# Properties of moments of random variables \*

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Let X and Y be real random variables, and let r and s be real positive constants (r > 0, s > 0). The distribution functions of X and Y are denoted  $F_X(x) = P[X \le x]$  and  $F_Y(x) = P[Y \le x]$ .

### 1. Existence of moments

- **1.1** EXISTENCE OF ABSOLUTE AND ORDINARY MOMENTS.  $\mathsf{E}(|X|)$  always exists in the extended real numbers  $\overline{\mathbb{R}} \equiv \mathbb{R} \cup \{\infty\} \cup \{-\infty\}$  and  $\mathsf{E}(|X|) \in [0,\infty]; \textit{i.e.}$ , either  $\mathsf{E}(|X|)$  is a non-negative real number or  $\mathsf{E}(|X|) = \infty$ .
- **1.2** E(X) exists and is finite  $\Leftrightarrow E(|X|) < \infty$ .
- **1.3**  $E(|X|) < \infty \Rightarrow |E(X)| \le E(|X|) < \infty$ .
- **1.4** If  $0 < r \le s$ , then

$$\mathsf{E}(|X|^s) < \infty \Rightarrow \mathsf{E}(|X|^r) < \infty. \tag{1.1}$$

- **1.5** Monotonicity of  $L_r$ .  $L_s \subseteq L_r$  for  $0 < s \le r$ .
- **1.6**  $E(|X|^r) < \infty \Rightarrow E(X^k)$  exists and is finite for all integers k such that  $0 < k \le r$ .

## 2. Moment inequalities

**2.1**  $c_r$ -INEQUALITY.

$$\mathsf{E}(|X+Y|^r) \le c_r [\mathsf{E}(|X|^r) + \mathsf{E}(|Y|^r)] \tag{2.1}$$

where

$$c_r = 1, \quad \text{if } 0 < r \le 1, = 2^{r-1}, \quad \text{if } r > 1.$$
 (2.2)

**2.2** MEAN FORM OF  $c_r$ -INEQUALITY.

$$\begin{split} \mathsf{E}(|\tfrac{1}{2}(X+Y)|^r) & \leq \left(\tfrac{1}{2}\right)^r \left[\mathsf{E}(|X|^r) + \mathsf{E}(|Y|^r)\right], & \text{if } 0 < r \leq 1, \\ & \leq \tfrac{1}{2} \left[\mathsf{E}(|X|^r) + \mathsf{E}(|Y|^r)\right], & \text{if } r > 1. \end{split} \tag{2.3}$$

**2.3** CLOSURE OF  $L_r$ . Let a and b be real numbers. Then

$$X \in L_r \text{ and } Y \in L_r \Rightarrow aX + bY \in L_r$$
. (2.4)

**2.4** HÖLDER INEQUALITY. If r > 1 and  $\frac{1}{r} + \frac{1}{s} = 1$ , then

$$\mathsf{E}(|XY|) \le [\mathsf{E}(|X|^r)]^{1/r} [\mathsf{E}(|Y|^s)]^{1/s} \,. \tag{2.5}$$

2.5 CAUCHY-SCHWARZ INEQUALITY.

$$\mathsf{E}(|XY|) \le [\mathsf{E}(X^2)]^{1/2} [\mathsf{E}(Y^2)]^{1/2} \,. \tag{2.6}$$

**2.6** MINKOWSKI INEQUALITY. If  $r \ge 1$ , then

$$\mathsf{E}(|X+Y|^r)^{1/r} \le [\mathsf{E}(|X|^r)]^{1/r} + [\mathsf{E}(|Y|^r)]^{1/r} \,. \tag{2.7}$$

**2.7** MOMENT MONOTONICITY.  $[\mathsf{E}(|X|^r)]^{1/r}$  is a non-decreasing function of r, *i.e.* 

$$0 < r \le s \Rightarrow [\mathsf{E}(|X|^r)]^{1/r} \le [\mathsf{E}(|X|^s)]^{1/s}. \tag{2.8}$$

**2.8 Theorem** LIAPUNOV THEOREM.  $\log[\mathsf{E}(|X|^r)]$  is a convex function of r, i.e. for any  $\lambda \in [0,1]$ ,

$$\log[\mathsf{E}(|X|^{\lambda r + (1 - \lambda)s})] \le \lambda \log[\mathsf{E}(|X|^r)] + (1 - \lambda) \log[\mathsf{E}(|X|^s)]. \tag{2.9}$$

**2.9** Lower bounds on the moments of a sum. If  ${\sf E}(|X|^r)<\infty$ ,  ${\sf E}(|Y|^r)<\infty$  and  ${\sf E}(Y\,|X)=0$ , then

$$E(|X + Y|^r) \ge E(|X|^r), \text{ for } r \ge 1.$$
 (2.10)

**2.10** JENSEN INEQUALITY. If g(x) is a convex function on  $\mathbb{R}$  and  $\mathsf{E}(|X|) < \infty$ , then, for any constant  $c \in \mathbb{R}$ ,

$$g(c) \le \mathsf{E}[g(X - EX + c)] \tag{2.11}$$

and, in particular,

$$g(EX) \le \mathsf{E}[g(X)] \,. \tag{2.12}$$

### 3. Markov-type inequalities

**3.1** MARKOV INEQUALITIES. Let  $g:\mathbb{R}\to\mathbb{R}$  be a function such that g(X) is a real random variable,  $\mathsf{E}(|g(X)|)<\infty$  and

$$P[0 \le g(X) \le M] = 1 \tag{3.1}$$

where  $M \in [0, \infty]$ . If g(x) is a non-decreasing function on  $\mathbb{R}$ , then, for all  $a \in \mathbb{R}$ ,

$$\frac{\mathsf{E}\left[g\left(X\right)\right] - g\left(a\right)}{M} \le \mathsf{P}[X \ge a] \le \frac{\mathsf{E}\left[g\left(X\right)\right]}{g\left(a\right)}.\tag{3.2}$$

If g(x) is a non-decreasing function on  $[0,\infty)$  and g(x)=g(-x) for any x, then, for all  $a\geq 0$ ,

$$\frac{\mathsf{E}\left[g\left(X\right)\right] - g\left(a\right)}{M} \le \mathsf{P}[|X| \ge a] \le \frac{\mathsf{E}\left[g\left(X\right)\right]}{g\left(a\right)} \tag{3.3}$$

where  $0/0 \equiv 1$ .

**3.2** CHEBYSHEV INEQUALITIES. If  $P[|X| \le M] = 1$ , where  $M \in [0, \infty]$ , then, for all  $a \ge 0$ ,

$$\frac{\mathsf{E}\left(\left|X\right|^{r}\right) - a^{r}}{M^{r}} \le \mathsf{P}[\left|X\right| \ge a] \le \frac{\mathsf{E}\left(\left|X\right|^{r}\right)}{a^{r}}.\tag{3.4}$$

**3.3 Theorem** REFINED MARKOV INEQUALITIES. Let  $g : \mathbb{R} \to \mathbb{R}$  be a function such that g(X) is a real random variable,  $\mathsf{E}(|g(X)|) < \infty$  and

$$0 \leq g(x) \leq M_U \text{ for } x \geq A_U, \tag{3.5}$$

$$0 \leq g(x) \leq M_L \text{ for } x \leq A_L, \tag{3.6}$$

where  $0 \le M_U \le \infty, 0 \le M_L \le \infty, 0 \le A_U \le \infty$  and  $0 \le A_L \le \infty$ . Let also

$$C_U(g, a) = \int_{[a, \infty)} g(x) dF_X(x), \quad C_L(g, a) = \int_{(-\infty, a]} g(x) dF_X(x).$$
 (3.7)

(a) If g(x) is nondecreasing on  $[A_U, \infty)$ , then, for  $a \ge A_U$ ,

$$\frac{C_U(g, a)}{M_U} \le \mathsf{P}[X \ge a] \le \frac{C_U(g, a)}{g(a)}. \tag{3.8}$$

(b) If g(x) is nonincreasing on  $(-\infty, A_L]$ , then, for  $a \leq A_L$ ,

$$\frac{C_L(g, a)}{M_L} \le \mathsf{P}[X \le a] \le \frac{C_L(g, a)}{g(a)}. \tag{3.9}$$

(c) If g(x) is nondecreasing on  $[A_U, \infty)$  and nonincreasing on  $(-\infty, A_L]$ , then, for  $a \ge 1$ 

 $\max\{|A_U|, |A_L|\},\$ 

$$P[|X| \ge a] \le \frac{C_U(g, a)}{g(a)} + \frac{C_L(g, a)}{g(-a)}$$

$$\le \frac{C_U(g, a) + C_L(g, a)}{\min\{g(a), g(-a)\}},$$
(3.10)

$$P[|X| \ge a] \ge \frac{C_U(g, a)}{M_U} + \frac{C_L(g, a)}{M_L} \\ \ge \frac{C_U(g, a) + C_L(g, a)}{\max\{M_U, M_L\}}.$$
(3.11)

### 4. Moments and behavior of tail areas

**4.1 Proposition** Moment existence and tail area decay. Let r>0. If  $\mathsf{E}(|X|^r)<\infty$ , then

$$\lim_{x \to \infty} \left\{ x^r \, \mathsf{P} \left[ X \ge x \right] \right\} \quad = \quad \lim_{x \to -\infty} \left\{ \left| x \right|^r \, \mathsf{P} \left[ X \le x \right] \right\}$$

$$= \quad \lim_{x \to \infty} \left\{ x^r \, \mathsf{P} \left[ \left| X \right| \ge x \right] \right\} = 0 \,. \tag{4.1}$$

In particular, if  $E(|X|) < \infty$ , then

$$\lim_{x \to \infty} \{ x \, \mathsf{P} \, [X \ge x] \} = \lim_{x \to -\infty} \{ |x| \, \mathsf{P} \, [X \le x] \}$$

$$= \lim_{x \to \infty} \{ x \, \mathsf{P} \, [|X| \ge x] \} = 0 \, . \tag{4.2}$$

**4.2 Theorem** DISTRIBUTION DECOMPOSITION OF r-MOMENTS. For any r > 0,

$$\int_0^\infty x^r dF_X(x) = r \int_0^\infty x^{r-1} [1 - F_X(x)] dx, \qquad (4.3)$$

$$\mathsf{E}(|X|^r) = r \int_0^\infty x^{r-1} \mathsf{P}(|X| \ge x) dx$$
$$= r \int_0^\infty x^{r-1} [1 - F_X(x) + F_X(-x)] dx \,, \tag{4.4}$$

and

$$\mathsf{E}(|X|^r) < \infty \Leftrightarrow x^{r-1}\mathsf{P}(|X| \ge x) \text{ is integrable on } (0, +\infty)$$

$$\Leftrightarrow |x|^{r-1} \left[1 - F_X(x) + F_X(-x)\right] \text{ is integrable on } (0, +\infty)$$

$$\Leftrightarrow \int_0^\infty x^{r-1} [1 - F_X(x)] dx < \infty \text{ and } \int_{-\infty}^0 |x|^{r-1} F_X(x) dx < \infty. \tag{4.5}$$

4.3 Proposition DISTRIBUTION DECOMPOSITION OF THE FIRST ABSOLUTE MOMENT.

$$\int_0^\infty x dF_X(x) = \int_0^\infty [1 - F_X(x)] dx,$$
 (4.6)

$$\mathsf{E}|X| = \int_0^\infty \mathsf{P}(|X| \ge x) dx \,, \tag{4.7}$$

and

$$\mathsf{E}(|X|) < \infty \Leftrightarrow \mathsf{P}(|X| \ge x) \text{ is integrable on } (0, +\infty)$$

$$\Leftrightarrow [1 - F_X(x) + F_X(-x)] \text{ is integrable on } (0, +\infty)$$

$$\Leftrightarrow \int_0^\infty [1 - F_X(x)] dx < \infty \text{ and } \int_{-\infty}^0 F_X(x) dx < \infty.$$

$$(4.8)$$

**4.4 Proposition** MOMENT-TAIL AREA INEQUALITIES. Let g(x) be a nonnegative strictly increasing function on  $[0, \infty)$  and let  $g^{-1}(x)$  be the inverse function of g. Then,

$$\sum_{n=1}^{\infty} P[|X| \ge g^{-1}(n)] \le E[g(X)] \le \sum_{n=0}^{\infty} P[|X| > g^{-1}(n)].$$
 (4.9)

In particular, for any r > 0,

$$\sum_{n=1}^{\infty} \mathsf{P}(|X| \ge n^{1/r}) \le \mathsf{E}(|X|^r) \le \sum_{n=0}^{\infty} \mathsf{P}(|X| > n^{1/r})$$

$$\le 1 + \sum_{n=1}^{\infty} \mathsf{P}(X > n^{1/r}). \tag{4.10}$$

**4.5** Corollary MEAN-TAIL AREA INEQUALITIES. If X is a positive random variable,

$$\sum_{n=1}^{\infty} P(X \ge n) \le E(X) \le 1 + \sum_{n=1}^{\infty} P(X > n).$$
 (4.11)

### 5. Moments of sums of random variables

In this section, we consider a sequence  $X_1, \ldots, X_n$  of random variables, and study the moments of the corresponding sum and average:

$$S_n = \sum_{i=1}^n X_i , \quad \bar{X}_n = S_n/n .$$
 (5.1)

**5.1 Proposition** BOUNDS ON THE ABSOLUTE MOMENTS OF A SUM OF RANDOM VARIABLES.

$$\mathsf{E}(|S_n|^r) \leq \sum_{i=1}^n \mathsf{E}(|X_i|^r), \quad \text{if } 0 < r \leq 1, 
\leq n^{r-1} \sum_{i=1}^n \mathsf{E}(|X_i|^r), \quad \text{if } r > 1,$$
(5.2)

and

$$\mathsf{E}(|\bar{X}_{n}|^{r}) \leq \left(\frac{1}{n}\right)^{r} \sum_{i=1}^{n} \mathsf{E}(|X_{i}|^{r}), \quad \text{if } 0 < r \leq 1, 
\leq \frac{1}{n} \sum_{i=1}^{n} \mathsf{E}(|X_{i}|^{r}), \quad \text{if } r > 1.$$
(5.3)

**5.2 Proposition** MINKOWSKI INEQUALITY FOR n VARIABLES. If  $r \ge 1$ , then

$$[\mathsf{E}(|S_n|^r)]^{1/r} \le \sum_{i=1}^n [\mathsf{E}(|X_i|^r)]^{1/r} \tag{5.4}$$

and

$$[\mathsf{E}(|\bar{X}_{n}|^{r})]^{1/r} \leq \frac{1}{n} \sum_{i=1}^{n} [\mathsf{E}(|X_{i}|^{r})]^{1/r}$$

$$\leq \left\{ \frac{1}{n} \sum_{i=1}^{n} \mathsf{E}(|X_{i}|^{r}) \right\}^{1/r}. \tag{5.5}$$

**5.3 Proposition** BOUNDS ON THE ABSOLUTE MOMENTS OF A SUM OF RANDOM VARIABLES UNDER CONDITIONAL SYMMETRY. If the distribution of  $X_{k+1}$  given  $S_i$  is symmetric about zero for  $k=1,\ldots,n-1$ , and  $\mathsf{E}(|X_i|^r)<\infty, i=1,\ldots,n$ , then

$$\mathsf{E}(|S_n|^r) \le \sum_{i=1}^n \mathsf{E}(|X_i|^r) \quad \text{for } 1 \le r \le 2,$$
 (5.6)

and

$$\mathsf{E}(|\bar{X}_n|^r) \le \left(\frac{1}{n}\right)^r \sum_{i=1}^n \mathsf{E}(|X_i|^r) \quad \text{for } 1 \le r \le 2,$$
 (5.7)

with equality holding when r = 2.

**5.4 Proposition** Bounds on the absolute moments of a sum of random variables under martingale condition. *If* 

$$\mathsf{E}(X_{k+1} \mid S_k) = 0$$
 a.s.,  $k = 1, \dots, n-1,$  (5.8)

and  $E(|X_i|^r) < \infty, i = 1, \ldots, n$ , then

$$\mathsf{E}(|S_n|^r) \le 2\sum_{i=1}^n \mathsf{E}(|X_i|^r), \quad \text{for } 1 \le r \le 2,$$
 (5.9)

and

$$\mathsf{E}(|\bar{X}_n|^r) \le 2\left(\frac{1}{n}\right)^r \sum_{i=1}^n \mathsf{E}(|X_i|^r), \quad \text{for } 1 \le r \le 2.$$
 (5.10)

Furthermore, for r=2,

$$\mathsf{E}(S_n^2) = \sum_{i=1}^n \mathsf{E}(X_i^2). \tag{5.11}$$

**5.5 Proposition** Bounds on the absolute moments of a sum of random variables under two-sided martingale condition. *Let* 

$$S_{m(k)} = \sum_{i=1, i \neq k}^{m+1} X_i, \quad 1 \le k \le m+1 \le n.$$
 (5.12)

If

$$\mathsf{E}(X_k \mid S_{m(k)}) = 0$$
 a.s., for  $1 \le k \le m + 1 \le n$ , (5.13)

and  $E(|X_i|^r) < \infty, i = 1, \ldots, n$ , then

$$\mathsf{E}(|S_n|^r) \le \left(2 - \frac{1}{n}\right) \sum_{i=1}^n \mathsf{E}(|X_i|^r) \,, \quad \text{for } 1 \le r \le 2 \,,$$
 (5.14)

and

$$\mathsf{E}(|\bar{X}_n|^r) \le \left(\frac{1}{n}\right)^r \left(2 - \frac{1}{n}\right) \sum_{i=1}^n \mathsf{E}(|X_i|^r), \quad \text{for } 1 \le r \le 2.$$
 (5.15)

**5.6 Proposition** BOUNDS ON THE ABSOLUTE MOMENTS OF A SUM OF INDEPENDENT RANDOM VARIABLES. Let the random variables  $X_1, \ldots, X_n$  be independent with  $\mathsf{E}(X_i) = 0$  and  $\mathsf{E}(|X_i|^r) < \infty, i = 1, \ldots, n$ , and let

$$D(r) = [13.52/(2.6\pi)^r] \Gamma(r) \sin(r\pi/2).$$
(5.16)

If D(r) < 1 and  $1 \le r \le 2$ , then

$$\mathsf{E}(|S_n|^r) \le [1 - D(r)]^{-1} \sum_{i=1}^n \mathsf{E}(|X_i|^r), \tag{5.17}$$

and

$$\mathsf{E}(|\bar{X}_n|^r) \le \left(\frac{1}{n}\right)^r [1 - D(r)]^{-1} \sum_{i=1}^n \mathsf{E}(|X_i|^r) \,, \quad \text{for } 1 \le r \le 2 \,. \tag{5.18}$$

### 6. Proofs

- **1.1** to **3.2**. See Loève (1977, Volume I, Sections 9.1 and 9.3, pp. 151-162). For Jensen inequality, see also Chow and Teicher (1988, Section 4.3, pp. 103-106). Hannan (1985), Lehmann and Shaffer (1988), Piegorsch and Casella (1988) and Khuri and Casella (2002) discussed conditions for the existence of the moments of 1/X.
- **2.9**. See von Bahr and Esseen (1965, Lemma 3).

PROOF OF THEOREM 3.3 (a) For  $x \ge a \ge A_U$ , we have  $g(x) \ge g(a)$  and  $g(x) \le M_U$ , hence

$$C_{U}(g, a) = \int_{[a, \infty)} g(x) dF_{X}(x) \ge g(a) \int_{[a, \infty)} dF_{X}(x) = g(a) P[X \ge a]$$

and

$$\int_{[a,\infty)} g(x) dF_X(x) \le M_U P[X \ge a],$$

from which we get the inequality

$$\frac{C_U(g, a)}{M_U} \le P[X \ge a] \le \frac{C_U(g, a)}{g(a)}.$$

(b) For  $x \le a \le A_L$ , we have  $g(x) \ge g(a)$  and  $g(x) \le M_L$ , hence

$$C_L(g, a) = \int_{[-\infty, a)} g(x) dF_X(x) \ge g(a) \int_{[-\infty, a)} dF_X(x) = g(a) P[X \le a]$$

and

$$\int_{[-\infty,a)} g(x) dF_X(x) \le M_L P[X \le a]$$

from which we get the inequality

$$\frac{C_L(g, a)}{M_L} \le P[X \le a] \le \frac{C_L(g, a)}{g(a)}.$$

(c) For  $a \ge \max(|A_U|, |A_L|)$ , we have  $a \ge A_U$  and  $-a \le A_L$ , hence

$$P[|X| \ge a] = P[X \ge a] + P[X \le -a]$$

$$\le \frac{C_U(g, a)}{g(a)} + \frac{C_L(g, a)}{g(a)}$$

$$\leq \frac{C_U(g, a) + C_L(g, a)}{\min\{q(a), q(-a)\}}$$

and

$$P[|X| \ge a] \ge \frac{C_U(g, a)}{M_U} + \frac{C_L(g, a)}{M_L}$$
$$\ge \frac{C_U(g, a) + C_L(g, a)}{\max(M_U, M_L)}.$$

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- **4.1** to **4.3**. See Feller (1966, Section V.6, Lemma 1), Chung (1974, Exercises 17-18), Serfling (1980, Section 1.14, pp. 46-47) and Chow and Teicher (1988, Section 4.3, pp. 103-106). For other inequalities involving absolute moments, the reader may consult Beesack (1984).
- **4.4**. See Chow and Teicher (1988, Section 4.1, Corollary 3, p. 90).
- **4.5**. The inequality (4.11) is given by Chung (1974, Theorem 3.2.1) and Serfling (1980, Section 1.3, p. 12).
- **5.1**. See von Bahr and Esseen (1965), Chung (1974, p. 48) and Chow and Teicher (1988, p. 108).
- **5.2**. See Chung (1974, p. 48).

PROOF OF PROPOSITION **5.2** The first inequality follows by recursion on applying the Minkowski inequality for two variables. The first part of the second inequality is obtained by multiplying both sides of the first one by (1/n). The second part follows on observing that the function  $x^{1/r}$  is concave in x for x > 0 when r > 1.

- **5.3**. See von Bahr and Esseen (1965, Theorem 1).
- **5.4**. See von Bahr and Esseen (1965, Theorem 2).
- **5.5**. See von Bahr and Esseen (1965, Theorem 3).
- **5.6**. See von Bahr and Esseen (1965, Theorem 4).

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