

HW2

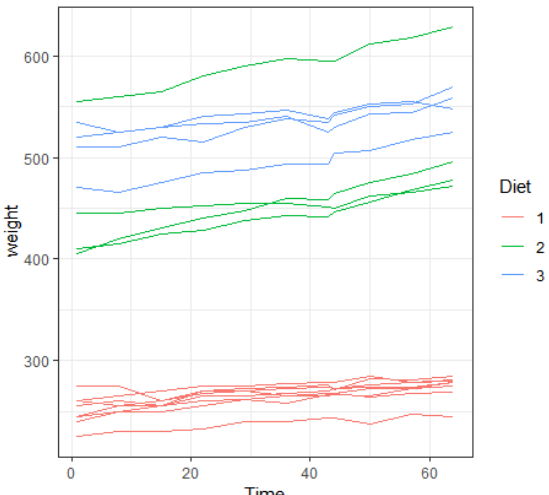
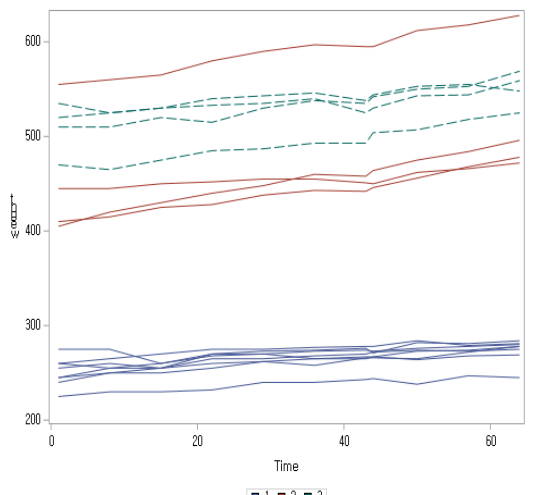
임상시험자료분석 II

182STG27

임지연

1. R의 nlme package 에 내장되어 있는 BodyWeight 자료를 이용하시오.

1) 적절한 그림으로 자료를 살펴보세요.

R	SAS
1. CODE	
<pre>library(nlme); library(faraway) library(ggplot2); library(gridExtra) library(gmodels); library(car) library(reshape); library(dplyr) str(BodyWeight) theme_set(theme_bw()) ggplot(BodyWeight, aes(x = Time, y = weight, colour = Diet)) + geom_line(aes(group = Rat))</pre>	<pre>data BodyWeight; infile "C:\Users\Wjeeyeon\Desktop\data\BodyWeight.csv" delimiter=',' firstobs=2; input weight Time Rat Diet; week = compress('W' Time, ''); run; proc sgplot data=BodyWeight; title 'BodyWeight'; series x=time y=weight / group=Rat groupcl=Diet name='grouping'; keylegend 'grouping' / type=linecolor; run;</pre>
2. PLOT	
	
3. RESULT	
<p>BodyWeight 자료는 세 그룹의 쥐를 각각 다른식이 요법으로 처리하여 1일 ~ 7 일까지 64 일 동안 측정 된 쥐의 체중 (44 일에는 추가 측정)을 나타낸 자료이다. 이 자료의 변수설명은 다음과 같다.</p> <ul style="list-style-type: none"> - weight : numeric vector -giving the body weight of the rat - time : numeric vector -giving time at which the measurement is made - Rat : ordered factor - Diet : a factor level 1 to 3 <p>즉, 이 데이터에서 Diet - Trt, Rat - Patient , Time - time, Weight - Score 에 해당한다. 우리의 목적은 각각 다른 식이요법간에 (Trt- Diet) 쥐의 Weight 차이가 있는지를 살펴보는 것이다. 따라서, 각 그룹을 Diet 1,2,3으로 나누어 그래프를 그려본 결과 1 그룹의 Weight는 300 미만으로 현저하게 낮게 나타난 반면, 2,3 그룹의 Weight는 400~ 600 사이이며, 2 그룹의 분산이 더 커보인다. 그래프로 보서는 Diet 1 약의 효과가 있을 것이라고 예상할 수 있다.</p>	

2) 적절한 방법으로 분석하여 결과를 해석하시오.

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<pre># Anova ----- lm1 <- lm(weight ~ factor(Diet) + factor(Rat) + factor(Time) +factor(Diet) * factor(Time), data = BodyWeight) temp <- anova(lm1); temp F.group <- temp[1,3] / temp[2,3] # F.value Pgroup <- 1-pf(F.group,temp[1,1], temp[2,1]) c(F.group, Pgroup) lme1 <- lme(weight ~ factor(Diet) + factor(Time) +factor(Diet) * factor(Time), random = ~1 factor(Rat),data = BodyWeight) anova(lme1) # Manova ----- head(BodyWeight) BodyWeight <- BodyWeight %>% mutate(Week = paste("W",BodyWeight\$Time, sep="")) BodyWeight_M <- cast(BodyWeight , Diet + Rat ~ Week, value="weight") lm2 <- lm(cbind(W1, W8, W15, W22, W29, W36, W43, W44, W50, W57, W64) ~ Diet ,data = BodyWeight_M) measure_time <- factor(c("W1", "W8", "W15", "W22", "W29", "W36", "W43", "W44", "W50","W57","W64")) measure_time_data <- data.frame(measure_time = measure_time) mv1 <- Anova(lm2, idata = measure_time_data, idesign = ~ measure_time) summary(mv1)</pre>	<pre>/* ANOVA -----*/ PROC GLM data=BodyWeight; CLASS Diet Rat Time; MODEL Weight = Diet Rat(Diet) Time Diet*Time / ss3; RANDOM RAT(Diet); TEST H = Diet E = Rat(Diet); QUIT; RUN; /* MANOVA -----*/ proc sort data=BodyWeight; by Diet Rat; run; proc transpose data=BodyWeight out=Bodyweight_M; by Diet Rat; id Week; var weight; run; ods exclude partialCorr ErrorSSCP; proc glm data = Bodyweight_M; class Diet; model W1 W8 W15 W22 W29 W36 W43 W44 W50 W57 W64 = Diet / ss3; repeated time profile /printe summary; quit; run;</pre>																																																																																																																													
2. TABLE																																																																																																																														
<pre># ANOVA TABLE > temp <- anova(lm1); temp Analysis of Variance Table Response: weight</pre> <table><tr><th></th><th>Df</th><th>Sum Sq</th><th>Mean Sq</th><th>F value</th><th>Pr(>F)</th></tr><tr><td>factor(Diet)</td><td>2</td><td>2604050</td><td>1302025</td><td>32738.2082</td><td>< 2.2e-16 ***</td></tr><tr><td>factor(Rat)</td><td>13</td><td>192186</td><td>14784</td><td>371.7177</td><td>< 2.2e-16 ***</td></tr><tr><td>factor(Time)</td><td>10</td><td>23400</td><td>2340</td><td>58.8378</td><td>< 2.2e-16 ***</td></tr><tr><td>factor(Diet):factor(Time)</td><td>20</td><td>4906</td><td>245</td><td>6.1679</td><td>2.883e-11 ***</td></tr><tr><td>Residuals</td><td>130</td><td>5170</td><td>40</td><td></td><td></td></tr></table> <pre>--- Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 > c(F.group, P.group) [1] 88.07276434383193 0.00000002763486 > anova(lme1)</pre> <table><tr><th></th><th>numDF</th><th>denDF</th><th>F-value</th><th>p-value</th></tr><tr><td>(Intercept)</td><td>1</td><td>130</td><td>1759.9049</td><td><.0001</td></tr><tr><td>factor(Diet)</td><td>2</td><td>13</td><td>88.0728</td><td><.0001</td></tr><tr><td>factor(Time)</td><td>10</td><td>130</td><td>58.8378</td><td><.0001</td></tr><tr><td>factor(Diet):factor(Time)</td><td>20</td><td>130</td><td>6.1679</td><td><.0001</td></tr></table>		Df	Sum Sq	Mean Sq	F value	Pr(>F)	factor(Diet)	2	2604050	1302025	32738.2082	< 2.2e-16 ***	factor(Rat)	13	192186	14784	371.7177	< 2.2e-16 ***	factor(Time)	10	23400	2340	58.8378	< 2.2e-16 ***	factor(Diet):factor(Time)	20	4906	245	6.1679	2.883e-11 ***	Residuals	130	5170	40				numDF	denDF	F-value	p-value	(Intercept)	1	130	1759.9049	<.0001	factor(Diet)	2	13	88.0728	<.0001	factor(Time)	10	130	58.8378	<.0001	factor(Diet):factor(Time)	20	130	6.1679	<.0001	<pre># ANOVA TABLE The GLM Procedure Dependent Variable: weight</pre> <table><tr><th>Source</th><th>DF</th><th>Sum of Squares</th><th>Mean Square</th><th>F Value</th><th>Pr > F</th></tr><tr><td>Model</td><td>45</td><td>2824541.744</td><td>62767.594</td><td>1578.23</td><td><.0001</td></tr><tr><td>Error</td><td>130</td><td>5170.205</td><td>39.771</td><td></td><td></td></tr><tr><td>Corrected Total</td><td>175</td><td>2829711.949</td><td></td><td></td><td></td></tr></table> <table><tr><th></th><th>R-Square</th><th>Coeff Var</th><th>Root MSE</th><th>weight Mean</th></tr><tr><td></td><td>0.998173</td><td>1.640231</td><td>6.306410</td><td>384.4830</td></tr></table> <table><tr><th>Source</th><th>DF</th><th>Type III SS</th><th>Mean Square</th><th>F Value</th><th>Pr > F</th></tr><tr><td>Diet</td><td>2</td><td>2604049.733</td><td>1302024.866</td><td>32738.2</td><td><.0001</td></tr><tr><td>Rat(Diet)</td><td>13</td><td>192185.670</td><td>14783.513</td><td>371.72</td><td><.0001</td></tr><tr><td>Time</td><td>10</td><td>27000.205</td><td>2700.020</td><td>67.89</td><td><.0001</td></tr><tr><td>Diet*Time</td><td>20</td><td>4906.080</td><td>245.304</td><td>6.17</td><td><.0001</td></tr></table> <pre>The GLM Procedure Source Type III Expected Mean Square Diet Var(Error) + 11 Var(Rat(Diet)) + Q(Diet,Diet*Time) Rat(Diet) Var(Error) + 11 Var(Rat(Diet)) Time Var(Error) + Q(Time,Diet*Time) Diet*Time Var(Error) + Q(Diet*Time) The GLM Procedure Dependent Variable: weight Tests of Hypotheses Using the Type III MS for Rat(Diet) as an Error Term Source DF Type III SS Mean Square F Value Pr > F Diet 2 2604049.733 1302024.866 88.07 <.0001</pre>	Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	Model	45	2824541.744	62767.594	1578.23	<.0001	Error	130	5170.205	39.771			Corrected Total	175	2829711.949					R-Square	Coeff Var	Root MSE	weight Mean		0.998173	1.640231	6.306410	384.4830	Source	DF	Type III SS	Mean Square	F Value	Pr > F	Diet	2	2604049.733	1302024.866	32738.2	<.0001	Rat(Diet)	13	192185.670	14783.513	371.72	<.0001	Time	10	27000.205	2700.020	67.89	<.0001	Diet*Time	20	4906.080	245.304	6.17	<.0001
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Diet , Time 변수의 p-value < 0.0001로 두 변수가 모두 유의하다고 볼 수 있다.																																																																																																																														

MANOVA TABLE

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Multivariate Tests: Diet
  Df test stat approx F num Df den Df Pr(>F)
Pillai 2 0.93127 88.07276 2 13 0.000000027635 ***
Wilks 2 0.06873 88.07276 2 13 0.000000027635 ***
Hotelling-Lawley 2 13.54966 88.07276 2 13 0.000000027635 ***
Roy 2 13.54966 88.07276 2 13 0.000000027635 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Multivariate Tests: measure_time
  Df test stat approx F num Df den Df Pr(>F)
Pillai 1 0.97831 18.04504 10 4 0.0066562 **
Wilks 1 0.02169 18.04504 10 4 0.0066562 **
Hotelling-Lawley 1 45.11260 18.04504 10 4 0.0066562 **
Roy 1 45.11260 18.04504 10 4 0.0066562 **
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Multivariate Tests: Diet:measure_time
  Df test stat approx F num Df den Df Pr(>F)
Pillai 2 1.82897 5.346895 20 10 0.0047364 **
Wilks 2 0.00327 6.590314 20 8 0.0050459 **
Hotelling-Lawley 2 50.23337 7.535006 20 6 0.0095215 **
Roy 2 44.52489 22.262445 10 5 0.0015751 **
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Mauchly Tests for Sphericity

          Test statistic              p-value
measure_time 0.000000068081 0.0000000005357
Diet:measure_time 0.000000068081 0.0000000005357

```

MANOVA TABLE

Sphericity Tests				
Variables	DF	Mauchly's Criterion	Chi-Square	Pr > ChiSq
Transformed Variates	54	0.0004138	73.747132	0.0383
Orthogonal Components	54	6.8081E-8	156.22426	<.0001

MANOVA Test Criteria and Exact F Statistics for the Hypothesis of no time Effect						
H = Type III SSCP Matrix for time						
E = Error SSCP Matrix						
		S=1	M=4	N=1		
Statistic	Value	F Value	Num DF	Den DF	Pr > F	
Wilks' Lambda	0.01448680	27.21	10	4	0.0030	
Pillai's Trace	0.98551320	27.21	10	4	0.0030	
Hotelling-Lawley Trace	68.02837375	27.21	10	4	0.0030	
Roy's Greatest Root	68.02837375	27.21	10	4	0.0030	

MANOVA Test Criteria and F Approximations for the Hypothesis of no time*Diet Effect						
H = Type III SSCP Matrix for time*Diet						
E = Error SSCP Matrix						
		S=2	M=3.5	N=1		
Statistic	Value	F Value	Num DF	Den DF	Pr > F	
Wilks' Lambda	0.00327436	6.59	20	8	0.0050	
Pillai's Trace	1.82896904	5.35	20	10	0.0047	
Hotelling-Lawley Trace	50.23337469	10.05	20	4	0.0186	
Roy's Greatest Root	44.52488954	22.26	10	5	0.0016	

NOTE: F Statistic for Roy's Greatest Root is an upper bound.
NOTE: F Statistic for Wilks' Lambda is exact.

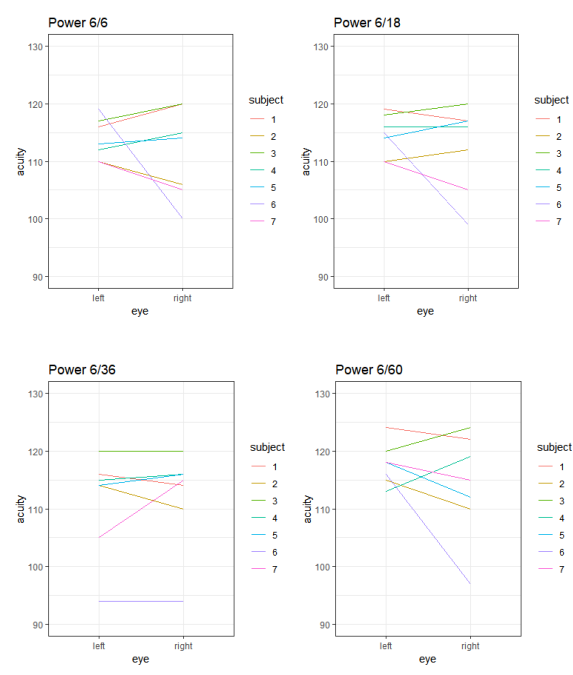
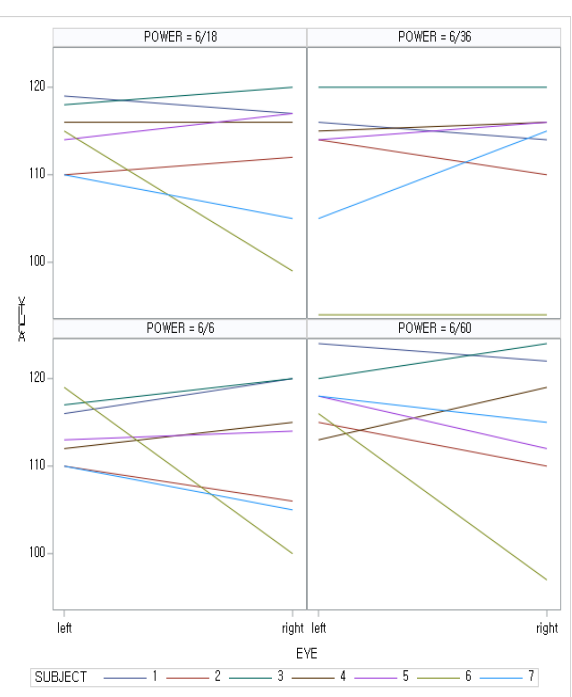
P-value = 0.0047로, Diet, Time 의 교호효과가 있다고 할 수 있다. 또한 Diet, Time 각각의 변수도 유의하다고 보여진다. 또한 구형성 검정을 해보았을 때, p-value 값은 매우 작아 구형성가정을 만족하지 않는다고 할 수 있다. 따라서 Anova 분석보다 Manova 분석이 더 적합하다.

3. RESULT

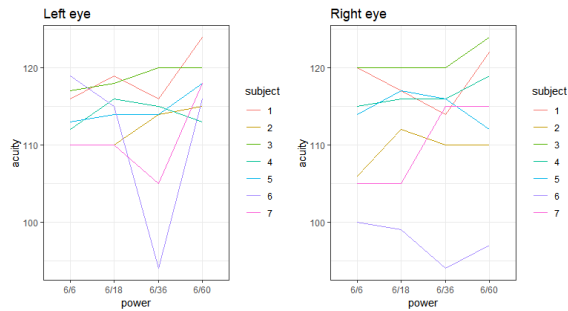
BodyWeight 자료분석 결과, p-value < 0.05 이므로 구형성가정을 만족하지 않아, Manova 분석이 더 적합하다고 판단하였다. Manova 분석 결과 Time * Diet 의 p-value < 0.05로, 효과가 유의하다고 할 수 있다. 따라서 Time, Diet 에 따라 Weight 에 효과가 있으며, Diet 그룹 1,2,3 에 대하여 사후분석을 통해 어떠한 그룹간에 차이가 있는지 살펴볼 수 있다.

2. R의 faraway package에 내장되어 있는 vision 자료를 이용하시오.

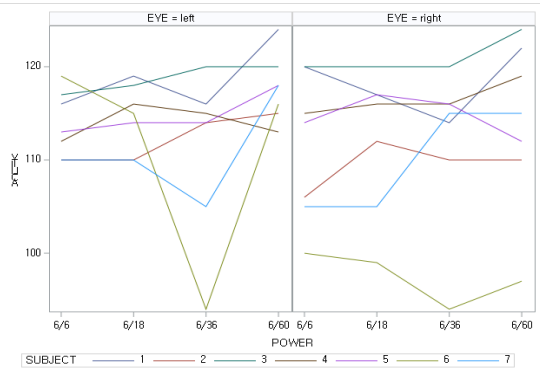
1) 적절한 그림으로 자료를 살펴보시오.

R	SAS
1. CODE	
<pre>str(vision) theme_set(theme_bw()) g1 <- ggplot(vision[vision\$power=='6/6',],aes(eye,acuity))+ geom_line(aes(color=subject,group=subject))+ ggtitle('Power 6/6')+ ylim(90,130) g2 <- ggplot(vision[vision\$power=='6/18',],aes(eye,acuity))+ geom_line(aes(color=subject,group=subject))+ ggtitle('Power 6/18')+ ylim(90,130) g3 <- ggplot(vision[vision\$power=='6/36',],aes(eye,acuity))+ geom_line(aes(color=subject,group=subject))+ ggtitle('Power 6/36')+ ylim(90,130) g4 <- ggplot(vision[vision\$power=='6/60',],aes(eye,acuity))+ geom_line(aes(color=subject,group=subject))+ ggtitle('Power 6/60')+ ylim(90,130) grid.arrange(g1,g2,ncol=2); grid.arrange(g3,g4,ncol=2) g1 <- ggplot(vision[vision\$eye=='left',], aes(x=power,y=acuity))+ geom_line(aes(color=subject,group=subject))+ggtitle('Left eye') g2 <- ggplot(vision[vision\$eye=='right',], aes(x=power,y=acuity))+ geom_line(aes(color=subject,group=subject))+ggtitle('Right eye') grid.arrange(g1,g2,ncol=2)</pre>	<pre>PROC SGPPANEL DATA=vision; panelby EYE; title 'vision'; series X=POWER Y=ACUITY / group=SUBJECT; RUN; PROC SGPPANEL DATA=vision; panelby POWER; title 'vision'; series X=EYE Y=ACUITY / group=SUBJECT; RUN;</pre>
2. PLOT	
<p># Power 수준을 나눈 right / left eye의 acuity graph</p> 	<p># Power 수준을 나눈 right / left eye의 acuity graph</p> 

Eye 수준을 나눠 각 power별 acuity graph



Eye 수준을 나눠 각 power별 acuity graph



3. RESULT

Vision 자료는 7명의 피실험자를 4가지 종류의 렌즈(power)를 왼쪽눈, 오른쪽눈으로 나누어 시력검사를 한 자료이다. 이 자료의 변수설명은 다음과 같다.

- acuity : a numeric vector (예민함)
- power : a factor with levels 6/6 6/18 6/36 6/60 : power of lens의미/ 물체의거리 Diet : a factor level 1 to 3
- eye : a factor with levels : left right
- subject : a factor with levels 1 2 3 4 5 6 7

즉, 이 데이터에서 power-Trt, subject - Subject, eye- Time, acuity - Score 에 해당한다. 우리의 목적은 각각 다른 렌즈간에 (Trt- power) 피실험자의 acuity 차이가 있는지를 살펴보는 것이다. 따라서, 각 subject를 power(6/6, 6/18, 6/36, 6/60), eye(left, right)로 나누어 그래프를 그려본 결과 power에 따라 개인별 차이를 알 수 있었다. 또한 개인별로 left, right 차이를 알아보기 위하여 그래프를 그려본 결과 거의 비슷하다는 것을 직관적으로 알 수 있었다.

2) 적절한 방법으로 분석하여 결과를 해석하시오.

R	SAS
1. CODE	
<pre># Anova lm1 <- lm(acuity ~ factor(power) + factor(subject) + factor(eye) +factor(power) * factor(eye), data = vision) temp <- anova(lm1); temp F.group <- temp[1,3] / temp[2,3] # F.value P.group <- 1-pf(F.group,temp[1,1], temp[2,1]) c(F.group, P.group) lme2 <- lme(acuity ~ factor(power) + factor(eye) + factor(power) * factor(eye), random = ~1 factor(subject),data = vision) anova(lme2) # Manova vision_M <- cast(vision , power + subject ~ eye, value="acuity") lm2 <- lm(cbind(left, right) ~ power ,data = vision_M) measure_time <- factor(c("left","right")) measure_time_data <- data.frame(measure_time = measure_time) mv2 <- Anova(lm2, idata = measure_time_data, idesign = ~ measure_time) summary(mv2)</pre>	<pre>/* ANOVA */ PROC GLM data=vision; CLASS power subject eye; MODEL acuity = power subject eye power*eye / ss3; RANDOM subject; TEST H = power E = subject; QUIT; RUN; /* MANOVA */ proc sort data=vision; by power subject; run; proc transpose data=vision out=vision_M; by power subject; id eye; var acuity; run; ods exclude partialCorr ErrorSSCP; proc glm data = vision_M; class power; model right left = power / ss3; repeated eye profile /printe summary; quit; run;</pre>

