Simulating Competing Risk Survival Settings

With High-Dimensional Data

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BioWiMium

Introduction

Simulation setup part of ongoing project

Motivation

- Variable selection in high-dimensional settings with competing risks
- Focus more on high-dimensional setting than sophisticated survival outcome

Outcome

- 2 competing events, censoring approx. equal prevalence

Generating Survival Times (I)

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Gene expression

Boosting for high-dimensional time-to-event data with competing risks

Harald Binder[1,2,*], Arthur Allignol[1,2], Martin Schumacher[2] and Jan Beyersmann[1,2]

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Binder et al. (2009)

Adapting Prediction Error Estimates for Biased Complexity Selection in High-Dimensional Bootstrap Samples

Harald Binder & Martin Schumacher

Universität Freiburg i. Br.

Nr. 100 December 2007

Binder and Schumacher (2008)

Generating Survival Times to Simulate Cox Proportional Hazards Models

Ralf Bender[1], Thomas Augustin[2], Maria Blettner[1]

[1]Dept. of Epidemiology and Medical Statistics, School of Public Health
University of Bielefeld, Germany

[2]Department of Statistics, University of Munich, Germany

Bender et al. (2005)

Predictors: Structure

- $N = 400$, $p = 5000$, 16 informative (12 per event)
- Organized in 4 blocks á 250 correlated variables + uncorrelated noise

Blocks

- Block 1: $\rho \approx 0.5$
- Block 2: $\rho \approx 0.35$
- Block 3: $\rho \approx 0.05$
- Block 4: $\rho \approx 0.32$
- Rest: $\rho \approx 0$

Predictors: Generation

- $j = 1, \dots, p$ and $i = 1, \dots, N$
- $\epsilon_{ij} \sim \mathcal{N}(0, 1)$ and $u_{i\{1,2,3\}} \sim \mathcal{U}(0, 1)$

$$x_{ij} = \begin{cases} -1 + \epsilon_{ij} & \text{for } i \leq 0.5n \text{ and } j \leq 0.05p \\ 1 + \epsilon_{ij} & \text{for } i > 0.5n \text{ and } j \leq 0.05p \\ 1.5 \cdot \mathbf{1}\{u_{i1} < 0.4\} + \epsilon_{ij} & \text{for } 0.05p < j \leq 0.1p \\ 0.5 \cdot \mathbf{1}\{u_{i2} < 0.7\} + \epsilon_{ij} & \text{for } 0.1p < j \leq 0.2p \\ 1.5 \cdot \mathbf{1}\{u_{i3} < 0.3\} + \epsilon_{ij} & \text{for } 0.2p < j \leq 0.3p \\ \epsilon_{ij} & \text{for } j > 0.3p \end{cases}$$

Effect Assignment

- Cause-specific hazards for events $k = 1, 2$ and coefficients $\boldsymbol{\beta}^{(k)}$
- $\beta_j^{(k)} = \pm 0.5$ for effect variables
- Block 1 (**Mutual**): 4 effect variables
  - same effect on both $\boldsymbol{\beta}^{(1,2)}$ cause-specific hazards
- Block 2 (**Reversed**): 4 effect variables
  - positive effect in $\boldsymbol{\beta}^{(1)}$ and negative in $\boldsymbol{\beta}^{(2)}$
- Block 3 (**Disjoint**): 8 variables
  - 3.1: 4 effect variables in $\boldsymbol{\beta}^{(1)}$ only
  - 3.2: 4 effect variables $\boldsymbol{\beta}^{(2)}$ only

Event and Censoring Times (I)

- Cox-exponential model for events $k = 1, 2$
- $\lambda_{1,2,C}$: Baseline hazards for event times $T_i^{(k)}$, and censoring times C_i
- Simple setting: $\lambda_1 = \lambda_2 = \lambda_C = 0.1$

$$T_i^{(k)} = \frac{-\log(U_i)}{\lambda_k \exp(\mathbf{x}_i^T \boldsymbol{\beta}^{(k)})}$$

$$C_i = \frac{-\log(U_i)}{\lambda_C}$$

$$U_i \sim \mathcal{U}(0, 1) \Rightarrow -\log(U_i) \sim \text{Exponential}(\lambda = 1)$$

Event and Censoring Times (II)

Assign observed event times t_i and censoring δ_i indicator accordingly

$$t_i = \min(T_i^{(1)}, T_i^{(2)}, C_i) \qquad \delta_i = \begin{cases} 0 & \text{if } t_i = C_i \\ 1 & \text{if } t_i = T_i^{(1)} \\ 2 & \text{if } t_i = T_i^{(2)} \end{cases}$$

Resulting Outcome

- $\boldsymbol{\beta}^{(k)}$ sparse with 12 non-zero entries, $\sum_{i=1}^{p} \beta_j^{(k)} = 2$
- $\exp(\mathbf{x}_i^T \boldsymbol{\beta}^{(k)})$ in range of $[10^{-4}, 10^3]$

Event prevalences

Mean event counts and prevalence after 100 replicates with $N = 400$:

δ	n (min - max)	% (min - max)
0	118.0 (97 - 136)	29.5% (24.2% - 34.0%)
1	165.1 (142 - 190)	41.3% (35.5% - 47.5%)
2	116.9 (93 - 141)	29.2% (23.2% - 35.2%)

Thanks for listening!

Buffer Slide

Next up: Implementation details nobody will want to see but I put them on here any just in case anyone asks because I don't have too much else to talk about in this regard sorry Mensa anyone?

Predictors: Code


```r
X <- matrix(rnorm(n * p), nrow = n, ncol = p)
ui1 <- runif(n); ui2 <- runif(n); ui3 <- runif(n)
j_seq <- seq_len(p)
block1 <- which(j_seq <= 0.05 * p)
X[seq_len(n/2), block1] <- -1 + X[seq_len(n/2), block1]
X[-seq_len(n/2), block1] <- 1 + X[-seq_len(n/2), block1]
block2 <- which((j_seq > (0.05 * p)) & (j_seq <= (0.1 * p)))
X[, block2] <- 1.5 * (ui1 < 0.4) + X[, block2]
block3 <- which((0.1 * p < j_seq) & (j_seq <= 0.2 * p))
X[, block3] <- 0.5 * (ui2 < 0.7) + X[, block3]
block4 <- which((0.2 * p < j_seq) & (j_seq <= 0.3 * p))
X[, block4] <- 1.5 * (ui3 < 0.3) + X[, block4]
```

Efect Assignment: Code


```r
ce <- 0.5
# first block
j_block1 <- which(j_seq <= 0.05 * p)
beta1[j_block1[1:4]] <- ce
beta2[j_block1[1:4]] <- ce
# second block
j_block2 <- which((j_seq > (0.05 * p)) & (j_seq <= (0.1 * p))
beta1[j_block2[1:4]] <- ce
beta2[j_block2[1:4]] <- -ce
# third block
j_block3 <- which((0.1 * p < j_seq) & (j_seq <= 0.2 * p))
beta1[j_block3[1:4]] <- -ce
beta2[j_block3[5:8]] <- ce # offset by 4
```

Event Times: Code


```r
lp1 <- X %*% beta1
lp2 <- X %*% beta2

Ti1 <- -log(runif(n)) / (lambda1 * exp(lp1))
Ti2 <- -log(runif(n)) / (lambda2 * exp(lp2))
Ci <- -log(runif(n)) / lambda_c

ti <- pmin(Ti1, Ti2, Ci)
di <- as.integer(Ti1 <= Ci | Ti2 <= Ci)
di[which(Ti2 <= Ti1 & Ti2 <= Ci)] <- 2
```

Thank you for your attention

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