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1 Introduction

Meta tests have been shown to be a powerful tool when testing for the null of non-cointegration. The distribution of their test statistic, however, is mostly not available in closed form. This might pose difficulties when implementing the meta tests in econometric software packages, as one has to include tables of critical values and p-values for each combination of the underlying tests. Software package size limitations are therefore quickly exceeded.

In this paper we propose supervised Machine Learning Algorithms to approximate the p-values of the meta test by Bayer and Hanck (2012) which tests for the null of non-cointegration. This approach might reduce the size of associated software packages considerably. The algorithms are trained on simulated data for various specifications of the aforementioned test.

[Ergebnis der Models \(1-2 Sätze\)](#)

[Inhalt Paper](#)

2 Bayer Hanck Test

The choice as to which of the available cointegration tests to use is a recurrent issue in econometric time series analysis. Bayer and Hanck (2012) propose powerful meta tests which provide unambiguous test decisions. They combine several residual- and system-based tests in the manner of Fisher's (1932) Chi-squared test.

Bayer and Hanck build their work on results from Pesavento (2004), who defines the underlying model as $z'_t = [x'_t, y_t]$. x_t , an $n_1 \times 1$ vector, describes the regressor dynamics, while y_t is a scalar which defines the cointegrating relation. They can be written as

$$\Delta x_t = \tau_1 + v_{1t}, \tag{2.1}$$

$$y_t = (\mu_2 - \gamma' \mu_1) + (\tau_2 - \gamma' \tau_1)t + \gamma' x_t + u_t, \tag{2.2}$$

$$u_t = \rho u_{t-1} + v_{2t}. \tag{2.3}$$

μ_1, μ_2, τ_1 and τ_2 are the deterministic parts of the model. They are subject to the following restrictions: (i) $\mu_2 - \gamma' \mu_1$ and $\tau = 0$ which translates to no deterministic, (ii) $\tau = 0$ which corresponds to a constant in the cointegrating vector, (iii) $\tau_2 - \gamma' \tau_1 = 0$, a constant plus trend.

$v_t = [v'_{1t} v_{2t}]'$ with Ω the long-run covariance matrix of v_t . For derivation of v_t see Pesavento (2004). Pesavento shows that $\{v_t\}$ satisfies an FCLT, i.e. $T^{-1/2} \sum_{t=1}^{[T\cdot]} v_t \Rightarrow \Omega^{1/2} W(\cdot)$. It is further assumed that the x_t are not cointegrated.

It clearly follows from (2.3) that z_t is cointegrated if $\rho < 1$. Hence the null hypothesis of no cointegration is $H_0 : \rho = 1$.

Furthermore, Pesavento introduces two other parameters. First, R^2 measures the squared correlation of v_{1t} and v_{2t} . It can be interpreted as the influence of the right-hand side variables in (2.2). It ranks between zero and one. When there is no long-run correlation between those variables and the errors from the cointegration regression, R^2 equals zero. Secondly, the number of lags is approximated by a finite number k .

Assumptions (BH S. 84)?

Bayer and Hanck's (2012) meta test combines the test statistics of four stand-alone tests. Namely, these are the tests of Engle and Granger (1987), Johansen (1988), Boswijk (1994) and Banerjee, Dolado, and Mestre (1998).

Engle-Granger

Johansen

Banerjee and Boswijk

"the test statistics are shown to converge under a local alternative to random variables whose distributions are functions of Brownian Motions and Ornstein–Uhlenbeck processes and of a single nuisance parameter"

To combine the results from the underlying tests Bayer and Hanck draw upon Fisher's combined probability test (Fisher, 1932). It merges the tests using the formula

$$\tilde{\chi}_{\mathcal{I}}^2 := -2 \sum_{i \in \mathcal{I}} \ln(p_i). \quad (2.4)$$

Let t_i be the i^{th} test statistic. If test i rejects for large values, take $\xi_i := t_i$. If test i rejects for small values, take $-\xi_i := t_i$. With $\Xi_i(x) := \Pr_{\mathcal{H}_i}(\xi_i \geq x)$ the p-value of the i^{th} test is $p_i := \Xi_i(\xi_i)$.

Fisher (1932) shows that under the assumption of independence the null distribution of $\tilde{\chi}_{\mathcal{I}}^2$ follows a chi-squared distribution with $2\mathcal{I}$ degrees of freedom. If this assumption is violated the null distribution is less evident. Here, the latter case occurs, as the ξ_i are not independent. The $\tilde{\chi}_{\mathcal{I}}^2$, however,

have well-defined asymptotic null distributions $F_{\mathcal{F}_{\mathcal{I}}}$, as $\tilde{\chi}_{\mathcal{I}}^2 \rightarrow_d \mathcal{F}_{\mathcal{I}}$ under \mathcal{H} , if $T \rightarrow \infty$, with $\mathcal{F}_{\mathcal{I}}$ some random variable. It is therefore feasible to simulate the joint null distribution of the ξ_i to obtain the distribution $F_{\mathcal{F}_{\mathcal{I}}}$ of 2.4. The $F_{\mathcal{F}_{\mathcal{I}}}$ depend on which and how many tests are combined. The distributions of the ξ_i depend on $K - 1$ and the deterministic case.

3 Simulation

4 Models

5 Package

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A Appendices

Better sorting of the Appendix

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Essen, den _____

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