

# Forecasting electricity demand – ensemble methods

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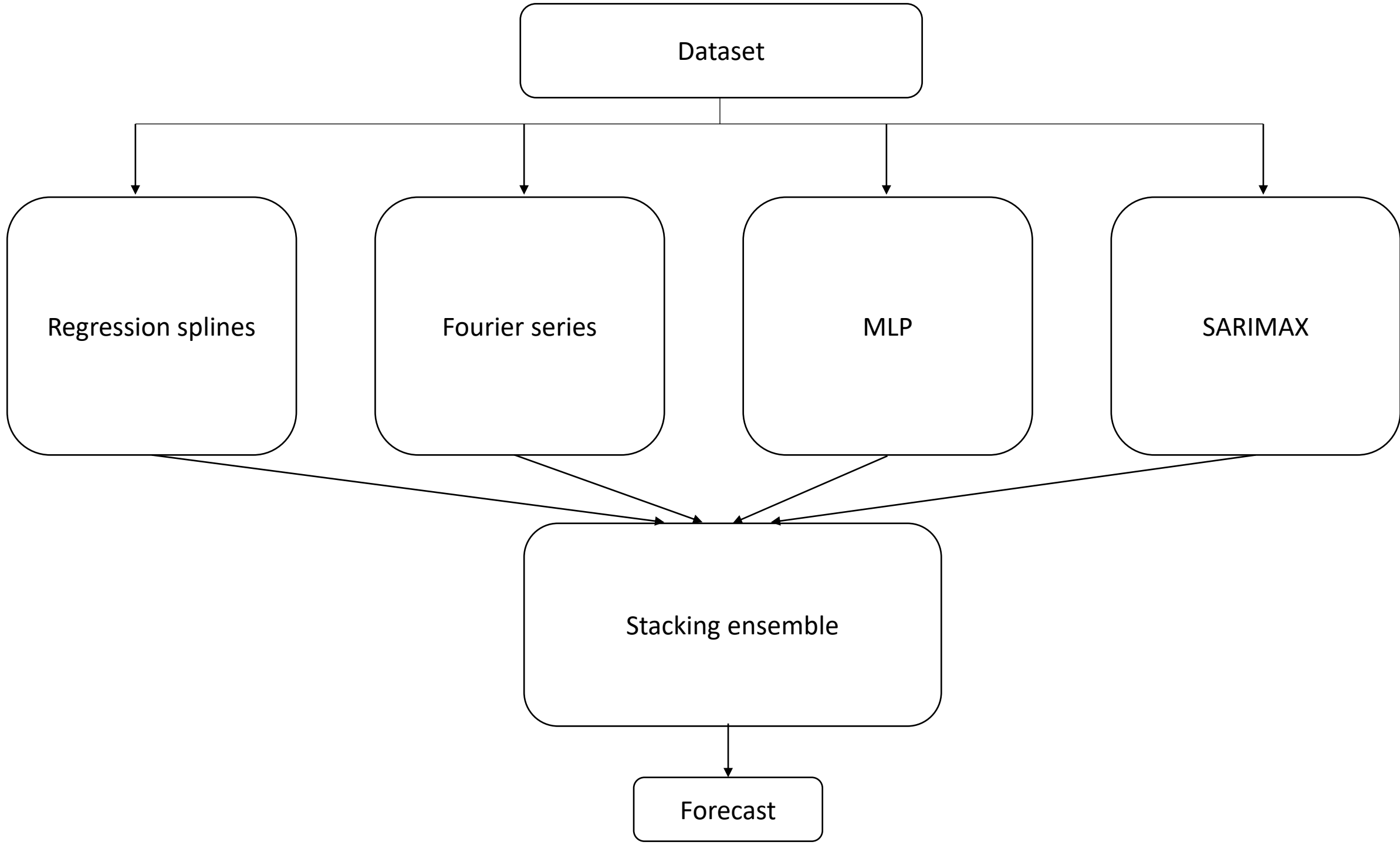
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# Ensemble methods - stacking

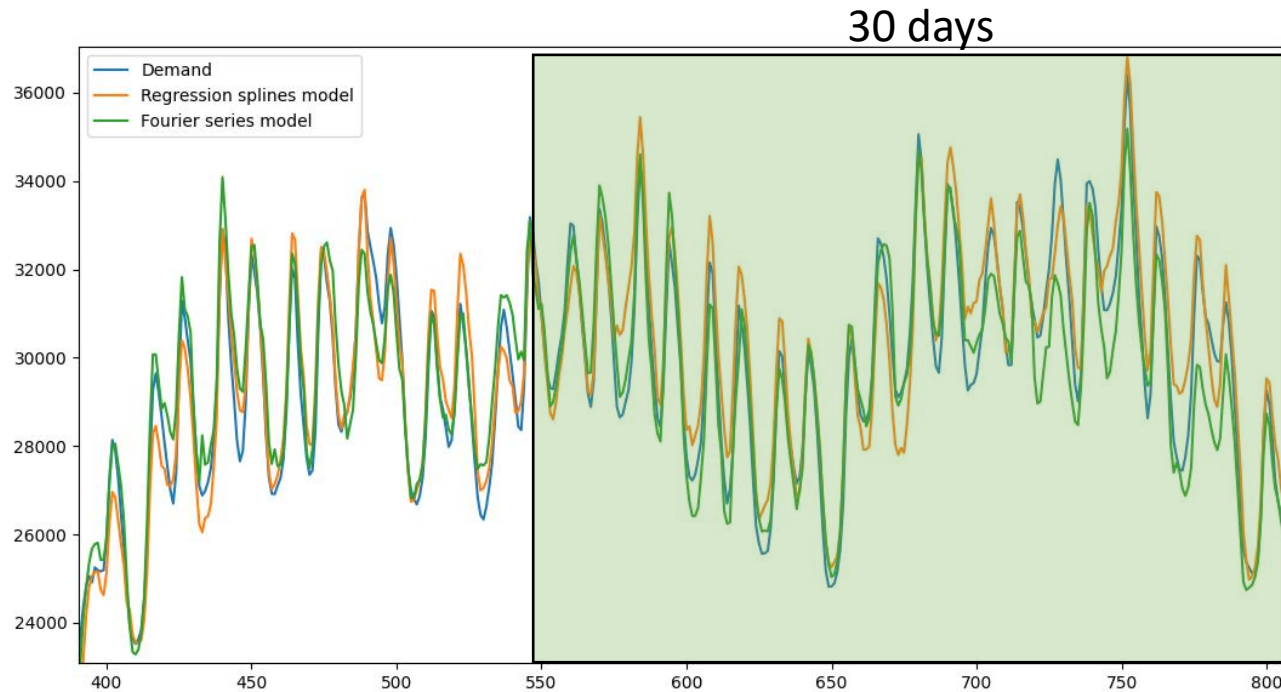
- Create many base models, average their predictions.

$$\left. \begin{array}{l} \bullet \hat{y}_t^1 \\ \bullet \hat{y}_t^2 \\ \bullet y_t \\ \bullet \hat{y}_t^3 \\ \bullet \hat{y}_t^4 \end{array} \right\} \hat{y}_t = \sum_{i=1}^4 \alpha_i \hat{y}_t^i$$

- Improve accuracy, reduce errors.
- How to determine weights for each forecast?



# Rolling window linear regression



$$\hat{y}_t = c + \sum_{i=1}^4 \alpha_i \hat{y}_t^i$$

- With  $c$  and  $\alpha_i$  determined by least squares regression

# Least squares regression

Linear regression model:  $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots \beta_k x_k + \epsilon$

Matrix form:  $Y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}, X = \begin{pmatrix} 1 & x_{11} & \cdots & x_{k1} \\ 1 & x_{12} & \cdots & x_{k2} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{1n} & \cdots & x_{kn} \end{pmatrix}, \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}$

Model:  $y = X\beta + \epsilon$

Solution:  $\hat{\beta}$ , least squares estimator of  $\beta$  given by:

$$\hat{\beta} = (X^T X)^{-1} X^T Y \text{ provided } X^T X \text{ is invertible.}$$

# Multicollinearity

Problems arise if predictors are correlated among themselves.

1. If predictors are correlated,  $X$  may not be of full rank and thus  $X^t X$  may not be invertible.
2. Inflation of variance of  $\hat{\beta}$ .
3. The least squares estimator  $\hat{\beta}$  may be far from true value of  $\beta$ .

# Multicollinearity

For simplicity, take a simple linear regression model with two predictors:

$$y = \beta_1 x_1 + \beta_2 x_2 + \epsilon$$

With  $x_1, x_2, y$  standardized.

Then,  $X^T X = n \begin{pmatrix} 1 & r \\ r & 1 \end{pmatrix}$  with  $r$  the correlation coefficient between predictors  $x_1$  and  $x_2$ .

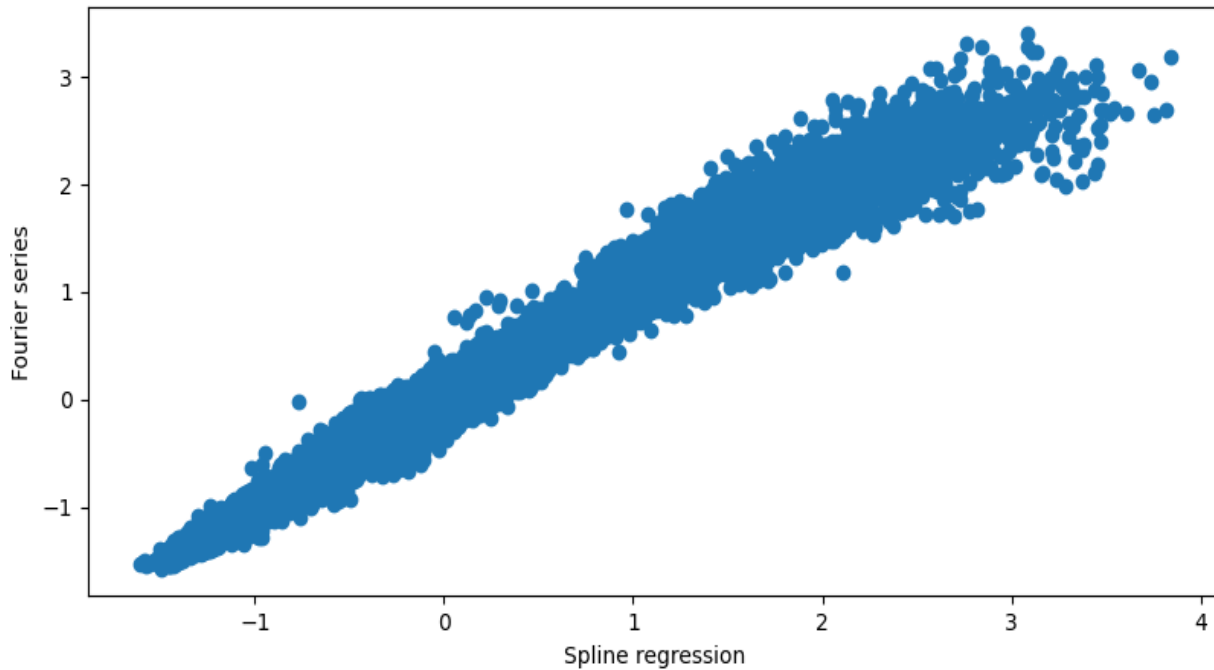
# Multicollinearity

1. We have  $|X^T X| = \begin{vmatrix} 1 & r \\ r & 1 \end{vmatrix} = 1 - r^2$ . For  $r \rightarrow 1$ ,  $|X^T X| \rightarrow 0$
2. We can show that  $\text{var}(\hat{\beta}_i) = \frac{\sigma^2}{n(1-r^2)}$ . For  $r \rightarrow 1$ ,  $\text{var}(\hat{\beta}_i) \rightarrow \infty$
3. We can show that  $E(\|\hat{\beta} - \beta\|^2) = \frac{2\sigma^2}{1-r^2}$   
As  $r \rightarrow 1$ ,  $\hat{\beta}$  is on average far from  $\beta$ .



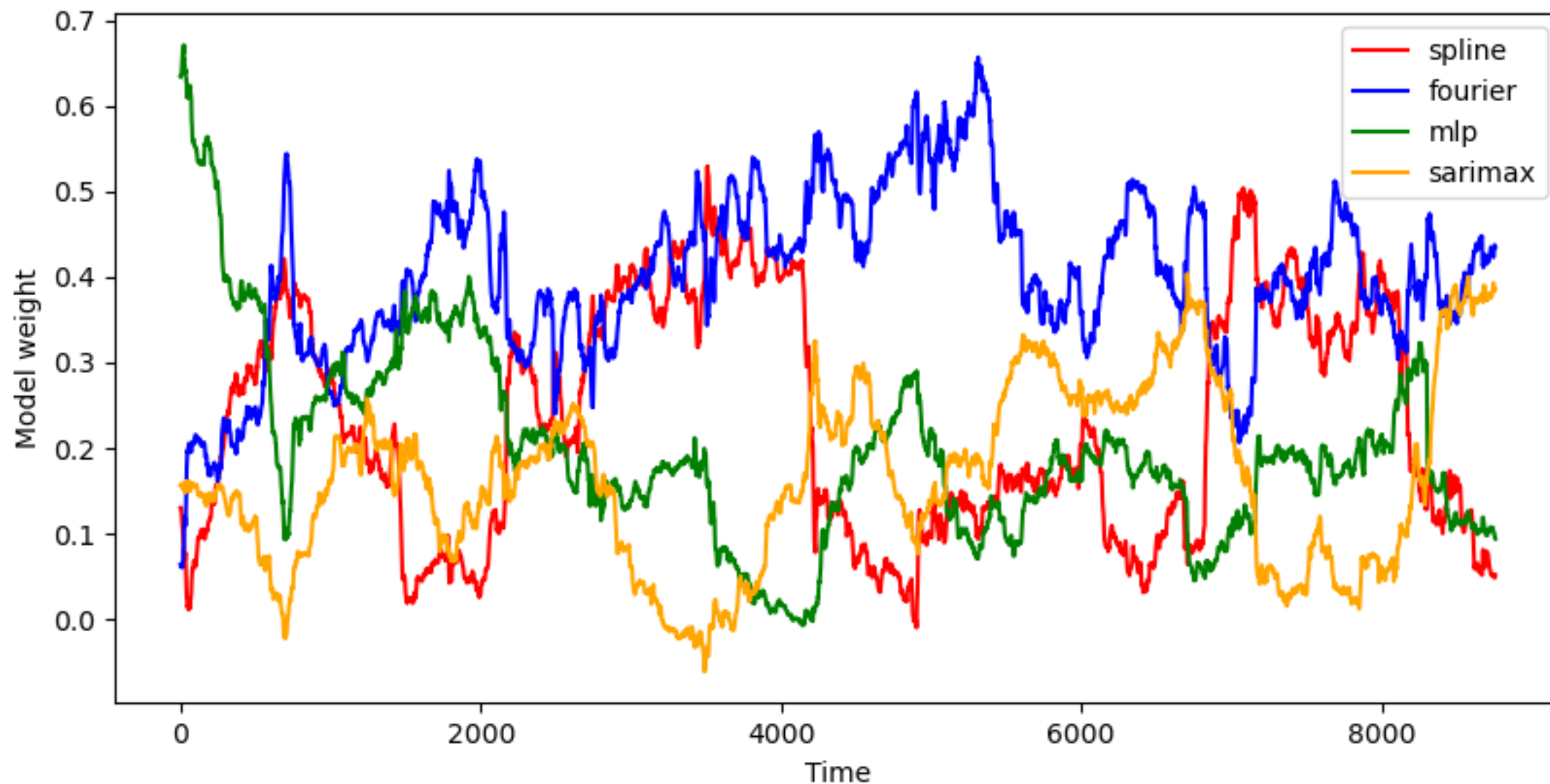
# Multicollinearity

Spline regression model forecast vs Fourier series model forecast



- For a time window of 30 days, correlation coefficient between any two models varies between 0.90 and 0.99.

# Model weights from OLS ensemble



# Regularization

Consider the regularized least squares problem:

$$\min_{\beta} \frac{1}{2} \|Y - X\beta\|_2^2 + \frac{\lambda}{2} \|\beta\|_2^2 = \min_{\beta} \left\| \begin{pmatrix} X \\ \sqrt{\lambda} I \end{pmatrix} \beta - \begin{pmatrix} Y \\ 0 \end{pmatrix} \right\|_2^2$$

Solve like OLS but for  $\lambda > 0$ ,  $\begin{pmatrix} X \\ \sqrt{\lambda} I \end{pmatrix}$  has full rank,

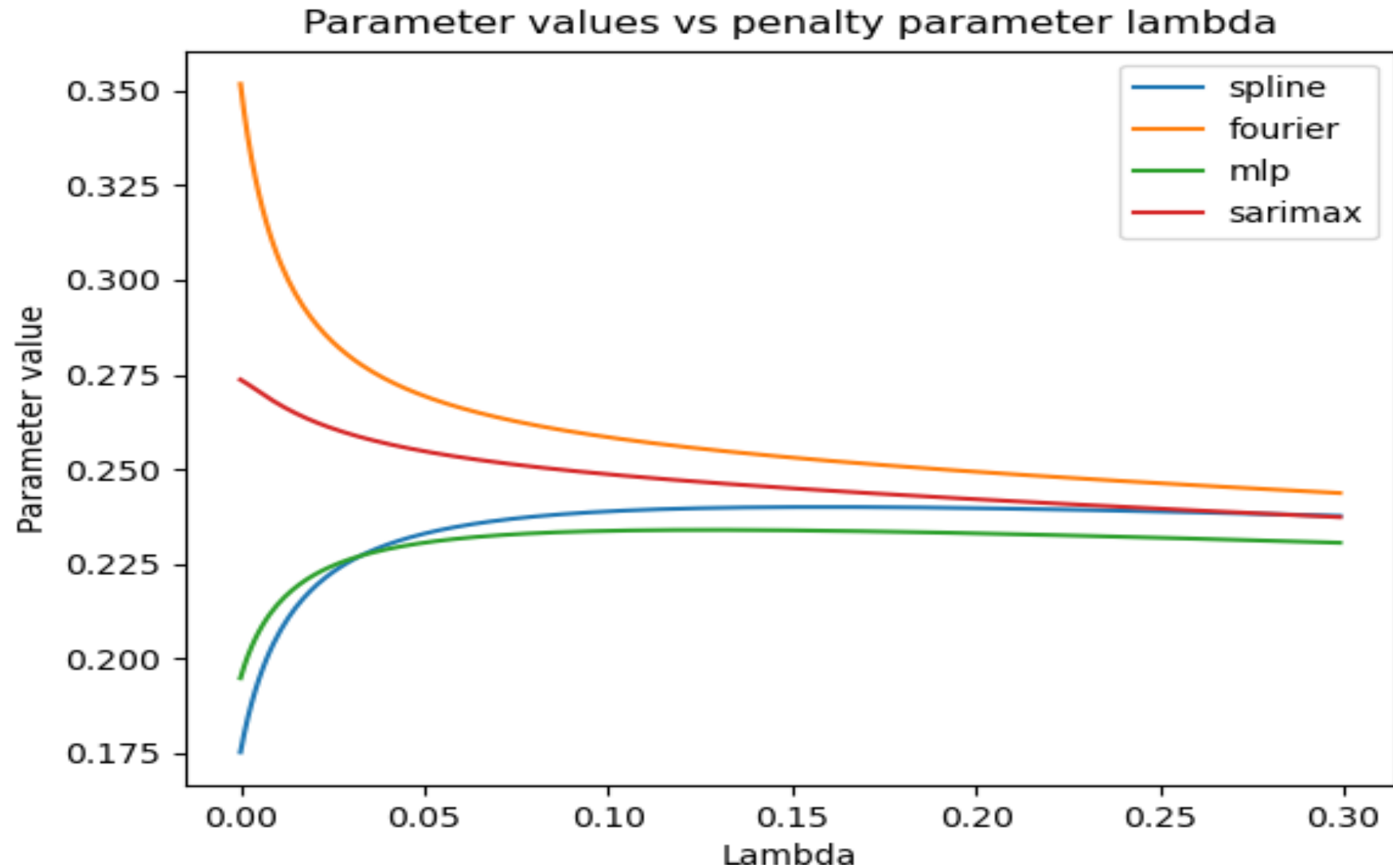
Thus, from normal equations,  $\hat{\beta}_r = (X^T X + \lambda I)^{-1} X^T Y$

# Regularization

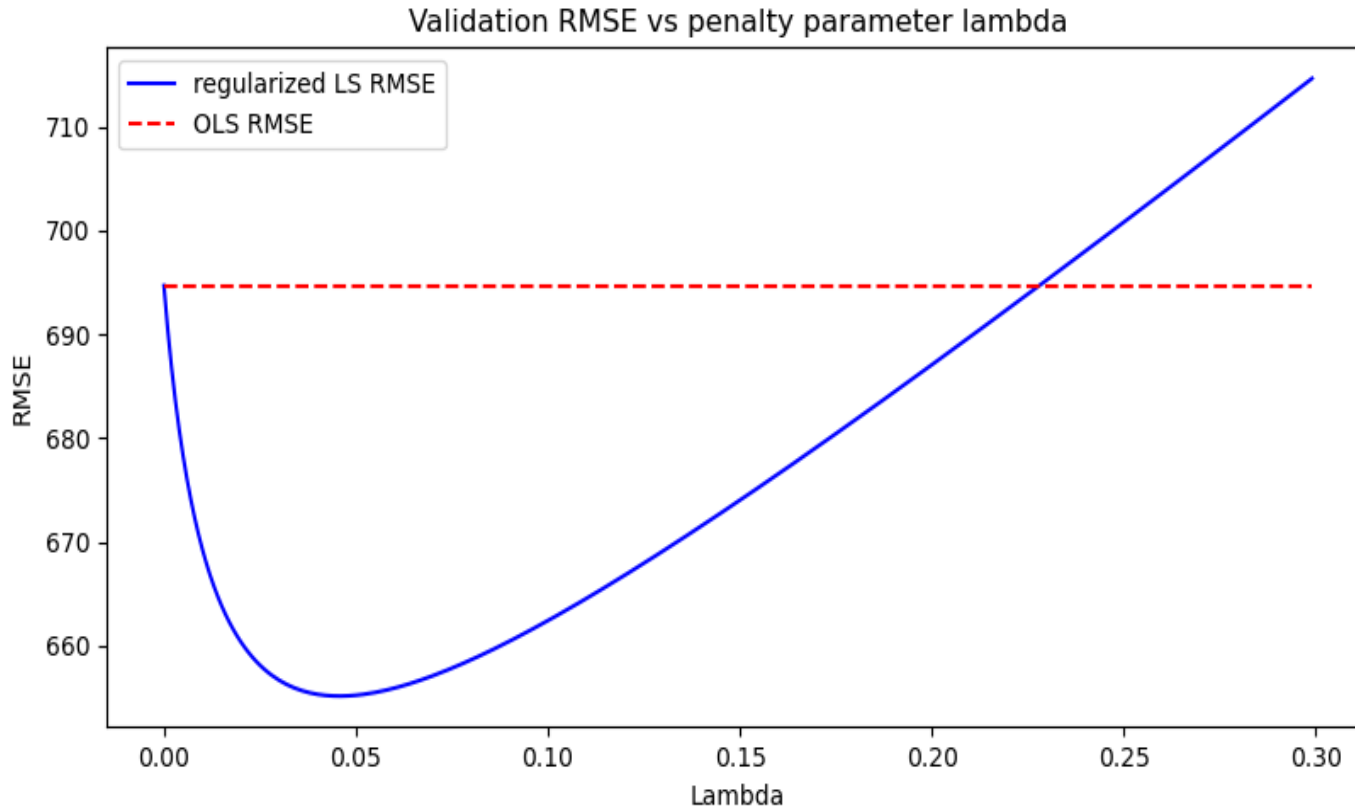
Thus, from normal equations,  $\hat{\beta}_r = (X^T X + \lambda I)^{-1} X^T Y$

1. Numerically easier to solve for  $\hat{\beta}_r$  since  $(X^T X + \lambda I)$  is always invertible.
2. We are essentially minimizing  $RSS + \lambda \|\beta\|_2^2$ , by choosing an appropriate  $\lambda$ , we get a tradeoff between magnitude of the parameters and fit.

# Regularization



# Regularization – Determining $\lambda$

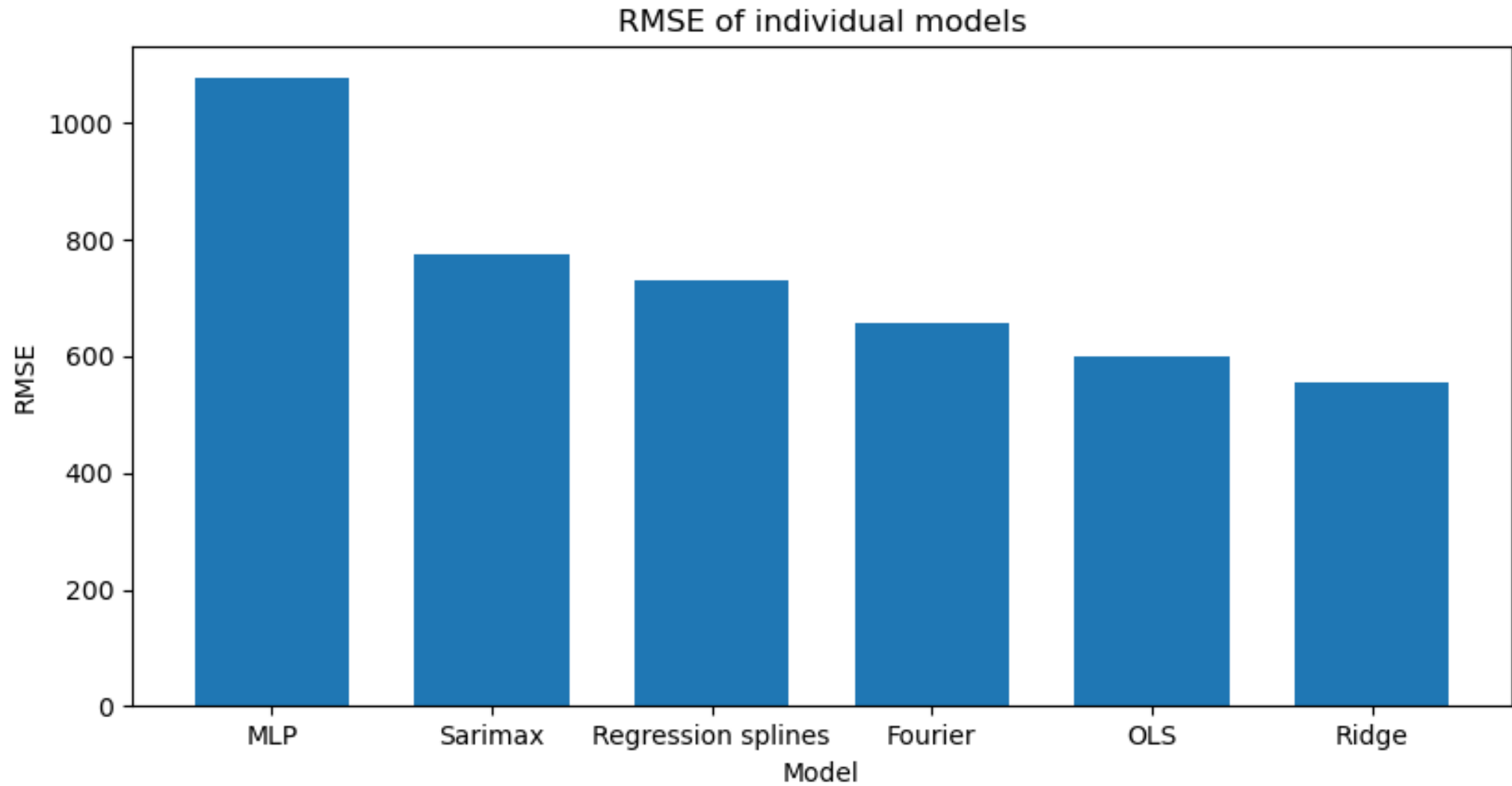


- We do not know a priori the optimal  $\lambda$ .
- Take mean optimal lambda from training set.

# Results – Regularized ensemble method

- Best model – Fourier series
  - RMSE: 656 MW
  - MAPE: 2.24 %
- OLS ensemble method
  - RMSE: 601 MW
  - MAPE: 1.95 %
- Ridge ensemble method
  - RMSE: 553 MW
  - MAPE: 1.82 %

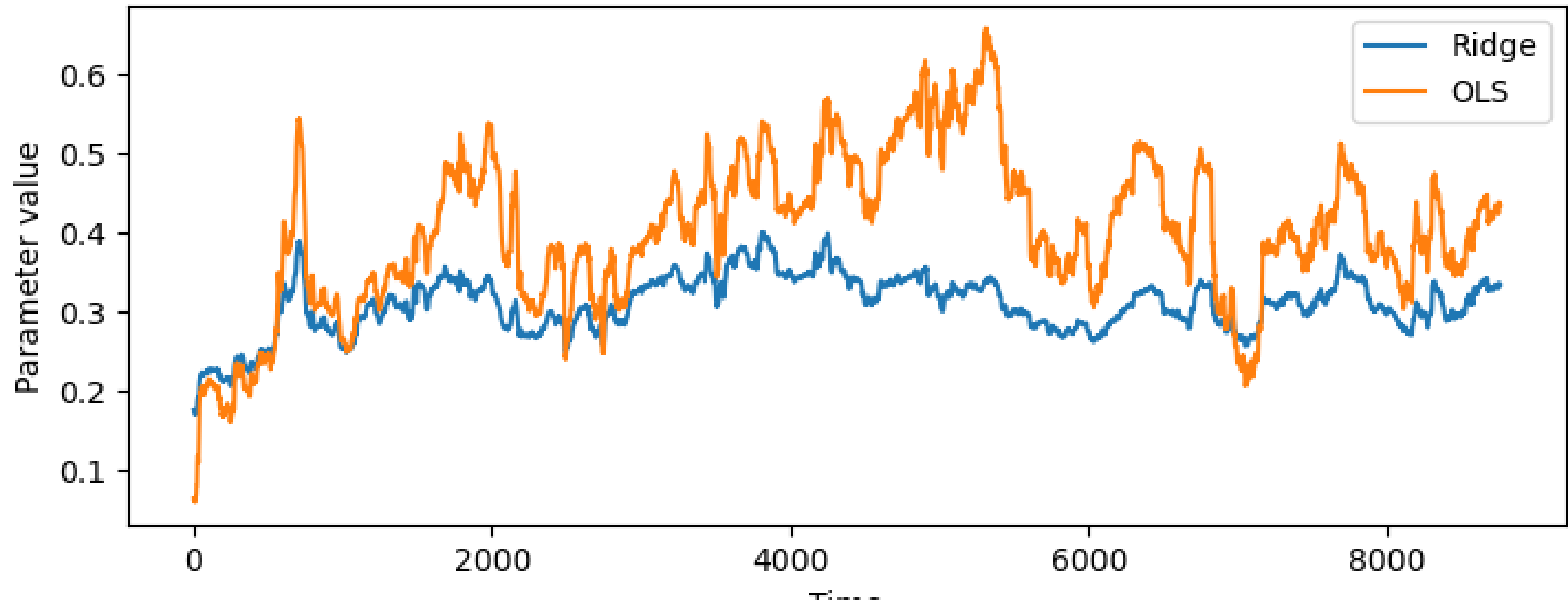
# Results



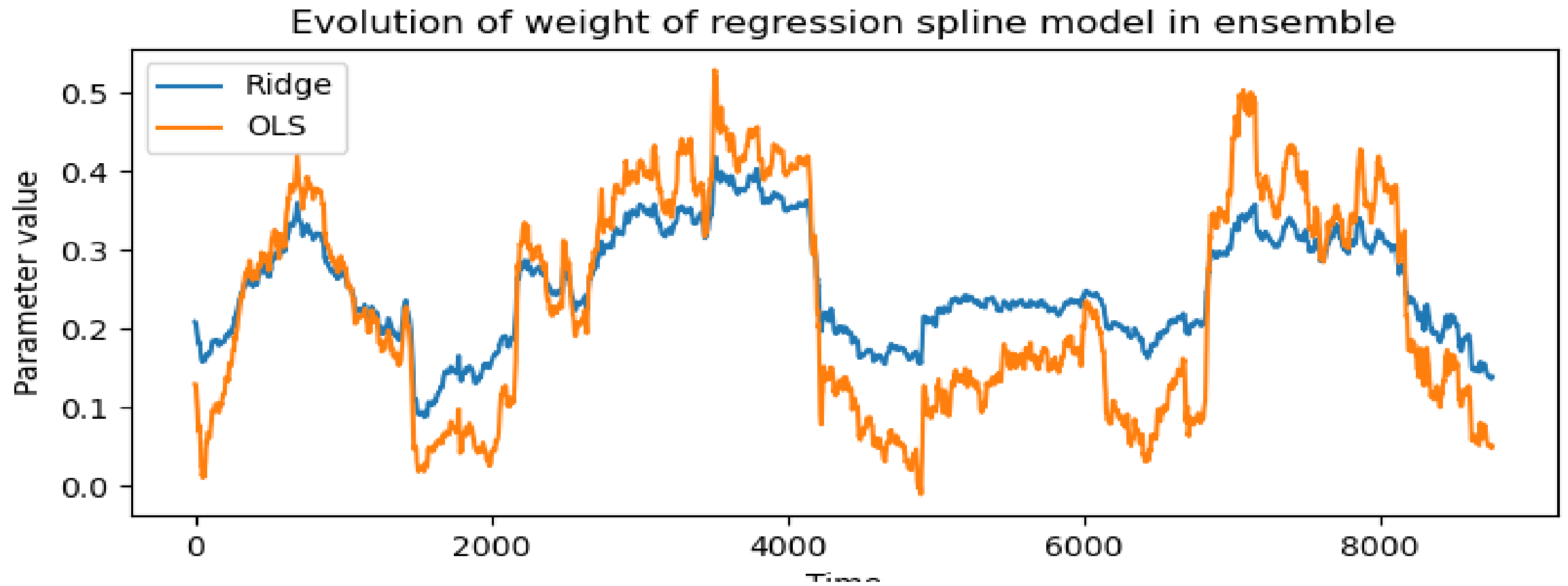


# Results

Evolution of weight of Fourier model in ensemble



# Results



# Errors

