity and task strategy was further supported by correlation analyses showing that BOLD signals in the prefrontal cortex predicted the learning parameter that encoded the relative extent of learning from confirmation and disconfirmation. These findings are consistent with previous brain imaging work suggesting that the prefrontal cortex maintains task rules during inductive tasks⁴²⁻⁴⁴ and that patients with (left) prefrontal damage are impaired in generating abstract rules⁴⁵.

Our simulations show that confirmatory hypothesis-testing is suboptimal. Why might humans have evolved such a strategy? A key challenge faced by any intelligent system is that the space of possible occurrences in the world is virtually unbounded, a state of affairs that places unrealistic demands on the cognitive apparatus we use to make inferences about the local environment. In philosophy, this is known as the ‘frame problem’⁴⁶. Hypothesis-testing may have evolved to act as a tractable inference strategy, allowing inferences to be updated about the most probable occurrences in the world, avoiding costly inference about a series of impossible or improbable events, and contributing to optimal data harvesting⁴⁷. In our experiment, we saw that symbolic cues reduced the bias to learn excessively from confirmation. The evolution of a system of symbolic representation, that permits hypotheses to be represented and communicated via language, is likely to have enhanced the tractability of inference by allowing humans to focus on a few probable theories without costly trial-and-error learning.

FIGURE 1

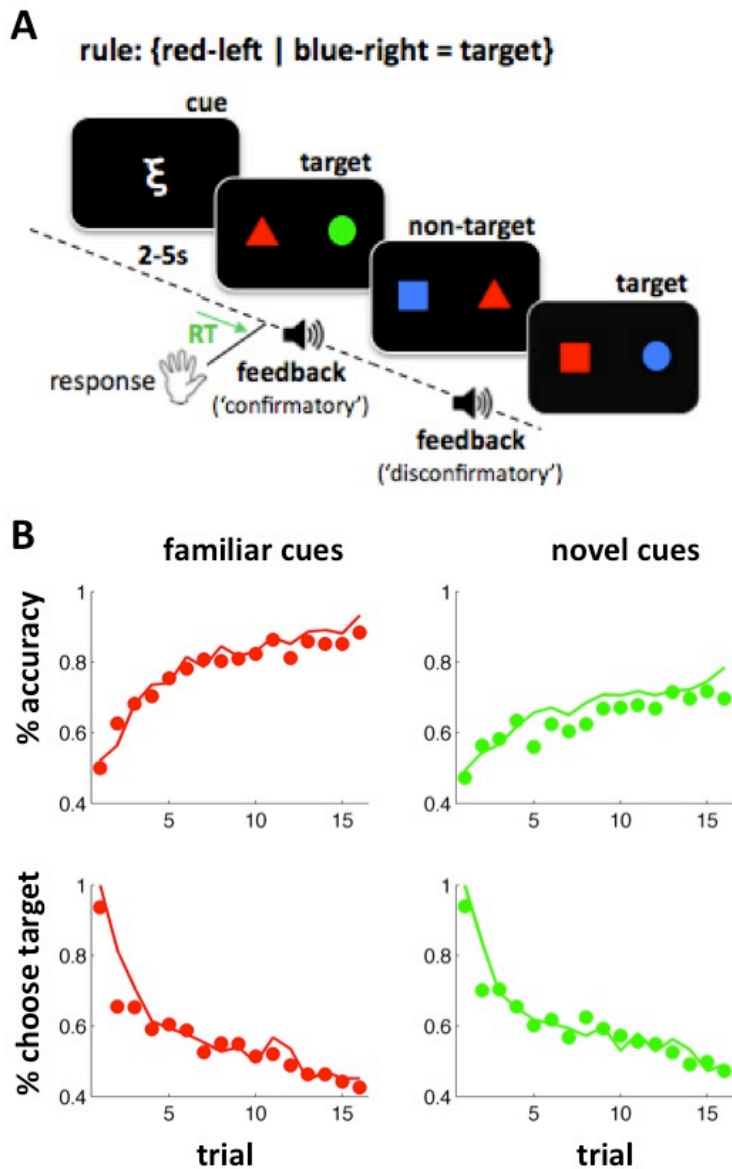


Figure 1. A. Behavioural task. On each trial participants view a pair of coloured shapes on the left and right of the screen. They responded ‘target’ or ‘nontarget’ and were provided with positive or negative auditory feedback. Each block of 16 trials was preceded by a Greek symbol which was either fully predictive of the relevant dimensions (e.g. {colour-left, colour-right}; familiar cues condition) or not at all predictive (novel cues condition). Associations between cues and dimensions were learned in a prior training session. **B.** Behavioural data from 18 human participants (dots) and predictions of the model (lines). % accuracy (top panels) and % target choice (bottom panels) over trials (1-16) are shown for familiar cues (left panels; red) and novel cues (right panels; green). Lines show the same data for the best-fitting parameterisation of the model.

FIGURE 2

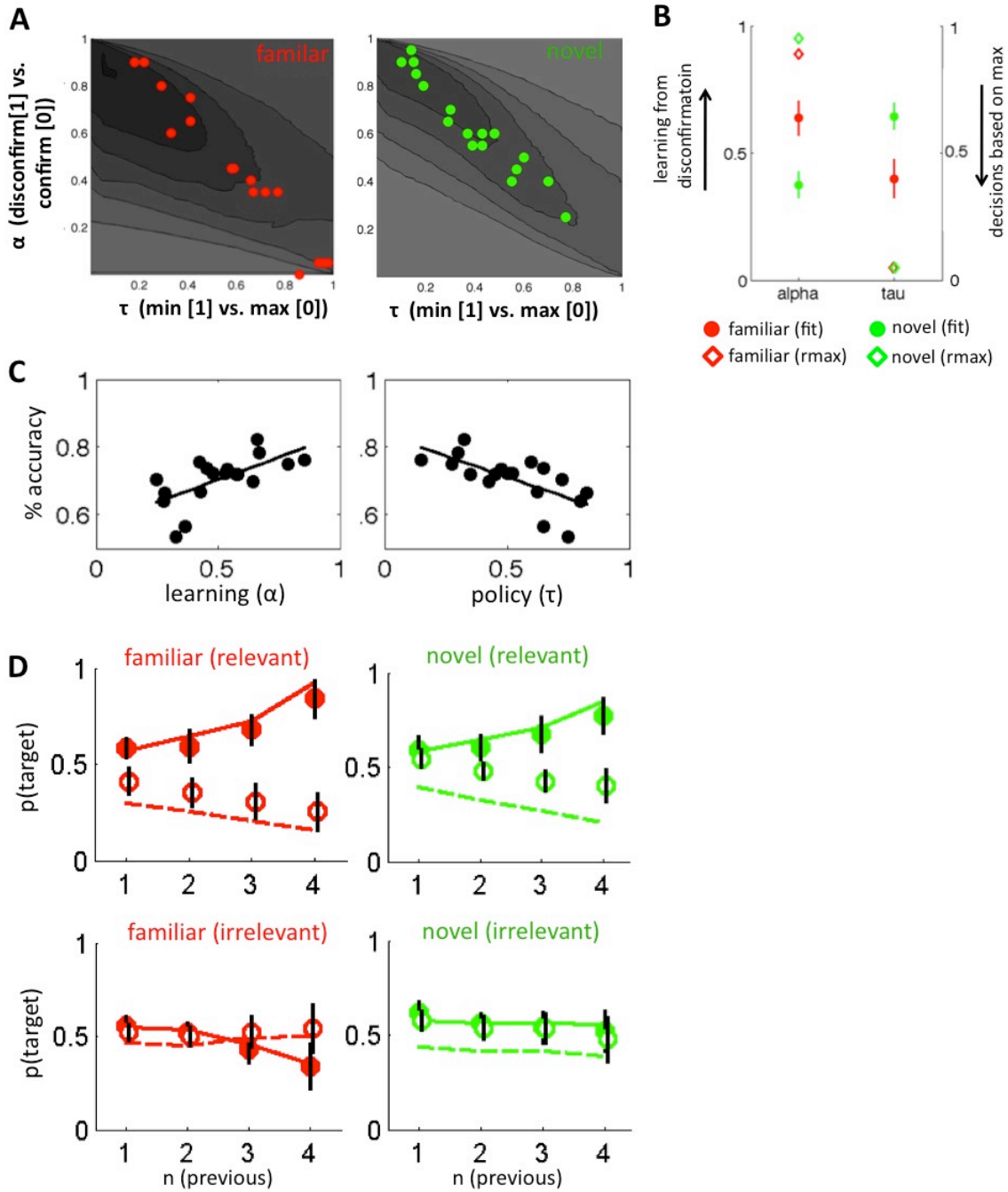


Figure 2. A. model performance as a function of model parameters α (steepness of disconfirmatory vs. confirmatory learning) and τ (decisions based on max vs. min target value) on blocks with familiar cues (left panel) and novel cues (right panel). Darker grey indicates higher performance (% correct). Red and green dots show individual human subjects in the novel and familiar cues conditions respectively. **B.** Best-fitting (circles) and reward-maximising (diamonds) values of α and τ in the familiar (red) and novel (green) cues condition. **C.** Correlations between α (left panel) and τ (right panel) and performance (% correct) for each subject. The line shows the best-fitting linear trend. **D.** % choose target responses as a function of the number of previous times the rule-relevant feature (top panels) and rule-irrelevant feature (bottom panels) was associated with confirmatory (filled circles) or disconfirmatory (open circles) feedback. Data are shown separately for familiar (left panels) and novel (right panels) cues conditions.

FIGURE 3

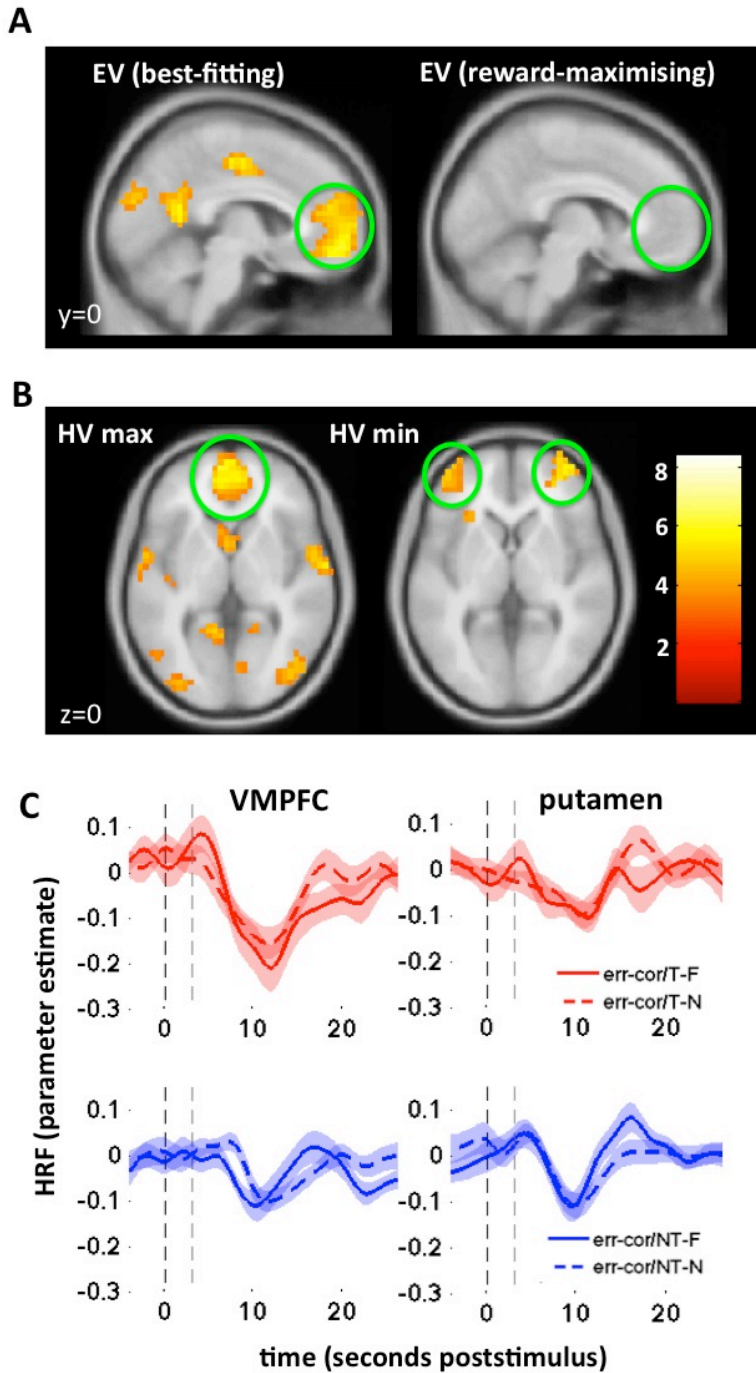


Figure 3. A. Voxels correlating with expected value predicted by the best-fitting parameterisation (left panel) and the parameterisation that maximises performance (right panels) rendered onto a sagittal slice of the MNI template brain. The green ring highlights the VMPFC. **B.** Voxels correlating with the maximum hypothesis value (left panel) and minimum hypothesis value (right panel). Green rings indicate the medial and lateral orbitofrontal cortex respectively. **C.** Relative peri-stimulus BOLD responses on error and correct trials (error minus correct) for targets (top panels) and non-targets (bottom panels) in the VMPFC (left panel) and putamen (right panel).

FIGURE 4

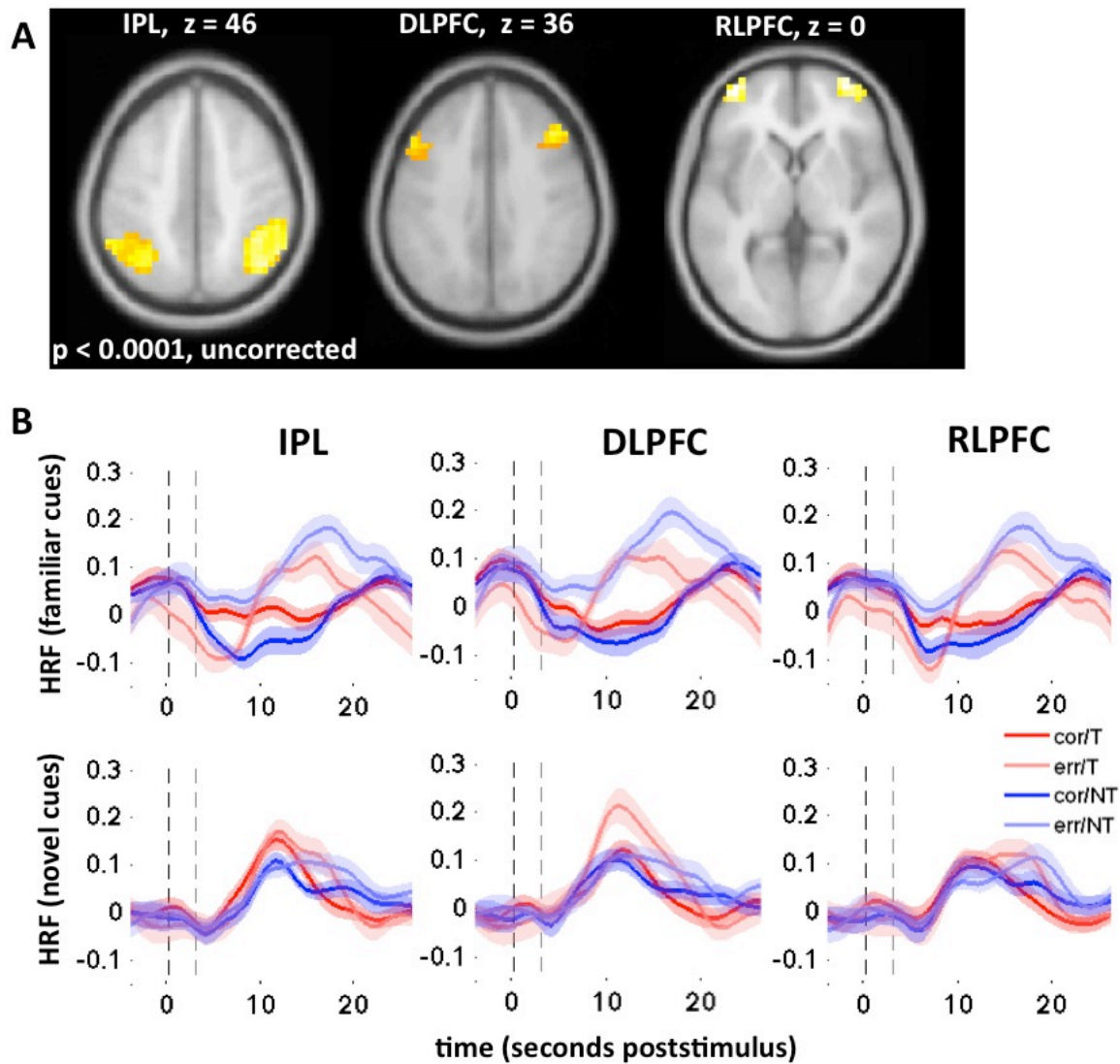


Figure 4. A. Voxels responding to the three-way interaction between stimulus (target vs. nontarget), cue (novel vs. familiar) and feedback (correct vs. error) at a threshold of $p < 0.0001$ (uncorrected). **B.** Haemodynamic responses estimated with an FIR filter in three brain regions: the inferior parietal lobule (IPL), dorsolateral prefronta cortex (DLPFC) and rostrolateral prefrontal cortex (RLPFC). Blue lines show nontarget trials and red lines show target trials; darker lines show correct trials and lighter lines error trials. Upper panels show the responses in familiar cues blocks, and lower panels the responses in novel cues blocks. The darker dashed line at time zero indicates stimulus onset and the subsequent lighter dashed lines shows when the auditory feedback signal occurred.

FIGURE 5

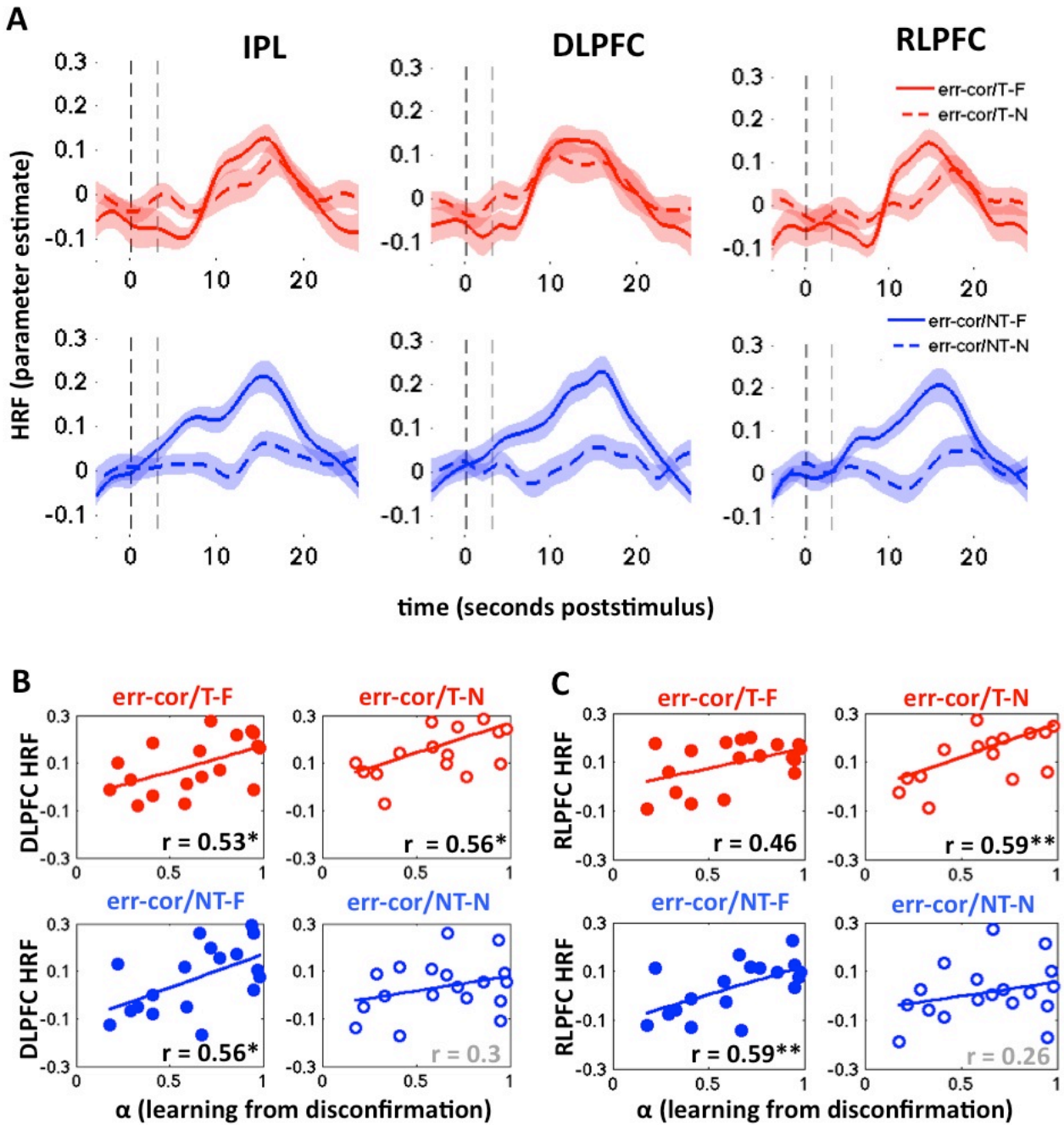


Figure 5. A. Data from figure 3b are replotted as difference between error and correct trials for IPL, DLPFC and RLPFC as a function of cues (familiar [full lines] vs. novel [dashed lines]) and stimulus (target [red; upper panels] vs. nontarget [blue; lower panels]). **B.** Correlations between relative BOLD signals in the DLPFC and model parameter α (indexing relative steepness of confirmatory and disconfirmatory learning). Larger relative (error-correct) BOLD signals following feedback led to higher values of α in all conditions except the novel cues, nontarget trials. **C.** As for B, but correlations with RLPFC BOLD.

Online Methods

Subjects. 18 healthy participants (6 female, 12 male; age 20-34, mean 25.0 years) were recruited into the study in accordance with local ethical guidelines. No participants reported a history of psychiatric or neurological illness, and all had normal or corrected-to-normal vision. They were paid £35 for participation in both a practice and a scanner session on two separate days.

Stimuli and task. Participants performed a rule-learning task that required pairs of shapes to be classified as “target” or “non-target” according to an unknown rule. Each block began with the presentation of one of 8 abstract symbolic cues (Greek letters) for 3s. After a period of 2-5s (jittered), a train of 16 pairs of stimuli appeared on the left or right of the screen at approximately 3° eccentricity. Each stimulus pair remained on the screen for 3s. Each member of the pair could be a square, circle or triangle coloured red, green or blue. Participants pressed a key (training task) or response button (scanner task) at any point during the 3s presentation period to indicate whether the stimulus was a target or nontarget. Stimulus-response contingencies were fully counterbalanced across participants. Responses were followed by fully informative auditory feedback consisting of a pair of tones with ascending (correct) or descending (incorrect) pitch (400Hz; 800z) that lasted 200ms in total. An interval of 2-6s was interposed between stimuli, during which the screen was blank. Participants completed a total of 48 blocks during training, and then a further 48 blocks for the scanner session, which occurred on a subsequent day. In the scanner, the experiment was divided into 4 runs of 12 blocks each, buttressed by lead-in and lead-out durations of 10s. Each block lasted ~17 minutes, bringing total scanning time to just over an hour.

Rules. Rules were disjunctive – for example, in one block the rule might be “if the shape on the left is red, OR the shape on the right is blue, the stimulus pair is a target”. We denote this {red-left, blue right}. We divided rules into 4 classes according to the feature that was relevant on each side (left-right): colour-colour, colour-shape, shape-colour and shape-shape. The same feature was never selected in both cases, for example {red-left, red-right} was disallowed, leaving 6 possible rules in each class – i.e. 24 rules total. Each rule was thus repeated twice during practice and twice in the main experiment. Pairs of shapes were selected pseudorandomly in each block so that 8 trials were targets and 8 were nontargets, but no combinations of coloured shapes were repeated.

Cues. Four symbolic cues were randomly assigned to the four rule classes for each participant. Half of the total 48 blocks were designated ‘novel cues’ blocks, and the remainder were ‘familiar cues’ blocks. In familiar cues blocks, the symbolic cue faithfully indicated the rule class that was relevant (but not the precise rule). Participants were fully briefed as to the meaning of the cues at the beginning of the practice session, and the cues remained unchanged during the later scanner session. For the novel cues blocks, one of the remaining four cues was chosen pseudorandomly at the start of the block (irrespective of the rule class) with the only constraint that each cue was selected an equal number of times.

Behavioural analyses. We analysed accuracy and choice (target vs. nontarget) with ANOVAs testing the influence of cue condition and trial number. We also verified the influence of the past history of ‘confirmatory’ and ‘disconfirmatory’ feedback indicating that stimulus features. For each feature (e.g. red on the left), we estimated the probability to respond target (as a proportion of

target responses) as a function of the dimension (relevant, irrelevant), the cue (familiar, novel) and the number of times it previously appeared associated with a target (one, two, three, or more). Probabilities shown in Fig. 2d were averaged across features and sides (left and right).

Computational modelling. Our model updated beliefs about a space of possible hypotheses concerning the rule on each side of the screen, using a delta rule (equations described in the main text). In total, there were 12 hypothesis computed in parallel, defined by the dimension/side combination (colour-left, colour-right, shape-left, shape-right) and the feature (red/green/blue or shape/circle/triangle). Only the values corresponding to the features currently shown on the screen are updated. This space is reduced to 6 values in the familiar case (where only the relevant dimensions are taken into account). The model chose target if $TV \geq 0$, and nontarget otherwise.

Fittings of the model were done using an exhaustive search through a 101x101 grid with uniformly distributed values over between 0 and 1 with a step of 0.01 for both parameters α and τ . The best-fitting model was deemed to be that which minimized the difference to choice and performance:

$$\epsilon = 0.5 * (SSE_{\text{choice}} + SSE_{\text{performance}}) \quad (S2)$$

where SSE_{choice} and $SSE_{\text{performance}}$ are the sum of squared errors of predictions on choice and performance respectively.

fMRI Data Acquisition and Preprocessing. Magnetic resonance images were acquired with a 3T Siemens VERIO scanner with a 32-channel head coil using a standard echo-planar imaging sequence. Whole-head T_2^* -weighted echo-planar images were continuously acquired with a repetition time of 2 s, echo time of 30 ms. We acquired 510 volumes per block, plus 3 dummy scans discarded before the analyses. Each volume included $64 \times 64 \times 36$ voxels of $3 \times 3 \times 3$ mm. A high-resolution T_1 -weighted structural image was also obtained (voxel size = $1 \times 1 \times 1$ mm). For standard preprocessing and univariate statistical analyses, we used SPM8 (Wellcome Department of Cognitive Neurology, London, United Kingdom). All other analyses were done with custom scripts for Matlab (Mathworks, Natick, MA, United States of America). We also used xjview (<http://www.alivelearn.net/xjview>) to visualize the data and to construct mask and conjunction images. For each participant, we first realigned all functional images, then we co-registered (rigid body transformation) the subject's anatomical scan to the mean functional image, and then co-registered the participant's data to the Montreal Neurological Institute (MNI) template brain. We then normalized each subject's data to the template brain space, using segmented probabilistic maps for grey matter, white matter, and cerebro-spinal fluid. Functional images were resampled ($4 \times 4 \times 4$ mm voxels) and spatially smoothed (8-mm full-width half-maximum (FWHM) Gaussian kernel).

fMRI analysis. Our univariate analyses used a generalized linear model (GLM) approach. A 128-s temporal high-pass filter was applied to remove low-frequency scanner artifacts. Temporal autocorrelation in the time series data was estimated using restricted maximum-likelihood estimates of variance components using a first-order autoregressive model (AR-1), and the resulting non-sphericity was used to form maximum-likelihood estimates of the activations. Our GLM included regressors coding for onsets and durations of stimuli or events, which were then convolved with the canonical hemodynamic response function (HRF) and regressed against the observed fMRI data. Experimental blocks were modelled using separate regressors, and constant

terms for each block were included. Additionally, motion parameters were included as nuisance variables.

For the model-based analyses, we constructed a design matrix with 4 regressors: the expected value (i.e. $|TV|$) predicted by the best-fitting parameterisation of the model, that predicted by the performance-maximising parameterisation, the maximum hypothesis value across the all rules [$\max(HV)$] and the minimum hypothesis value across all rules [$\min(HV)$]. For the standard analyses, we included a design matrix with regressors aligned to stimulus onset, encoding the main effect of stimulus (target vs. nontarget), cues (novel vs. familiar) and feedback (correct vs. error), alongside their two- and three-way interactions. We report voxels that responded to these regressors at thresholds that were corrected for multiple comparisons, using a false discovery rate of $p < 0.05$. Voxelwise statistics were rendered onto the MNI template brain using xjview (<http://www.alivelearn.net/xjview8/>). Subsequently, to interrogate the complex interactions among factors, we extracted raw data from cluster-corrected regions of interest (ROIs) and resubmitted them to a new finite impulse response (FIR) regression analysis that estimated the parameter estimates associated with each of 16 peri-stimulus time bins. Haemodynamic response data were temporally upsampled by a factor of 10 for plotting. Statistics are reported for an early time window (~6-8s post-stimulus) and a late time window (~8-10s post-feedback) where the BOLD signal peaked overall. All ANOVAs were carried out with Greenhouse-Geisser correction for sphericity (reporting adjusted d.f. rounded to the nearest integer) using an alpha of $p < 0.05$.

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SUPPLEMENTARY MATERIALS

Supplementary section 1 - Hierarchical Bayesian model.

Notation

For simplicity will note $P(t)$, $P(x)$ and $P(r)$ the probabilities $P(T=1)$, $P(X=x)$ and $P(R=r)$ respectively, where T is the random binary variable describing if any given trial is target ($T=1$) or nontarget ($T=0$); R is a random variable associated to the block rule; and X is a variable associated with the stimulus set presented in a trial. A subscript $P(t_n)$, $P(x_n)$ will be used to refer to this probability for trial n . We will use **bold** notation $\mathbf{x}_n=(x_1...x_n)$ and $\mathbf{t}_n=(t_1...t_n)$ to refer to the history of previous trials (stimulus and target respectively). The same notation will be used to refer to conditional probabilities.

Assumptions

Our hierarchical Bayesian model (HBM) is built under a few simple assumptions. Like humans, the model is given the space of possible rules and the output (target or nontarget) that each rule predicts for each pair of stimuli presented during a trial. The HBM doesn't take into account the fact that each block had 50% target trials and 50% nontarget trials (neither were humans informed of this). The HBM was simulated independently on each different block. Thus, the history of rules in previous blocks wasn't exploited to predict the current rule. Another prior was that the underlying rule r and any observation x_i were uniformly distributed across their respective spaces. The probability for the current trial of being target was thus dependent on the probability of the underlying rule and the history of trials within the block, but not the type of block (i.e., which were the relevant dimensions). In familiar blocks, rules associated with the irrelevant dimensions were discarded (see Implementation). The hierarchical Bayesian model is optimal based on the following assumptions:

- independence of block rules across blocks
- independence of the stimulus presented with the history of previous stimuli
- independence of the stimulus presented with the underlying rule
- knowledge about the space of possible rules
- knowledge about the constraint due to the block cues
- uniform distribution across rules
- uniform distribution across observations

Structure

The HBM uses two layers: the target layer and the rule layer.

The first «target» layer estimates the probability for a certain trial of being target: $P(t_n|\mathbf{x}_n,\mathbf{t}_{n-1})$. This is calculated as the marginal probability across all the possible rules:

$$P(t_n|\mathbf{x}_n,\mathbf{t}_{n-1}) = \sum_r \{ P(t_n|r,x_n) P(r|\mathbf{x}_{n-1},\mathbf{t}_{n-1}) \} \quad (\text{eq. SE1})$$

where \sum_r is the sum across all possible rules (i.e., the 6x6 combinations of features between left and right sides). Note that the target probability $P(t_n|r,x_n)$ was only dependent on the underlying rule and the current observation, while the probability of a certain rule $P(r|\mathbf{x}_{n-1},\mathbf{t}_{n-1})$ depended on the history of previous trials.

The second «rule layer» estimates, from the history of previous trials, the probability for each rule of being the underlying one: $P(r_n | \mathbf{x}_{n-1}, \mathbf{t}_{n-1})$. For instance, it can be shown that

$$\begin{aligned}
 P(r | \mathbf{x}_{n-1}, \mathbf{t}_{n-1}) &= P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) / \sum_r \{ P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) \} \\
 &= P(r) P(\mathbf{x}_{n-1}, \mathbf{t}_{n-1} | r) / P(\mathbf{x}_{n-1}, \mathbf{t}_{n-1}) \\
 &= P(r) P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) / P(\mathbf{t}_{n-1} | \mathbf{x}_{n-1}) \\
 &= P(r) P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) / \sum_r \{ P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) P(r | \mathbf{x}_{n-1}) \} \\
 &= P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) / \sum_r \{ P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) \} \\
 &= \prod_{i=1 \dots n-1} \{ P(t_i | r, x_i) \} / \sum_r \{ P(\mathbf{t}_{n-1} | r, \mathbf{x}_{n-1}) \} \quad (\text{eq. SE2})
 \end{aligned}$$

where $\prod_{i=1 \dots n-1}$ is the product operator, and under the additional assumptions that \mathbf{x}_n and r are independent, and x_i is independent across different trials i .

This probability is 0 if the rule is inconsistent with the history of previous trials, and is equally distributed across all remaining (consistent) rules. For example, if four possible rules were consistent with the history of trials, the probability for each of these would be of 0.25, while 0 for any other rule.

This model is shown to be optimal based on the previous priors.

Implementation

We created an instance of this model by keeping track of all possible rules in a 6*6 binary matrix – called the *candidates matrix* **C**. This matrix allowed us to keep track of the rules that were consistent in previous trials. Each cell (i,j) in this matrix was associated with each possible rule on each side, where $i,j \in \{1 \dots 6\}$ corresponded to the feature on the left and right sides respectively (e.g., red, green, blue, square, triangle, circle).

The *candidates matrix* **C** was used in the following way. If $C_{ij}=1$, then the rule (i,j) was consistent with the history of previous trials – and was a candidate for the underlying rule. If $C_{ij}=0$, then the rule (i,j) was inconsistent with the history of previous trials and could not candidate as the underlying rule for the current block.

The *candidates matrix* corresponds to the «rule layer» in our hierarchical Bayesian model.

At the beginning of the block, all cells in the candidates matrix were set to 1, meaning that all rules were possibly the underlying rule for the current block. In familiar blocks, the 27 rules corresponding to the irrelevant dimensions were set to 0, leaving only 9 possible rules left.

On each trial and for each candidate rule (i.e., for each rule $r=(i,j)$ such that $C_{ij}=1$), we calculated the conditional probability of being target: $P(t_n | r, \mathbf{x}_n)$. This probability could be 1 (i.e., the current trial is target) or 0 (trial is nontarget) following the prediction given by each rule. From equation (SE1), the probability of a target trial, $P(t_n | \mathbf{x}_n, \mathbf{t}_{n-1})$, was the average across these probabilities. This calculation corresponds to the «target layer» of the hierarchical Bayesian model.

We used a greedy policy for the model, where the response was target if $P(t_n) \geq 0.5$ and nontarget in any other case.

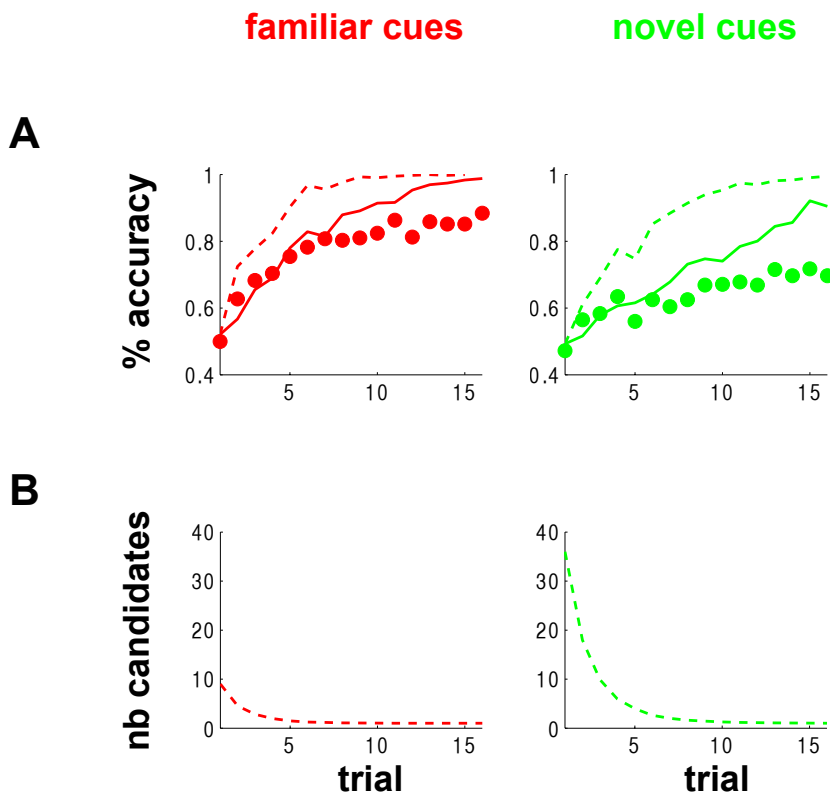


Figure SF1. A. Behavioural data from 18 human participants (dots) and predictions of the reinforcement model (continued lines) and the hierarchical Bayesian model (dotted lines). % accuracy over trials (1-16) is shown for familiar cues (left panels; red) and novel cues (right panels; green). Lines show the same data for the best-performance parametrisation of the model. **B.** Average number of candidate rules in the hierarchical Bayesian model across trials, both for familiar and novel blocks (red; green). At the beginning of the block, 9 and 36 candidate rules were used in familiar and novel blocks respectively. These values decreased to 1 by the end of the block.

As can be seen in Fig SF1 panel A, the model largely outperforms both the reinforcement model presented in equations (1–3) and % accuracy achieved by humans, by inferring the probability for each trial of being target in a Bayesian fashion. The number of consistent rules explaining the history of previous trials (see Fig SF1, panel B) decreases from 9 and 36 candidates (familiar and novel blocks, respectively) to 1, when the model can conclude with complete certainty what is the underlying rule for that block. This model thus allows us to estimate an upper boundary on the accuracy of responses given by either human behaviour or any other model predictions.

Figure S2. BOLD responses to the 3-way interaction between stimulus, cue and feedback.

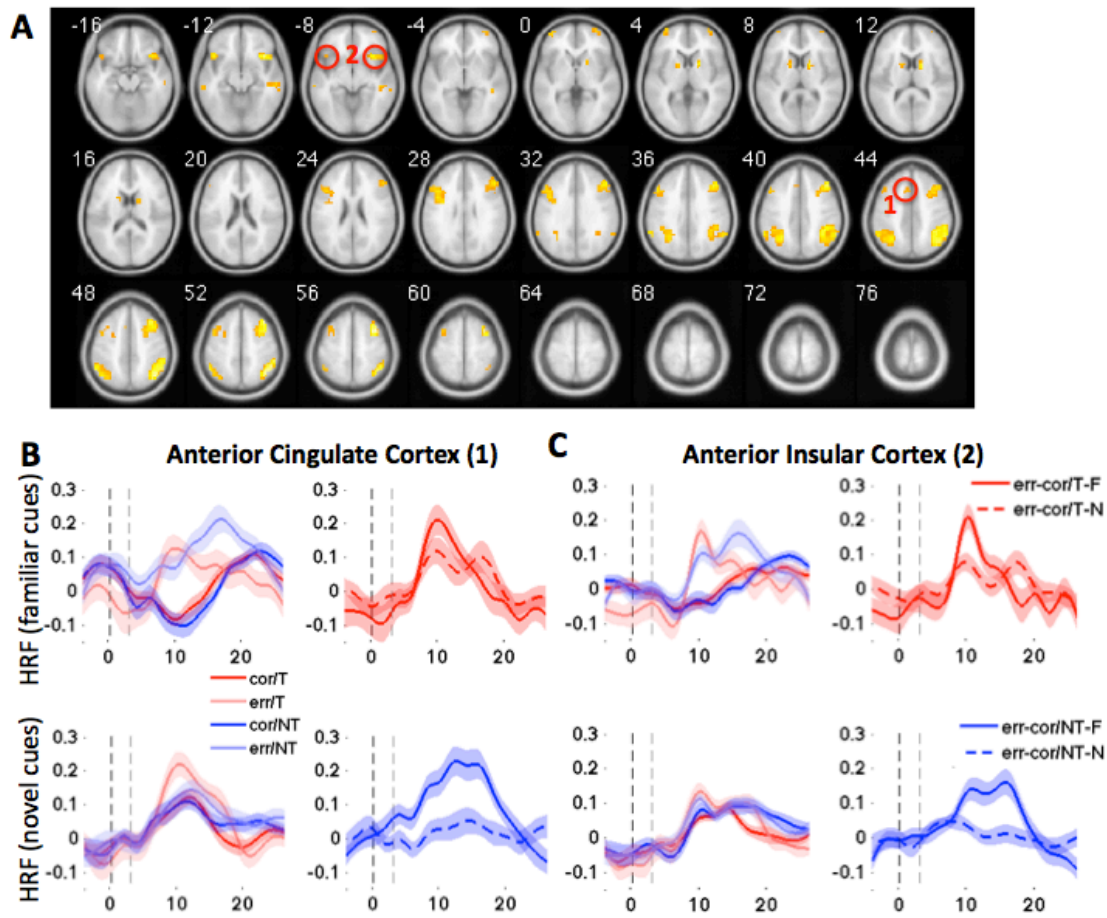


Figure S2. A. All voxels responding to the three-way interaction between stimulus, cue and feedback, rendered onto a template brain at a threshold of $p < 0.001$ uncorrected. All voxels shown survive false discovery rate correction at $p < 0.05$. Red circles and numbering indicate the locations of the dACC and AINS. **B.** Left panels: BOLD responses in the dorsal anterior cingulate cortex (ACC) to correct and error target and nontarget trials, for novel cues (bottom panel) and familiar cues (top panel) blocks. Right panels: replotting the same data as error-correct BOLD for targets (top panel) and nontargets (bottom panel). **C.** As B, but for the anterior insular cortex (AIns).

Table S1. Voxels sensitive to three-way interaction between stimulus, cue and feedback.

STATISTICS: p-values adjusted for search volume

cluster	voxel	voxel	voxel	voxel	voxel	x,y,z {mm}			
p(cor)	k	p(FDR)	T	Z	p(unc)				
0.000	395	0.000	10.03	5.66	0.000	38	8	58	right DLPFC
		0.000	8.84	5.34	0.000	38	24	54	
		0.000	7.96	5.07	0.000	46	28	42	
0.000	329	0.000	9.05	5.40	0.000	42	-64	50	right IPL
		0.000	8.89	5.36	0.000	50	-48	54	
		0.000	7.61	4.96	0.000	42	-44	42	
0.000	85	0.000	8.26	5.17	0.000	38	20	-10	
0.000	220	0.000	8.06	5.10	0.000	-10	-80	-26	
		0.000	6.60	4.59	0.000	18	-84	-34	
		0.001	5.58	4.15	0.000	14	-80	-26	
0.000	266	0.000	7.40	4.88	0.000	-30	-60	42	left IPL
		0.000	7.17	4.80	0.000	-54	-52	50	
		0.000	6.89	4.70	0.000	-42	-64	50	
0.000	63	0.000	7.19	4.81	0.000	18	8	10	right caudate
		0.006	4.24	3.45	0.000	10	-8	6	
0.000	341	0.000	6.62	4.59	0.000	-42	20	30	left DLPFC
		0.000	6.47	4.54	0.000	-42	8	30	
		0.001	6.43	4.52	0.000	-50	20	38	
0.003	35	0.000	6.55	4.57	0.000	-42	20	-10	
0.000	49	0.001	5.97	4.33	0.000	-38	60	2	left RLPFC
0.000	83	0.001	5.96	4.32	0.000	58	-24	-10	
		0.001	5.81	4.25	0.000	46	-28	-6	
		0.001	5.46	4.09	0.000	46	-36	-2	
0.000	90	0.001	5.92	4.30	0.000	34	60	2	right RLPFC
		0.001	5.57	4.15	0.000	50	36	-22	
		0.004	4.55	3.63	0.000	42	44	-10	
0.000	52	0.001	5.74	4.22	0.000	-14	8	10	left caudate
		0.003	4.87	3.80	0.000	-18	-4	22	
0.002	38	0.002	5.31	4.02	0.000	-62	-28	-6	
0.000	49	0.002	5.22	3.98	0.000	-10	20	46	
0.007	29	0.002	5.08	3.91	0.000	42	-72	-30	
		0.004	4.69	3.71	0.000	30	-64	-34	

Table S1. Voxels sensitive to the main effect of correct > error.

STATISTICS: p-values adjusted for search volume

cluster	voxel	voxel	voxel	voxel	voxel		
p(cor)	k	p(FDR)	T	Z	p(unc)	x,y,z {mm}	
0.000	189	0.022	6.79	4.66	0.000	14 -40 62	
		0.025	5.54	4.13	0.000	30 -44 70	
		0.031	4.54	3.63	0.000	6 -16 42	
0.000	90	0.022	6.51	4.55	0.000	-34 -88 26	
		0.022	6.05	4.36	0.000	-22 -92 30	
		0.025	5.39	4.06	0.000	-22 -96 22	
0.000	78	0.022	5.83	4.26	0.000	30 -8 2	right putamen
		0.031	4.48	3.59	0.000	38 -16 18	
0.000	102	0.025	5.51	4.12	0.000	46 -80 22	
		0.039	3.89	3.25	0.001	42 -68 6	
0.000	51	0.025	5.23	3.98	0.000	-34 0 10	
		0.033	4.33	3.51	0.000	-30 -16 10	
		0.035	4.21	3.44	0.000	-34 -8 -2	left putamen
0.006	30	0.025	5.02	3.88	0.000	-14 -44 66	
		0.037	4.07	3.35	0.000	-22 -48 58	
0.084	15	0.025	4.94	3.84	0.000	26 -40 -22	
		0.031	4.59	3.65	0.000	34 -48 -10	
0.000	73	0.027	4.83	3.78	0.000	-14 48 -2	VMPFC
		0.032	4.43	3.56	0.000	6 56 14	
		0.033	4.35	3.52	0.000	-6 56 18	
0.005	31	0.029	4.75	3.74	0.000	-18 -44 -18	
		0.032	4.39	3.54	0.000	-26 -40 -22	
		0.035	4.22	3.44	0.000	-26 -36 -10	