

University of Duisburg-Essen  
Faculty of Business Administration and  
Economics  
Chair of Econometrics



# P-Approximation

Seminar in Econometrics

Term Paper

Submitted to the Faculty of  
Business Administration and Economics  
at the  
University of Duisburg-Essen

from:

Jens Klenke and Janine Langerbein

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Reviewer: Christoph Hanck

Deadline: Jan. 17th 2020

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Name:	Jens Klenke	Janine Langerbein
Matriculation Number:	3071594	307
E-Mail:	jens.klenke@stud.uni-due.de	janine.langerbein@stud.uni-due.de
Study Path:	M.Sc. Economics	M.Sc. Economics
Semester:	5 <sup>th</sup>	5 <sup>th</sup>
Graduation (est.):	Winter Term 2020	Winter Term 2020

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# 1 Introduction

Meta tests have been shown to be a powerful tool when testing for the null of non-cointegration. The distribution of their test statistic, however, is mostly not available in closed form. This might pose difficulties when implementing the meta tests in econometric software packages, as one has to include tables of critical values and p-values for each combination of the underlying tests. Software package size limitations are therefore quickly exceeded.

In this paper we propose supervised Machine Learning Algorithms to approximate the p-values of the meta test by Bayer and Hanck (2012) which tests for the null of non-cointegration. This approach might reduce the size of associated software packages considerably. The algorithms are trained on simulated data for various specifications of the aforementioned test.

Ergebnis der Models (1-2 Sätze)

Inhalt Paper

## 2 Bayer Hanck Test

The choice as to which of the available cointegration tests to use is a recurrent issue in econometric time series analysis. Bayer and Hanck (2012) propose powerful meta tests which provide unambiguous test decisions. They combine several residual- and system-based tests in the manner of Fisher's (1932) Chi-squared test.

Bayer and Hanck build their work on results from Pesavento (2004), who defines the underlying model as  $z'_t = [x'_t, y_t]$ .  $x_t$ , an  $n_1 \times 1$  vector, describes the regressor dynamics, while  $y_t$  is a scalar which defines the cointegrating relation. They can be written as

$$\Delta x_t = \tau_1 + v_{1t}, \tag{2.1}$$

$$y_t = (\mu_2 - \gamma' \mu_1) + (\tau_2 - \gamma' \tau_1)t + \gamma' x_t + u_t, \tag{2.2}$$

$$u_t = \rho u_{t-1} + v_{2t}. \tag{2.3}$$

$\mu_1$ ,  $\mu_2$   $\tau_1$  and  $\tau_2$  are the deterministic parts of the model. They are subject to the following restrictions: (i)  $\mu_2 - \gamma' \mu_1$  and  $\tau = 0$  which translates to no deterministics, (ii)  $\tau = 0$  which corresponds to a constant in the cointegrating vector, (iii)  $\tau_2 - \gamma' \tau_1 = 0$ , a constant plus trend.

$v_t = [v'_{1t} v_{2t}]'$  with  $\Omega$  the long-run covariance matrix of  $v_t$ . For derivation of  $v_t$  see Pesavento (2004). Pesavento shows that  $\{v_t\}$  satisfies an FCLT, i.e.  $T^{-1/2} \sum_{t=1}^{[T]} v_t \Rightarrow \Omega^{1/2} W(\cdot)$ . It is further assumed that the  $x_t$  are not cointegrated.

It clearly follows from (2.3) that  $z_t$  is cointegrated if  $\rho < 1$ . Hence the null hypothesis of no cointegration is  $H_0 : \rho = 1$ .

Furthermore, Pesavento introduces two other parameters. First,  $R^2$  measures the squared correlation of  $v_{1t}$  and  $v_{2t}$ . It can be interpreted as the influence of the right-hand side variables in (2.2). It ranks between zero and one. When there is no long-run correlation between those variables and the errors from the cointegration regression,  $R^2$  equals zero. Secondly, the number of lags is approximated by a finite number  $k$ .

### Assumptions (BH S. 84)?

Bayer and Hanck's (2012) meta test combines the test statistics of four stand-alone tests. Namely, these are the tests of Engle and Granger (1987), Johansen (1988), Boswijk (1994) and Banerjee, Dolado, and Mestre (1998).

### Engle-Granger

### Johansen

### Banerjee and Boswijk

To combine the results from the underlying tests Bayer and Hanck draw upon Fisher's combined probability test (Fisher, 1932). It merges the tests using the formula

$$\tilde{\chi}_{\mathcal{I}}^2 := -2 \sum_{i \in \mathcal{I}} \ln(p_i). \quad (2.4)$$

Let  $t_i$  be the  $i^{th}$  test statistic. If test  $i$  rejects for large values, take  $\xi_i := t_i$ . If test  $i$  rejects for small values, take  $-\xi_i := t_i$ . With  $\Xi_i(x) := \Pr_{\mathcal{H}_0}(\xi_i \geq x)$  the p-value of the  $i^{th}$  test is  $p_i := \Xi_i(\xi_i)$ .

Fisher (1932) shows that under the assumption of independence the null distribution of  $\tilde{\chi}_{\mathcal{I}}^2$  follows a chi-squared distribution with  $2\mathcal{I}$  degrees of freedom. If this assumption is violated the null distribution is less evident. Here, the latter case occurs, as the  $\xi_i$  are not independent. The  $\tilde{\chi}_{\mathcal{I}}^2$ , however, have well-defined asymptotic null distributions  $F_{\mathcal{F}_{\mathcal{I}}}$ , as  $\tilde{\chi}_{\mathcal{I}}^2 \rightarrow_d \mathcal{F}_{\mathcal{I}}$  under  $\mathcal{H}_0$  if  $T \rightarrow \infty$ , with  $\mathcal{F}_{\mathcal{I}}$  some random variable. It is therefore feasible to simulate the joint null distribution of the  $\xi_i$  to obtain the distribution  $F_{\mathcal{F}_{\mathcal{I}}}$  of 2.4. The

$F_{\mathcal{F}_I}$  depend on which and how many tests are combined. The distributions of the  $\xi_i$  depend on  $K - 1$  and the deterministic case.

### **3   Simulation**

### **4   Models**

### **5   Package**

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# A Appendices

Better sorting of the Appendix

### **Eidesstattliche Versicherung**

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Essen, den \_\_\_\_\_

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Jens Klenke and Janine Langerbein