Seasonal Models

Pure Seasonal Models

simple pure seasonal AR model

$$y_t = \phi_s y_{t-s} + a_t$$

ACF: spike at each multiple of s PACF: single spike at lag s

simple pure seasonal MA model

$$y_t = a_t + \theta_s a_{t-s}$$

PACF: spike at each multiple of *s* ACF: single spike at lag *s*

in practice most time series contain a seasonal AR or MA component at the same time as regular AR or MA component



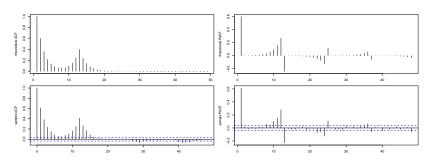
Additive Seasonal AR model

▶ AR model with an additive seasonal MA component

$$(1 - \phi)x_t = (1 + \Theta B^s)a_t$$

so that
$$x_t = \phi x_{t-1} + a_t + \Theta a_{t-s}$$

• example: $\phi = 0.6$, $\Theta = 0.5$



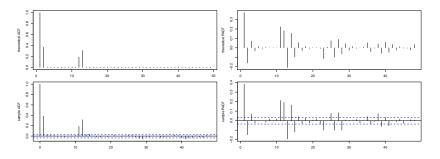
Additive Seasonal MA model

▶ MA model with an additive seasonal MA component

$$x_t = (1 + \theta B + \Theta B^s)a_t$$

so that $x_t = a_t + \theta a_{t-1} + \Theta a_{t-s}$

- ▶ ACF: $\rho_1 \neq 0$, $\rho_{s-1} \neq 0$, $\rho_s \neq 0$
- example: $\theta = 0.6$, $\Theta = 0.5$



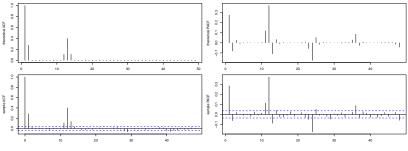
Multiplicative Seasonal MA model

Multiplicative seasonal MA model

$$x_t = (1 + \theta B)(1 + \Theta B^s)a_t$$

so that $x_t = a_t + \theta a_{t-1} + \Theta a_{t-s} + \theta \Theta a_{t-s-1}$

- ACF: $\rho_1 \neq 0$, $\rho_{s-1} \neq 0$, $\rho_s \neq 0$, $\rho_{s+1} \neq 0$
- compared to the additive model, multiplicative model allows for interaction of regular and seasonal components
- example: $\theta = 0.3$, $\Theta = 0.5$



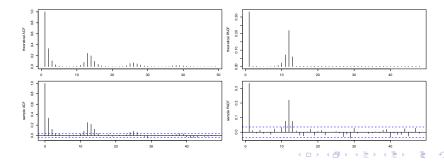
Multiplicative Seasonal AR model

multiplicative AR model with seasonal component

$$(1 - \phi B)(1 - \Phi B^s)x_t = a_t$$

so that
$$x_t = \phi x_{t-1} + \Phi x_{t-s} + \phi \Phi x_{t-s-1} + a_t$$

- ▶ ACF: if $\phi > 0$, $\Phi > 0$ exponential decay interrupted by increasing autocorrelations around *each* multiple of s PACF: large spikes at lag 1 and lag s and multiple smaller spikes between lag 2 and lag s+1
- example: $\phi_1 = 0.3$, $\Phi_1 = 0.5$.



Seasonal Differencing

 for economic data that is nonstationarity due to economic growth a common approach is to transform data using a logartihm and apply regular differencing

$$w_t = \Delta \log y_t$$

where $\Delta = 1 - B$, so that $w_t = (1 - B) \log y_t$

for economic data that is both nonstationarity due to economic growth and shows seasonal pattern the approach is to transform data using a logartihm and apply both regular and seasonal differencing

$$w_t = \Delta_s \Delta \log y_t$$

where $\Delta=1-B$ and $\Delta_s=1-B^s$, so we have $w_t=(1-B^s)(1-B)\log y_t$

- ocassionally data has to be diffferenced more than once by applying $\Delta^d = (1 B)^d$, or $\Delta^s_s = (1 B^s)^D$
- multiplicative models are written in the form ARIMA $(p, d, q)(P, D, Q)_s$
- in practice ARIMA(1,1,0)(0,1,1) $_s$ and ARIMA(0,1,1)(0,1,1) $_s$ occur routinely

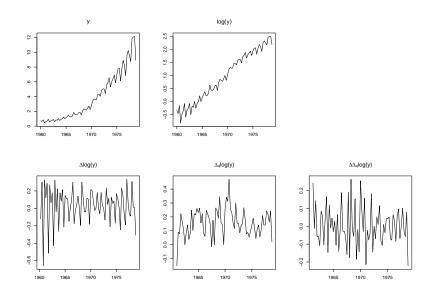
Example: Johnson & Johnson quarterly earnings per share

```
str(y)
## Time-Series [1:84] from 1960 to 1981: 0.71 0.63 0.85 0.44 0.61 0.69 0.92 0.55 0.72 0.77 ...
# split sample into two parts - estimation sample and prediction sample
vall <- v
v1 <- window(vall, end=c(1978,4))
y2 <- window(yall, start=c(1979,1))
# first part used to identify and estimate the model
y <- y1
# log, log-change, seasonal log change
lv \leftarrow log(y)
dly1 <- diff(ly)
dlv4 \leftarrow diff(lv,4)
dly4_1 <- diff(diff(ly),4)
```

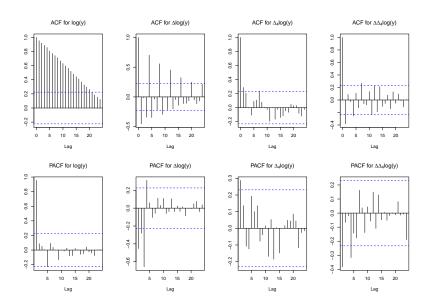
Original and transformed data

```
par(mfrow=c(2,3))
plot(y, main=expression(y))
plot(ly, main=expression(log(y)))
plot.new()
plot(dly1, main=expression(paste(Delta, "log(y)")))
plot(dly4, main=expression(paste(Delta[4], "log(y)")))
plot(dly4_1, main=expression(paste(Delta, Delta[4], "log(y)")))
```

Original and transformed data



ACF and PACF



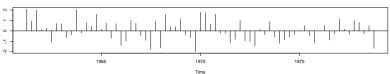
Estimate model m1 for $\Delta_s \Delta \log y_t$

```
# estimate model - twice differenced data
m1 \leftarrow arima(dly4_1, order=c(0,0,1), seasonal=list(order=c(0,0,1), period=4))
m1
##
## Call:
\#\# arima(x = dly4_1, order = c(0, 0, 1), seasonal = list(order = c(0, 0, 1), period = 4))
##
## Coefficients:
                     sma1 intercept
##
             ma1
##
         -0.6604 -0.3492
                              0.0013
## s.e. 0.1084 0.1101
                              0.0026
##
## sigma^2 estimated as 0.008374: log likelihood = 68.42, aic = -128.84
```

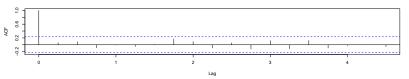
Check model m1 for adequacy

tsdiag(m1,gof.lag=36)

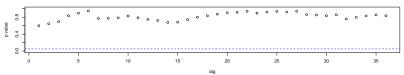




ACF of Residuals



p values for Ljung-Box statistic



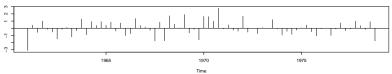
Estimate model m2 for $\Delta_s \log y_t$

```
# estimate model - seasonally differenced data
m2 <- arima(dly4,order=c(1,0,0))</pre>
m2
##
## Call:
## arima(x = dly4, order = c(1, 0, 0))
##
## Coefficients:
            ar1 intercept
##
##
         0.3412
                    0.1557
## s.e. 0.1214
                 0.0166
##
## sigma^2 estimated as 0.008689: log likelihood = 68.62, aic = -131.24
```

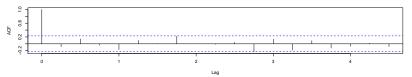
Check model m2 for adequacy

tsdiag(m2,gof.lag=36)

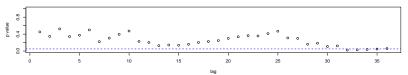




ACF of Residuals



p values for Ljung-Box statistic



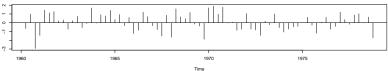
Estimate model m3 for $\Delta \log y_t$

```
# estimate model - regularly differenced data
m3 <- arima(dly1,order=c(0,0,1),seasonal=list(order=c(1,0,1),period=4))
m3
##
## Call:
## arima(x = dly1, order = c(0, 0, 1), seasonal = list(order = c(1, 0, 1), period = 4))
##
## Coefficients:
##
            ma1
                    sar1
                            sma1 intercept
##
         -0.7047 0.9386 -0.3166
                                     0.0309
## s.e. 0.1047 0.0443 0.1285
                                     0.0220
##
## sigma^2 estimated as 0.008275: log likelihood = 70.16, aic = -130.32
```

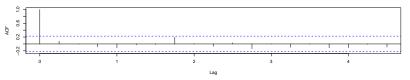
Check model m3 for adequacy

tsdiag(m3,gof.lag=36)

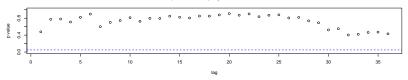




ACF of Residuals



p values for Ljung-Box statistic



Estimate model m4 for $\log y_t$

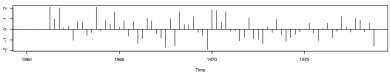
```
# estimate model - data not differenced
m4 <- arima(ly,order=c(0,1,1),seasonal=list(order=c(0,1,1),period=4))
m4

## ## Call:
## arima(x = ly, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1), period = 4))
##
## Coefficients:
## mal smal
## -0.6559 -0.3492
## s.e. 0.1094 0.1104
##
## sigma^2 estimated as 0.008409: log likelihood = 68.28, aic = -130.57</pre>
```

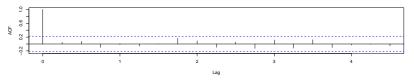
Check model m4 for adequacy

tsdiag(m4,gof.lag=36)

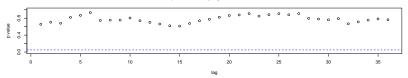




ACF of Residuals



p values for Ljung-Box statistic



Forecasts

```
library("forecast")

# construct eight quarters ahead forecasts
m1.fcast <- forecast(m1, h=8)
m2.fcast <- forecast(m2, h=8)
m3.fcast <- forecast(m3, h=8)
m4.fcast <- forecast(m4, h=8)</pre>
```

Forecasts

```
par(mfrow=c(2,2), cex=0.7, mar=c(2,4,3,1))
plot(m1.fcast, xlim=c(1970,1981))
lines(diff(diff(log(yal1),4)))
plot(m2.fcast, xlim=c(1970,1981))
lines(diff(log(yal1),4))
plot(m3.fcast, xlim=c(1970,1981), ylim=c(-0.4,0.6))
lines(diff(log(yal1))))
plot(m4.fcast, xlim=c(1970,1981), ylim=c(0.5,3.5))
lines(log(yal1))
```

