

# Exercise Sheet 5

1 From Theorem 4.4:

$$C(\ell) = \log(\sum_a(\ell)) + \boxed{\frac{\ell}{T} \cdot C_T}$$

i)  $\lim_{T \rightarrow \infty} C_T \rightarrow \infty$

ii)  $\lim_{T \rightarrow \infty} \frac{C_T}{T} \rightarrow 0$

iff i) and ii) hold,  $C(\ell)$  chooses the optimal/correct model.

AIC:  $C_T = 2K^2$

$$\lim_{T \rightarrow \infty} C_T = 2K^2 \not\rightarrow \infty \quad \neq$$

$$BIC: \zeta_T = \log(T) \cdot K^2$$

$$\lim_{T \rightarrow \infty} \zeta_T = \lim_{T \rightarrow \infty} \log(T) \cdot K^2 \rightarrow \infty \checkmark$$

$$\lim_{T \rightarrow \infty} \frac{\zeta_T}{T} = \lim_{T \rightarrow \infty} \frac{\log(T)}{T} \cdot K^2 \rightarrow 0 \checkmark$$

$\Rightarrow$  consistent

$$HQ: \zeta_T = 2 \cdot \log(\log(T)) K^2$$

$$\lim_{T \rightarrow \infty} \zeta_T \rightarrow \infty \checkmark$$

$$\lim_{T \rightarrow \infty} \frac{\zeta_T}{T} \rightarrow 0 \checkmark$$

Note:  $p \cdot K^2$  is the number of freely estimated parameters. Restrictions can change that number even

holding the lag order constant.  
(except: intercepts)

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a)  $\rightarrow R$  + Exercise Sheet 9,  
Task 2

b) Yes. The test rejects even for large 'm', so there must be a dynamic pattern for the VAR to model.

c) # coefficients  
 $= \sum_{\phi_i: \forall i \in \{1, \dots, p\}} K^2 \cdot p + \sum_{\text{intercept}} K$

# data points (observations times dimension)  
 $= T \cdot K$

$$\Rightarrow K \cdot T \geq K^2 \cdot \rho + K$$

$$\Rightarrow \frac{K \cdot (T-1)}{K^2} \geq \rho$$

d) - AIC and HQ are flat  
around the minima  $\rightarrow$  no  
distinct optimum visible.

- Conceivable reasons: Persistence, omitted variables, wrong functional form.  $\rightarrow$  The VAR may just work as an approximation.
- BIC is the most conservative IC.

$$\text{Minima at } \rho = \begin{cases} 1 & \text{BIC} \\ 2 & \text{HQ} \\ 4 & \text{AIC} \end{cases}$$

e) Adjust with  $K^2 \cdot p$  degrees of freedom! Adjustment for the intercept is not necessary, since  $z_t$  needs to be demeaned anyways ( $P_0 = (z_t - Mxz_t \cdot \mathbf{1})'$ )!

The Ljung-Box test has  $m \cdot K^2$  degrees of freedom ( $K^2$  per lagged  $CM$ ), so after adjustments, we have  $(m-p)K^2$  degrees of freedom. Since the test rejects everywhere else, the VARs do not explain the dynamics entirely.

f) i)  $\rightarrow R$  ( $K^2 \cdot p + K$ )  
ii) = Residuals' (co-)variances are higher for the VAR(1)  
 $\Rightarrow$  VAR(4) predicts better (in-sample)

iii) No, it is exactly the other way around. Estimating more coefficients with the same information leads to less information per coefficient.

g) Again:  $K^2 \cdot p + K$

$$+h)i) = 2^2 \cdot \begin{Bmatrix} 1 \\ 4 \end{Bmatrix} + 2 = \begin{Bmatrix} 6 \\ 18 \end{Bmatrix}$$

h) ii+iii) Same pattern as in f), but not that pronounced this time.

$\frac{\text{\# coefficients}}{\text{\# data points}}$  is now larger and does increase less sharply if  $K=2$  compared to  $K=5$  if  $p \uparrow$ . Furthermore, there are less coefficients to model the same magnitude of correlation.



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a) There seems to be few correlation across time except lags 20 to 24. Hence there is not much for a VAR to exploit.

b) O. The criteria support the Ljung-Box test.

c) No important difference. There is almost no variation taken away by the VAR(1).

d) No. (c) has shown  
that nothing has changed  
at all.

e) — Plenty of lagged cross-  
auto correlations.

— Criteria suggest high  
lag orders

$\Rightarrow$  Fitting a VAR appears  
sensible.



f) Finding: e) implies conditional heteroscedasticity.

$\Rightarrow$  Can  $z_t$  still be weakly stationary?

$$\Rightarrow a_t = \sigma_t \cdot \varepsilon_t, \quad \varepsilon_t \stackrel{\text{iid}}{\sim} (0, \sigma_\varepsilon^2)$$

and  $\sigma_t$  dependent on  $\varphi(L)\sigma_t$ .

$\Rightarrow E(\sigma_t^2 | \sigma_{t-1}^2, \sigma_{t+1}^2, \dots)$  does not necessarily equal  $E(\sigma_t^2)$  !

BUT: For stationarity, only

the unconditional moments matter!

If  $\varphi(L)$  is a stable process  
and  $\varepsilon_t \stackrel{iid}{\sim} (0, \sigma_\varepsilon^2)$  (hence no  
unconditional heteroscedasticity), the  
time series  $z_t$  is still weakly  
stationary.

(analogous to a VAR(p) process

when  $E(z_t | z_{t-1}, z_{t-2}, \dots, z_1)$

does not necessarily equal  $\mu_z = E(z_t)$ !

Still, we have proved those models to

be stationary under some  
conditions )

