# Eco 5316 Time Series Econometrics

Lecture 6 Forecasting

#### Forecasting

three main components needed to produce a forecast

- ▶ information set  $\mathcal{I}_t = \{y_0, y_1, \dots, y_t\}$  at forecast origin t
- ▶ forecast horizon h
- ▶ loss function  $L(y_{t+h} f_{t,h})$  or  $L(e_{t,h})$

where  $f_{t,h}$  is the h-step ahead forecast at forecast origin t given information set  $\mathcal{I}_t$  and  $e_{t,h}=y_{t+h}-f_{t,h}$  is the forecast error

 $\mbox{\bf optimal forecast}:$  forecaster wants to construct a forecast  $f_{t+h}^*$  that minimizes the expected loss

$$E[L(y_{t+h}-f_{t,h})|\mathcal{I}_t] = \int L(y_{t+h}-f_{t,h})f(y_{t+h}|\mathcal{I}_t)dy_{t+h}$$

thus

$$f_{t,h}^* = \arg\min_{f_{t,h}} E\left[L(y_{t+h} - f_{t,h})|\mathcal{I}_t\right]$$

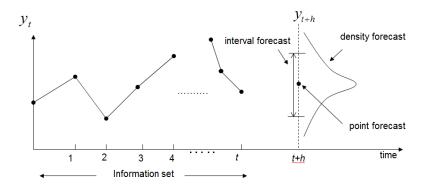
# Point, Interval and Density Forecasts

first, we need conditional distribution and moments for  $y_{t+h}$  given information set  $\mathcal{I}_t$ 

- conditional probability density function  $f(y_{t+h}|\mathcal{I}_t)$
- $\begin{array}{l} \blacktriangleright \text{ conditional mean } \mu_{t+h|t} = E_t(y_{t+h}|\mathcal{I}_t) \\ \blacktriangleright \text{ conditional variance } \sigma^2_{t+h|t} = var_t(y_{t+h}|\mathcal{I}_t) \end{array}$

these will be used to build the point, interval and density forecasts

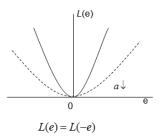
# Point, Interval and Density Forecasts



# Symmetric Loss Function

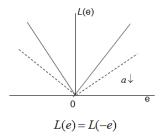
#### Quadratic loss function

$$L(e) = ae^2$$
,  $a > 0$ 



#### Absolute value loss function

$$L(e) = a |e|, \quad a > 0$$



# Point, Interval and Density Forecasts

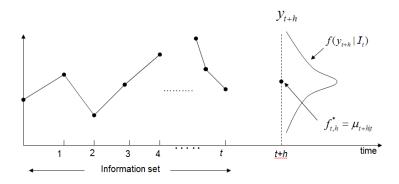
suppose that conditional density  $f(y_{t+h}|\mathcal{I}_t)$  is  $N(\mu_{t+h|t}, \sigma_{t+h|t}^2)$  then density forecast is  $N(\mu_{t+h|t}, \sigma_{t+h|t}^2)$  and

- 1. if loss function is quadratic  $L(e_{t,h}) = ae_{t,h}^2$
- optimal point forecast is  $f_{t+h}^* = \mu_{t+h|t}$
- ▶ 95% interval forecast is  $\mu_{t+h|t} \pm 1.96\sigma_{t+h|t}$

- 2. if loss function is absolute value  $L(e_{t,h}) = a|e_{t,h}|$
- lacktriangle optimal point forecast is the conditional median  $f_{t+h}^* = median(y_{t+h}|\mathcal{I}_t)$

note: if  $f(y_{t+h}|\mathcal{I}_t)$  is symmetric then mean and median coincide

# Quadratic Loss Function



suppose that  $y_t$  follows an AR(1) model  $y_t = \phi_0 + \phi_1 y_{t-1} + \varepsilon_t$  with  $\varepsilon_t \sim N(0, \sigma_\varepsilon^2)$  and that  $L(e_{t,h}) = a e_{t,h}^2$  then:

for conditional mean we have

$$\mu_{t+1|t} = E_t(y_{t+1}|\mathcal{I}_t) = \phi_0 + \phi_1 y_t$$

for conditional variance

$$\sigma_{t+1|t}^2 = var_t(y_{t+1}|\mathcal{I}_t) = var(\varepsilon_{t+1}) = \sigma_{\varepsilon}^2$$

▶ thus the 1 step ahead point forecast of  $y_{t+1}$  is

$$f_{t,1} = \mu_{t+1|t} = \phi_0 + \phi_1 y_t$$

- ▶ the conditional density forecast for  $y_{t+1}$  is  $N(\phi_0 + \phi_1 y_t, \sigma_{\varepsilon}^2)$
- ▶ the 95% interval forecast is  $\mu_{t+1|t}\pm 1.96\sigma_{t+1|t}$  that is  $\phi_0+\phi_1y_t\pm 1.96\sigma_{\varepsilon}$

for forecast step  $h \in \{1, 2, 3, \ldots\}$ 

for conditional mean we have

$$\mu_{t+1|t} = E_t(y_{t+1}|\mathcal{I}_t) = \phi_0 + \phi_1 y_t$$

$$\mu_{t+2|t} = E_t(y_{t+2}|\mathcal{I}_t) = \phi_0 + \phi_1 E_t(y_{t+1}|\mathcal{I}_t) = (1+\phi_1)\phi_0 + \phi_1^2 y_t$$

$$\mu_{t+3|t} = E_t(y_{t+3}|\mathcal{I}_t) = \phi_0 + \phi_1 E_t(y_{t+2}|\mathcal{I}_t) = (1+\phi_1+\phi_1^2)\phi_0 + \phi_1^3 y_t$$

$$\vdots$$

and so  $\mu_{t+h|t} o rac{\phi_0}{1-\phi_1}$  as  $h o \infty$ 

for conditional variance

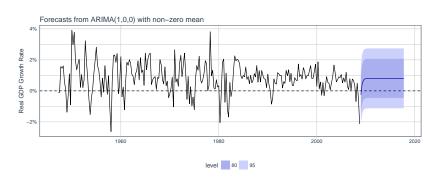
$$\begin{split} \sigma_{t+1|t}^2 &= var_t(y_{t+1}|\mathcal{I}_t) = var(\varepsilon_{t+1}) = \sigma_\varepsilon^2 \\ \sigma_{t+2|t}^2 &= var_t(y_{h+2}|\mathcal{I}_t) = var(\phi_1 y_{t+1} + \varepsilon_{t+2}|\mathcal{I}_t) = (1+\phi_1^2)\sigma_\varepsilon^2 \\ \sigma_{t+3|t}^2 &= var_t(y_{h+3}|\mathcal{I}_t) = var(\phi_1 y_{h+2} + \varepsilon_{t+3}|\mathcal{I}_t) = (1+\phi_1^2 + \phi_1^4)\sigma_\varepsilon^2 \\ &\vdots \\ \text{and so } \sigma_{t+h|t}^2 &\to \frac{\sigma_\varepsilon^2}{1-\phi^2} \text{ as } h \to \infty \end{split}$$

conditional mean thus converges to the unconditional mean, conditional variance converges to the unconditional variance

```
library(tidyquant)
library(timetk)
# obtain data on real GDP, construct its log change
data.tbl <-
   tq_get("GDPC1",
           get = "economic.data",
           from = "1947-01-01",
           to = "2017-12-31") \%
    rename(v = price) %>%
    mutate(dly = log(y) - lag(log(y)))
# split sample - estimation subsample dates
fstQ <- 1947.00 # 1947Q1
1stQ <- 2008.75 # 2008Q4
# convert data into ts, which is the format that Acf, auto.arima and forecast expect
data.ts <- data.tbl %>%
    tk_ts(select = dly, start = fstQ, frequency = 4)
# split sample - estimation and prediction subsamples
data.ts.1 <- data.tbl %>%
    tk_ts(select = dly, start = fstQ, end = lstQ, frequency = 4)
data.ts.2 <- data.ts %>%
    window(start = 1stQ + 0.25)
```

```
# create 1,2,..,h step ahead forecasts, with 2008Q4 as forecast origin
library(forecast)
m1 <- Arima(data.ts.1, order=c(1,0,0))
m1.f.1.to.hmax <- forecast(m1, length(data.ts.2))

# plot the forecast
library(ggplot2)
theme_set(theme_tq())
autoplot(m1.f.1.to.hmax) +
    geom_hline(yintercept = 0, linetype="dashed") +
    scale_y_continuous(labels = scales::percent) +
    labs(x = "", y = "Real GDP Growth Rate")</pre>
```



#### Example: MA(2) model

suppose that  $y_t$  follows an MA(2) model  $y_t = \phi_0 + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2}$  with  $\varepsilon_t \sim N(0, \sigma_\varepsilon^2)$  and that  $L(e_{t,h}) = a e_{t,h}^2$  then:

for conditional mean we have

$$\mu_{t+1|t} = E_t(y_{t+1}|\mathcal{I}_t) = \phi_0 + \theta_1 \varepsilon_t + \theta_2 \varepsilon_{t-1}$$

$$\mu_{t+2|t} = E_t(y_{t+2}|\mathcal{I}_t) = \phi_0 + \theta_2 \varepsilon_t$$

$$\mu_{t+3|t} = E_t(y_{t+3}|\mathcal{I}_t) = \phi_0$$

▶ for conditional variance

$$\begin{split} \sigma_{t+1|t}^2 &= var_t(y_{t+1}|\mathcal{I}_t) = var(\varepsilon_{t+1}) = \sigma_{\varepsilon}^2 \\ \sigma_{t+2|t}^2 &= var_t(y_{h+2}|\mathcal{I}_t) = var(\varepsilon_{t+2} + \theta_1\varepsilon_{t+1}) = (1 + \theta_1^2)\sigma_{\varepsilon}^2 \\ \sigma_{t+3|t}^2 &= var_t(y_{h+3}|\mathcal{I}_t) = var(\varepsilon_{t+3} + \theta_1\varepsilon_{t+2} + \theta_2\varepsilon_{t+1}) = (1 + \theta_1^2 + \theta_2^2)\sigma_{\varepsilon}^2 \end{split}$$

▶ the 1, 2, and 3 step ahead point forecasts are thus

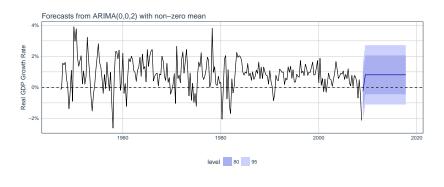
$$f_{t,1} = \mu_{t+1|t} = \phi_0 + \theta_1 \varepsilon_t + \theta_2 \varepsilon_{t-1}$$

$$f_{t,2} = \mu_{t+2|t} = \phi_0 + \theta_2 \varepsilon_t$$

$$f_{t,3} = \mu_{t+3|t} = \phi_0$$

## Example: MA(2) model

```
m2 <- Arima(data.ts.1, order=c(0,0,2))
m2.f.1.to.hmax <- forecast(m2, length(data.ts.2))
autoplot(m2.f.1.to.hmax) +
    geom_hline(yintercept = 0, linetype="dashed") +
    scale_y_continuous(labels = scales::percent) +
    labs(x = "", y = "Real GDP Growth Rate")</pre>
```



# Forecasting using ARMA(p,q) models

 $\mathsf{ARMA}(p,q)$  models are mostly suitable for forecasts with a small step h, forecasts of distant future are not particularly accurate

forecast based on an AR(p) model:

- conditional mean converges to unconditional mean gradually
- conditional variance converges to unconditional variance gradually

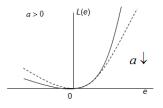
forecast based on an MA(q) model:

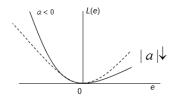
- lacktriangleright once h>q the conditional mean jumps straight to unconditional mean
- lacktriangle once h>q the conditional variance jumps straight to unconditional variance

#### Asymmetric Loss Function

#### Linex function

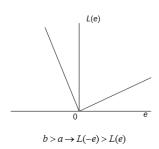
$$L(e) = \exp(ae) - ae - 1$$
,  $a \neq 0$ 





#### Lin-lin function

$$L(e) = \begin{cases} a \mid e \mid & e > 0 \\ b \mid e \mid & e \le 0 \end{cases}$$



# Point, Interval and Density Forecasts

suppose that conditional density  $f(y_{t+h}|\mathcal{I}_t)$  is  $N(\mu_{t+h|t}, \sigma_{t+h|t}^2)$  so that density forecast is  $N(\mu_{t+h|t}, \sigma_{t+h|t}^2)$  and

- 1. if loss function is linex  $L(e_{t,h}) = exp(ae_{t,h}) ae_{t,h} 1$
- $\blacktriangleright$  optimal point forecast is  $f_{t+h}^* = \mu_{t+h|t} + \frac{a}{2} \sigma_{t+h|t}^2$

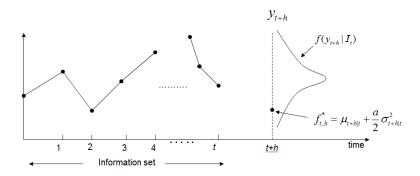
2. if loss function is linlin

$$L(e_{t,h}) = \begin{cases} a|e_{t,h}| & \text{if } e_{t,h} < 0\\ (1-a)|e_{t,h}| & \text{if } e_{t,h} \ge 0 \end{cases}$$

lacktriangle optimal point forecast is conditional quintile  $f_{t+h}^* = q_a(y_{t+h}|\mathcal{I}_t)$ 

thus for asymmetric loss function optimal forecast is actually biased - on average forecast error is either positive or negative

#### Linex Loss Function



suppose that  $y_t$  follows an AR(1) model  $y_t = \phi_0 + \phi_1 y_{t-1} + \varepsilon_t$  with  $\varepsilon_t \sim N(0, \sigma_\varepsilon^2)$  and that  $L(e_{t,h}) = exp(ae_{t,h}) - ae_{t,h} - 1$  then:

for conditional mean we have

$$\mu_{t+1|t} = E_t(y_{t+1}|\mathcal{I}_t) = \phi_0 + \phi_1 y_t$$

for conditional variance

$$\sigma_{t+1|t}^2 = var_t(y_{t+1}|\mathcal{I}_t) = var(\varepsilon_{t+1}) = \sigma_{\varepsilon}^2$$

▶ thus the 1 step ahead point forecast of  $y_{t+1}$  is

$$f_{t,1} = \mu_{t+1|t} + \frac{a}{2}\sigma_{t+1|t}^2 = \phi_0 + \phi_1 y_t + \frac{a}{2}\sigma_{\varepsilon}^2$$

▶ the conditional density forecast for  $y_{t+1}$  is  $N(\phi_0+\phi_1y_t,\sigma_{\varepsilon}^2)$ 

## **Evaluating Accuracy of Forecasts**

#### general idea:

- ▶ split sample into two parts: estimation sample  $y_1, \dots, y_t$ prediction sample  $y_{t+1}, \dots, y_T$
- estimate the model using the first subsample
- evaluate in-sample accuracy compare fitted values  $\hat{y}_1, \dots, \hat{y}_t$  with actual values  $y_1, \dots, y_t$
- use the second subsample to construct set of h step ahead forecasts  $f_{t,h}, f_{t+1,h}, \dots, f_{T-h,h}$
- evaluate **out-of-sample accuracy** compare forecasts  $f_{t,h}, f_{t+1,h}, \ldots, f_{T-h,h}$  with actual values  $y_{t+h}, y_{t+1+h}, \ldots, y_T$

#### In-Sample Evaluation of Accuracy

given the fitted values  $\hat{y}_j$  from the model, and in sample residuals  $e_j = y_j - \hat{y}_j$ 

Mean Error - measure of the average bias

$$ME = \frac{1}{t} \sum_{j=0}^{t} e_j$$

Mean Squared Error - sample average loss for quadratic loss function

$$MSE = \frac{1}{t} \sum_{i=1}^{t} e_j^2$$

Mean Absolute Error - sample average loss for absolute value loss function

$$MAE = \frac{1}{t} \sum_{j=0}^{t} |e_j|$$

Mean Absolute Percentage Error

$$MAPE = \frac{1}{t} \sum_{i=1}^{t} \left| \frac{e_j}{y_j} \right|$$

**Mean Absolute Scaled Error** - compares in sample MAE of the model forecast with in sample MAE for one-step naive forecast method  $\hat{y}_{j+1}=y_j$ 

$$MASE = \frac{\frac{1}{t} \sum_{j=0}^{t} |e_j|}{\frac{1}{t-1} \sum_{j=1}^{t-1} |y_{j+1} - y_j|}$$

# In-Sample Evaluation of Accuracy

```
m1 <- Arima(data.ts.1, order = c(1,0,0))
accuracy(m1)
```

## ME RMSE MAE MAPE MAPE MASE ACF1
## Training set 1.429256e-05 0.009159989 0.006768226 -44.3292 166.0401 0.6358981 -0.03028121

# Out-of-Sample Evaluation of Accuracy

given out of sample forecast errors  $e_{t,h}, e_{t+1,h}, \dots, e_{T-h,h}$ 

Mean Error

$$ME = \frac{1}{T - t - h + 1} \sum_{j=0}^{T - h - t} e_{t+j,h}$$

Mean Squared Error

$$MSE = \frac{1}{T - t - h + 1} \sum_{i=0}^{T - h - t} e_{t+j,h}^2$$

Mean Absolute Error

$$MAE = \frac{1}{T - t - h + 1} \sum_{i=0}^{T - h - t} |e_{t+j,h}|$$

Mean Absolute Percentage Error

$$MAPE = \frac{1}{T - t - h + 1} \sum_{j=0}^{T - h - t} \left| \frac{e_{t+j,h}}{y_{t+j+h}} \right|$$

Mean Absolute Scaled Error

$$MASE = \frac{\frac{1}{T - l - h + 1} \sum_{j=0}^{T - h - t} |e_{t+j,h}|}{\frac{1}{t - h} \sum_{j=1}^{t-1} |y_{j+h} - y_{j}|}$$

## Out-of-Sample Evaluation of Accuracy - Forecasting Schemes

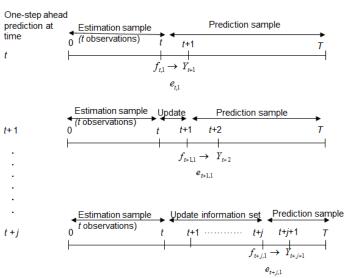
out of sample forecasts and forecast errors used to calculate ME, MSE, MAE, MPE, MAPE, . . . can be constructed using one of the three schemes:

- fixed scheme
- recursive scheme
- rolling scheme

#### Forecasting Schemes

Fixed scheme example for one step ahead forecast:

model is estimated only once, each one step ahead forecast is constructed using same parameters



#### Out-of-Sample Evaluation of Accuracy - Fixed Scheme

```
# estimate AR(1) model 1947Q2 to 2008Q4 m1 <- Arima(y = data.ts.1, order = c(1,0,0)) # create 1-step ahead forecasts - forecast origin is moving from 2008Q4 to 2017Q3 # but always use same estimated model m1 so this is a fixed forecasting scheme m1.f.1 <- Arima(y = data.ts, model = m1) # evaluate accuracy of 1-step ahead forecast throughout the whole sample 1947Q2 to 2016Q4 accuracy(fitted(m1.f.1), data.ts)
```

```
## Test set -0.0002269938 0.008719256 0.006378762 -51.41388 166.8595 -0.02181353 0.9346668
```

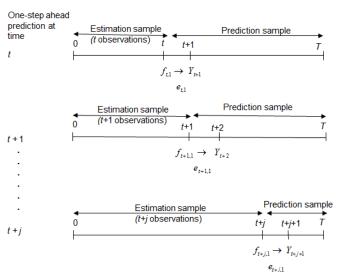
```
# evaluate accuracy of out-of-sample 1-step ahead forecasts
accuracy(fitted(m1.f.1), data.ts.2)
```

```
## ME RMSE MAE MPE MAPE ACF1 Theil's U
## Test set -0.001882486 0.004686156 0.003706608 -100.0227 172.4811 -0.3309034 0.3467054
```

#### Forecasting Schemes

Recursive scheme example for one step ahead forecast:

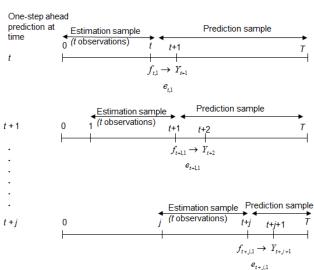
estimation sample keeps expanding and model is re-estimated again when each new observation is added to the estimation sample



#### Forecasting Schemes

Rolling scheme example for one step ahead forecast:

estimation sample always contains the same number of observation and model is re-estimated again within each rolling sample



# Out-of-Sample Evaluation of Accuracy - Rolling scheme

```
library(tibbletime)
library(sweep)
# window size
window.length <- length(data.ts.1)
# create rolling ARMA function with rollifu from tibbletime package
roll_Arima <- rollify(~Arima(.x, order=c(1,0,0)), window = window.length, unlist = FALSE)
# estimate rolling ARMA model, create 1 period ahead rolling forecasts
h <- 1
results <-
   data.tbl %>%
    as_tbl_time(index = date) %>%
    mutate(arma.model = roll Arima(dly)) %>%
   filter(!is.na(arma.model)) %>%
    mutate(arma.f = map(arma.model, (. %>% forecast(h) %>% sw_sweep())))
# extract the 1 period ahead rolling forecasts
m1.f.1.rol <-
    results %>%
    select(date, arma.f) %>%
    unnest(arma.f) %>%
   filter(kev == "forecast") %>%
    mutate(date = date %m+% months(3))
```

# Out-of-Sample Evaluation of Accuracy - Rolling Scheme

```
# plot the 1 period ahead rolling forecasts
m1.f.1.rol '\'>'\'
ggplot(aes(x = date, y = value)) +
    geom_ribbon(aes(ymin = lo.95, ymax = hi.95), color = NA, fill = "steelblue", alpha = 0.2) +
    geom_ribbon(aes(ymin = lo.80, ymax = hi.80), color = NA, fill = "steelblue", alpha = 0.3) +
    geom_line(size = 0.7, col = "blue") +
    geom_line(data = (data.tbl '\'>'\', filter(year(date) > 1999)), aes(x = date, y = dly)) +
    geom_hline(yintercept = 0, linetype="dashed") +
    scale_y_continuous(labels = scales::percent) +
    scale_color_manual(values = c("black", "darkblue")) +
    labs(x = "", y = "", title = "Real GDP Growth Rate") +
    theme(legend.position = "none")
```



## Forecasting Schemes - Comparison

advantages and disadvantages of the three schemes:

#### fixed scheme

- ▶ fast and convenient because there is one and only one estimation
- does not allow for parameter updating, so again problem with structural breaks and model's stability

#### recursive scheme

- incorporates as much information as possible in the estimation of the model
- advantageous if the model is stable over time
- if the data have structural breaks, model's stability is compromised and so is the forecast

#### rolling scheme

- avoids the potential problem with the model's stability
- more robust against structural breaks in the data
- does not make use of all the data

#### Comparison

```
# multistep forecast
accuracy(m1.f.1.to.hmax$mean, data.ts.2)

## ME RMSE MAE MPE MAPE ACF1 Theil's U
```

```
# 1 step ahead fixed scheme forecast
accuracy(fitted(m1.f.1), data.ts.2)
```

```
## Test set -0.001882486 0.004686156 0.003706608 -100.0227 172.4811 -0.3309034 0.3467054
```

## Test set -0.002916416 0.004895071 0.003811905 -144.8972 222.8371 -0.03828393 0.3542732

```
# 1 step ahead rolling scheme forecast
accuracy(m1.f.1.rol$value, data.ts.2)
```

## Test set -0.001667329 0.00463774 0.00366208 -94.00967 166.0779 -0.335279 0.3705797