

Homework 2

(1)

Suppose that the random variable Y has a gamma distribution with parameters $\alpha = 2$ and an unknown β . In Exercise 6.46, you used the method of moment-generating functions to prove a general result implying that $2Y/\beta$ has a χ^2 distribution with 4 degrees of freedom (df). Using $2Y/\beta$ as a pivotal quantity, derive a 90% confidence interval for β .

$$Y \sim \Gamma(\alpha, \beta) \rightarrow \alpha = 2, \beta \text{ unknown}$$

$$U = 2Y/\beta \sim \chi^2_4 \rightarrow \beta = 2Y/U$$

$$\rightarrow Y = \frac{1}{2} \cdot \beta \cdot U$$

$$\rightarrow \left| \frac{dy}{du} \right| = \frac{1}{2} \cdot \beta$$

i. Finding $f_u(u)$: 2 methods

$$\textcircled{1} f_u(u) = f_Y(y) \cdot \left| \frac{dy}{du} \right|$$

$$= \left[\frac{1}{\Gamma(\alpha) \cdot \beta^\alpha} \right] y^{\alpha-1} e^{-y/\beta} \times \frac{1}{2} \beta$$

$$= \frac{1}{\Gamma(2) \beta} \cdot y \cdot e^{-y/\beta} \cdot \frac{1}{2}$$

$$= \frac{1}{\Gamma(2) \beta} \cdot \frac{\beta u}{2} \cdot \frac{1}{2} \cdot e^{-u/2}$$

$$= \frac{u}{4} \cdot e^{-u/2}$$

$$\textcircled{2} \text{ p.d.f. for } \chi^2_v = \frac{u^{(v/2)-1} \cdot e^{-u/2}}{2^{v/2} \cdot \Gamma(v/2)}$$

$$= \frac{u \cdot e^{-u/2}}{2^2 \Gamma(2)}$$

$$= \frac{u}{4} \cdot e^{-u/2}$$

$$\Gamma(2) = ?$$

$$\Gamma(x+1) = x \cdot \Gamma(x)$$

$$= 1 \cdot \Gamma(1)$$

$$= 1$$

$$x+1=2 \\ x=1$$

ii. We need: $P(a \leq U \leq b) = 0.9$ (i.e., $P(U < a) = P(U > b) = 0.05$)

$$\int f_u(u) \cdot du = \int \frac{u e^{-u/2}}{4} \cdot du$$

$$= \int \frac{1}{4} \cdot (-2v) \cdot e^v \cdot (-2) dv$$

$$= \int v \cdot e^v \cdot dv$$

$$= v \cdot e^v - e^v \Big|_?$$

$$= -\frac{u}{2} \cdot e^{-u/2} - e^{-u/2} \Big|_?$$

It will be hard to solve for a and b directly from this

instead, from table:

$$\chi^2_{0.05, 4} = 9.48773$$

$$\chi^2_{0.95, 4} = 0.710721$$

$$(0.710721 \leq 2Y/\beta \leq 9.48773)$$

$$(2Y/9.4877 \leq \beta \leq 2Y/0.7107)$$

(2)

The administrators for a hospital wished to estimate the average number of days required for inpatient treatment of patients between the ages of 25 and 34. A random sample of 500 hospital patients between these ages produced a mean and standard deviation equal to 5.4 and 3.1 days, respectively. Construct a 95% confidence interval for the mean length of stay for the population of patients from which the sample was drawn.

$n = 500$ hospitals — large enough

$\bar{x} = 5.4$ days

$s = 3.1$ days

$\alpha = 0.05, \alpha/2 = 0.025$

$z_{\alpha/2} = 1.96$

$$-1.96 \leq \frac{5.4 - \mu}{\frac{3.1}{\sqrt{500}}} \leq 1.96$$

$$(-1.96) \frac{3.1}{\sqrt{500}} - 5.4 \leq -\mu \leq 1.96 \frac{3.1}{\sqrt{500}} - 5.4$$

$$(1.96) \frac{3.1}{\sqrt{500}} + 5.4 \geq -\mu \geq -1.96 \frac{3.1}{\sqrt{500}} + 5.4$$

$$(-1.96) \frac{3.1}{\sqrt{500}} + 5.4 \leq \mu \leq 1.96 \frac{3.1}{\sqrt{500}} + 5.4$$

$$(5.1283, 5.6717)$$

(3)

For a comparison of the rates of defectives produced by two assembly lines, independent random samples of 100 items were selected from each line. Line A yielded 18 defectives in the sample, and line B yielded 12 defectives.

- a Find a 98% confidence interval for the true difference in proportions of defectives for the two lines.
- b Is there evidence here to suggest that one line produces a higher proportion of defectives than the other?

$$\begin{array}{ll} n_A = 100 & p_A = 18/100 = 0.18 \\ n_B = 100 & p_B = 12/100 = 0.12 \end{array}$$

$$\begin{array}{ll} (a) \quad d = 0.02 & s_A = (\hat{p}_A \hat{q}_A) / n_A = (0.18)(0.82) / 100 = 0.001476 \\ \alpha/2 = 0.01 & s_B = (\hat{p}_B \hat{q}_B) / n_B = (0.12)(0.88) / 100 = 0.001056 \end{array}$$

$$-z_{0.01} \leq \frac{(\hat{p}_A - \hat{p}_B) - (p_A - p_B)}{\sqrt{(\hat{p}_A \hat{q}_A) / n_A + (\hat{p}_B \hat{q}_B) / n_B}} \leq z_{0.01}$$

$$-2.326 \leq \frac{(0.18 - 0.12) - (p_A - p_B)}{\sqrt{0.001476 + 0.001056}} \leq 2.326$$

$$C.I.: \quad 0.06 \pm (2.326)(0.050319)$$

$$(-0.057, 0.177)$$

- (b) Since the confidence interval includes zero, we are unable to rule out the possibility that the difference is zero.

No, there is no evidence.

(4)

Problem 4 We sample n observations Y_1, \dots, Y_n from a uniform distribution on $[0, \theta]$, and θ is unknown. Use the pivotal method to find a 90% lower confidence bound for θ .

pivotal quantity: $u = \frac{Y_{(n)}}{\theta} \rightarrow \theta = Y_{(n)}/u, Y_{(n)} = \theta \cdot u, \text{ and } \left| \frac{dy}{du} \right| = \theta$

$$\begin{aligned} F(Y_{(n)}) &= P(Y_{(n)} \leq y) \\ &= P(Y_1 \leq y, \dots, Y_n \leq y) \\ &= \left(\frac{y}{\theta}\right)^n \end{aligned}$$

$$f_{Y_{(n)}}(y) = \begin{cases} 1 & y > \theta \\ n \cdot \left(\frac{y}{\theta}\right)^{n-1} \cdot \frac{1}{\theta} & 0 \leq y \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

$$f_u(u) = f_{Y_{(n)}}(y) \cdot \left| \frac{dy}{du} \right|$$

$$= \begin{cases} 1 & \theta u > \theta \\ n \cdot \left(\frac{y}{\theta}\right)^{n-1} \cdot \frac{1}{\theta} \cdot \theta & 0 \leq \theta \cdot u \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} 1 & u > 1 \\ n u^{n-1} & 0 \leq u \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$P(\theta \geq L) = 0.9$$

equivalent: $P(Y_{(n)}/u \geq L)$

$$= P(u \leq Y_{(n)}/L)$$

$$= \int_0^{Y_{(n)}/L} n \cdot u^{n-1} du$$

$$= u^n \cdot \theta \Big|_0^{Y_{(n)}/L}$$

$$= \left(\frac{Y_{(n)}}{L}\right)^n = 0.9$$

$$Y_{(n)}/L = 0.9^{1/n}$$

$$L = \boxed{Y_{(n)}/0.9^{1/n}}$$

(5) **Problem 5** We sample one observation from a normal population $N(0, \sigma^2)$ and σ^2 is unknown.

- 1) Find a 95% confidence interval for σ^2 .
- 2) Find a 95% confidence interval for σ .

$$W = \frac{(n-1) \cdot s^2}{\sigma^2} \sim \chi^2_{n-1}$$

$$\begin{aligned} P\left(\chi^2_{n-1, 1-\alpha/2} \leq W \leq \chi^2_{n-1, \alpha}\right) &= P\left(\chi^2_{n-1, 1-\alpha/2} \leq \frac{(n-1) \cdot s^2}{\sigma^2} \leq \chi^2_{n-1, \alpha}\right) \\ &= P\left(\frac{1}{\chi^2_{n-1, 1-\alpha/2}} \geq \frac{\sigma^2}{(n-1) \cdot s^2} \geq \frac{1}{\chi^2_{n-1, \alpha}}\right) \\ &= P\left(\frac{(n-1) \cdot s^2}{\chi^2_{n-1, \alpha/2}} \leq \sigma^2 \leq \frac{(n-1) \cdot s^2}{\chi^2_{n-1, 1-\alpha/2}}\right) \\ &= P\left(\frac{(1-1) \cdot s^2}{\chi^2_{n-1, \alpha/2}} \leq \sigma^2 \leq \frac{(1-1) \cdot s^2}{\chi^2_{n-1, 1-\alpha/2}}\right) \\ &= P(0 \leq \sigma^2 \leq 0) = P(\sqrt{0} \leq \sqrt{\sigma^2} \leq \sqrt{0}) = P(0 \leq \sigma \leq 0) = 0.95 \end{aligned}$$

(1) (0,0)

(2) (0,0)

(6) **Problem 6** Suppose that we are to obtain a single observation Y from an exponential distribution with mean θ . Use the single observation to form a 90% confidence interval for θ by using two different pivotal quantities:

- 1) Use Y/θ as a pivotal quantity.
- 2) Use $2Y/\theta$ as a pivotal quantity.

$$Y_1, \dots, Y_n \sim \text{exponential}(\lambda = \frac{1}{\theta})$$

$$\begin{aligned} u &= cY/\theta \\ \theta &= cY/u \\ Y &= u \cdot \theta / c \\ \left| \frac{dy}{du} \right| &= \theta / c \end{aligned}$$

C.I: Find A and B such that $(A \leq u \leq B) = 0.9$

$$\begin{aligned} f_Y(y) &= \lambda e^{-\lambda y} \\ &= \frac{1}{\theta} e^{-y/\theta} \end{aligned}$$

$$\begin{aligned} \therefore P(u > B) &= P(u < A) = 0.05 \\ \downarrow & \quad \quad \downarrow \\ &= 1 - F_Y(B) \quad = F_Y(A) \end{aligned}$$

$$\begin{aligned} f_u(u) &= f_Y(y) \cdot \left| \frac{dy}{du} \right| \\ &= \frac{1}{\theta} e^{-(u\theta/c)/\theta} \times \left| \frac{\theta}{c} \right| \\ &= e^{-u/c} \times \left| \frac{1}{c} \right| \end{aligned}$$

$$\begin{aligned} F_u(u) &= \int_A^B e^{-u/c} \times \left| \frac{1}{c} \right| du \\ &= -e^{-u/c} \Big|_A^B \end{aligned}$$

$$\begin{aligned} (1) \quad P(u > B) &= 1 - F_Y(B) \\ &= 1 - \left[-e^{-u/\theta} \Big|_0^B \right] \\ &= 1 - \left(-e^{-B/\theta} - (-1) \right) \\ &= e^{-B/\theta} = 0.05 \\ \hookrightarrow B &= -\ln(0.05) = 2.996 \end{aligned}$$

$$\begin{aligned} P(u < A) &= F_Y(A) \\ &= -e^{-u/\theta} \Big|_0^A \\ &= -e^{-A/\theta} - (-1) \\ &= -e^{-A/\theta} + 1 = 0.05 \\ \hookrightarrow e^{-A/\theta} &= 0.95 \\ \hookrightarrow A &= -\ln(0.95) = 0.0513 \end{aligned}$$

$$\begin{aligned} 0.0513 &\leq u \leq 2.996 \\ 0.0513 &\leq Y/\theta \leq 2.996 \\ \frac{Y}{2.996} &\leq \theta \leq \frac{Y}{0.0513} \end{aligned}$$

$$\left(Y/2.996, Y/0.0513 \right)$$

$$\begin{aligned} (2) \quad P(u > B) &= 1 - F_Y(B) \\ &= 1 - \left[-e^{-u/2\theta} \Big|_0^B \right] \\ &= 1 - \left(-e^{-B/2\theta} + 1 \right) \\ \hookrightarrow e^{-B/2\theta} &= 0.05 \\ \hookrightarrow B/2 &= -\ln(0.05) = 2.996 \\ \hookrightarrow B &= 2 \times 2.996 \end{aligned}$$

$$\begin{aligned} P(u < A) &= F_Y(A) \\ &= -e^{-u/2\theta} \Big|_0^A \\ &= -e^{-A/2\theta} - (-1) \\ &= -e^{-A/2\theta} + 1 = 0.05 \\ e^{-A/2\theta} &= 0.95 \\ A/2 &= -\ln(0.95) \\ A &= 2 \times 0.0513 \end{aligned}$$

$$\begin{aligned} 2 \times 0.0513 &\leq u \leq 2 \times 2.996 \\ 2 \times 0.0513 &\leq 2Y/\theta \leq 2 \times 2.996 \\ \frac{Y}{2.996} &\leq \theta \leq \frac{Y}{0.0513} \end{aligned}$$

$$\left(Y/2.996, Y/0.0513 \right)$$