Appendix A: Models and Simulations to Evaluate the Consequences of Model Structure for Omitted Variable Bias

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The System and the Study

To evaluate the impact of Omitted Variable Bias (OVB) on different models, consider a system where oceanography drives site temperature and recruitment over time. Temperature also fluctuates over time within each site. Recruitment and temperature influence snail abundance, as do other uncorrelated drivers. You have conducted a study in this system measuring 10 sites, each site sampled once per year over 10 years. In this study, you have recorded snail abundance and temperature. But have no measure of recruitment. Note, the results of models will be the same if you had instead sampled in one year across 10 sites with 10 plots per site if there were plot-level drivers of temperature that behaved the same as below.

To parameterize our simulations, consider the following:

- Our Oceanography latent variable has a mean of 0 and a SD of 1.
- Site temperature is twice the oceanography variable and transformed to have a mean of 15C.
- Site recruitment is -2 multiplied by the oceanography variable and transformed to have a mean of 15 individuals per plot.
- There is additional random variation between sites with a mean of 0 and SD of 1.
- Sites are sampled over 10 years.
- Within a site, the temperature varies over time according to a normal distribution with a mean of 1.
- There is a 1:1 relationship between temperature and snail abundance and recruitment and snails.
- Other non-correlated drivers in the system influence snail abundance with a mean influence of 0 and a SD of 1.

Thus, the system looks like this:

Functions to Create The System

To simulated data, let's begin by loading some libraries

```
library(tidyverse)
library(lme4)
library(broom)
library(broom.mixed)
library(DiagrammeR)
library(glue)

theme_set(theme_bw(base_size = 14))
```

Next, we need a function that will create a template of simulated sites based on oceanography and the sampling design described above.

```
make_environment <- function(n_sites = 10,</pre>
                              ocean_temp = 2,
                              temp sd = 0,
                              ocean_recruitment = -2,
                              recruitment sd = 0,
                              temp_mean = 15,
                              rec_mean = 10,
                              site_sd = 1,
                              seed = NULL){
  if(!is.null(seed)) set.seed(seed)
  tibble(
    site = as.character(1:n_sites)) %>%
      oceanography = rnorm(n sites),
      site temp = temp mean +
        rnorm(n_sites, ocean_temp * oceanography, temp_sd),
      site recruitment = rec mean +
        rnorm(n_sites, ocean_recruitment * oceanography, recruitment_sd),
      site_int = rnorm(n_sites, 0, sd = site_sd)
}
Great. Now, we need to add that year-to-year or plot-to-plot variability.
make_plots <- function(sites_df,</pre>
                        n_plots_per_site = 10,
                        plot temp sd = 1,
                        temp effect = 1,
                        recruitment_effect = 1,
                        sd_plot = 2,
                        seed = NULL){
```

```
if(!is.null(seed)) set.seed(seed)
  sites df %>%
    rowwise() %>%
    mutate(
      plot_temp_dev_actual = list(rnorm(n_plots_per_site,
                                        0, plot temp sd))) %>%
    unnest(plot temp dev actual) %>%
    mutate(
      plot temp = site temp + plot temp dev actual,
      snails = rnorm(n(),
                     temp_effect*plot_temp +
                       recruitment effect*site recruitment +
                       site int,
                     sd plot)) %>%
    ungroup() %>%
    group_by(site) %>%
    mutate(year = 1:n(),
           site mean snails = mean(snails),
           site mean temp = mean(plot temp),
           plot_snail_dev = snails - site_mean_snails,
           plot_temp_dev = plot_temp - site_mean_temp,
           site mean_snail = mean(snails),
           site_snail_dev = snails - mean(snails),
           delta snails = snails - lag(snails),
           delta temp = plot temp - lag(plot temp)) %>%
    ungroup()
}
```

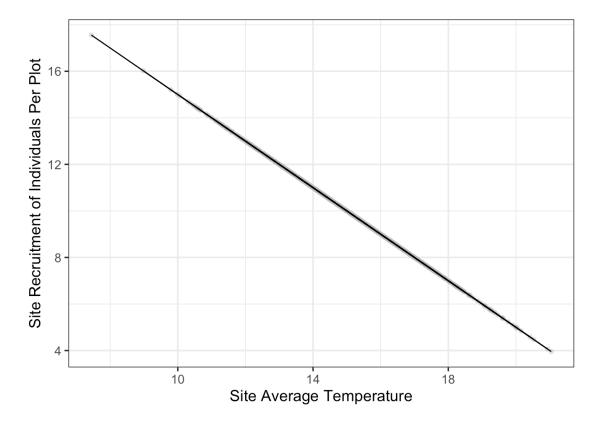
To analyze our data, we will compare several different fit models.

- A naive linear model with no site term
- A random effects using site as a random intercept
- A fixed effects model where site is a fixed effect (i.e., turned into 1/0 dummy variables)
- A model where we include the site mean temperature as a covariate and site is a random effect
- A model where we include site mean temperature as a covariate and site mean centered temperature (i.e., temperature at a site in a year minus it's mean over the entire data set). Site is included as a random effect

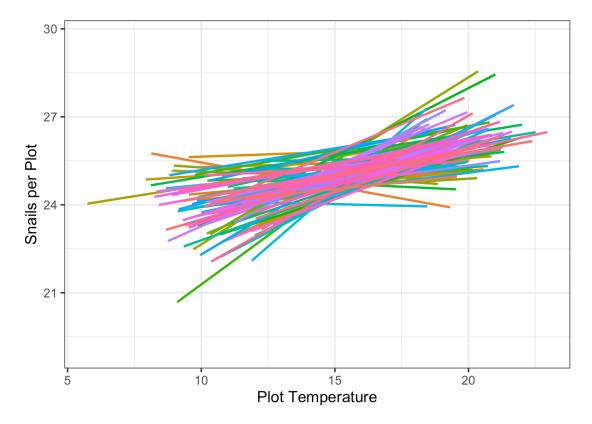
• A panel model where we look at change in snails between years versus change in temperature between years with a random site effect

```
analyze plots <- function(plot df){</pre>
  m <- tribble(</pre>
    ~model_type, ~fit,
    "Naive", lm(snails ~ plot_temp, data = plot_df),
    "RE", lmer(snails ~ plot_temp + (1|site), data = plot_df),
    "FE Using Mean Differencing", lm(plot_snail_dev ~ plot_temp_dev, data = p
lot df), #fix SE?
    "FE with Dummy Variables", lm(snails ~ plot temp + site, data = plot df),
    "Group Mean Covariate", lmer(snails ~ plot temp + site mean temp + (1|sit
e), data = plot df),
    "Group Mean Centered", lmer(snails ~ plot_temp_dev + site_mean_temp + (1
site), data = plot_df),
    "Group Mean Covariate, no RE", lm(snails ~ plot temp + site mean temp, da
ta = plot df),
    "Group Mean Centered, no RE", lm(snails ~ plot temp dev + site mean temp,
data = plot df),
    "First Differences", lm(delta_snails ~ delta_temp,data = plot_df) #fix SE
?
  ) %>%
    mutate(coefs = map(fit, tidy), #get coefficients with broom
           out stats = map(fit, glance),
           temp_effect = map(coefs, get_temp_coef),
           model type = fct inorder(model type))
 m
}
get_temp_coef <- function(a_tidy_frame){</pre>
  a_tidy_frame %>%
    filter(term %in% c("plot temp", "plot temp dev", "delta temp")) %>%
    select(estimate, std.error)
}
Simulations and Results
Let's begin by setting up 100 replicate simulations.
set.seed(31415)
n_sims <- 100
envt <- tibble(</pre>
  sims = 1:n sims
) %>%
  mutate(sites = map(sims, ~make environment()))
```

Just for a sanity check, here's the relationship between temperature and recruitment at the site level across all simulations.



Great! Now, let's setup our sampling over time.



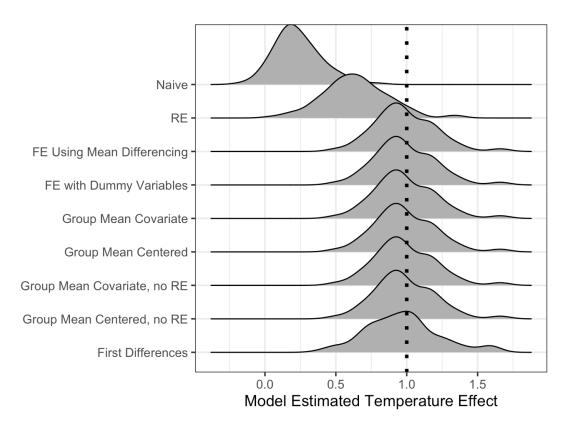
So we can see that in this setup, the snail-temperature relationship is positive in nearly all of the simulations. But, how positive is it? What would our coefficients show on average across simulations?

Let's fit models to each set of data

```
analysis_df <- plots_df %>%
  mutate(analysis = map(site_year, analyze_plots)) %>%
  unnest(analysis)
```

And now let's look at the distribution of coefficients that would describe the relationship between temperature and snails from each mode.

```
analysis_df %>%
  unnest(temp_effect) %>%
  ggplot(aes(y = fct_rev(model_type), x = estimate)) +
  ggridges::stat_density_ridges() +
  labs(y="", x = "Model Estimated Temperature Effect") +
  geom_vline(xintercept = 1, linewidth = 1.5, lty = 3)
```

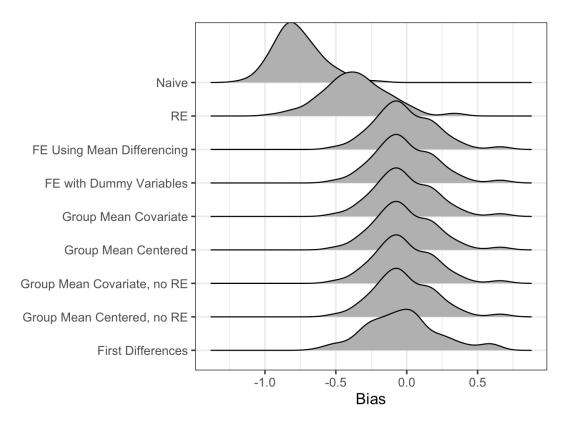


Eyeballing it, we can see of course the naive model is too low, as is the RE model. How bad is the bias for the RE model?

Model Type	Mean Estimate	SD Estimate
Naive	0.231	0.165
RE	0.640	0.232
FE Using Mean Differencing	0.985	0.215
FE with Dummy Variables	0.985	0.215
Group Mean Covariate	0.985	0.215
Group Mean Centered	0.985	0.215
Group Mean Covariate, no RE	0.985	0.215
Group Mean Centered, no RE	0.985	0.215
First Differences	0.971	0.259

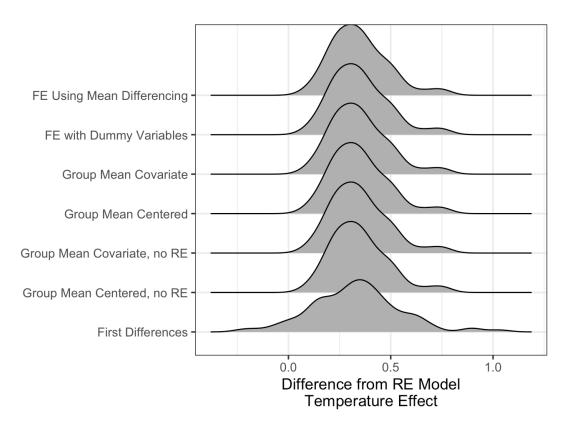
The downward bias produced by poor model choice is clear. We can see it if we plot the coefficient minus one, which if unbiased should reveal a distribution centered on 0.

```
analysis_df %>%
  unnest(temp_effect) %>%
  ggplot(aes(y = fct_rev(model_type), x = estimate - 1)) +
  ggridges::stat_density_ridges() +
  labs(y="", x = "Bias")
```



We can also look at the distribution of, for each simulated data set, how different the RE model is from each other model.

```
analysis_df %>%
  filter(model_type != "Naive") %>%
  unnest(temp_effect) %>%
  group_by(sims) %>%
  mutate(diff_from_re = estimate - estimate[1]) %>%
  ungroup() %>%
  filter(model_type != "RE") %>%
  ggplot(aes(y = fct_rev(model_type), x = diff_from_re)) +
  ggridges::stat_density_ridges() +
  labs(y="", x = "Difference from RE Model\nTemperature Effect")
```



Note, in all cases, we can see the effects of downward bias. This is clear, but, to put it in numbers -

model_type	Mean Diff from RE	SD Diff from RE
FE Using Mean Differencing	0.346	0.138
FE with Dummy Variables	0.346	0.138
Group Mean Covariate	0.346	0.138
Group Mean Centered	0.346	0.138
Group Mean Covariate, no RE	0.346	0.138
Group Mean Centered, no RE	0.346	0.138
First Differences	0.331	0.218

If we were doing straight hypothesis testing, how often would our estimate of the temperature coefficient either overlap 0 or not have 1 within its confidence interval?

Model Type	95% CI Contains 0	95% CI does Not Contain 1
RE	0.08	0.54
FE Using Mean Differencing	0.00	0.05
FE with Dummy Variables	0.00	0.05
Group Mean Covariate	0.00	0.05
Group Mean Centered	0.00	0.05
Group Mean Covariate, no RE	0.01	0.04
Group Mean Centered, no RE	0.01	0.04
First Differences	0.01	0.12

Here we see the RE model is more likely to be subject to type II error. Further, it is far more likely than any other technique to not have the true coefficient value within 2 CI of its estimand.

A Wrapper Function for Simulation

This has been useful, but, if we want to automate the process for further exploration, let's wrap the code above into a function.

```
seed = NULL) {
  #should we set a seed?
  if (!is.null(seed))
    set.seed(seed)
  # make an envt data frame
  out df <- tibble(sims = 1:n sims) %>%
    mutate(
      sites = map(
        sims,
        ~make_environment(
         n_sites = n_sites,
          ocean temp = ocean temp,
          temp_sd = temp_sd,
          ocean recruitment = ocean recruitment,
          recruitment_sd = recruitment_sd,
          temp_mean = temp_mean,
          rec mean = rec mean,
          site_sd = site_sd)
      )
    ) %>%
    #now add plots
    mutate(
      site year = map(
        sites,
        make plots,
        n_plots_per_site = n_plots_per_site,
        plot_temp_sd = plot_temp_sd,
        temp_effect = temp_effect,
        recruitment effect = recruitment effect,
        sd plot = sd plot
      )
    ) %>%
    #and analysis
    mutate(analysis = map(site_year, analyze_plots)) %>%
    unnest(analysis)
}
Is that Random Effect Needed?
If we look at the correlated random effects models, what's the RE?
analysis df %>%
  filter(model_type %in% c("Group Mean Covariate", "Group Mean Centered")) %>
  unnest(coefs) %>%
  filter(group == "site") %>%
  group_by(`Model Type` = model_type) %>%
```

Model Type	Term	Mean Site SD	SD in Site SD
Group Mean Covariate	Site Random Effect	0.977	0.438
Group Mean Centered	Site Random Effect	0.977	0.438

Both are the same, which makes sense given the formulation of the model. What if there was no additional site-level variation uncorrelated with temperature, though?

Model Type	Term	Mean Site SD	SD in Site SD
Group Mean Covariate	Site Random Effect	0.273	0.261
Group Mean Centered	Site Random Effect	0.273	0.261

Both REs are the same - again - but both overlap 0. In the system we simulated, there is no uncorrelated site-level variability. It's all at the site-year (or site-plot) level. We have indeed run these models with *no site random effect*. They produce the same answers for coefficients - both in this simulation with no site-level random effect as well as above when there *was* a site-level random effect.

Model Type	Mean Temp Effect	SD Temp Effect
Group Mean Covariate	0.969	0.203
Group Mean Centered	0.969	0.203

Model Type	Mean Temp Effect	SD Temp Effect	
Group Mean Covariate, no RE	0.969	0.203	
Group Mean Centered, no RE	0.969	0.203	
# With random site variation	on		
analysis_df %>%			
filter(grepl("Group Mean'	', model_type)) %	> %	
<pre>unnest(temp_effect) %>%</pre>			
<pre>group_by(`Model Type` = model_type) %>%</pre>			
summarize(`Mean Temp Effe	•		
`SD Temp Effect	i = sd(estimate))	%>%	
knit_table			

Model Type	Mean Temp Effect	SD Temp Effect
Group Mean Covariate	0.985	0.215
Group Mean Centered	0.985	0.215
Group Mean Covariate, no RE	0.985	0.215
Group Mean Centered, no RE	0.985	0.215

So should we just not worry about a site-level random effect? Not necesarily. First, if we consider a site RE, we can look at residual variation due to both between site differences as well as residual replicate-level variation. This can be a useful analysis when attempting to tease apart variation that is versus is not correlated with a driver of interest. Second, while models without a site random effect can be fit and used, these models will be structurally incorrect - replicates are not IID, as there is correlated error within sites. Last, mixed models can also provide advantages when handling models with unbalanced data between sites (see below).

Practically, however, the difference comes in with respect to whether you are interested in between site variation or not. The residual inflates with no RE, as it is now the combination of between site residual and within site residual terms. This could make a difference for various statistical tests down the line as well, but in terms of parameter estimates, we are still estimating a clean causal effect.

What About Unbalanced Data?

One of the advantages to mixed modeling approaches is how they handle unbalanced data. What if, in the above example, we had lost samples from each site generating unbalanced data?

How does this lack of balance affect estimation of causal coefficients?

Model Type	Mean Estimate	SD Estimate
Naive	0.245	0.195
RE	0.500	0.297
FE Using Mean Differencing	0.986	0.275
FE with Dummy Variables	0.990	0.281
Group Mean Covariate	0.986	0.281
Group Mean Centered	0.986	0.281
Group Mean Covariate, no RE	0.982	0.299
Group Mean Centered, no RE	0.982	0.299
First Differences	0.961	0.316

We can see that point estimation here is improved somewhat. Although this could vary. More telling would be an exploration of those group means in the mixed versus fixed models.