An Alternate Method for Minimizing χ^2

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ABSTRACT

In this paper, we describe an algorithm and associated software package (sfit_minimize) for maximizing the likelihood function of a set of parameters by minimizing χ^2 . The key element of this method is that the algorithm estimates the second derivative of the χ^2 function using first derivatives of the function to be fitted. These same derivatives can also be used to calculate the uncertainties in each parameter. We test this algorithm against several standard minimization algorithms in SciPy.optimize.minimize() by fitting point lens models to light curves from the 2018 Korea Microlensing Telescope Network event database. We show that for fitting microlensing events, SFit works faster than the Nelder-Mead simplex method and is more reliable than the BFGS gradient method.

1. THE DERIVATION

Our method builds upon the discussion in " χ^2 and Linear Fits" (Gould 2003), which noted that the approach to non-linear models could be expanded beyond the scope of that work. Suppose that the function we want to minimize is a function F(x) that is described by n parameters A_i (where we use A_i (instead of a_i) as a reminder that in the general case, they are non-linear). Considering the general (nonlinear) case, we can Taylor expand χ^2 , in terms of the n parameters:

$$\chi^2 = \chi_0^2 + \sum_i \frac{\partial \chi^2}{\partial A_i} A_i + \frac{1}{2} \sum_{i,j} \frac{\partial^2 \chi^2}{\partial A_i \partial A_j} A_i A_j \tag{1}$$

$$= \chi_0^2 + \sum_i D_i * A_i + \sum_{i,j} B_{ij} * A_i A_j + \dots , \qquad (2)$$

where

$$D_i \equiv \frac{\partial \chi^2}{\partial A_i} \tag{3}$$

$$B_{ij} \equiv (1/2) \frac{\partial^2 \chi^2}{\partial A_i \partial A_j} \quad . \tag{4}$$

Then,

$$\frac{\partial \chi^2}{\partial A_i} = -2 \sum_k \frac{(y_k - F(x_k))}{\sigma_k^2} \frac{\partial F(x_k)}{\partial A_i}$$
 (5)

and

$$\frac{\partial^2 \chi^2}{\partial A_i \partial A_j} = -2 \sum_k \left[\frac{1}{\sigma_k^2} \frac{\partial F(x_k)}{\partial A_i} \frac{\partial F(x_k)}{\partial A_j} + \frac{(y_k - F(x_k))}{\sigma_k^2} \frac{\partial^2 F(x_k)}{\partial A_i \partial A_j} \right] \quad . \tag{6}$$

In the special case of a linear function, $F(x) = \sum_i a_i f_i(x)$ then

$$\frac{\partial F(x)}{\partial a_i} = f_i(x) \quad \text{and} \quad \frac{\partial^2 F(x)}{\partial a_i \partial a_j} = 0,$$
 (7)

so the second term disappears, and we find that the solution (derived from second derivative of χ^2) can be expressed in terms of products of the FIRST derivatives of the general functional form. For the general case, we simply make the approximation that the second derivative term can be neglected; i.e.,

$$\frac{\partial^2 \chi^2}{\partial A_i \partial A_j} \approx -2 \sum_k \frac{1}{\sigma_k^2} \frac{\partial F(x_k)}{\partial A_i} \frac{\partial F(x_k)}{\partial A_j} \quad . \tag{8}$$

Hence, there are three ways to generalize Newton's method (actually discovered by Simpson) to multiple dimensions:

- 1. Use only first derivatives of the χ^2 function (which is what Simpson did in 1-D), the so-called gradient method.
- 2. Taylor expand χ^2 and truncate at second term, then solve this (very inexact equation) exactly by inversion of the matrix of second derivatives (Hessian).
- 3. First generalize Simpson's idea that a 1-D function is well described by its first derivative (which can easily be solved exactly) to several dimensions (i.e., assume the function is well described by a tangent plane) and solve this exactly, as is done here.

Because first derivatives are more stable than second derivatives, this algorithm could potentially be significantly more stable for situations in which the derivatives are derived numerically, which we investigate in the next section.

2. IMPLEMENTATION

We have implemented the above algorithm in the sfit_minimizer package. The goal was to make the calling sequence similar to that of SciPy.optimize.minimize():

where my_func is an object of the type sfit_minimizer.SFitFunction() in which the user defines either the model, $F(x_k)$ or the residual $y_k - F(x_k)$ calculation (i.e., the method my_func.model() or my_func.residuals()) and the partial derivatives of the function to be minimized, $\partial F(x_k)/\partial A_i$ (i.e., the method my_func.df()). The package includes a simple example (example_00_linear_fit.py) for fitting a linear model to demonstrate this usage.

The sfit_minimizer.SFitFunction() class contains methods that use the partial derivative function to calculate the next step from the D_i and B_{ij} following the method in Gould (2003) for linear functions. That is, D_i and B_{ij} are calculated from Equations 3 and 4, respectively. Then, the step size for each parameter, Δ_i , is

$$\Delta_i = \sum_j C_{ij} D_j \quad \text{where} \quad C \equiv B^{-1} \quad ,$$
(10)

which is returned by $sfit_minimizer.SFitFunction.get_step()$. The new value of A_i is calculated by $sfit_minimizer.minimize()$ to be

$$A_i = A_{i,0} + \epsilon \Delta_i \quad . \tag{11}$$

In sfit minimizer.minimize(), the user has the option to specify the value of ϵ or to make use of an adaptive step size, which starts at $\epsilon = 0.001$ and becomes larger as the minimum is approached.

Ultimately, sfit_minimizer.minimize() returns an sfit_minimizer.SFitResults() object that contains attributes similar to the object returned by SciPy.optimize.minimize(). These include the best-fit values of the parameters, x, and their uncertainties sigma (i.e., $\sigma_i = \sqrt{C_{ii}}$). For the rest of this paper, we will refer to our algorithm as SFit for brevity.

3. PERFORMANCE TEST

To test the performance of SFit, we use the package to fit point-source-point-lens models (Paczynski 1986) to a sample of microlensing events from the Korea Microlensing Telescope Network (KMT-Net; Kim et al. 2016). For comparison, we also perform the fitting using the Nelder-Mead (Gao & Han 2012), Newton-CG (Nocedal & Wright 2006), and BFGS (Nocedal & Wright 2006) algorithms in SciPy.optimize.minimize(). The Nelder-Mead algorithm is a simplex algorithm, so it only relies on evaluating the χ^2 . In contrast to our algorithm, the Newton-CG and BFGS algorithms use the jacobian of the likelihood function for the minimization. In all cases, we set tol = 1e-5.

We select our sample from microlensing events discovered in 2018 by KMTNet (Kim et al. 2018a,b,c). We use only "clear" microlensing events with reported fit parameters. We eliminate any events that were flagged as anomalous in the 2018 AnomalyFinder search (although possible finite source or buried host events were left in the sample; Gould et al. 2022; Jung et al. 2022). These cuts left 1822 events in the sample.

For this sample, we use the online, *I*-band, pySIS (Albrow et al. 2009) data from the KMTNet website (https://kmtnet.kasi.re.kr/ulens/). KMTNet takes data from three different sites and has multiple, sometimes overlapping, fields of observations. We treat data from different sites and different fields as separate datasets. For each dataset, we calculate the mean sky background and standard deviation as well as the mean and standard deviation of the full-width-at-half-max (FWHM) for each observation. We eliminate points with sky background more than 1 standard deviation above the mean or FWHM more than 3 standard deviations above the mean. This removes a large fraction of the outliers from the data.

We fit each event with a standard, point-lens model A, which is parameterized by the three Paczyński parameters: t_0 , u_0 , and $t_{\rm E}$ (for the definitions of these parameters see, e.g., Gaudi 2012). In addition, there are two flux parameters used to scale each dataset, k, to the model, A: $f_{\rm mod,k} = f_{\rm S,k}A + f_{\rm B,k}$.

We use MulensModel (Poleski & Yee 2019) to calculate the model, its derivatives, and the jacobians. We perform a linear fit to the flux parameters at each iteration using built in functions from MulensModel and use the fitting algorithm to optimize t_0 , u_0 , and t_E .

Before doing the fitting, we determined the starting values for t_0 , u_0 , and t_E using the following procedure. The KMTNet website reports an estimate for the values of these parameters. We use the value of t_0 as reported by KMTNet. To ensure an optimum starting point for the fits, we tested a series of values of $u_{0,i} = [0.01, 0.3, 0.7, 1.0, 1.5]$ and calculated $t_{E,i} \equiv t_E u_0/u_{0,i}$, where u_0 and t_E are the

values reported in the KMTNet table. We performed a linear fit of the flux parameters $(f_{S,k}, f_{B,k})$ for each set of parameters i and chose the values of $u_{0,i}$ and $t_{E,i}$ that produced the smallest χ^2 to initialize the fits.

We calculated several metrics to evaluate the performance of each algorithm. First, for a given event, we compared the χ^2 of the best-fit reported by each algorithm to the best (minimum) value reported out of the four fits. The results are given in Table 1 for several values of $\Delta \chi^2$ classified by whether or not the algorithm reported that the fit was successful ("reported success").

Each fit may be classified in one of four ways:

- True positives: both true and reported success,
- False positives: algorithm reported success, but did not find the minimum,
- True negatives: algorithm reported failure and did not find the minimum,
- False negatives: algorithm reported failure, but did find the minimum.

For the purpose of these comparisons, we define $\Delta \chi^2 < 0.1$ as a "true success." **REWRITTEN FROM HERE:** The results are visualized in Figure 1.

We also calculated the number of χ^2 function evaluations required by each algorithm for fitting each event. The maximum number allowed was 999; if the number of evaluations exceeded this, the fit was marked as a failure. Table 2 provides statistical summaries of this metric.

4. TEST RESULTS

Table 1 shows that BFGS, Nelder-Mead, and SFit all had very low false positive rates; i.e., when the algorithm reported success, it also tended to find the best-fit solution. Nelder-Mead produced the most true successes 85% out of all the algorithms, although it also required the most function evaluations. SFit was the most reliable of the algorithms; in addition to the low false positive rate, it also had a low false negative rate. Of all the algorithms, BFGS was arguably the most successful, getting within $\Delta \chi^2 = 0.1$ of the minimum 98.5% of the time, but it only reported success $\sim 60\%$ of the time (it had a high false negative rate).

The Newton-CG algorithm performed the least well in our tests. It had both an extremely high false positive rate (65%) and false negative rate (66%).

Figure 1 and the middle and bottom sections of Table 1 compare the performance of the other algorithms to that of SFit. The table shows that there is not complete overlap in performance between SFit and any other algorithm. Nearly all BFGS fits reached the minimum, but some were reported as successes and some as failures. The reported success of BFGS was uncorrelated with the performance of SFit; none of the four classifications for BFGS completely overlapped with an SFit classification. There were also a handful of BFGS fits that failed for which SFit succeeded. This was also true for the Newton-CG fits.

The vast majority of the SFit successes were also true successes for Nelder-Mead. In addition, Nelder-Mead successfully fit $\sim 40\%$ of events that SFit failed to fit. Nevertheless, there were a handful of Nelder-Mead failures that were successfully fit by SFit.

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5. CONCLUSIONS

Given these tests, the best method for fitting microlensing events depends on the desired outcome. The Nelder-Mead algorithm produces the highest success rate (85%) with no false positives, but requires more function evaluations than the other algorithms. The most efficient method would be to simply use BFGS and accept whatever results it produced as the best-fit. Our tests show that this fit will be correct 99% of the time. However, there is no way to identify the 1% of failures because they will be mixed up with the 40% of fits that are reported as failures but actually succeeded (false negatives).

To achieve a better balance of efficiency, accuracy, and reliability, we combine the SFit and BFGS algorithms, both of which have extremely low false positive rates with minimal function evaluations. Most (73%) events can be fit successfully with SFit. The remaining 27% of events can be fit using BFGS, which will increase the number of true positives to 85%. Finally, the BFGS fit for last 15% of events is probably correct, but this can be verified by re-fitting them using SFit, but initializing the fit with the best-fit parameters from the ("failed") BFGS fit.

In testing this procedure, we find that 94% of the time it successfully finds the best-fit solution and reports success (true positives) with only five false positives. Of these false positives, four had identifiable light curve characteristics that could lead to a fit failure, and none of the algorithms performed well on the fifth. Of the remaining 6% of events, for which this procedure reports a failed fit, more than half had identifiable light curve characteristics that could lead to a fit failure. Thus, the false negative rate is only 3%.

ADD a concluding sentence?

The Python implementation of this algorithm, including its specific application to microlensing events, can be found on GitHub.

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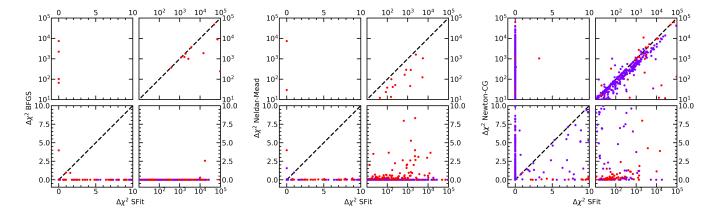


Figure 1. $\Delta\chi^2$ relative to the best-fit for BFGS, Nelder-Mead, and Newton-CG (the comparison algorithms) vs. SFit. Fits reported as successes by the comparison algorithm are plotted in purple, while those reported as failures are shown in red. Events that were fit successfully by both algorithms appear at (0, 0). In each set of four panels, the axes are split so that [0, 10] is on a linear scale, which $[10, 10^5]$ is on a log scale. The vertical bands of points at x=0 are fits that were successfully fit by SFit but failed to be fit by the comparison algorithm; purple points in those bands are false positives for the comparison algorithm. The horizontal bands of points at y=0 are points for which the comparison algorithm successfully found the minimum but SFit did not; red points in those bands are false negatives.

Table 1. Number of Fits with $\Delta \chi^2 \geq X$ of the Best-Fit

$\Delta \chi^2 <$												
Algorithm	Total	0.	1	1.0		10.0		100.0				
		N	%	N	%	N	%	N	%			
All 1822 Events:												
Algorithm Reported Success:												
BFGS	1074	1074	100	1074	100	1074	100	1074	100			
Nelder-Mead	1556	1551	100	1553	100	1554	100	1554	100			
Newton-CG	1630	572	35	727	45	888	54	1101	68			
SFit	1325	1324	100	1325	100	1325	100	1325	100			
Algorithm Reported Failure:												
BFGS	748	722	97	725	97	727	97	728	97			
Nelder-Mead	266	155	58	225	85	247	93	256	96			
Newton-CG	192	65	34	111	58	125	65	137	71			
SFit	497	1	0	7	1	54	11	255	51			
1325 Events for which SFit reported success:												
Algorithm Reported Success:												
BFGS	864	864	100	864	100	864	100	864	100			
Nelder-Mead	1321	1319	100	1320	100	1321	100	1321	100			
Newton-CG	1265	536	42	675	53	781	62	880	70			
Algorithm Reported Failure:												
BFGS	461	455	99	456	99	457	99	458	99			
Nelder-Mead	4	1	25	1	25	2	50	3	75			
Newton-CG	60	27	45	28	47	29	48	34	57			
497 Events for which SFit reported failure:												
Algorithm Reported Success:												
BFGS	210	210	100	210	100	210	100	210	100			
Nelder-Mead	235	232	99	233	99	233	99	233	99			
Newton-CG	365	36	10	52	14	107	29	221	61			
Algorithm Reported Failure:												
BFGS	287	267	93	269	94	270	94	270	94			
Nelder-Mead	262	154	59	224	85	245	94	253	97			
Newton-CG	132	38	29	83	63	96	73	103	78			

Table 2. Number of χ^2 Function Evalutions

Algorithm	Mean	Median	StdDev	Max^a
BFGS	122.1	44	217.2	830
Nelder-Mead	348.8	303	136.6	603
Newton-CG	44.5	21	80.2	633
SFit	142.5	125	121.1	999

 $[\]overline{^a}$ Fitting terminated after 999 function evaluations.