Properties of Expectations

Covariance, Variance of Sums [Ross S7.4]

Proposition 31.1 If X and Y are independent, then for any functions g(x) and h(y):

- $i) \quad E[g(X)h(Y)] = E[g(X)]E[h(Y)]$
- ii) g(X) and h(Y) are independent.

Why?

$$i) \quad E[g(X)h(Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f_{XY}(x,y)dxdy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f_X(x)f_Y(y)dxdy$$

$$= \int_{-\infty}^{\infty} g(x)f_X(x)dx \int_{-\infty}^{\infty} h(y)f_Y(y)dy$$

$$= E[h(Y)]E[g(X)]$$

ii) Let $A = \{x \mid g(x) \le a\}$ and $B = \{y \mid h(y) \le b\}$. Then:

$$\begin{split} P[g(X) &\leq a, h(Y) \leq b] \\ &= P[X \in A, Y \in B] \\ &= P[X \in A] \ P[Y \in B] \\ &= P[g(X) \leq a] \ P[h(Y) \leq b] \end{split}$$
 since X and Y are independent

For a single random variable X, its mean and variance give us some information about X.

For two random variables X and Y, **covariance** (and **correlation**) will give us information about the relationship between the pair X and Y.

Definition 31.1: The **covariance** between X and Y, denoted Cov[X,Y], is defined as

$$Cov[X, Y] = E[(X - E[X])(Y - E[Y])]$$

Just as $Var[X] = E[X^2] - (E[X])^2$, we also have:

$$\begin{split} Cov[X,Y] &= E\left[\; (X-E[X])(Y-E[Y]) \; \right] \\ &= E\left[\; XY-E[X]Y-E[Y]X+E[X]E[Y]) \; \right] \\ &= E[XY]+E[\; -E[X]Y\;]+E[\; -E[Y]X\;]+E[\; E[X]E[Y]\;] \\ &= E[XY]-E[X]E[Y]-E[Y]E[X]+E[X]E[Y] \\ &= E[XY]-E[X]E[Y] \end{split}$$

Note: If X and Y are independent, then E[XY] = E[X]E[Y] so Cov[X,Y] = 0.

Example 31.1: Does Cov[X, Y] = 0 imply X and Y are independent?

Solution: No!

Let
$$P[X = 0] = P[X = 1] = P[X = -1] = 1/3$$
.

Let
$$Y = \begin{cases} 0 & X \neq 0 \\ 1 & X = 0 \end{cases}$$

a) X and Y are not independent since

$$\frac{1}{3} = P[X = 0, Y = 0] \neq P[X = 0]P[Y = 0] = \frac{1}{3}\frac{2}{3}$$

b) Since XY = 0

$$Cov[X,Y] = \underbrace{E[XY]}_{=0} - \underbrace{E[X]}_{=0} E[Y] = 0$$

Proposition 31.2

- i) Cov[X, Y] = Cov[Y, X]
- ii) Cov[X, X] = Var[X]

$$iii)$$
 $Cov[aX, Y] = aCov[X, Y] = Cov[X, aY]$

$$\mathit{iv)} \ \mathit{Cov} \left[\sum_{i=1}^n X_i, \ \sum_{j=1}^m Y_j \right] = \sum_{i=1}^n \sum_{j=1}^m \mathit{Cov}[X_i, Y_j]$$

Why? i), ii) and iii) follow from

$$Cov[X,Y] = E[XY] - E[X]E[Y]$$

For iv), let
$$U=\sum_{i=1}^n X_i \qquad V=\sum_{j=1}^m Y_i$$

$$E[X_i]=\mu_i \qquad E[Y_j]=\nu_j$$
 Then:
$$E\left[U\right]=\sum_{i=1}^n \mu_i \qquad E\left[V\right]=\sum_{j=1}^m \nu_j$$

So,
$$Cov[U, V] = E\left[\left(\sum_{i=1}^{n} X_i - \sum_{i=1}^{n} \mu_i\right) \left(\sum_{j=1}^{m} Y_j - \sum_{j=1}^{m} \nu_j\right)\right]$$

$$= E \left[\sum_{i=1}^{n} (X_i - \mu_i) \sum_{j=1}^{m} (Y_j - \nu_j) \right]$$

$$= E \left[\sum_{i=1}^{n} \sum_{j=1}^{m} (X_i - \mu_i) (Y_j - \nu_j) \right]$$

$$= \sum_{i=1}^{n} \sum_{j=1}^{m} E \left[(X_i - \mu_i) (Y_j - \nu_j) \right]$$

$$= \sum_{i=1}^{n} \sum_{j=1}^{m} Cov[X_i, Y_j]$$

Now, say we pick n = m and $Y_i = X_i$ in part iv) of Proposition 31.2:

$$\begin{split} Var\left[\sum_{i=1}^{n}X_{i}\right] &= Cov\left[\sum_{i=1}^{n}X_{i}, \sum_{j=1}^{n}X_{j}\right] \\ &= \sum_{i=1}^{n}\sum_{j=1}^{n}Cov[X_{i},X_{j}] \\ &= \sum_{i=1}^{n}\left(Cov[X_{i},X_{i}] + \sum_{\substack{j=1\\j\neq i}}^{n}Cov[X_{i},X_{j}]\right) \\ &= \sum_{i=1}^{n}Var[X_{i}] + \sum_{\substack{i,j\\j\neq i}}^{n}Cov[X_{i},X_{j}] \\ &= \sum_{i=1}^{n}Var[X_{i}] + 2\sum_{\substack{i,j\\j\neq i}}^{n}Cov[X_{i},X_{j}] \end{split}$$

In the special case that each pair X_i, X_j are independent when $i \neq j$, then:

$$Var\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} Var[X_i]$$

Example 31.2: Recall (from Example 30.3) that $\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$ is called the **sample mean**. Let

$$S^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}$$

be the sample variance.

Let X_1, \ldots, X_n be iid with (common) mean μ and variance σ^2 .

Find a) $Var[\bar{X}]$ and b) $E[S^2]$. [b) is hard]

Solution: a)

$$Var[\bar{X}] = Var \left[\frac{1}{n} \sum_{i=1}^{n} X_i \right]$$
$$= \frac{1}{n^2} Var \left[\sum_{i=1}^{n} X_i \right]$$
$$= \frac{1}{n^2} \sum_{i=1}^{n} Var \left[X_i \right]$$
$$= \frac{\sigma^2}{n}$$

b) First some algebra:

$$(n-1)S^{2} = \sum_{i=1}^{n} (X_{i} - \mu + \mu - \bar{X})^{2}$$

$$= \sum_{i=1}^{n} (X_i - \mu)^2 + \sum_{i=1}^{n} (\bar{X} - \mu)^2 - 2\sum_{i=1}^{n} (\bar{X} - \mu)(X_i - \mu)$$

$$= \sum_{i=1}^{n} (X_i - \mu)^2 + \sum_{i=1}^{n} (\bar{X} - \mu)^2 - 2(\bar{X} - \mu)\sum_{i=1}^{n} (X_i - \mu)$$

$$= \sum_{i=1}^{n} (X_i - \mu)^2 - n(\bar{X} - \mu)^2$$

Hence
$$(n-1)E[S^2] = E\left[\sum_{i=1}^n (X_i - \mu)^2\right] - nE[(\bar{X} - \mu)^2]$$

 $= \sum_{i=1}^n E[(X_i - \mu)^2] - nVar[\bar{X}]$
 $= n\sigma^2 - n\frac{\sigma^2}{n}$
 $= (n-1)\sigma^2$

Example 31.3: Compute the variance of $X \sim \mathsf{Binomial}(n, p)$.

Solution: $X = X_1 + \cdots + X_n$ where $X_1, ..., X_n$ are iid and $\sim \text{Bernoulli}(p)$.

$$Var[X] = Var[X_1 + \dots + X_n]$$

$$= \underbrace{Var[X_1]}_{p(1-p)} + \dots + \underbrace{Var[X_n]}_{p(1-p)}$$
 [since X_i are independent]
$$= np(1-p)$$