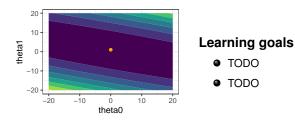
Optimization

Smooth unconstrained problems



GENERAL DEFINITION

Consider the optimization problem

$$\min_{\mathbf{x} \in \mathcal{S} \subseteq \mathbb{R}^d} f(\mathbf{x})$$

with objective function

$$f: \mathcal{S} \to \mathbb{R}$$
.

The problem is called

• **unconstrained**, if the domain S is not restricted:

$$S = \mathbb{R}^d$$

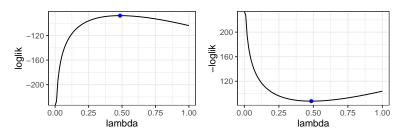
- smooth if f is smooth.
- univariate if d = 1, and multivariate if d > 1.

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NOTE: A CONVENTION IN OPTIMIZATION

W.l.o.g., we always **minimize** functions f.

Maximization results from minimizing -f.

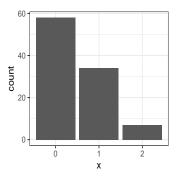


Poisson example: Maximizing the log-likelihood (left) is equivalent to minimizing the negative log-likelihood (right).

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EXAMPLE 1: MAXIMUM LIKELIHOOD ESTIMATION

Assume an i.i.d. sample $\mathcal{D} = (x^{(1)}, ..., x^{(n)})$ from a distribution with density $f(x \mid \theta)$. We want to find λ which makes the observed data most likely.



Example: Histogram of a sample drawn from a Poisson distribution $f(k \mid \lambda) := \mathbb{P}(x = k) = \frac{\lambda^k \cdot \exp(-\lambda)}{k!}.$

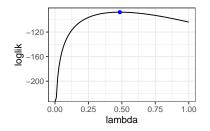
EXAMPLE 1: MAXIMUM LIKELIHOOD ESTIMATION

We operationalize this as **maximizing** the log-likelihood function (or equivalently: minimizing the negative log-likelihood) with respect to λ :

$$\begin{split} \hat{\lambda} &= & \arg\min_{\lambda} \ - \ell(\lambda, \mathcal{D}) = \arg\min_{\lambda} - \log \mathcal{L}(\lambda, \mathcal{D}) = \arg\min_{\lambda} - \log \prod_{i=1}^{n} f\left(\mathbf{x}^{(i)} \mid \lambda\right) \\ &= & \arg\min_{\lambda} - \sum_{i=1}^{n} f\left(\mathbf{x}^{(i)} \mid \lambda\right) = \arg\min_{\lambda} \sum_{i=1}^{n} \frac{-\lambda^{\mathbf{x}^{(i)}} \cdot \exp(-\lambda)}{\mathbf{x}^{(i)}!} \end{split}$$

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EXAMPLE 1: MAXIMUM LIKELIHOOD ESTIMATION



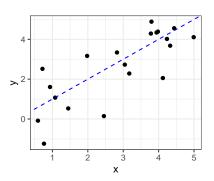
Example: The log-likelihood of a Poisson distribution for data example above. The objective function is univariate and differentiable, and the domain is unconstrained.

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EXAMPLE 2: NORMAL REGRESSION

Assume a dataset $\mathcal{D} = ((\mathbf{x}^{(1)}, y^{(1)}), \dots, (\mathbf{x}^{(n)}, y^{(n)}))$ generated according to

$$\mathbf{y}^{(i)} = \mathbf{\theta}^{\top} \mathbf{x}^{(i)} + \epsilon^{(i)}, \qquad \epsilon^{(i)} \stackrel{\textit{iid}}{\sim} \mathcal{N}(0, 1).$$

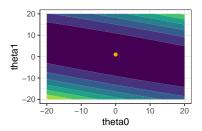


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EXAMPLE 2: NORMAL LINEAR REGRESSION

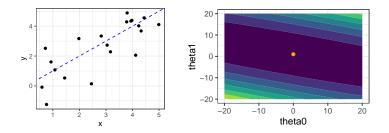
In normal linear regression the goal is to find a vector $\boldsymbol{\theta}$ which minimizes the sum of squared errors (SSE):

$$\min_{\boldsymbol{\theta} \in \mathbb{R}^d} \sum_{i=1}^n \left(\boldsymbol{\theta}^\top \mathbf{x}^{(i)} - y^{(i)} \right)^2$$



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EXAMPLE 2: NORMAL REGRESSION



- The problem is multivariate, smooth, and unconstrained
- Since the problem is a quadratic form, we easily obtain a geometric interpretation of the problem
- The problem has a closed-form solution, which is given by $\theta = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y}$, where **X** is the design matrix

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