

# A short introduction on how to fit a **SPDE** model with **INLA**

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This document illustrates how to do a geostatistical fully Bayesian analysis through the **Stochastic Partial Differential Equation** approach, [Lindgren et al., 2011], with **Integrated Nested Laplace Approximation**, [Rue et al., 2009], using the **INLA** package, <http://www.r-inla.org>.

## 1 Data simulation

**Locations** and the Random Field (RF) **covariance** matrix, exponential correlation function

```
n = 200; coo = matrix(runif(2*n), n)
k <- 10; s2s <- 0.7 ## RF params.
R <- s2s*exp(-k*as.matrix(dist(coo)))
```

**RF** sample, a multivariate Normal realization

```
s <- drop(rnorm(n)%*%chol(R))
```

A **covariate** effect and a noise can be added

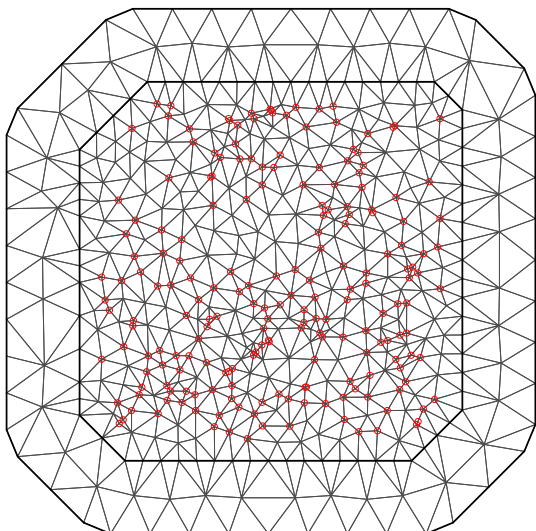
```
x <- runif(n); beta <- 1:2; s2e <- 0.3
lin.pred <- beta[1] + beta[2]*x + s
y <- lin.pred + rnorm(n, 0, sqrt(s2e))
```

## 2 Model fitting: steps

1. **Mesh**: a triangulation to discretize the random field (RF) at  $m$  nodes.

```
mesh <- inla.mesh.2d(
  coo, ## provide locations or domain
  max.edge=c(1/k, 2/k), ## mandatory
  cutoff=0.1/k) ## good to have >0
plot(mesh, asp=1); points(coo, col='red')
```

Constrained refined Delaunay triangulation



A little warning about the mesh. The additional triangles outer domain is to avoid boundary effects. Is good to have approximately isosceles triangles. And, to avoid tiny triangles. We need to have edges lengths of the inner mesh triangles less than the range of the process. Of course, if it is too small, there might not be any spatial effect.

2. Define the  $n \times m$  **projector matrix** to project the process at the mesh nodes to locations

```
dim(A <- inla.spde.make.A(
  mesh=mesh, loc=coo)) ## 'n' by 'm'
## [1] 200 505
```

3. **Build the SPDE model** on the mesh. Exponential correlation function,  $\alpha = 3/2$

```
spde <- inla.spde2.matern(
  mesh=mesh, alpha=1.5)
```

4. **Create a stack data** for the estimation. This is a way to allow models with complex linear predictors. In our case, we have a SPDE model defined on  $m$  nodes. It must be combined with the covariate (and the intercept) effect at  $n$  locations. We do it using different projector matrices.

```
stk.e <- inla.stack(tag='est', ## tag
  data=list(y=y), ## response
  A=list(A, 1), ## two projector matrices
  effects=list(## two elements:
    s=1:spde$n.spde, ## RF index
    data.frame(b0=1, x=x)))
```

5. **Fit** the posterior marginal distributions for all model parameters

```
formula <- y ~ 0 + b0 + x + ## fixed part
  f(s, model=spde) ## RF term
res <- inla(formula,
  data=inla.stack.data(stk.e),
  control.predictor=list(A =
    inla.stack.A(stk.e)) # projector
```

### 3 Posterior marginal distributions - PMDs

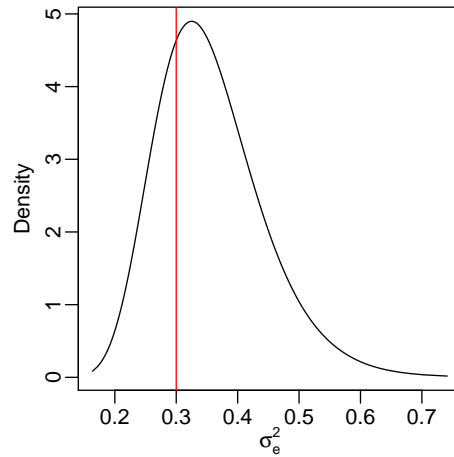
Summary of the regression coefficients PMDs

```
round(res$summary.fixed, 4)
```

##	mean	sd	0.025quant	0.5quant	0.975quant	mode	kld
## b0	1.041	0.2928	0.4223	1.050	1.607	1.064	0
## x	1.903	0.1895	1.5286	1.903	2.273	1.905	0

**INLA** works with precisions. We have to transform the precision PMD to have the variance PMD. It can be done and visualized by

```
post.s2e <-
  inla.tmarginal(# transformation function
    function(x) 1/x, ## inverse transf.
    res$marginals.hyperpar$
    'Precision for the Gaussian observations')
plot(post.s2e, type='l', ylab='Density',
      xlab=expression(sigma[e]^2))
abline(v=s2e, col=2) ## add true value
```

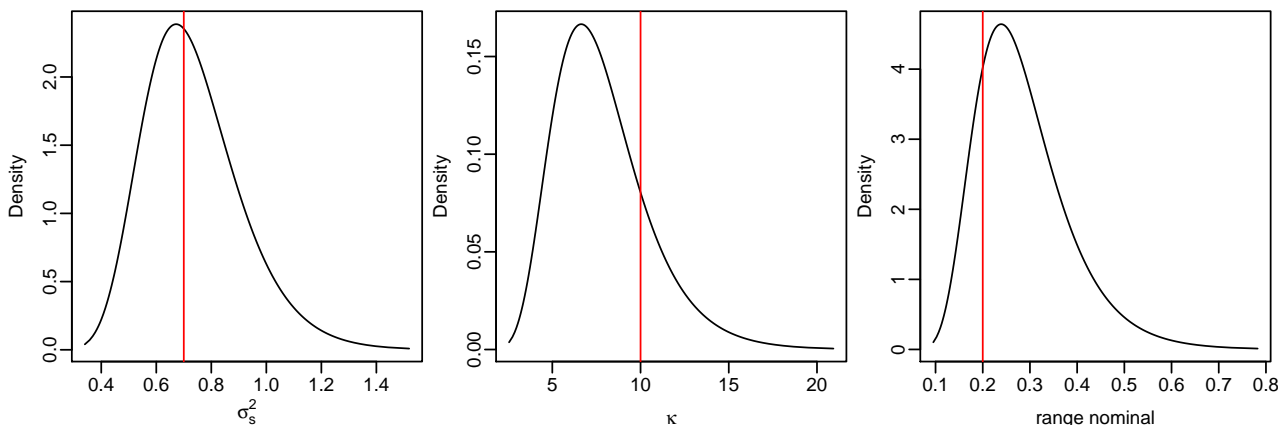


The SPDE approach uses a local variance,  $\tau^2$ , such that  $\sigma_s^2 = 1/(2\pi\kappa^2\tau^2)$ . On **INLA** we work  $\log(\tau^2)$  and  $\log(\kappa)$ . So, especially for  $\sigma_s^2$ , we have to do an additional computation. The PMDs for all RF parameters on user scale are computed by

```
rf <- inla.spde.result(
  inla=res, ## the inla() output
  name='s', ## name of RF index set
  spde=spde, ## SPDE model object
  do.transf=TRUE) ## to user scale
```

It can be visualized by

```
plot(rf$marginals.var[[1]], ty = "l", xlab = expression(sigma[s]^2), ylab = "Density")
abline(v = s2s, col = 2) ## add the true value
plot(rf$marginals.kap[[1]], type = "l", xlab = expression(kappa), ylab = "Density")
abline(v = k, col = 2) ## add the true value
plot(rf$marginals.range[[1]], type = "l", xlab = "range nominal", ylab = "Density")
abline(v = sqrt(4)/k, col = 2) ## add the 'true' value
```



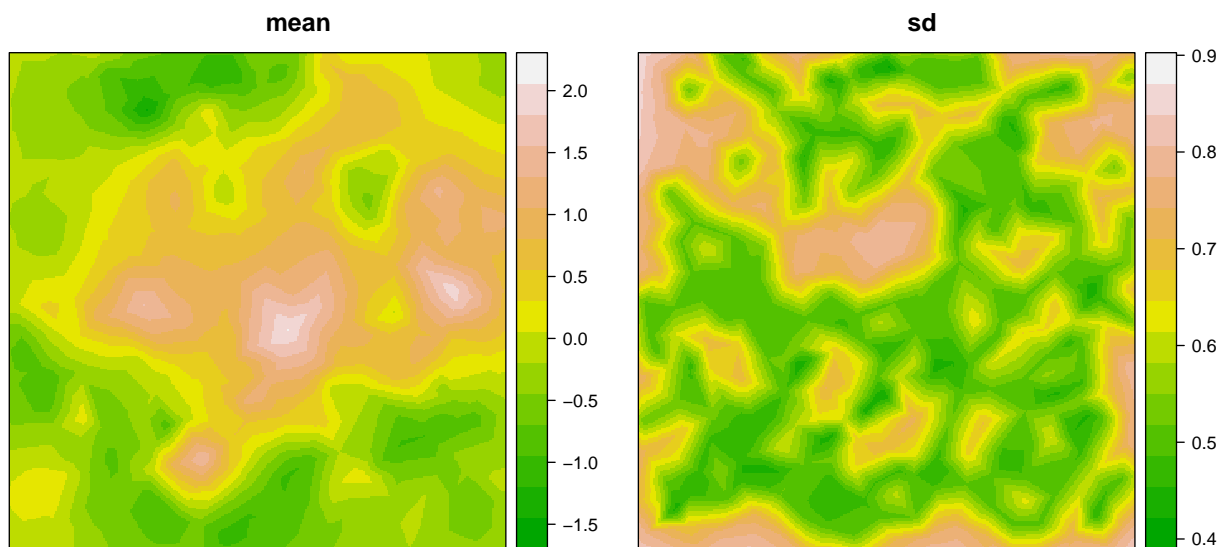
## 4 Projection on a grid

An interesting result is the map of the RF on a grid. The simplest way to have it is by projection. We just have to define the projector matrix and project, for example, the posterior mean and posterior standard deviation on the grid.

```
gproj <- inla.mesh.projector(mesh, xlim = 0:1, ylim = 0:1, dims = c(300, 300))
g.mean <- inla.mesh.project(gproj, res$summary.random$s$mean)
g.sd <- inla.mesh.project(gproj, res$summary.random$s$sd)
```

We can visualize it by

```
library(lattice); library(gridExtra)
trellis.par.set(regions=list(col=terrain.colors(16)))
grid.arrange(levelplot(g.mean, scales=list(draw=F), xlab='', ylab='', main='mean'),
              levelplot(g.sd, scal=list(draw=F), xla='', yla='', main='sd'), nrow=1)
```



## 5 Prediction

Define target locations, the corresponding projector matrix and covariate values at target locations

```
tcoo <- rbind(c(0.3, 0.3), c(0.5, 0.5), c(0.7, 0.7))
dim(Ap <- inla.spde.make.A(mesh = mesh, loc = tcoo))

## [1] 3 505

x0 <- c(0.5, 0.5, 0.5)
```

To do a fully Bayesian analysis, we include the target locations on the estimation process by assigning NA for the response at these locations. Defining the prediction stack

```
stk.pred <- inla.stack(tag='pred', A=list(Ap, 1), data=list(y=NA), ## response as NA
                      effects=list(s=1:spde$n.spde, data.frame(x=x0, b0=1)))
```

Fit the model again with the full stack

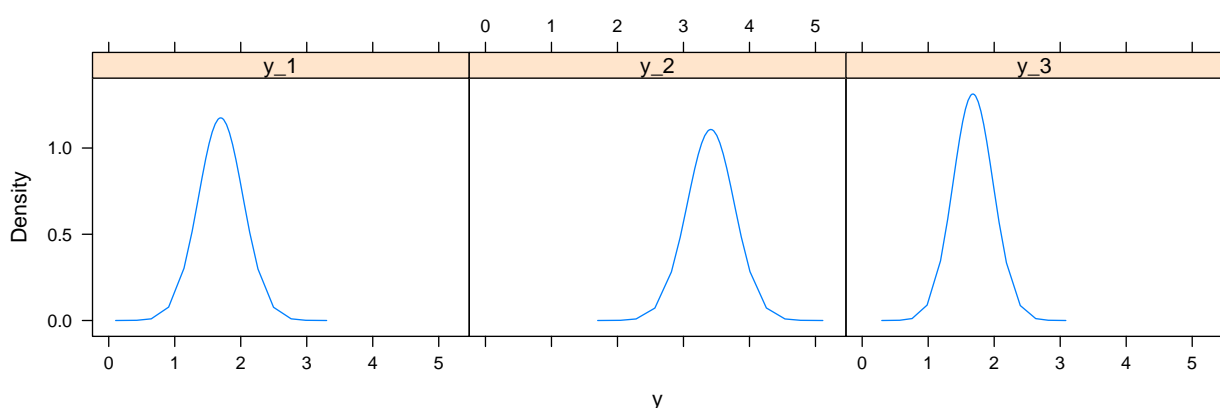
```
stk.full <- inla.stack(stk.e, stk.pred)
p.res <- inla(formula, data=inla.stack.data(stk.full), ## full stack
              control.predictor=list(compute=TRUE, ## compute the predictor
                                    A=inla.stack.A(stk.full))) ## using full stack data
```

Get the prediction data index and collect the linear predictor PMDs to work with

```
pred.ind <- inla.stack.index(stk.full, tag = "pred")$data
ypost <- p.res$marginals.fitted.values[pred.ind]
```

Visualize with commands bellow

```
names(ypost) <- paste('y', seq_along(ypost), sep='_'); library(plyr)
xyplot(y~x | .id, ldply(ypost), panel='llines', xlab='y', ylab='Density')
```



In **INLA** we have some functions to work with marginals distributions

```
apropos("marginal")

## [1] "inla.dmarginal"    "inla.emarginal"    "inla.hpdmarginal"
## [4] "inla.mmarginal"    "inla.pmarginal"    "inla.qmarginal"
## [7] "inla.rmarginal"    "inla.smarginal"    "inla.tmarginal"
```

```
inla.mmarginal(ypost[[1]]) ## mode

## [1] 1.697

inla.qmarginal(c(0.15, 0.7), ## quantiles
               ypost[[1]])

## [1] 1.345 1.875

inla.pmarginal(inla.qmarginal(
  0.3, ypost[[1]]), ypost[[1]])

## [1] 0.3
```

## References

- [Lindgren et al., 2011] Lindgren, F., Rue, H., and Lindström, J. (2011). An explicit link between gaussian fields and gaussian markov random fields: the stochastic partial differential equation approach (with discussion). *J. R. Statist. Soc. B*, 73(4):423–498.
- [Rue et al., 2009] Rue, H., Martino, S., and Chopin, N. (2009). Approximate bayesian inference for latent gaussian models using integrated nested laplace approximations (with discussion). *Journal of the Royal Statistical Society, Series B*, 71(2):319–392.