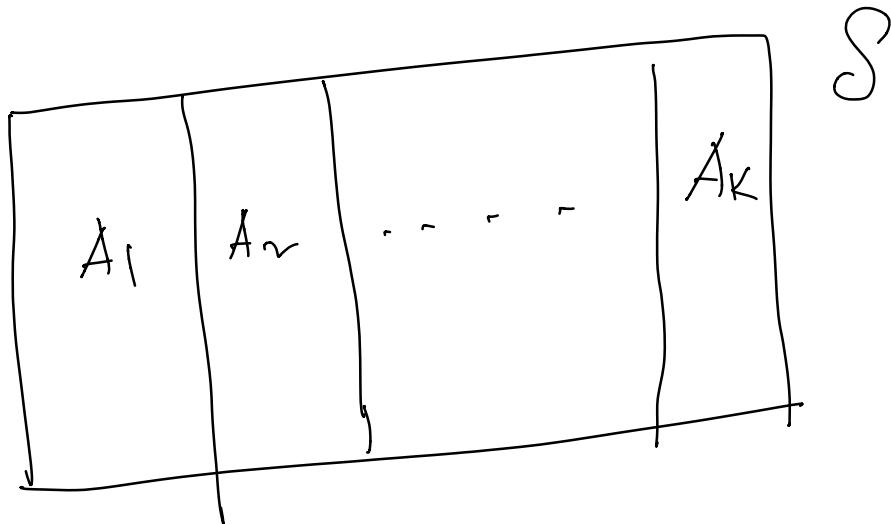


Partitioning of the sample space:

Events A_1, A_2, \dots, A_K are said to partition the sample space S if:

1. $A_i^\circ \cap A_j^\circ = \emptyset$ for $i \neq j$

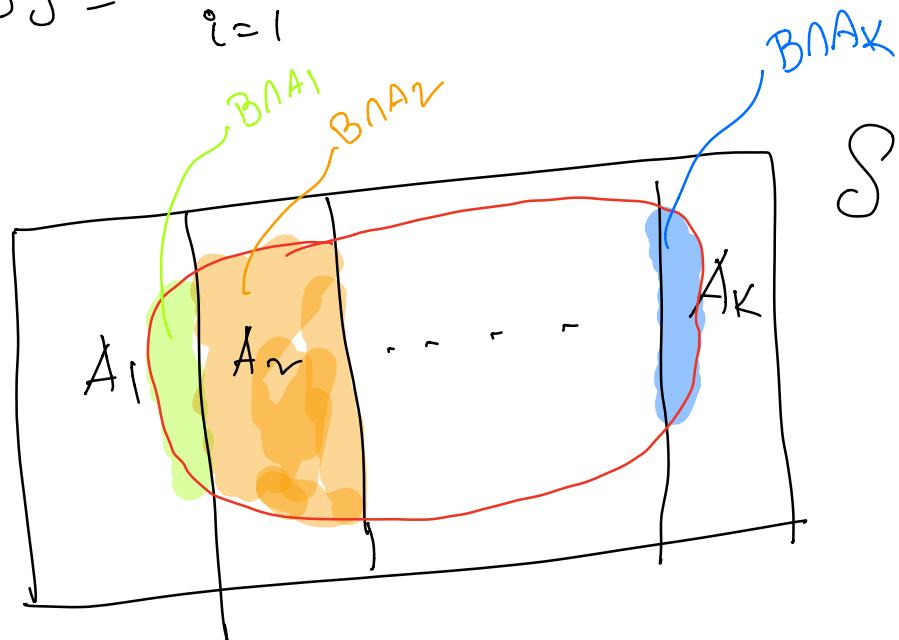
2. $\bigcup_{i=1}^K A_i^\circ = S$



Law of total probability:

Suppose A_1, \dots, A_K forms a partition of the sample space. Then for any event B in the sample space, we have

$$P[B] = \sum_{i=1}^K P[B|A_i] P[A_i]$$



Bayes Rule:

Let $\{A_1, A_2, \dots, A_K\}$ be a partition of the sample space S , and suppose each of the events A_1, A_2, \dots, A_K has nonzero probability. Let B be any event for which $P(B) > 0$. Then for each integer n ($1 \leq n \leq K$), we have Bayes formula:

$$P(A_n | B) = \frac{P(A_n) P(B|A_n)}{\sum_{j=1}^K P(A_j) P(B|A_j)}$$

Chain Rule:

For any events A and B, we have

$$P(A \cap B) = P(A) P(B|A)$$

more generally, for any events A_1, \dots, A_n ,

$$P\left(\bigcap_{i=1}^n A_i\right) = P(A_1) P(A_2 | A_1) P(A_3 | A_1, A_2) \dots P(A_n | \bigcap_{i=1}^{n-1} A_i)$$

To prove the above, we will use induction on n . The base case is $n=1$. For the base case,

$$P(A_1) = P(A_1)$$

which is trivially true. For the inductive step, let $n > 1$ and assume (the inductive hypothesis)

that

$$P\left(\bigcap_{i=1}^{n-1} A_i\right) = P(A_1) P(A_2 | A_1) \dots P(A_{n-1} | \bigcap_{i=1}^{n-2} A_i)$$

Then,

$$P(\cap_{i=1}^n A_i) = P\left(A_n \mid \cap_{i=1}^{n-1} A_i\right)$$

By the definition of conditional probability,

$$P(\cap_{i=1}^n A_i) = P(A_n \mid \cap_{i=1}^{n-1} A_i) P(\cap_{i=1}^{n-1} A_i)$$

Now, by the induction hypothesis,

$$P(\cap_{i=1}^n A_i) = P(A_n \mid \cap_{i=1}^{n-1} A_i) P(A_1) P(A_2 \mid A_1) \cdots P(A_{n-1} \mid \cap_{i=1}^{n-2} A_i)$$

This completes the proof by induction.

Practice Problems on probability basics:

1) Let's define the following events:

Z_D : A randomly selected chip is defective

Z_A : A randomly selected chip was manufactured by A.

Z_B : A randomly selected chip was manufactured by B.

Z_C : A randomly selected chip was manufactured by C.

Now, we want to compute

$$P(Z_A | Z_D), P(Z_C | Z_D)$$

From the problem statement, we know

$$P(Z_D|Z_A) = 0.005, P(Z_D|Z_B) = 0.001$$

$$P(Z_D|Z_C) = 0.01.$$

By chain rule of probability

$$P(Z_A|Z_D) = \frac{P(Z_D|Z_A) P(Z_A)}{P(Z_D)}$$

Since Z_A , Z_B and Z_C forms a partition
of the sample space, so by law of
total probability

$$P(Z_D) = P(Z_D|Z_A)P(Z_A) + P(Z_D|Z_B)P(Z_B)$$

$$+ P(Z_D|Z_C)P(Z_C)$$

$$= 0.005 \times 0.5 + 0.001 \times 0.1 + 0.01 \times 0.4$$

$$= 0.0025 + 0.0001 + 0.004$$

$$P(Z_D) = 0.0066$$

Hence,

$$P(z_A | z_D) = \frac{0.005 \times 0.5}{0.0066} = 0.3788$$

$$P(z_c | z_D) = \frac{0.01 \times 0.4}{0.0066} = 0.606$$

2

Suppose we define the following events:

T: A two headed coin is flipped

F: A fair coin is flipped

B: A biased coin is flipped

H: The flipped coin shows a head

a) Since T, F, and B forms a partition of the sample space, so by law of total probability

$$P(H) = P(H|T)P(T) + P(H|F)P(F) + P(H|B)P(B)$$

$$= P(T) + \frac{1}{2}P(F) + P(P(B))$$

$$= \frac{1}{3} + \frac{1}{6} + \frac{1}{3}P$$

$$\therefore P(H) = \frac{1}{2} + \frac{1}{3}P$$

since $0 \leq P \leq 1$, so

$$\max P(H) = \frac{1}{2} + \frac{1}{3} = \frac{5}{6}$$

b) By the definition of conditional probability

$$P(F|H) = \frac{P(H|F) \cdot P(F)}{P(H)}$$

$$P(F|H) = \frac{\frac{1}{2} \cdot \frac{1}{3}}{\frac{1}{2} + \frac{1}{3}P} = \frac{\frac{1}{6}}{\frac{1}{2} + \frac{1}{3}P}$$

Since $0 \leq P \leq 1$, so $P(F|H)$ is maximized when $P = 0$

$$\max P(F|H) = 1/3$$

$P(F|H)$ is minimized when $P = 1$

$$\min P(F|H) = \frac{1/6}{5/6} = 1/5$$

3

Suppose we define the following events

E_1 : Person selected has tuberculosis

E : Person's chest X-Ray is

diagnosed as positive i.e. as

showing tuberculosis

Since E_1 and E_1^C forms a partition
of the sample space, so by
Bayes rule we have

$$P(\hat{E}_1 | \bar{E}) = \frac{P(E_1) P(E | \hat{E}_1)}{P(E | E_1) \cdot P(E) + P(E | E_1^c) \cdot P(E_1^c)}$$

$$= \frac{(0.001)(0.9)}{(0.001)(0.9) + (0.01)(0.999)}$$

$$= 0.023$$

$\therefore P(E_1 | \bar{E}) = 0.023$

Although the X-Ray test is fairly reliable we have found that only slightly more than 2% of those with positive X Rays turn out to have tuberculosis. The results of such calculations must be taken into account when large-scale medical diagnostic tests are planned.

Practice problem on linear algebra basics

a) Let λ_i^* be an eigenvalue of A with corresponding eigenvector v_i^* . Then,

$$Av_i^* = \lambda_i^* v_i^*$$

multiplying both sides on the left by v_i^{*T} . we get

$$v_i^{*T} Av_i^* = \lambda_i^* v_i^{*T} v_i^*$$

$$\Rightarrow v_i^{*T} Av_i^* = \lambda_i^* \|v_i^*\|_2^2$$

Since $b^T Ab > 0$ for all $b \in \mathbb{R}^n$, so

$$v_i^{*T} Av_i^* = \lambda_i^* \|v_i^*\|_2^2 > 0$$

Since $\|v_i^*\|_2^2 > 0$, so $\lambda_i^* > 0$.

\therefore All eigenvalues of A are positive.

⑥ Let λ_i^* be an eigenvalue of A
with corresponding eigenvector v_i^* . Then,

$$Av_i^* = \lambda_i^* v_i^*$$

Multiplying both sides to the left by
 A^T , we get

$$A^T A v_i^* = \lambda_i^* A^T v_i^*$$

since A is an orthogonal matrix, so

$$A^T A = A A^T = I$$

Hence,
 $v_i^* = \lambda_i^* A^T v_i^*$

Taking the 2-norm of both sides

$$\|v_i^*\|_2^2 = |\lambda_i^*|^2 \|A^T v_i^*\|_2^2$$

Now,
 $\|A^T v_i^*\|_2^2 = (A^T v_i^*)^T (A^T v_i^*)$
 $= v_i^{*T} A A^T v_i^* = \|v_i^*\|_2^2$

Hence, we have

$$\|v_i\|_2^2 = |\lambda_i|^2 \|v_i\|_2^2$$

$\therefore |\lambda_i| = 1$. Hence, all eigenvalues of
A have norm 1.

c) $A \in \mathbb{R}^{m \times n}$ has the following SVD

$$A = U \Sigma V^T$$

Now,

$$\begin{aligned} A A^T &= (U \Sigma V^T) (U \Sigma V^T)^T \\ &= U \Sigma V^T V \Sigma^T U^T \\ A A^T &= U \Sigma \Sigma^T U^T \end{aligned}$$

Now, $\Sigma = \begin{bmatrix} \sigma_1 & & & & & \\ & \sigma_2 & & & & \\ & & \ddots & & & \\ & & & \sigma_r & & \\ & & & & 0 & \dots \\ & & & & & 0 \end{bmatrix}$

so, $\Sigma \Sigma^T = \begin{bmatrix} \sigma_1^2 & & & & & \\ & \sigma_2^2 & & & & \\ & & \ddots & & & \\ & & & \sigma_r^2 & & \\ & & & & 0 & \dots \\ & & & & & 0 \end{bmatrix}$

since $\Sigma \Sigma^T$ is a diagonal matrix and U is an orthogonal matrix so the

eigen decomposition of the symmetric matrix AA^T is

$$u \Sigma \Sigma^T u^T$$

Hence,

$$\lambda_i(AA^T) = \sigma_i^2(A)$$

2

a) $(A^T)^T A^T$

$$= A A^T$$

$$= I \text{ (since } A \text{ is orthogonal)}$$

Also,

$$A^T (A^T)^T$$

$$= A^T A$$

$$= I \text{ (since } A \text{ is orthogonal)}$$

Hence, A^T is also orthogonal

$$\begin{aligned}
 b) \quad & (AB)^T(AB) \\
 &= B^T A^T A B \\
 &= B^T I B (\text{since } A \text{ is orthogonal}) \\
 &= B^T B \\
 &= I (\text{since } B \text{ is orthogonal})
 \end{aligned}$$

Also,

$$\begin{aligned}
 & AB(A^T B^T) \\
 &= AB B^T A^T \\
 &= A I A^T (\text{since } B \text{ is orthogonal}) \\
 &= A A^T \\
 &= I (\text{since } A \text{ is orthogonal})
 \end{aligned}$$

$\therefore AB$ is orthogonal

c) Let $A = B = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

Then A and B are both orthogonal.

but

$$A+B = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$$

is not orthogonal.

d) Suppose the column vectors of A are orthonormal. Hence

$$a_i^T a_i = 1 \quad \forall i$$

$$a_i^T a_j = 0 \quad i \neq j$$

which implies

$$A^T A = I$$

Since $A^{-1} = A^T$, so A is an orthogonal matrix. From (a), we know that A^T is also orthogonal. Since A^T is also orthogonal, so

$$(A^T)^T A^T = I$$

The above relation implies A^T has orthonormal columns, meaning that

A has orthonormal rows.