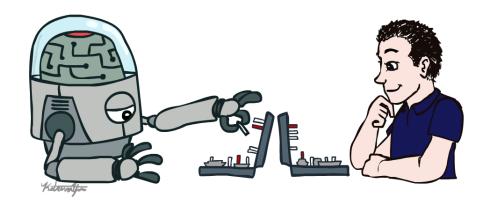
CSE 3521: Introduction to Artificial Intelligence





Regression (curve fitting): what model?

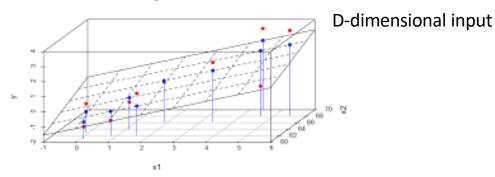
- Relationship (function):
 - Linear

•
$$f(x) = a x + b$$
 or $f(x) = w^T x + b = w[1]x[1] + ... + w[D]x[D] + b$

1-dimensional input

Non-linear

$$f(x) = a x^2 + bx + c$$



- Parameter Estimation
 - \circ Given data points and desired labels/target values $\{(x_1, y_1), ..., (x_N, y_N)\}$
 - Find values for <u>parameters</u> a, b, c, w, etc.
 - \circ Aim to find parameters that make $f(x_n)$ as close as possible to y_n

Regression (curve fitting): close?

- Difference or <u>error</u> between predictions and labels
 - Total error: error over all training data instances
- Different types of total error
 - Sum of Absolute Error (SAE)

$$E_1 = \sum_i |y_i - f(x_i)|$$

Maximum Error

$$E_{\infty} = \max_{i} |y_i - f(x_i)|$$

Sum of Squared Error (SSE) or residual sum of squares (RSS)

$$E_2 = \sum_i [y_i - f(\mathbf{x}_i)]^2$$

Linear least squares

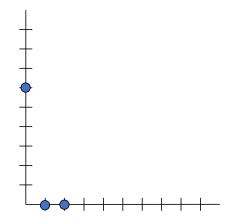
- f is a linear function (linear combination of feature variables)
 - o <u>Linear in parameters</u>: The function is a sum of terms x[d], where parameters are *only* coefficients of those terms $f(x) = \mathbf{w}^T x + b = w[1]x[1] + \dots + w[D]x[D] + b$
 - Thus, only works when we expect relationship is a line/plane/hyperplane
- Minimize sum of squared error
- Why SSE?
 - O Short answer: The math is easier!
 - Other answers: You will see more after probabilistic modeling
- Estimation of model parameters derived from
 - Calculus
 - Linear Algebra

Example: 1-dimensional data

- Find a straight line through points (0,6), (1,0), and (2,0)
 Model equation: y=wx + b
- Need values for parameters w, b that satisfy

$$6 = w \cdot 0 + b$$
$$0 = w \cdot 1 + b$$

$$0 = w \cdot 2 + b$$



Linear system of equations

$$y = w \cdot x + b$$

$$6 = w \cdot 0 + b$$

$$0 = w \cdot 1 + b$$

$$0 = w \cdot 2 + b$$

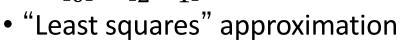
$$\begin{bmatrix} 6 \\ 0 \\ 0 \end{bmatrix} = w \cdot \begin{bmatrix} 0 \\ 1 \\ 2 \end{bmatrix} + b \cdot \begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix} \longrightarrow \begin{bmatrix} 6 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 & 1 \\ 1 & 1 \\ 2 & 1 \end{bmatrix} \begin{bmatrix} w \\ b \end{bmatrix}$$
Expand x with 1

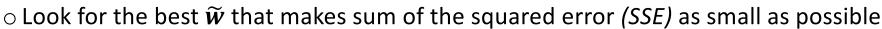
$$\mathbf{y} = \widetilde{\mathbf{X}}^T \widetilde{\mathbf{w}}$$
Combine w with b

 $\widetilde{\boldsymbol{w}} = [w, b]$ unsolvable for these actual points! But...

- Find a set of parameters (w, b) that "best" fit the data
 - \circ Minimize the difference between \mathbf{y} and $\widetilde{\mathbf{X}}^T\widetilde{\mathbf{w}}$

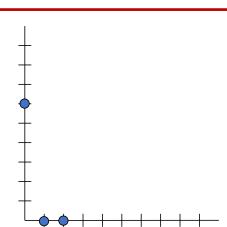
$$\circ \begin{bmatrix} 6 \\ 0 \\ 0 \end{bmatrix} - \begin{bmatrix} 0 & 1 \\ 1 & 1 \\ 2 & 1 \end{bmatrix} \begin{bmatrix} w \\ b \end{bmatrix}$$





$$E = \sum_{i} (wx_{i} + b - y_{i})^{2} = \sum_{i} ([w, b] \begin{bmatrix} x_{i} \\ 1 \end{bmatrix} - y_{i})^{2} = \sum_{i} (\widetilde{\mathbf{w}}^{T} \widetilde{\mathbf{x}}_{i} - y_{i})^{2} = \sum_{i} (\widetilde{\mathbf{x}}_{i}^{T} \widetilde{\mathbf{w}} - y_{i})^{2}$$

$$= [(\widetilde{\mathbf{x}}_{1}^{T} \widetilde{\mathbf{w}} - y_{1}), \dots, (\widetilde{\mathbf{x}}_{N}^{T} \widetilde{\mathbf{w}} - y_{N})] \begin{bmatrix} \widetilde{\mathbf{x}}_{1}^{T} \widetilde{\mathbf{w}} - y_{1} \\ \vdots \\ \widetilde{\mathbf{x}}_{N}^{T} \widetilde{\mathbf{w}} - y_{N} \end{bmatrix} = ||\widetilde{\mathbf{X}}^{T} \widetilde{\mathbf{w}} - \mathbf{y}||_{2}^{2}$$





Derivatives (gradients) are zero when function at a (local) minimum

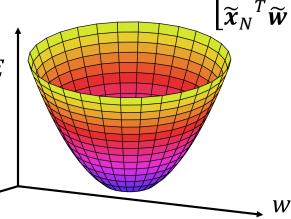
Error function

$$E = \sum_{i} (wx_{i} + b - y_{i})^{2} = \sum_{i} ([w, b] \begin{bmatrix} x_{i} \\ 1 \end{bmatrix} - y_{i})^{2} = \sum_{i} (\widetilde{w}^{T} \widetilde{x}_{i} - y_{i})^{2}$$

$$= [(\widetilde{x}_{1}^{T} \widetilde{w} - y_{1}), \dots, (\widetilde{x}_{N}^{T} \widetilde{w} - y_{N})] \begin{bmatrix} \widetilde{x}_{1}^{T} \widetilde{w} - y_{1} \\ \vdots \\ \widetilde{x}_{N}^{T} \widetilde{w} - y_{N} \end{bmatrix} = ||\widetilde{X}^{T} \widetilde{w} - y_{N}||_{2}^{2}$$

$$= \left[\left(\widetilde{\boldsymbol{x}}_{1}^{T} \widetilde{\boldsymbol{w}} - \boldsymbol{y}_{1} \right), \dots, \left(\widetilde{\boldsymbol{x}}_{N}^{T} \widetilde{\boldsymbol{w}} - \boldsymbol{y}_{N} \right) \right] \begin{vmatrix} \widetilde{\boldsymbol{x}}_{1}^{T} \widetilde{\boldsymbol{w}} - \boldsymbol{y}_{1} \\ \vdots \\ \widetilde{\boldsymbol{x}}^{T} \widetilde{\boldsymbol{w}} - \boldsymbol{y} \end{vmatrix} = \left\| \widetilde{\boldsymbol{X}}^{T} \widetilde{\boldsymbol{w}} - \boldsymbol{y} \right\|_{2}^{2}$$

- E has the shape of a parabola
 - Due to linear the function and SSE
 - Local minimum = global minimum



- We have two unknowns (w, b), therefore we have two derivatives
 - Both derivatives are zero at the function minimum
- Compute the <u>partial derivatives</u> and set to 0

$$\frac{\partial E}{\partial w} = \frac{\partial}{\partial w} \sum_{i} (w \cdot x_i + b - y_i)^2 = 0$$

$$= 2 \sum_{i} (w \cdot x_i + b - y_i) \cdot x_i = 0$$

$$\frac{\partial E}{\partial b} = \frac{\partial}{\partial b} \sum_{i} (w \cdot x_i + b - y_i)^2 = 0$$

$$= 2 \sum_{i} (w \cdot x_i + b - y_i) = 0$$

Substitute in data

Recall: (0,6), (1,0), (2,0)

$$\frac{\partial E}{\partial w} = 2\sum_{i} (w \cdot x_{i} + b - y_{i}) \cdot x_{i} = 0$$

$$= (w \cdot 0 + b - 6) \cdot 0 + (w \cdot 1 + b - 0) \cdot 1 + (w \cdot 2 + b - 0) \cdot 2 = 0$$

$$= 5w + 3b = 0$$

$$\frac{\partial E}{\partial b} = 2 \sum_{i} (w \cdot x_i + b - y_i) = 0$$

$$= (w \cdot 0 + b - 6) + (w \cdot 1 + b - 0) + (w \cdot 2 + b - 0) = 0$$

$$= 3w + 3b - 6 = 0$$

$$3w + 3b = 6$$

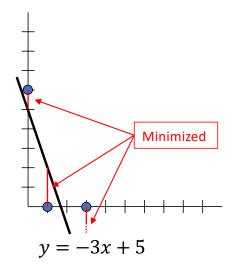
Solve for model parameters

$$5w + 3b = 0$$

$$3w + 3b = 6$$

$$w = -3$$

$$b = 5$$



Solving systems of equations can also be done with linear algebra:

$$\begin{bmatrix} 5 & 3 \\ 3 & 3 \end{bmatrix} \begin{bmatrix} w \\ b \end{bmatrix} = \begin{bmatrix} 0 \\ 6 \end{bmatrix}$$
$$\begin{bmatrix} w \\ b \end{bmatrix} = \begin{bmatrix} 5 & 3 \\ 3 & 3 \end{bmatrix}^{-1} \begin{bmatrix} 0 \\ 6 \end{bmatrix} = \begin{bmatrix} -3 \\ 5 \end{bmatrix}$$

• Recall we originally formulated our problem by \mathbf{y} and $\widetilde{\mathbf{X}}^T$ $\widetilde{\mathbf{w}}$. Is there some relationship with the above?

Yes!

$$\widetilde{X}\widetilde{X}^{T} = \begin{bmatrix} 0 & 1 & 2 \\ 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} 0 & 1 \\ 1 & 1 \\ 2 & 1 \end{bmatrix} = \begin{bmatrix} 5 & 3 \\ 3 & 3 \end{bmatrix}$$

$$\widetilde{X}y = \begin{bmatrix} 0 & 1 & 2 \\ 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} 6 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ 6 \end{bmatrix}$$

$$\widetilde{w} = \begin{bmatrix} w \\ b \end{bmatrix}$$

We can rewrite the previous solution as:

$$\begin{bmatrix} w \\ b \end{bmatrix} = \begin{bmatrix} 5 & 3 \\ 3 & 3 \end{bmatrix}^{-1} \begin{bmatrix} 0 \\ 6 \end{bmatrix} = \begin{bmatrix} -3 \\ 5 \end{bmatrix}$$

$$(\widetilde{X}\widetilde{X}^T)\widetilde{w} = \widetilde{X}\mathbf{y}$$

$$\widetilde{w} * = (\widetilde{X}\widetilde{X}^T)^{-1}\widetilde{X}\mathbf{y}$$
Optimal solution

$$\widetilde{\boldsymbol{w}} * = \left(\widetilde{\boldsymbol{X}}\widetilde{\boldsymbol{X}}^T\right)^{-1}\widetilde{\boldsymbol{X}}\boldsymbol{y}$$

This is equivalent to the partial derivative formulation we derived earlier.

In other words, linear least squares is as simple as correctly forming your \widetilde{X} matrix and y vector and then applying standard linear algebra operations to obtain a closed-form solution

Another way of derivation (for D-dimensional data)

$$f(\mathbf{x}) = \mathbf{w}^T \mathbf{x} + \mathbf{b} = \mathbf{w}[1]\mathbf{x}[1] + \dots + \mathbf{w}[D]\mathbf{x}[D] + \mathbf{b} = [\mathbf{w}^T, \mathbf{b}] \begin{bmatrix} \mathbf{x} \\ 1 \end{bmatrix} = \widetilde{\mathbf{w}}^T \widetilde{\mathbf{x}}$$

$$\widetilde{X} = [\widetilde{x}_1, \dots, \widetilde{x}_N] = \begin{bmatrix} | & & | \\ \widetilde{x}_1 & \dots & \widetilde{x}_N \\ | & & | \end{bmatrix}$$

$$E = \|\widetilde{X}^T \widetilde{w} - y\|_2^2 = (\widetilde{X}^T \widetilde{w} - y)^T (\widetilde{X}^T \widetilde{w} - y) = (\widetilde{w}^T \widetilde{X} - y^T) (\widetilde{X}^T \widetilde{w} - y)$$

= $\widetilde{w}^T \widetilde{X} \widetilde{X}^T \widetilde{w} - 2y^T \widetilde{X}^T \widetilde{w} + y^T y = \widetilde{w}^T \widetilde{X} \widetilde{X}^T \widetilde{w} - 2y^T \widetilde{X}^T \widetilde{w} + \mathbf{Const.}$

$$\nabla_{\widetilde{\boldsymbol{w}}}E = \begin{bmatrix} \frac{\partial E}{\partial \widetilde{\boldsymbol{w}}[1]} \\ \vdots \\ \frac{\partial E}{\partial \widetilde{\boldsymbol{w}}[D]} \end{bmatrix} = 2\widetilde{\boldsymbol{X}}\widetilde{\boldsymbol{X}}^T\widetilde{\boldsymbol{w}} - 2\widetilde{\boldsymbol{X}}\boldsymbol{y} = 0$$

$$\widetilde{\boldsymbol{w}} * = (\widetilde{\boldsymbol{X}}\widetilde{\boldsymbol{X}}^T)^{-1}\widetilde{\boldsymbol{X}}\boldsymbol{y}$$

- Works for any model with <u>linear</u> combination of <u>parameters</u>
- Need enough examples for unknowns (i.e., parameters)
 - \circ D+1 unknowns \rightarrow D+1 equations/points (unique)
 - Often need many more than *D* to get a good result
 - ➤ Noise
 - ➤ Overfitting
- Another explanation:
 - $\circ \widetilde{X}\widetilde{X}^T$ is (D+1)-by-(D+1)
 - $\circ (\widetilde{X}\widetilde{X}^T)^{-1}$ exists only if $\widetilde{X}\widetilde{X}^T$ is of full rank: $N \geq D + 1$

Extension: non-linear models (1-dimension)

Consider polynomials...

$$y = g + cx + bx^2 + ax^3 = g \cdot 1 + \cdots$$

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix} \qquad \widetilde{\mathbf{X}}^T = \begin{bmatrix} 1 & x_1 & x_1^2 & x_1^3 \\ 1 & x_2 & x_2^2 & x_2^3 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_N & x_N^2 & x_N^3 \end{bmatrix} \qquad \widetilde{\mathbf{w}} = \begin{bmatrix} g \\ c \\ b \\ a \end{bmatrix}$$

$$N \ge 4$$

Though g is a constant term, you can think of it as being the coefficient for 1 (equivalently x^0 in polynomials)

Note that the order in $\widetilde{\boldsymbol{w}}$ doesn't matter, as long as the rows of $\widetilde{\boldsymbol{w}}$ correspond with the columns of $\widetilde{\boldsymbol{X}}^T$

Extension: non-linear models (2-dimension)

Consider 2nd-order polynomials...

$$y = b + w[1]x[1] + w[2]x[2] + w[3](x[1]x[2]) + w[4](x[1]^2) + w[5](x[2]^2)$$

How many training data are needed?

Extension: non-linear models

Or non-polynomial terms...

 $N \ge 3$

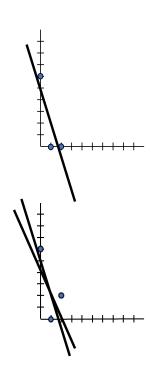
o But parameters are still linear in terms of these terms

$$y = a\frac{1}{x} + b\log x + c e^x$$

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix} \qquad \widetilde{\mathbf{X}}^T = \begin{bmatrix} 1/x_1 & \log x_1 & e^{x_1} \\ 1/x_2 & \log x_2 & e^{x_2} \\ \vdots & \vdots & \vdots \\ 1/x_N & \log x_N & e^{x_N} \end{bmatrix} \qquad \widetilde{\mathbf{w}} = \begin{bmatrix} a \\ b \\ c \end{bmatrix}$$

Problems

- Requires relationship to be *linear*
- Matrix inverse is expensive for large numbers of parameters (complex models)
 - \circ In general, $O((D+1)^3)$
 - Sometime, matrix inverse might even not exist
- Outliers (noise in the "label" or "input features")
 - Data point(s) which don't match the pattern
 - Squared error is large, large influence on SSE
 - Thus, outliers can significantly affect results



Notes

- In different textbooks or papers
 - \circ *b* is sometimes denoted as w[0];
 - o **w** is called weights or parameters or parameter vector, and **b** as bias;
 - \circ sometimes, $\widetilde{\boldsymbol{w}}$ is called parameters, too;
 - \circ sometimes, people write \boldsymbol{w} , but they actually mean $\widetilde{\boldsymbol{w}}$

$$\circ$$
 sometimes, people write \widetilde{X} (or X) to denote $\begin{bmatrix} -\widetilde{x}_1^T - \\ \vdots \\ -\widetilde{x}_N^T - \end{bmatrix}$ (or $\begin{bmatrix} -x_1^T - \\ \vdots \\ -x_N^T - \end{bmatrix}$)

 So please pay attention to contexts when you read papers, textbooks, or assigned reading material.

Summary

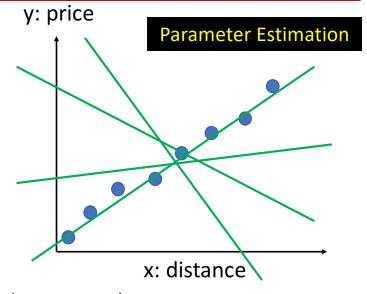
- Linear regression (LR)
 - o Fit a linear line to labeled data:
 - For one-dimensional cases, f(x) = wx + b
 - For multi-dimensional cases, $f(x) = \sum_{d=1}^{D} w[d]x[d] + b$ = $f(x) = w^{T}x + b$
 - Loss/error: absolute, square, etc.
- Parameter estimation for LR
 - \circ For the sum of square error: $E = \sum_{i} [y_i f(x_i)]^2$
 - \circ The closed-form solution that minimizes E (if we have the inverse)

$$\bullet \widetilde{w} * = (\widetilde{X}\widetilde{X}^T)^{-1}\widetilde{X}y,$$

$$\widetilde{\boldsymbol{w}} * = \begin{bmatrix} \boldsymbol{w} \\ \boldsymbol{b} \end{bmatrix} , \widetilde{\boldsymbol{X}} = [\widetilde{\boldsymbol{x}}_1, \dots, \widetilde{\boldsymbol{x}}_N] = \begin{bmatrix} \boldsymbol{x}_1 \\ 1 \end{bmatrix} \dots \begin{bmatrix} \boldsymbol{x}_N \\ 1 \end{bmatrix}$$

o Careful about how the data is represented

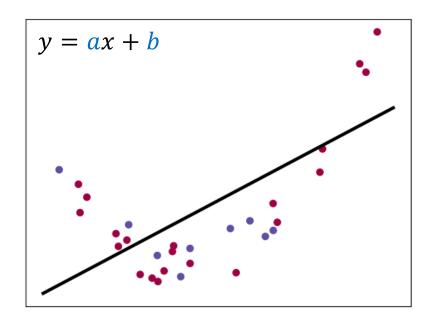
Concatenate b at the end of the column vector **w**

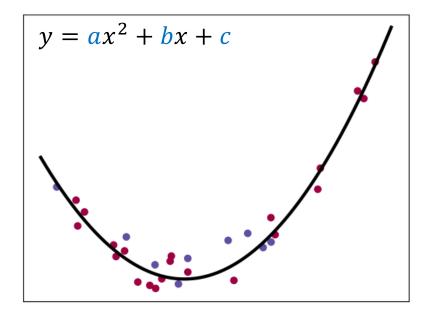


Concatenate 1 at the end of each column vector

Recap: 1-dimensional case

- The closed-form solution is applicable to some non-linear regression
 - \circ Polynomial, or when $f(x_i)$ is the linear combination of terms derived by x.
 - o a, b, c just are just to indicate what to estimate. You can change the notations.





Recap: 1-dimensional case

$$y = g + cx + bx^2 + ax^3$$

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix}$$

$$\widetilde{X} = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_1 & x_2 & \cdots & x_N \\ x_1^2 & x_2^2 & \cdots & x_N^2 \\ x_1^3 & x_2^3 & \cdots & x_N^3 \end{bmatrix}$$

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix} \qquad \widetilde{\mathbf{X}} = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_1 & x_2 & \cdots & x_N \\ x_1^2 & x_2^2 & \cdots & x_N^2 \\ x_1^3 & x_2^3 & \cdots & x_N^3 \end{bmatrix} \qquad \widetilde{\mathbf{X}}^T = \begin{bmatrix} 1 & x_1 & x_1^2 & x_1^3 \\ 1 & x_2 & x_2^2 & x_2^3 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_N & x_N^2 & x_N^3 \end{bmatrix} \qquad \widetilde{\mathbf{w}} = \begin{bmatrix} g \\ c \\ b \\ a \end{bmatrix}$$

$$\widetilde{\boldsymbol{w}} * = (\widetilde{\boldsymbol{X}}\widetilde{\boldsymbol{X}}^T)^{-1}\widetilde{\boldsymbol{X}}\boldsymbol{y}$$

 x_i : the i-th data instance

x[d]: the d-th dimension of a data instance

 x_i [d]: the d-th dimension of the i-th data instance

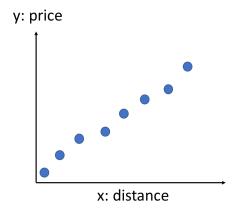
Topics

- Linear w.r.t. x or w
- Non-linear least squares
 - o Estimate parameters for non-linear functions w/o closed form solutions
- Gradient descent
 - Alternate solution methods
 - Not limited to SSE error functions

We will focus on 1-dimensional cases first!

In finding the solution

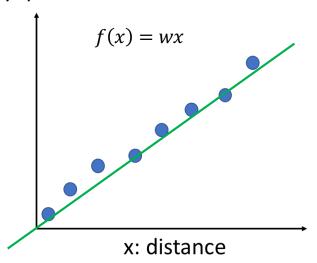
- The training data instances (x_i, y_i) are fixed
- The variables are the parameters to estimate
- A machine learning algorithm finds the parameters to minimize E
- For the users, we care about given a future x, what the prediction will be



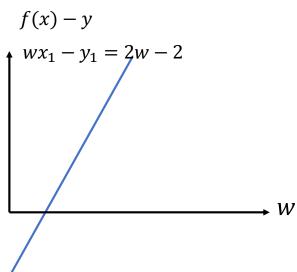
$$E = \sum_{i} [y_i - f(\mathbf{x}_i)]^2$$

Different figures

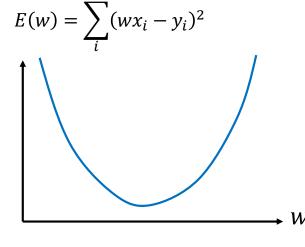




For simplicity, only estimate w



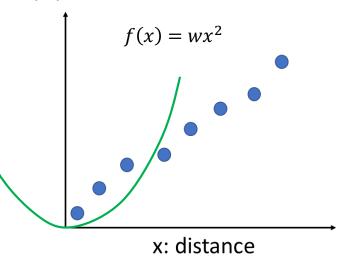
Let
$$(x_1, y_1) = (2, 2)$$



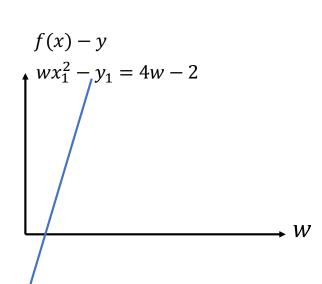
Parabolic!

Different figures

y: price

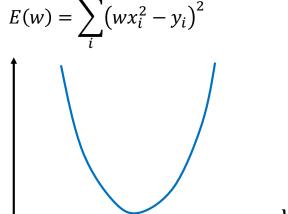


Nonlinear regression as $f(x) = wx^2$ is not a straight line w.r.t. x.



Let
$$(x_1, y_1) = (2, 2)$$

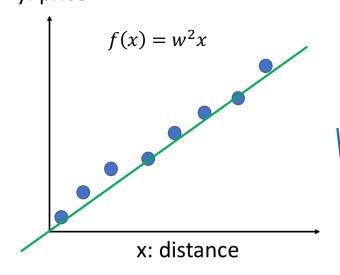
Still a straight line w.r.t. w!

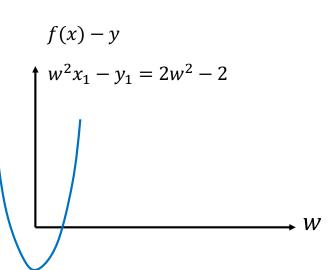


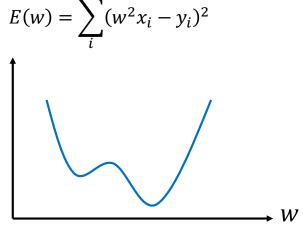
Parabolic!

Different figures









Linear regression as $f(x) = w^2x$ is a straight line w.r.t. x.

Let $(x_1, y_1) = (2, 2)$

Not a straight line w.r.t. w!

Not parabolic!

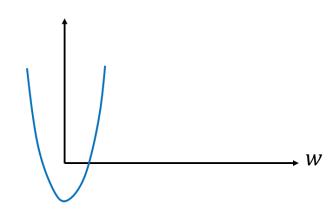
When ...

To indicate that w is the parameter!

- Given a random (x, y), when the prediction function f(x; w) or f(x; w) yis not a straight line (i.e., nonlinear) w.r.t. w
 - o there may be no closed form solution
 - o E may not be parabolic anymore
- The closed form solution $\widetilde{\boldsymbol{w}} * = \left(\widetilde{\boldsymbol{X}}\widetilde{\boldsymbol{X}}^T\right)^{-1}\widetilde{\boldsymbol{X}}\boldsymbol{y}$ exists of for the SSE loss $E(w) = \sum_i (f(x;w) y_i)^2$

 - \circ for f(x; w) linear w.r.t. w
- We say linear or nonlinear regression, w.r.t. x

$$w^2x_1 - y_1 = 2w^2 - 2$$



Let
$$(x_1, y_1) = (2, 2)$$

How to check if f(x; w) linear w.r.t. x or w?

If you want to check w.r.t. x

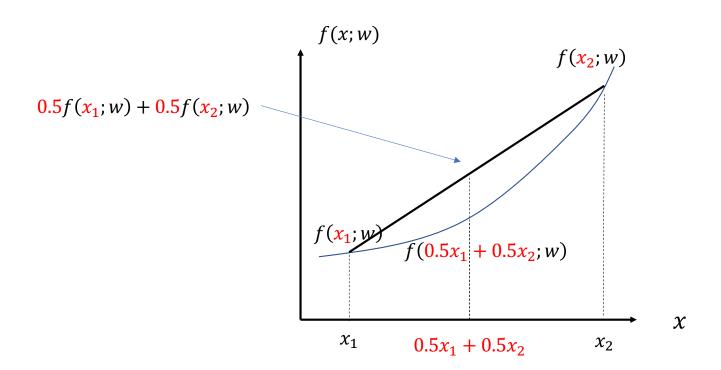
$$\circ$$
 Ex. $f(x; w) = wx$

- Find 2 random x: x_1 and x_2
 - \circ Ex. $x_1 = 1, x_2 = 2$
- Find 2 random scalars (even for *D*-dimensional cases): v_1 and v_2

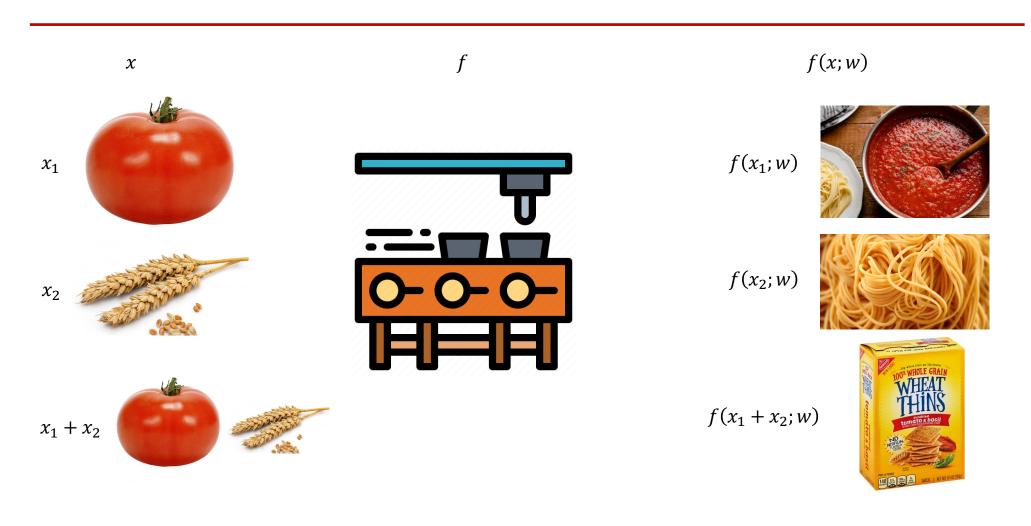
$$\circ$$
 Ex. $v_1 = 1$, $v_2 = 2$

- If $f(v_1x_1 + v_2x_2; w) = v_1f(x_1; w) + v_2f(x_2; w)$
 - $\circ f(x; w)$ is linear w.r.t. x; otherwise, not
 - \circ Need to hold for all possible x_1, x_2, v_1, v_2
- Practice:
 - \circ Is f(x; w) linear w.r.t. x?

Interpretation 1



Interpretation 2



How to check if f(x; w) linear w.r.t. x or w?

If you want to check w.r.t. w

$$\circ$$
 Ex. $f(x; w) = wx^2$

- Find 2 random w: w_1 and w_2
 - \circ Ex. $w_1 = 1, w_2 = 2$
- Find 2 random scalars (even for *D*-dimensional cases): v_1 and v_2

$$\circ$$
 Ex. $v_1 = 1$, $v_2 = 2$

- If $f(x; v_1w_1 + v_2w_2) = v_1f(x; w_1) + v_2f(x; w_2)$
 - $\circ f(x; w)$ is linear w.r.t. w; otherwise, not
 - \circ Need to hold for all possible w_1, w_2, v_1, v_2
- Practice:
 - \circ Is f(x; w) linear w.r.t. w?

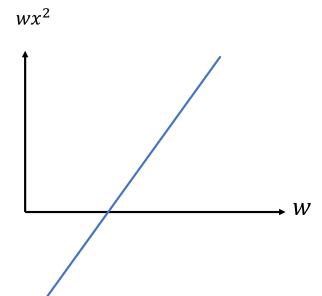
Practice

- wx^2
 - Is it linear w.r.t. w?

$$\circ$$
 Let $w_1 = 1$, $w_2 = 2$, $v_1 = 1$, $v_2 = 2$

$$\circ (v_1 w_1 + v_2 w_2) x^2 = 5x^2$$

$$\circ v_1(w_1x^2) + v_2(w_2x^2) = 5x^2$$



Practice

- wx^2
 - \circ Is it linear w.r.t. x?

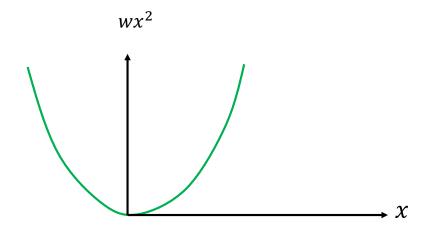
Practice

- wx^2
 - \circ Is it linear w.r.t. x?

$$\circ$$
 Let $x_1 = 1$, $x_2 = 2$, $v_1 = 1$, $v_2 = 2$

$$\circ w(v_1 x_1 + v_2 x_2)^2 = 25w$$

$$\circ v_1(wx_1^2) + v_2(wx_2^2) = 9w$$



Topics

- Linear w.r.t. x or w
- Non-linear least squares
 - o Estimate parameters for non-linear functions
- Gradient descent
 - Alternate solution methods
 - Not limited to SSE error functions

Non-linear regression

- For many cases when f(x; w) is not linear w.r.t. x, f(x; w) is not linear w.r.t. w
- Examples:

$$\circ f(x; w) = e^{wx}$$

$$\circ f(x; w) = \sin(wx)$$

- ullet For non-linear regression, $oldsymbol{x}$ and $oldsymbol{w}$ very likely have different dimensionality
- Example:

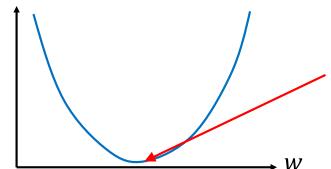
$$\circ f(x; w) = e^{w[1]x} + \sin(w[2]x)$$

• We will say $oldsymbol{x}$ is D-dimensional and $oldsymbol{w}$ is K-dimensional

When f(x; w) is nonlinear w.r.t. w

- Squared-error function is no longer parabolic w.r.t. the parameters
 - o Can not assume local minimums of error functions are always global minimums
- Often no closed-form solution
 - Need numerical approximation

$$E(w) = \sum_{i} (wx_i - y_i)^2$$

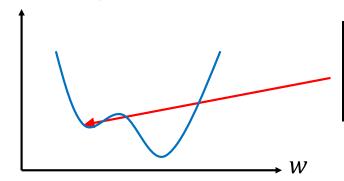


- $\frac{\partial E}{\partial w} = 0$ means global minimums
- Closed-forms to find w^* for $\frac{\partial E}{\partial w} = 0$

When f(x; w) is nonlinear w.r.t. w

- Squared-error function is no longer parabolic w.r.t. the parameters
 - o Can not assume local minimums of error functions are always global minimums
- Often no closed-form solution
 - Need numerical approximation

$$E(w) = \sum_{i} (f(x_i; w) - y_i)^2$$

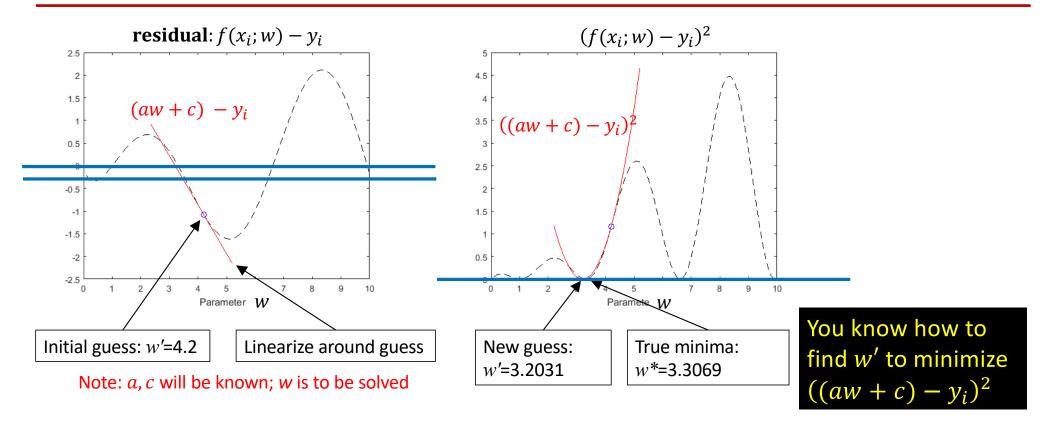


- $\frac{\partial E}{\partial w} = 0$ does not mean global minimums
- No closed-forms to find w^* for $\frac{\partial E}{\partial w} = 0$

How to solve? Approximation via linearization

- The problem is the nonlinear f(x; w) w.r.t. parameters w
 - So non-parabolic squared error function E w.r.t. parameters w
- Idea: force the model f be linear w.r.t. parameters
 - Local around <u>current guess</u> for parameters
 - This means it must be possible to approximate a function with a line locally
 - Use <u>truncated Taylor series</u>
 - Solve linear system as with <u>linear least squares</u>
 - Update <u>current guess</u> and repeat
- Note: ${m w}$ may have a very different dimensionality from ${m x}$

Local Linearization (1-dim cases)



Want to find where the **red line** intersect the **blue line**!

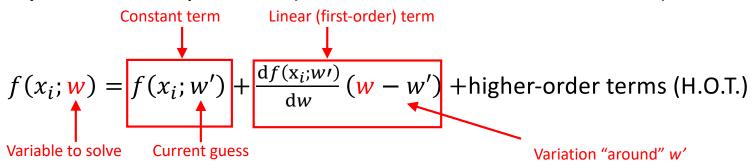
Very close even after just one iteration!

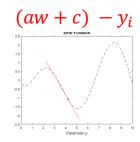
Local Linearization (a 1-dim case for w and x)

Recall the model (per data point i)

$$f(x_i; w)$$
 Read: f of x_i given parameters w

Taylor series expansion (assume w is one-dimensional)





Linearize: drop higher-than-linear terms

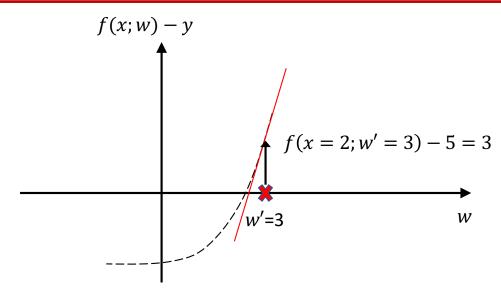
$$f(x_i; \mathbf{w}) - y_i \approx f(x_i; \mathbf{w}') + \frac{\mathrm{d}f(\mathbf{x}_i; \mathbf{w}')}{\mathrm{d}w} (\mathbf{w} - \mathbf{w}') - y_i = J(\mathbf{w}') \Delta \mathbf{w} - \Delta y_i(\mathbf{w}')$$
$$\Delta y_i(\mathbf{w}') = y_i - f(x_i; \mathbf{w}') \qquad \Delta \mathbf{w} = \mathbf{w} - \mathbf{w}' \qquad \underline{\mathsf{Jacobian}} : J_i(\mathbf{w}') = \frac{\mathrm{d}f(x_i; \mathbf{w}')}{\mathrm{d}w}$$

Local Linearization (a 1-dim case for w and x)

• Example:

$$\circ f(x;w) = x^w$$

$$\circ (x,y) = (2,5)$$



Initial guess

$$\circ$$
 $w'=3$

$$0 \frac{df(x=2; w'=3)}{dw} = \log(2) \times 2^3 = 5.545$$

$$0 J(w') \Delta w - \Delta y_i(w') = 5.545(w-3) - (5-8) = 5.545(w-3) + 3$$

Minimizing the "sum of square error"

• Then, sum of squared error:

$$E'(\Delta w) = \sum_{i} [J_i(w')\Delta w - \Delta y_i(w')]^2 = \sum_{i} [J_i\Delta w - \Delta y_i]^2$$

Note, to simplify, we are not stating the dependencies on the current guess w', but it is always there!

• "Least squares" approximation (take derivative and set it to 0)

$$0 = \frac{\mathrm{d}}{\mathrm{d}\Delta w} E'(\Delta w) = \sum_{i} \frac{\mathrm{d}}{\mathrm{d}\Delta w} [J_i \Delta w - \Delta y_i]^2$$

$$=2\sum_{i}\left[J_{i}\Delta w-\Delta y_{i}\right]\frac{\mathrm{d}}{\mathrm{d}\Delta w}\left[J_{i}\Delta w-\Delta y_{i}\right]$$

$$=2\sum_{i}[J_{i}\Delta w-\Delta y_{i}]J_{i}$$

Minimizing the sum of square error (cont.)

• "Least squares" approximation (cont.)

$$\sum_{i} J_{i} \, \Delta y_{i} = \sum_{i} J_{i} \, J_{i} \Delta w$$

Rewriting in matrix form gives:

$$J^{T} \Delta y = (J^{T}J) \Delta w \qquad \Delta w = (J^{T}J)^{-1}J^{T} \Delta y$$

$$J = \begin{bmatrix} J_{1} \\ \vdots \\ J_{N} \end{bmatrix}, \qquad \Delta y = \begin{bmatrix} \Delta y_{1} \\ \vdots \\ \Delta y_{N} \end{bmatrix} \qquad \text{Reminder: These should be } J_{i}(w') \text{ and } \Delta y_{i}(w') \text{ as they change with the current guess } w'!$$

Note: Compare with linear least squares

$$\circ$$
 LLSQ reminder: $\widetilde{X}y = (\widetilde{X}\widetilde{X}^T)\widetilde{w}$

Local Linearization (a 1-dim case for w and x)

• Example:

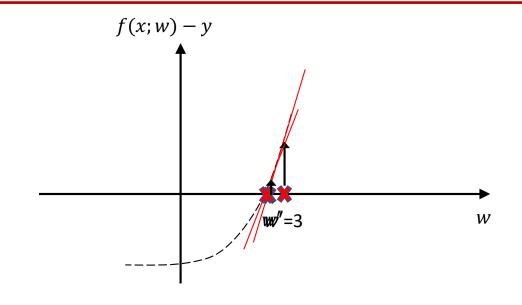
$$\circ f(x; w) = x^w$$

$$\circ (x, y) = (2, 5)$$

Algorithm

- \circ Initial: w'=3
- If not converge:

•
$$w' := w' + \Delta w$$



Gauss-Newton Algorithm

- 1. Get initial guess for parameters $\widehat{w}^{(1)}$
- 2. For $t=1...\infty$ (or until we decide to stop)
 - 1. Find error for current guess $(\widehat{w}^{(t)})$ $\Delta y_i = y_i - f(\mathbf{x}_i; \widehat{w}^{(t)})$
 - 2. Find Jacobian around current guess

$$J_i = \frac{df(\mathbf{x}_i:\widehat{w}^{(t)})}{dw}$$

3. Calculate change in guess

$$\Delta \boldsymbol{w} = (\boldsymbol{J}^T \boldsymbol{J})^{-1} \boldsymbol{J}^T \Delta \boldsymbol{y}$$

4. Make new guess

$$\widehat{w}^{(t+1)} = \widehat{w}^{(t)} + \Delta \widehat{w}$$

Example

Log-linear model

$$f(x; \mathbf{w}) = e^{wx}$$

Data points

$$\circ x_1 = 0.5, y_1 = 1$$

$$0 x_2 = 1.5, y_2 = 2$$

$$\circ x_3 = 3, y_3 = 3$$

Example - Jacobian

• Equation is simple, so closed-form Jacobian is possible

$$\frac{df(x; \mathbf{w})}{dw} = x e^{wx}$$

$$J(\mathbf{w}) = \begin{bmatrix} \frac{df(x_1; \mathbf{w})}{dw} \\ \frac{df(x_2; \mathbf{w})}{dw} \\ \frac{df(x_3; \mathbf{w})}{dw} \end{bmatrix} = \begin{bmatrix} 0.5 e^{0.5w} \\ 1.5 e^{1.5w} \\ 3 e^{3w} \end{bmatrix}$$

Example

• Initial $\widehat{w}^{(1)} = 1$

$$J(\hat{w}^{(1)}) = \begin{bmatrix} \frac{df(x_1; \hat{w}^{(1)})}{dw} \\ \frac{df(x_2; \hat{w}^{(1)})}{dw} \\ \frac{df(x_3; \hat{w}^{(1)})}{dw} \end{bmatrix} = \begin{bmatrix} 0.5 e^{0.5 \hat{w}^{(1)}} \\ 1.5 e^{1.5 \hat{w}^{(1)}} \\ 3 e^{3 \hat{w}^{(1)}} \end{bmatrix} = \begin{bmatrix} 0.5 e^{0.5} \\ 1.5 e^{1.5} \\ 3 e^{3w} \end{bmatrix}$$

$$x_1 = 0.5, y_1 = 1$$

 $x_2 = 1.5, y_2 = 2$
 $x_3 = 3, y_3 = 3$

$$\Delta \mathbf{y}(\widehat{w}^{(1)}) = \begin{bmatrix} y_1 - f(\mathbf{x}_1; \widehat{w}^{(1)}) \\ y_2 - f(\mathbf{x}_2; \widehat{w}^{(1)}) \\ y_3 - f(\mathbf{x}_3; \widehat{w}^{(1)}) \end{bmatrix} = \begin{bmatrix} 1 - e^{0.5} \widehat{w}^{(1)} \\ 2 - e^{1.5} \widehat{w}^{(1)} \\ 3 - 3e^{3} \widehat{w}^{(1)} \end{bmatrix} = \begin{bmatrix} 1 - e^{0.5} \\ 2 - e^{1.5} \\ 3 - 3e^{3} \end{bmatrix}$$

Local Linearization (let w be K-dim)

Recall the model (per data point i)

Taylor series expansion

$$f(\boldsymbol{x}; \boldsymbol{w}) = f(\boldsymbol{x}_i; \boldsymbol{w}') + \sum_{k=1}^K \frac{\partial f(\boldsymbol{x}_i; \boldsymbol{w}')}{\partial w[k]} (w[k] - w'[k]) + \text{higher-order terms (H.O.T.)}$$

$$= f(\boldsymbol{x}_i; \boldsymbol{w}') + \left[\frac{\partial f(\boldsymbol{x}_i; \boldsymbol{w}')}{\partial w[1]} \dots \frac{\partial f(\boldsymbol{x}_i; \boldsymbol{w}')}{\partial w[K]}\right] \begin{bmatrix} (w[1] - w'[1]) \\ \vdots \\ (w[K] - w'[K]) \end{bmatrix} + \text{H.O.T.}$$

$$= f(\boldsymbol{x}_i; \boldsymbol{w}') + \nabla_w f(\boldsymbol{w}')^T (\boldsymbol{w} - \boldsymbol{w}') + \text{H.O.T.}$$

Local Linearization (let w be K-dim)

Linearize: drop higher-than-linear terms

$$f(\mathbf{x}; \mathbf{w}) - y_i \approx f(\mathbf{x}; \mathbf{w}') + \left\{ \sum_{k=1}^K \frac{\partial f(\mathbf{x}_i; \mathbf{w}')}{\partial w[k]} (w[k] - w'[k]) \right\} - y_i$$
$$= \left\{ \sum_{k=1}^K J_{ik}(\mathbf{w}') \Delta w[k] \right\} - \Delta y_i(\mathbf{w}')$$

$$\Delta y_{i}(w') = y_{i} - f(x; w')$$

$$\Delta w = w - w'$$

$$\exists \text{Jacobian matrix: } J = \begin{bmatrix} \frac{\partial f(x_{1}; w')}{\partial w[1]} & \cdots & \frac{\partial f(x_{1}; w')}{\partial w[K]} \\ \vdots & \ddots & \vdots \\ \frac{\partial f(x_{N}; w')}{\partial w[1]} & \cdots & \frac{\partial f(x_{N}; w')}{\partial w[K]} \end{bmatrix}$$

$$J_{ik} = J[i, k] = \frac{\partial f(x_{i}; w')}{\partial w[k]}$$

Minimizing the "sum of square error" (let w be K-dim)

• Then, sum of squared error:

$$E'(\Delta \mathbf{w}) = \sum_{i} \left[\left\{ \sum_{k=1}^{K} J_{ik}(\mathbf{w}') \Delta \mathbf{w}[k] \right\} - \Delta y_{i}(\mathbf{w}') \right]^{2} = \sum_{i} \left[\left\{ \sum_{k=1}^{K} J_{ik} \Delta \mathbf{w}[k] \right\} - \Delta y_{i} \right]^{2}$$

• "Least squares" approximation (take partial derivatives and set them to 0)

$$0 = \frac{\partial}{\partial \Delta w[k]} E'(\Delta w) = \sum_{i} \frac{\partial}{\partial \Delta w[k]} \left[\left\{ \sum_{k=1}^{K} J_{ik} \Delta w[k] \right\} - \Delta y_{i} \right]^{2}$$

$$= 2 \sum_{i} \left[\left\{ \sum_{j=1}^{K} J_{ij} \Delta w[k] \right\} - \Delta y_{i} \right] \frac{\partial}{\partial \Delta w[k]} \left[\left\{ \sum_{k=1}^{K} J_{ik} \Delta w[k] \right\} - \Delta y_{i} \right]$$

$$= 2 \sum_{i} \left[\left\{ \sum_{j=1}^{K} J_{ij} \Delta w[k] \right\} - \Delta y_{i} \right] J_{ik}$$

Minimizing the "sum of square error" (let w be K-dim)

• "Least squares" approximation (cont.)

$$\sum_{i} J_{ik} \, \Delta y_i = \sum_{i} \sum_{j=1}^{K} J_{ik} J_{ij} \Delta w[k]$$

Rewriting in matrix form gives:

$$\boldsymbol{J}^T \Delta \boldsymbol{y} = (\boldsymbol{J}^T \boldsymbol{J}) \ \Delta \boldsymbol{w}$$

Reminder: These should be $J_i(\mathbf{w}')$ and $\Delta y_i(\mathbf{w}')$ as they change with the current guess \mathbf{w}' !

Gauss-Newton Algorithm (let w be K-dim)

- 1. Get initial guess for parameters $\widehat{\boldsymbol{w}}^{(1)}$
- 2. For $t=1...\infty$ (or until we decide to stop)
 - 1. Find error for current guess ($\hat{\boldsymbol{w}}^{(t)}$) $\Delta y_i = y_i f(\mathbf{x}_i; \hat{\boldsymbol{w}}^{(t)})$
 - 2. Find Jacobian around current guess

$$J_{ik} = \frac{\partial f(\mathbf{x}_i; \widehat{\mathbf{w}}^{(t)})}{\partial w[k]}$$

3. Calculate change in guess

$$\Delta \mathbf{w} = (\mathbf{J}^T \mathbf{J})^{-1} \mathbf{J}^T \Delta \mathbf{y}$$

4. Make new guess

$$\widehat{\boldsymbol{w}}^{(t+1)} = \widehat{\boldsymbol{w}}^{(t)} + \Delta \widehat{\boldsymbol{w}}$$

Example

Log-linear model

$$f(x; \mathbf{w}) = e^{w[1]x + w[2]}$$

Data points

$$\circ x_1 = 0.5, y_1 = 1.6487$$

$$\circ x_2 = 1.5, y_2 = 0.6065$$

$$\circ x_3 = 3, y_3 = 0.1353$$

Example - Jacobian

• Equation is simple, so closed-form Jacobian is possible

$$\frac{\partial f(x; \mathbf{w})}{\partial w[1]} = x e^{w[1]x + w[2]}$$
$$\frac{\partial f(x; \mathbf{w})}{\partial w[2]} = e^{w[1]x + w[2]}$$

$$J(\mathbf{w}) = \begin{bmatrix} \frac{\partial f(x_1; \mathbf{w})}{\partial w[1]} & \frac{\partial f(x_1; \mathbf{w})}{\partial w[2]} \\ \frac{\partial f(x_2; \mathbf{w})}{\partial w[1]} & \frac{\partial f(x_2; \mathbf{w})}{\partial w[2]} \\ \frac{\partial f(x_3; \mathbf{w})}{\partial w[1]} & \frac{\partial f(x_3; \mathbf{w})}{\partial w[2]} \end{bmatrix} = \begin{bmatrix} 0.5 e^{0.5w[1] + w[2]} & e^{0.5w[1] + w[2]} \\ 1.5 e^{1.5w[1] + w[2]} & e^{1.5w[1] + w[2]} \\ 3 e^{3w[1] + w[2]} & e^{3w[1] + w[2]} \end{bmatrix}$$

Problems

- Similar problems to linear least squares
 - Matrix inverse is expensive

$$\Delta \boldsymbol{w} = (\boldsymbol{J}^T \, \boldsymbol{J})^{-1} \boldsymbol{J}^T \, \Delta \boldsymbol{y}$$

- Outliers!
- Initial guess
 - Very sensitive to choice of values
 - Cannot guarantee global minimum
 - O Where do we get it from?
 - Expert knowledge
 - OR, try random values (multiple times!)

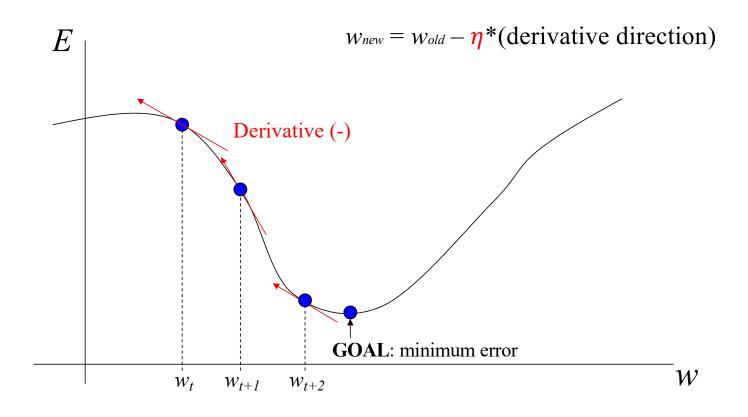
Topics

- Brief overview of machine learning: Part 3
 - Training vs. testing
 - Parameter estimation
- Non-linear least squares
 - Estimate parameters for non-linear functions
- Gradient descent
 - Alternate solution methods
 - Not limited to SSE error functions

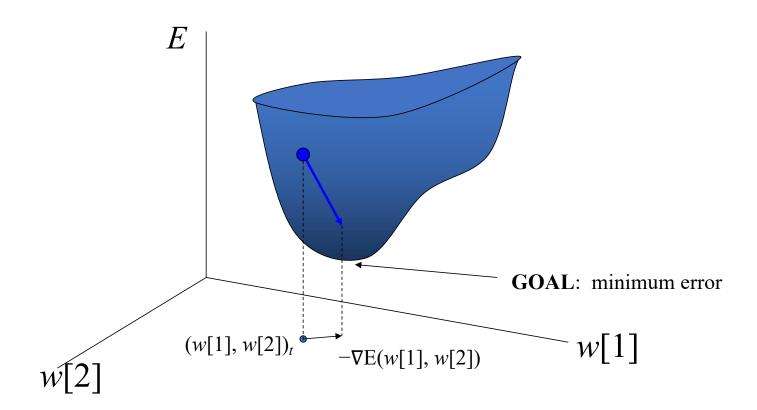
Gradient Descent (GD)

- Simple (and faster?) method for optimization
 - $\circ f(x; w)$ linear or nonlinear w.r.t. parameters w
 - Sum or square losses or other losses
- A general method to find local minimum for a multi-parameter error function
 - Search through "parameter space" to find minimum error
- Main idea
 - Evaluate error function for <u>current guess</u> of parameters
 - Determine change in parameters that <u>decrease error</u>
 - From gradient of error function
 - Update parameters in this new change
 - Repeat process until converge (or hit max iterations)

Problem Visualization



Problem Visualization



Why does gradient descent work?

- Recall the Taylor expansion
 - \circ But now apply it to the loss function E(w) directly, not to $f(x_i; w)$

$$E(\mathbf{w}) = E(\mathbf{w}') + \sum_{k=1}^{K} \frac{\partial E(\mathbf{w}')}{\partial w[k]} (w[k] - w'[k]) + \text{higher-order terms (H.O.T.)}$$

$$= E(\mathbf{w}') + \nabla_{\mathbf{w}} E(\mathbf{w}')^{T} (\mathbf{w} - \mathbf{w}') + \text{H.O.T.} = E(\mathbf{w}') + \nabla_{\mathbf{w}} E(\mathbf{w}')^{T} \Delta \mathbf{w} + \text{H.O.T.}$$

- Approximate E(w) by the linear terms: $E'(\Delta w) = E(w') + \nabla_w E(w')^T \Delta w$
- Which direction, i.e., Δw with $||\Delta w||_2^2 = 1$, can minimize $E'(\Delta w)$?
 - o Answer: direction of $\Delta w = -\frac{\nabla_w E(w')}{\|\nabla_w E(w')\|_2}$
 - \circ Interpretation: move in the direction of $-\nabla_{w}E(w')$ is the most efficient for minimization
 - \circ Only work for small Δw (so a scale η) due to the nature of linear Taylor expansion

Algorithm

Initialize parameters:

$$\widehat{\mathbf{w}}^{(1)} = [w[1], w[2], ..., w[K]]^T$$

Error function:

 $E(\mathbf{w})$

Compute gradients:

$$\nabla E(\hat{\boldsymbol{w}}^{(t)}) = \begin{bmatrix} \frac{\partial E(\hat{\boldsymbol{w}}^{(t)})}{\partial w[1]} \\ \vdots \\ \frac{\partial E(\hat{\boldsymbol{w}}^{(t)})}{\partial w[K]} \end{bmatrix}$$

Update parameters:

$$\widehat{\boldsymbol{w}}^{(t+1)} = \widehat{\boldsymbol{w}}^{(t)} - \underbrace{\eta} \cdot \nabla E(\widehat{\boldsymbol{w}}^{(t)})$$

[Repeat until converges]

Learning rate

Compute Gradients: Least Squares (w is 1-dim)

Error function:
$$E(\mathbf{w}) = \sum_{i} [f(\mathbf{x}_i; \mathbf{w}) - y_i]^2$$

Error gradient: $\nabla E(w)$

$$\frac{dE(\widehat{\boldsymbol{w}}^{(t)})}{dw} = \frac{\partial}{\partial w} \sum_{i} [f(\boldsymbol{x_i}; \widehat{\boldsymbol{w}}^{(t)}) - y_i]^2$$

$$=2\sum_{i}(f(\boldsymbol{x_i}; \widehat{\boldsymbol{w}}^{(t)})-y_i)\frac{df(\boldsymbol{x_i}; \widehat{\boldsymbol{w}}^{(t)})}{dw}$$

Recall: Gauss-Newton

$$2\sum_{i} (f(\mathbf{x}_{i}; \widehat{\mathbf{w}}^{(t)}) - y_{i}) \frac{df(\mathbf{x}_{i}; \widehat{\mathbf{w}}^{(t)})}{\partial w} = 2\sum_{i} -\Delta y_{i} J_{i}(\widehat{\mathbf{w}}^{(t)}) \qquad \qquad \nabla E(\mathbf{w}) = -2 J^{T}(\mathbf{w}) \Delta y(\mathbf{w})$$

Compute Gradients: Least Squares (w is K-dim)

Error function:
$$E(\mathbf{w}) = \sum_{i} [f(\mathbf{x}_i; \mathbf{w}) - y_i]^2$$

Error gradient:
$$\nabla E(\mathbf{w})$$

$$\frac{\partial E(\widehat{\boldsymbol{w}}^{(t)})}{\partial w[k]} = \frac{\partial}{\partial w[k]} \sum_{i} [f(\boldsymbol{x_i}; \widehat{\boldsymbol{w}}^{(t)}) - y_i]^2$$

$$=2\sum_{i}(f(\mathbf{x}_{i};\widehat{\mathbf{w}}^{(t)})-y_{i})\frac{\partial f(\mathbf{x}_{i};\widehat{\mathbf{w}}^{(t)})}{\partial w[k]}$$

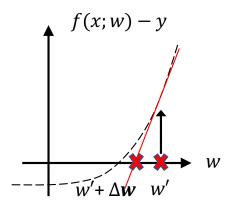
Recall: Gauss-Newton

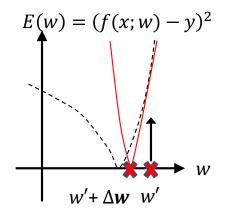
$$2\sum_{i} (f(\mathbf{x}_{i}; \widehat{\mathbf{w}}^{(t)}) - y_{i}) \frac{\partial f(\mathbf{x}_{i}; \widehat{\mathbf{w}}^{(t)})}{\partial w[k]} = 2\sum_{i} -\Delta y_{i} J_{ik}(\widehat{\mathbf{w}}^{(t)}) \qquad \qquad \nabla E(\mathbf{w}) = -2 J^{T}(\mathbf{w}) \Delta y(\mathbf{w})$$

Gradient descent vs Gauss-Newton work

Gauss-Newton

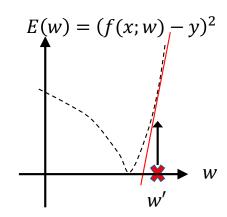
- \circ For square loss $\sum_{i} [f(x_i; w) y_i]^2$
- \circ Linearization on f(x; w) at w'
- \circ Solve the approximated least square with the closed-form solution to get Δw
- \circ Update w' by $w' + \Delta w$





Gradient descent

- \circ For any kind of loss E(w)
- \circ Linearization in E(w) at w'
- \circ Set Δw to the direction of $-\nabla_w E(w')$
- \circ Update w' by $w' + \eta \Delta w$



Gradient descent vs Gauss-Newton work

• Gauss-Newton can be seen as "quadratic approximation" of E(w)

$$E'(\Delta \mathbf{w}) = \sum_{i} [\nabla_{\mathbf{w}} E(\mathbf{w}')^{T} \Delta \mathbf{w} - \Delta y_{i}(\mathbf{w}')]^{2}$$

$$= \Delta \mathbf{w}^{T} \left(\sum_{i} \nabla_{\mathbf{w}} E(\mathbf{w}') \nabla_{\mathbf{w}} E(\mathbf{w}')^{T} \right) \Delta \mathbf{w} - 2 \sum_{i} \Delta y_{i}(\mathbf{w}')^{T} \nabla_{\mathbf{w}} E(\mathbf{w}')^{T} \Delta \mathbf{w} + \sum_{i} \Delta y_{i}(\mathbf{w}')^{T} \Delta y_{i}(\mathbf{w}')$$

- As we use the quadratic term, the approximation is good for a larger Δw
 - \circ Thus we can directly update w' by $w' + \Delta w$ (without a learning rate η)
 - \circ In practice, you can also do w' by $w' + \eta \Delta w$

Direct 2-order Taylor expansion approximation

• Taylor series expansion ($\Delta w = w - w'$)

$$E(\Delta w) = E(w') + \nabla_w E(w')^T \Delta w + \frac{1}{2} \Delta w^T H(w') \Delta w + \text{H.O.T.}$$

• Hessian matrix
$$H(w')$$

$$H(w') = \begin{bmatrix} \frac{\partial^2 E(\widehat{w}_t)}{\partial w[1]\partial w[1]} & \cdots & \frac{\partial^2 E(\widehat{w}_t)}{\partial w[1]\partial w[K]} \\ \vdots & \ddots & \vdots \\ \frac{\partial^2 E(\widehat{w}_t)}{\partial w[K]\partial w[1]} & \cdots & \frac{\partial^2 E(\widehat{w}_t)}{\partial w[K]\partial w[K]} \end{bmatrix}$$
• Newton's method

- Newton's method
 - \circ For any kind of loss E(w)
 - \circ Quadratic approximation on E(w) at w'
 - \circ Solve the quadratic approximation to get $\Delta w = -H^{-1}(w') \nabla_w E(w')$
 - \circ Update w' by $w' + \Delta w$
- Pros: faster convergence; Cons: H(w') and $H^{-1}(w')$ are time consuming

Benefits, Issues, Helpers

- Benefits of gradient descent
 - No expensive matrix operations (i.e., inversion)
 - Non-SSE error functions (as long as gradient is well formed)
 - Easy to parallelize (supercomputers!)
- Issues
 - Can be misled by local minima
 - Get stuck at smaller valleys
 - Can get stuck on a flat plain
 - Flat region in state-space function
 - Can overshoot back and forth
 - Oscillate from side to side
 - Learning rate too large
- Helpers
 - Use random-restart to handle some problems
 - Do process multiple times with random initial parameters
 - Methods exist for selecting appropriate (best) learning rate at each iteration

Summary

- Gauss-Newton non-linear least squares
 - Iterative approximation
 - Linearize model around current parameter guess
 - Apply Linear Least Squares methods to linearized model
- Gradient descent
 - "Valley descending" (opposed to "hill climbing")
 - Make the best update in parameters
 - In direction of negative gradient of error function