The Effect of Children on Women's Labor Supply: A Bayesian Replication Analysis

Jennah Gosciak

May 16, 2022

```
## load data
df <- read_dta("../00_data/sample1.dta")</pre>
df_samp <- df %>%
  sample_n(1000)
df_samp <- df_samp %>%
 mutate(across(c("age", "age_fbirth"), ~ . - mean(., na.rm = T))) %>%
  mutate(l_incwage = if_else(incwage <= 0, log(1), log(incwage)),</pre>
         l_wkswork1 = if_else(wkswork1 <= 0, log(1), log(wkswork1)),</pre>
         incwage_mod = if_else(incwage <= 0, 1, incwage))</pre>
df_samp %>%
  group_by(samesex) %>%
 summarize(n = n())
## # A tibble: 2 x 2
     samesex
       <dbl> <int>
        0 473
## 1
               527
           1
```

Run outcome with Gamma distribution

- Uses log link
- Similar to implementation in rstanarm

```
writeLines(readLines("linear_gamma.stan"))
```

```
## #include quantile_functions.stan
## data {
##
     int<lower = 0> N; // number of observations
     int<lower = 0> K; // number of predictors
##
##
    matrix[N, K] X; // matrix of predictors
                      // outcomes
##
    vector[N] y;
##
     int<lower = 0, upper = 1> prior_only; // ignore data?
##
    vector[K + 2] m;
                                             // prior medians
##
    vector<lower = 0>[K + 2] scale;
                                                 // prior IQRs
##
    real r;
```

```
## }
## parameters {
##
    vector[K] beta;
##
     real alpha;
##
     real<lower = 0> shape;
## }
##
## model { // log likelihood, equivalent to target += normal_lpdf(y | alpha + X * beta, sigma)
##
     if (!prior_only) {
##
       vector[N] mu = alpha + X * beta;
##
       for (i in 1:N) target += gamma_lpdf(y | shape, shape/exp(mu[i]));
##
    target += normal_lpdf(beta | m[1], scale[1]); // ^ important
##
     target += normal_lpdf(alpha | m[2:K + 1], scale[2:K + 1]);
##
     target += exponential_lpdf(shape | r); // exponential
##
## }
##
## generated quantities {
    vector[N] log_lik;
##
##
    vector[N] yrep;
##
##
       vector[N] mu = alpha + X * beta;
##
       for (n in 1:N) {
         log_lik[n] = gamma_lpdf(y[n] | shape, shape / exp(mu[n]));
##
##
         yrep[n] = gamma_rng(shape, shape / exp(mu[n]));
##
       }
##
     }
## }
m \leftarrow rep(-0.1, 10)
s < - rep(1, 10)
stan_data \leftarrow list(N = nrow(df_samp), K = 8,
                                        y = df_samp$incwage,
                                        X = df_samp[, c("cnum_mt2", "age", "age_fbirth", "f_boy",
                                                    "s_boy", "r_black", "hisp", "r_oth")],
                                        prior_only = FALSE, m = m,
                                        scale = s, r = 0.5)
post_gamma <- stan("linear_gamma.stan", data = stan_data,</pre>
                                         seed = 12345)
# print output
print(post_gamma, pars = c("alpha", "beta", "shape"))
## Inference for Stan model: linear_gamma.
## 4 chains, each with iter=2000; warmup=1000; thin=1;
## post-warmup draws per chain=1000, total post-warmup draws=4000.
##
##
                          sd 2.5%
                                     25% 50% 75% 97.5% n_eff Rhat
           mean se_mean
                      0 0.00 8.87 8.88 8.88 8.88 8.89 2528
## alpha
           8.88
                      0 0.00 -0.01 0.00 0.00 0.00 0.01
## beta[1] 0.00
                                                          3268
## beta[2] 0.00
                      0 0.00 0.00 0.00 0.00 0.00 0.00
## beta[3] 0.00
                     0 0.00 0.00 0.00 0.00 0.00
                                                          4196
                                                                   1
## beta[4] 0.00
                     0 0.00 -0.01 0.00 0.00 0.00 0.01 3383
                    0 0.00 -0.01 0.00 0.00 0.00 0.01 3162
## beta[5] 0.00
```

```
## beta[6] 0.00
                     0 0.01 -0.01 0.00 0.00 0.01 0.01
## beta[7] 0.00
                   0 0.01 -0.02 -0.01 0.00 0.01 0.02
                                                         2787
## beta[8] 0.00
                     0 0.01 -0.03 -0.01 0.00 0.01 0.03
                                                         3029
## shape
          0.18
                     0 0.00 0.18 0.18 0.18 0.18 0.18
                                                        4836
## Samples were drawn using NUTS(diag_e) at Mon May 16 08:52:30 2022.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
pp_check(as.numeric(stan_data$y),
 rstan::extract(post_gamma, par = "yrep")$yrep[sample(1:length(stan_data$y), size = 150), ],
 ppc_dens_overlay
                                                                               -y_{rep}
```

le+00 1e+05 2e+05 3e+05

```
##
## Computed from 4000 by 1000 log-likelihood matrix
##
## Estimate SE
## elpd_loo -7618.3 125.7
## p_loo 0.0 0.0
```

15236.6 251.5

looic ## -----

loo(post_gamma)

```
## Monte Carlo SE of elpd_loo is 0.0.
##
## All Pareto k estimates are good (k < 0.5).
## See help('pareto-k-diagnostic') for details.</pre>
```