# Texas Tech University Department of Economics Spring 2018

Eco 4306: Economic and Business Forecasting

## Midterm 2

Short questions (45 points)
Q1. 7.5 points
Q2. 7.5 points
Q3. 7.5 points
Q4. 7.5 points
Q4. 7.5 points
Q5. 7.5 points
Q6. 7.5 points
Applied problems (60 points)
Q7. 10 points
Q8. 10 points
Q9. 10 points
Q10. 10 points
Q11. 10 points
Q12. 10 points

Name:

ID:

Good luck!

## Question 1 (7.5 points)

Explain the difference between in-sample evaluation and out-of-sample evaluation.

See slide 2 in lec13slides.pdf and slide 10 in lec06slides.pdf.

## **Question 2** (7.5 points)

Explain how Mean Squared Error and Mean Loss are used in the assessment of forecasts.

See slides 9 and 14 in lec13slides.pdf.

## Question 3 (7.5 points)

Give an example of a deterministic trend g(t) other than a linear trend and plot its graph. Write the equation of a model with this trend.

See slides 10 to 12 in lec16slides.pdf.

### Question 4 (7.5 points)

Explain the difference between a trend stationary time series and a difference stationary time series.

See slide 15 in lec18slides.pdf.

& account o (1.0 points	Question	<b>5</b>	(7.5)	points	)
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Explain what it means for a time series process to be I(1), and what it means for a process to be I(0). See slide 14 in lec18slides.pdf.

## Question 6 (7.5 points)

Write down the equation for a pure seasonal S-AR(1) model. Describe how its AC and PAC functions look like.

See slides 9 and 10 in lec11slides.pdf.

#### Question 7 (10 points)

Consider two candidate models for change in monthly private residential construction spending, AR(1) and AR(2)+SAR(1), the results for which are below. Evaluate the adequacy of these models based on the correlograms of residuals, AIC and BIC, and statistical significance of coefficients.

Dependent Variable: DCONST Method: ARMA Maximum Likelihood (BFGS) Date: 04/11/18 Time: 16:28 Sample: 1993M02 2013M12 Included observations: 251 Convergence achieved after 3 iterations

Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	34.02301	284.5681	0.119560	0.9049
AR(1)	0.503787	0.082472 6.108569		0.0000
SIGMASQ	4263658.	311155.6	13.70266	0.0000
R-squared	0.254386	Mean dependent var		49.44223
Adjusted R-squared	0.248373	S.D. dependent var		2396.078
S.E. of regression	2077.314	Akaike info criterion		18.12859
Sum squared resid	1.07E+09	Schwarz criterion		18.17072
Log likelihood	-2272.138	Hannan-Quinn criter.		18.14554
F-statistic	42.30579	Durbin-Watson stat		2.030264
Prob(F-statistic)	0.000000			

Date: 04/11/18 Time: 16:30 Sample: 1993M01 2013M12 Included observations: 251

Q-statistic probabilities adjusted for 1 ARMA term

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
40	l do	1	-0.016	-0.016	0.0654	
· 🗀		2	0.196	0.196	9.8361	0.002
<b>=</b> -	<b> </b>  -	3	-0.156	-0.156	16.045	0.000
<b>□</b> !	<b>□</b> '	4	-0.152	-0.202	21.954	0.000
ı 🗓 ı	1 10	5	-0.073	-0.014	23.343	0.000
<u> </u>	<b>=</b>	6	-0.437		72.819	0.000
<b>₁</b> Щ •	<b>=</b>  -	7	-0.077		74.368	0.000
<u> </u>	ļ @ -	8	-0.167		81.652	0.000
<b>■</b> '	<b> </b>		-0.141		86.868	0.000
' <b>!=</b>	ļ ( <b>4</b> )	10		-0.055	94.847	0.000
1/1	! <b>□</b> !	11	-0.004		94.852	0.000
1		12	0.928	0.873	323.74	0.000
111	<u> </u>	13	-0.022	0.004	323.87	0.000
' 🟴	! ■'	14		-0.179	333.60	0.000
<u> </u>	ļ ' <b>Ū</b> '	:	-0.170		341.37	0.000
<u>"</u> '	']'		-0.138	0.000	346.48	0.000
'"	'"['			-0.027	347.93	0.000
· ·	'['		-0.426	0.036	397.29	0.000
<u>"</u> "	'] '		-0.074	0.040	398.77	0.000
<u>"</u>	'1[:	20	-0.181	-0.046	407.75	0.000
<b>Q</b> !	' <b> </b>	21	-0.107	0.110	410.89	0.000
' <b>!</b> !!	'¶'	22		-0.070	416.75	0.000
1111	'  '	23		-0.017	416.81	0.000
'	' '	24		-0.003	623.42	0.000
'[_	']'		-0.023	0.005	623.56	0.000
<u> </u>	' '	26		-0.009	633.38	0.000
<u>"</u>	']'		-0.183		642.88	0.000
9'	'¶'	:	-0.128		647.55	0.000
<u>"</u> " '	'['		-0.074		649.09	0.000
<u> </u>	' <b>]</b>		-0.408	0.052	696.98	0.000
<u>"</u> "	<u> </u>		-0.070		698.38	0.000
<u>"</u> '	! <b>"</b> !'		-0.203		710.36	0.000
'만	' <b> </b>  '	33		0.028	711.77	0.000
'.₽	']'	34		-0.001	715.82	0.000
1111	1 11	35		-0.014	716.18	0.000
	(1)	36	0.783	-0.102	897.23	0.000

Dependent Variable: DCONST Method: ARMA Maximum Likelihood (BFGS) Date: 03/08/19 Time: 17:29 Sample: 1993M02 2013M12 Included observations: 251 Convergence achieved after 7 iterations

Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(1) AR(2) SAR(12) SIGMASQ	214.4219 0.497140 0.116143 0.944592 373960.6	1374.109 0.054595 0.052147 0.013249 26485.07	0.156044 9.105947 2.227211 71.29646 14.11968	0.8761 0.0000 0.0268 0.0000 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.934603 0.933540 617.7066 93864109 -1980.137 878.9103 0.000000	Mean depend S.D. depende Akaike info cri Schwarz criter Hannan-Quin Durbin-Watso	nt var iterion rion n criter.	49.44223 2396.078 15.81783 15.88805 15.84609 1.975678

Date: 03/08/19 Time: 17:29 Sample: 1993M01 2013M12 Included observations: 251

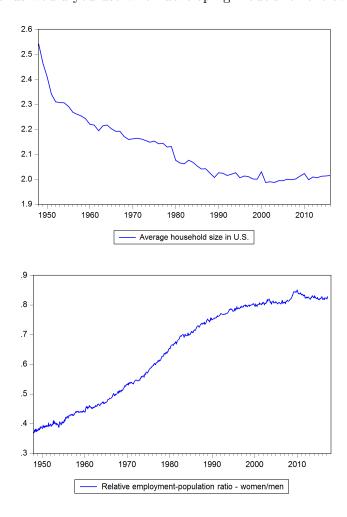
Q-statistic probabilities adjusted for 3 ARMA terms

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
1)1	())	1	0.011	0.011	0.0322	
ı <b>j</b> i	<u> </u>	2	0.054	0.054	0.7743	
ı <b>ğ</b> ı	i[  i	3	-0.054		1.5282	
1 <b>(</b> 1	<b>  [[</b>	4	-0.032	-0.034	1.7972	0.180
ı <b>j</b> ir	i  Ii	5	0.042	0.049	2.2539	0.324
<b>□</b> '	<u> </u>	6	-0.091	-0.092	4.3953	0.222
<b>₁</b> Щ ₁	<b>□</b>   ·	7	-0.081	-0.089	6.1091	0.191
ı þi	·  =	8	0.112	0.132	9.3965	0.094
ι <b>α</b> ι	III	9	-0.073	-0.078	10.799	0.095
1 <b>(</b> 1	<b>       </b>	10	-0.031	-0.065	11.048	0.137
<b>- 1</b>	1(1		-0.041		11.501	0.175
1 1	1 1		-0.005		11.507	0.243
1(1	' <b>[</b>   '		-0.013		11.553	0.316
1 1	1 11		-0.007	0.014	11.565	0.397
1 1	1 11	15	0.006	0.021	11.575	0.480
1   1	'  '	16		-0.024	11.670	0.555
- i <b>j</b> i	' <b> </b>    '	17	0.040	0.045	12.109	0.598
·¶ ·	<u>'</u>	18	-0.095		14.580	0.482
· þ·	ļ ' <u>l</u> l'	19	0.042	0.037	15.068	0.520
· 🏴	י ויי	20	0.070	0.082	16.416	0.495
- III -	' <b>[</b>   '		-0.041		16.875	0.532
1 11	'  '	22		-0.006	16.904	0.596
'[[	'11'		-0.005	0.036	16.912	0.659
' <b>P</b>	' <u> </u>	24	0.158	0.146	23.933	0.296
<u>'</u>	'[]'	25		-0.038	23.952	0.350
'¶'	'¶'		-0.075		25.527	0.324
'[['	'['		-0.033		25.837	0.362
1]1	']'	28	0.024	0.018	25.999	0.408
'['	'¶'		-0.044		26.564	0.432
111	'  :		-0.043		27.097	0.459
<u> </u>	'. .'		-0.037		27.498	0.491
י וווי	<u> </u>	32	0.064	0.024	28.673	0.482
111	<u>                                   </u>	33	0.022	0.028	28.816	0.527
:11:		34	0.044	0.058	29.388	0.549
111	'∐'.		-0.016 -0.071		29.459 30.955	0.596
<u>""</u>	<u> </u>	30	-0.071	-0.119	30.905	0.009

While coefficients are statistically significant in both AR(1) and AR(2)+S-AR(1) model, both the AIC and the BIC favor AR(2)+S-AR(1) model, and the correlograms of residuals shows that there is a serious problem with seasonality that is not accounted for in the AR(1). Overall, AR(2)+SAR(1) is a much better model for change in monthly private residential construction spending.

## Question 8 (10 points)

Which deterministic trends would you use when developing models for the two series plotted below?

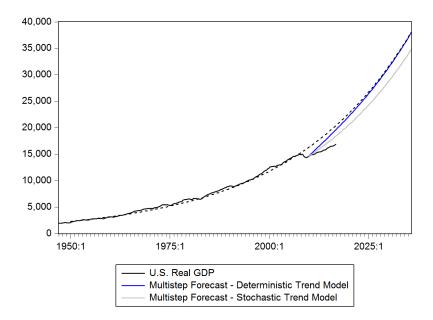


Exponential or logistic trend look like reasonable candidates to capture the time path of the average household size.

Logistic trend would be a good candidate to model relative employment-population ratio of women vs men.

#### Question 9 (10 points)

The following figure shows the multistep forecasts for the U.S. real GDP, from the deterministic model and from the stochastic trend model, both for the period 2010Q1-2035Q4. Discuss the main difference in the behavior of the two forecasts and explain the reason for this difference.



The effect of a one time shock is only temporary in the deterministic trend model, multistep forecast thus converges back to the original trend.

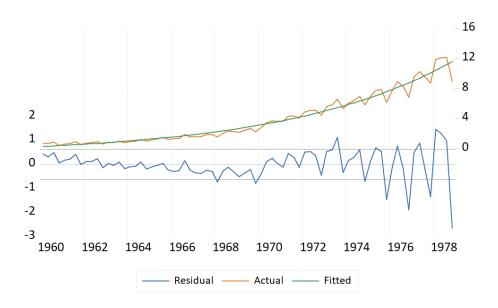
The effect of a one time shock is permanent in the stochastic trend model, multistep forecast thus does not converge back to the original trend, but rather stay permanently below the original trend.

#### Question 10 (10 points)

Consider a model for quarterly earnings per share of the Johnson and Johnson company

$$JNJ_t = \beta_0 + \beta_1 e^{\beta_2 t} + \varepsilon_t$$

Given the plot with actual values, fitted values, and residuals below, explain how you would proceed with modifying/developing the model further.



There are two problems with the estimated model:

- 1. it can match the trend, but not the seasonal pattern
- 2. the variance of residuals does not appear to be constant, it is increasing over time

To deal with the issue of variance of residuals increasing over time we reestimate the model using log transformed data  $\log JNJ_t$  as the dependent variable instead of  $JNJ_t$ .

To deal with he seasonal pattern in the residuals seasonal AR or MA model needs to be specified, based on the correlogram of residuals.

#### Question 11 (10 points)

Below are the results for the Augmented Dickey-Fuller unit root test for log transformed earnings per share  $\log JNJ_t$ , and for the first difference of the log transformed earnings per share  $\Delta \log JNJ_t$ .

Interpret the results, and determine whether  $\log JNJ_t$  is I(0) or I(1).

Null Hypothesis: LJNJ has a unit root Exogenous: Constant, Linear Trend

Lag Length: 3 (Automatic - based on SIC, maxlag=11)

		t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic		-1.696535	0.7428
Test critical values:	1% level	-4.090602	
	5% level	-3.473447	
	10% level	-3.163967	

<sup>\*</sup>MacKinnon (1996) one-sided p-values.

Null Hypothesis: D(LJNJ) has a unit root Exogenous: Constant, Linear Trend

Lag Length: 2 (Automatic - based on SIC, maxlag=11)

		t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic		-19.93554	0.0001
Test critical values:	1% level	-4.090602	
	5% level	-3.473447	
	10% level	-3.163967	

<sup>\*</sup>MacKinnon (1996) one-sided p-values.

The p-value in the unit root test for  $\log JNJ_t$  is 0.7428 we thus can not reject the hypothesis that  $\log JNJ_t$  has a unit root.

The p-value in the unit root test for first difference  $\Delta \log JNJ_t$  is 0.0001 we thus strongly reject the hypothesis that  $\Delta \log JNJ_t$  has a unit root.

These two results combined imply that  $\log JNJ_t$  is an I(1) time series - it is non-stationary but taking first difference transforms it into a stationary time series.

#### Question 12 (10 points)

Consider two models for U.S. real GDP, used to construct forecast for the period 2010Q1-2016Q4:

- model A: deterministic trend model for which the sequence of 1-step ahead forecasts has RMSE=103.45 and the multistep forecast has RMSE=1649.06
- model B: stochastic trend model for which the sequence of 1-step ahead forecasts has RMSE=77.32 and the multistep forecast has RMSE=905.18.

The 1-step ahead forecasts are then used to perform the test of equal predictive ability be estimating

$$\Delta L_{t+j,1} = \beta_0 + u_{t+j}$$
 with  $j = 0, 1, 2, \dots, T - t - 1$ 

where  $\Delta L_{t+j,1} = (e_{t+j,1}^A)^2 - (e_{t+j,1}^B)^2$ , and  $e_{t+j,1}^k$  is the one step ahead forecast error for forecast from model k in period t+j. Explain the idea behind this test and interpret its results below. Discuss how we would use it together with above RMSE values in model selection process.

Dependent Variable: DL\_TREND
Method: Least Squares
Date: 04/09/17 Time: 18:34
Sample (adjusted): 2010Q1 2016Q4
Included observations: 28 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	5454.311	1293.939 4.215275		0.0002
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood Durbin-Watson stat	0.000000 0.000000 6846.884 1.27E+09 -286.5045 2.683486	Mean depend S.D. depende Akaike info cri Schwarz criter Hannan-Quin	nt var terion ion	5454.311 6846.884 20.53604 20.58361 20.55058

The main idea behind the test of equal predictive ability is to test whether the difference in MSE or RMSE between two two competing forecasts is statistically significant or not. The null hypothesis of equal predictive ability of forecasts from two models A and B is that the difference in the MSE and RMSE is not statistically significant, and thus the estimated coefficient  $\beta_0$  in the test regression is not statistically significant. If  $\beta_0$  is statistically significant we reject the hypothesis of equal predictive ability of forecasts A and B. This is exactly the case here when comparing model A: a deterministic trend model for real GDP growth vs model B: stochastic trend model since the p-value for  $\beta_0$  is only 0.0002. Model B is thus a much better model for real GDP growth rate forecasting.