

Homework 12

Reading

- Section 8.3 covers undirected graphical models.
- Section 1.2.3 provides a general discussion of Bayesian probability. Bishop uses Bayesian models throughout the text, we have just not emphasized them previously.

1. The file `HW12_problem1.txt` contains 50 iid samples, $\hat{X}_1, \hat{X}_2, \dots, \hat{X}_{50}$, from a one-dimensional random variable X . Assume that $X \sim \mathcal{N}(\mu, \sigma^2)$ where σ^2 is known with $\sigma^2 = 1$. Our goal is to estimate μ . Let \bar{x} be the sample mean and $N = 50$.

- (a) Show that the maximum likelihood estimate (MLE) of μ is given by the sample mean \bar{x} .
- (b) Taking a Bayesian approach, assume a normal prior on μ , $\mu \sim \mathcal{N}(0, \beta^2)$ with $\beta = 10$. Let $f(\mu)$ be the pdf of the prior. Let $p(\mu)$ be the posterior. Show,

$$p(\mu) = \frac{1}{Z} P(\hat{X}_1, \hat{X}_2, \dots, \hat{X}_{50} \mid \mu) f(\mu), \quad (1)$$

where

$$Z = \int_{-\infty}^{\infty} P(\hat{X}_1, \hat{X}_2, \dots, \hat{X}_{50} \mid \mu) f(\mu) d\mu. \quad (2)$$

Then show

$$p(\mu) \sim \mathcal{N}\left(\frac{\bar{x}}{1 + \frac{\sigma^2}{\beta^2 N}}, \frac{\sigma^2 \beta^2}{N \beta^2 + \sigma^2}\right) \quad (3)$$

Graph $p(\mu)$ using the data. Hint: use completing the squares to combine a product of exponentials into a single exponential. (Here you're showing that for random mean and fixed variance, the normal distribution is a conjugate prior to itself)

- (c) Again take a Bayesian approach, but this time assume a prior $f(\mu)$,

$$f(\mu) = \begin{cases} \frac{1}{10} & \text{if } x \in [5, 7] \\ 4 & \text{if } x \in [9, 9.2] \\ 0 & \text{otherwise} \end{cases} \quad (4)$$

In this case the posterior $p(\mu)$ is not normally distributed. Graph $p(\mu)$ and compare to the graph in (b). (Hint: you can compute Z using a numerical integration function. In R, **integrate**).

2. This problem relates to the figure below, which is Figure 17.3 in the book Elements of Statistical Learning. Let W be a 3-dimensional r.v. Typically we write $W = (W_1, W_2, W_3)$, but to match the figure, let $W = (X, Y, Z)$.

Suppose that X, Y, Z are each in $\{-1, 1\}$. Consider three parametrizations of a probabilities distribution for W , i.e. a joint probability distribution for X, Y, Z . Let $\eta = (\eta_1, \eta_2, \eta_3)$.

$$P(X = x, Y = y, Z = z) = \alpha \exp[\eta_1 x + \eta_2 y + \eta_3 z - w_{12}xy - w_{13}xz] \quad (5)$$

$$P(X = x, Y = y, Z = z) = \alpha \exp[\eta_1 x + \eta_2 y + \eta_3 z - w_{12}xy - w_{13}xz - w_{23}yz] \quad (6)$$

$$P(X = x, Y = y, Z = z) = \alpha \exp[\eta_1 x + \eta_2 y + \eta_3 z - w_{12}xy - w_{13}xz - w_{23}yz - w_{123}xyz], \quad (7)$$

where the α is a normalizing constant that differs between the three distributions. (In the lecture I used $1/Z$ for the normalization, but here Z is one of the random variables.)

- (a) Show that for distribution (5), $Y \perp Z|X$ so that the edge between Y and Z in the graph is not consistent with the distribution. Show that for distribution (5), we can write

$$P(X = x, Y = y, Z = z) = \alpha \psi_1(x, y) \psi_2(x, z) \quad (8)$$

corresponding to the maximal clique form of the Hammersley-Clifford Theorem. What are the cliques in this case?

- (b) Show that for distributions (6) and (7), no pair of the three r.v. X, Y, Z is conditionally independent given the third r.v., making each edge in the graph necessary for consistency with the probability distribution. (Note: this reflects the comment directly below the figure: *A graphical model does not always uniquely specify the higher-order structure of a joint probability distribution.*)
- (c) Consider distribution (5). Let $\eta_i = 1/2$ for $i = 1, 2, 3$. Let $w_{12} = 1$, $w_{13} = -1$. Use a Monte-Carlo approach based on a Metropolis-Hastings sampler to estimate the correlation between X and Y . (Since the graph is so small, in this case we could directly compute the correlation by summing through all possible values of W , but this would not be the case if the dimension of W was, say, 20.)

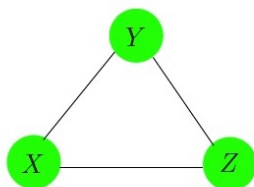


FIGURE 17.3. *A complete graph does not uniquely specify the higher-order dependence structure in the joint distribution of the variables.*