

Prueba de hipótesis para una media

* Independencia de observ: -> observ. 250 cistos -> nett

VV t -> pr desc, or desconocida

$$t_{calc} = \frac{\overline{x} - \mu_0}{\sqrt{n}} \sim t_{(n-1)}$$

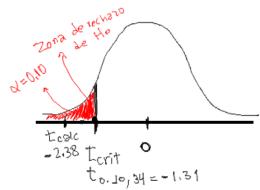
Decisióni

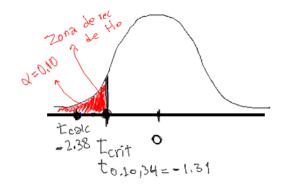
Ejemplo

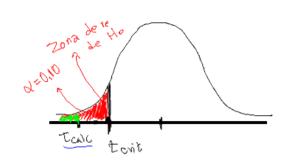
Una universidad afirma que sus estudiantes dedican en promedio 4 horas diarias al estudio. Se sospecha que el promedio es realmente menor, por lo que se toma una muestra aleatoria de 35 estudiantes, quienes reportan las siguientes horas de estudio:

Verificar la afirmación con un nivel de significancia del 10%.

$$H_0: \mu \geq 4$$
 $H_1: \mu \leq 4$ $\alpha = 0.10$







One Sample t-test

alternative hypothesis: true mean is less than 4 95 percent confidence interval:

-Inf 3.88878 sample estimates: mean of x 3.617143

$$> pt(q = -2.3833, df = 34)$$
 [1] 0.01144406

" greater"
>
" greater"

versos que pv < d > Rea. Ho

Prueba de hipótesis para una varianza

$$\chi^2_{calc} = \frac{(n-1)s^2}{\sigma_0^2} \sim \chi^2_{(n-1)}$$

Ejemplo

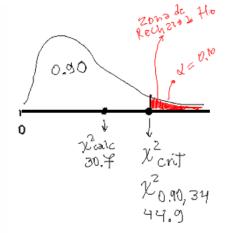
Para los mismos datos de tiempo de estudio, se sospecha que la varianza es mayor a 1 hora. ¿Se puede verificar dicha afirmación con un nivel de significancia del 10%?

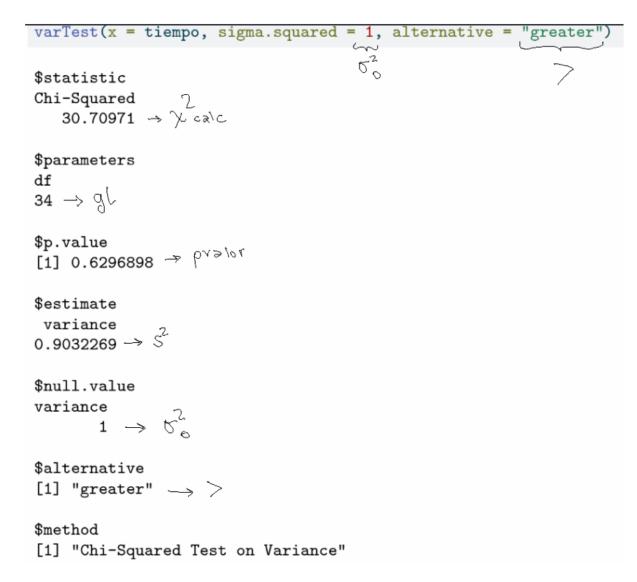
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(chicalc = (length(tiempo)-1)*var(tiempo)/1)

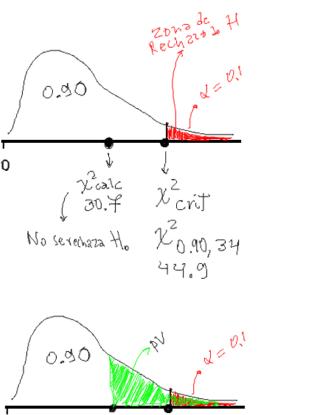
[1] 30.70971

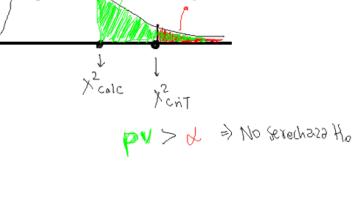
(chicrit = qchisq(p = 0.90, df = length(tiempo)-1))

[1] 44.90316
```









Prueba de hipótesis para una proporción

Prueba de hipótesis para una proporción
$$Z_{calc} = \frac{p - \pi_0}{\sqrt{\frac{\pi_0(1 - \pi_0)}{n}}} \sim N(0, 1) \qquad \qquad \begin{array}{c} \gamma = 100 \\ \gamma_0 = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma_0 = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma_0 = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma_0 = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.80 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 0.20 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array} \qquad \begin{array}{c} \gamma = 100 \\ \gamma = 100 \end{array}$$

Ejemplo

Una universidad sostiene que el 80% de los estudiantes están satisfechos con el servicio de la biblioteca. Se encuesta a 100 estudiantes al azar y 71 dicen estar satisfechos. Verificar si la proporción real difiere de 0.80, con un nivel de significancia del 5%.

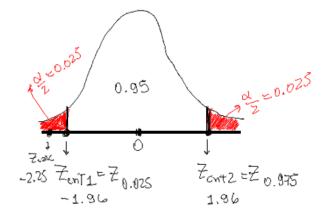
$$H_0: \pi = 0.80$$
 $H_1: \pi \neq 0.80$ $\alpha = 0.05$

$$(Z_crit1 \leftarrow qnorm(0.025))$$

[1] -1.959964

$$(Z_{crit2} \leftarrow qnorm(0.975))$$

[1] 1.959964



prop.test(x=71, n=100, p=0.80, alternative = "two.sided", correct = F)

1-sample proportions test without continuity correction

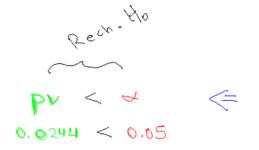
data: 71 out of 100, null probability 0.8

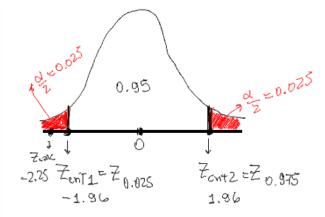
X-squared = 5.0625, df = 1, p-value = 0.02445 alternative hypothesis: true p is not equal to 0.8 95 percent confidence interval:

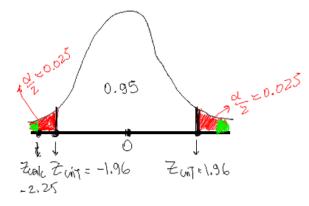
0.6146111 0.7898516 sample estimates:

p 0.71

$$Z_{calc} = -2.25 \rightarrow Z_{calc} = 5.0625$$
PROPIEDAD: $Z^2 \sim \chi_{alc}^2$





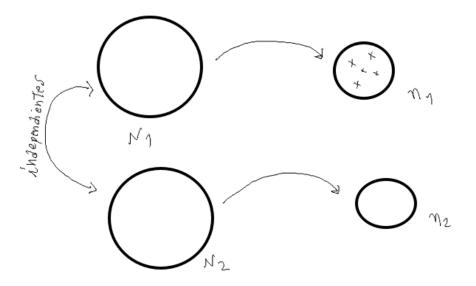


$$PN = 2P(Z<-2.25) = 0.0244$$

> 2*pnorm(-2.25)

[1] 0.02444895

Antes de comparar dos medias, se debe verificar si las varianzas son iguales o no, es decir si existe homogeneidad de varianzas o no.



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1	4000 <>	40∞	0
2	4500	0	4500
3	0	5000	-5000
4	2500	3500	-7000

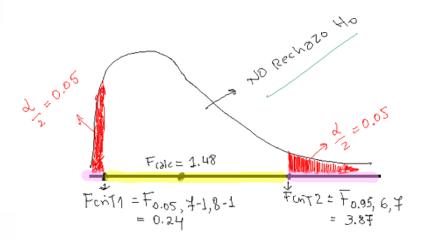
Prueba de hipótesis de homogeneidad de dos varianzas

$$F_{calc} = \frac{s_1^2}{s_2^2} \sim F_{(n_1-1,n_2-1)}$$

Ejemplo

Un laboratorio desea determinar si la variabilidad de concentración (en mg/L) de un fármaco en la sangre es igual para dos fabricantes distintos, considerando un nivel de significancia del 10%. Los datos de concentración con el fabricante A son: 8.1, 7.9, 8.3, 7.8, 8.0, 8.2, 7.7, mientras que con el fabricante B: 7.5, 7.2, 7.1, 7.4, 7.3, 7.6, 7.2, 7.5.

$$H_0: \frac{\sigma_A^2}{\sigma_B^2} = \boxed{1} \qquad H_1: \frac{\sigma_A^2}{\sigma_B^2} \neq 1 \qquad \alpha = 0.10$$
 Var. he have



A <-
$$c(8.1, 7.9, 8.3, 7.8, 8.0, 8.2, 7.7)$$

B <- $c(7.5, 7.2, 7.1, 7.4, 7.3, 7.6, 7.2, 7.5)$
(Fcalc <- $var(A)/var(B)$)

[1] 0.2377184

$$(Fcrit2 \leftarrow qf(0.95, 6, 7))$$

[1] 3.865969

~ \ \

var.test(A, B, alternative = "two.sided", ratio = 1)

F test to compare two variances

data: A and B

F = 1.4848, num df = 6, denom df = 7, p-value = 0.6136

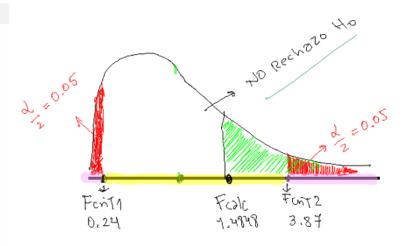
alternative hypothesis: true ratio of variances is not equal to 1 95 percent confidence interval:

0.290089 8.456911 sample estimates:

ratio of variances

1.484848

PV > d 0.61 > 0.10 No se rechaza Ho



$$> 1-pf(1.4848,6,7)$$

$$> 2*(1-pf(1.4848,6,7))$$

[1] 0.6136236

Prueba de hipótesis para dos medias independientes

$$H_0: \frac{\sigma_A^2}{\sigma_B^2} = 1 \qquad H_1: \frac{\sigma_A^2}{\sigma_B^2} \neq 1 \implies \quad t = \frac{\bar{X}_1 - \bar{X}_2 - \mu_0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_{\nu-Welch} \quad \text{(t do Welch)}$$

$$t = \frac{\bar{X}_1 - \bar{X}_2 - \mu_0}{\sqrt{s_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \sim t_{(n_1 + n_2 - 2)} \quad \text{donde} \quad s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2} \qquad (\text{to de Student})$$

Prueba de hipótesis de homogeneidad de dos varianzas

Prueba de hipótesis para dos medias pareadas

```
H_0: \mu_D \leq 3 \qquad H_1: \mu_D \geqslant 3 \qquad \alpha = 0.05 antes = c(120, 122, 121, 119, 118, 123, 121, 120, 122, 119, 115, 123) despues = c(125, 127, 126, 124, 123, 129, 126, 125, 128, 123, 116, 129) t.test(despues, antes, mu = 3, alternative = "greater", paired = T) \frac{\partial^2 f_{12} \partial f_{23}}{\partial f_{14} \partial f_{14}} = \frac{\partial^2 f_{14} \partial f_{14}}{\partial f_{14}} = \frac{\partial^2 f_{14}}{\partial f_{14}} = \frac{\partial^
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```
> antes = c(120,122,121,119,118,123,121,120,122,119,115,123)
> despues = c(125, 127, 126, 124, 123, 129, 126, 125, 128, 123, 116, 129)
> t.test(despues, antes, mu = 3, alternative = "greater", paired = T)
        Paired t-test
data: despues and antes
t = 4.7497, df = 11, p-value = 3e-04
                                                                  H_0: \mu_D \leq 3 \qquad H_1: \mu_D > 3 \qquad \alpha = 0.05
alternative hypothesis: true mean difference is greater than 3
95 percent confidence interval:
 4.140136
              Inf
sample estimates:
mean difference
       4.833333
                                                                                               pv = 0.0003
> diferencia = despues - antes
                                                                                             pv<d
> t.test(diferencia, mu = 3, alternative = "greater")
        One Sample t-test
                                                                                           Se rechaza Ho
data: diferencia
t = 4.7497, df = 11, p-value = 3e-04
alternative hypothesis: true mean is greater than 3
95 percent confidence interval:
4.140136
               Inf
```

sample estimates:

mean of x 4.833333

