## Random Variables (rv)

## Mean and Variance of Poisson [Ross S4.7]

Intuition: Say  $X \sim \text{Binomial}(n, p)$  with  $\lambda = np$ , n large, and p small

Then:

$$E[X] = np = \lambda$$

$$Var[X] = np(1 - p)$$

$$= \lambda(1 - p)$$

$$\approx \lambda$$

Exact: Let  $X \sim \mathsf{Poisson}(\lambda)$ . Then

$$\begin{split} E[X] &= \sum_{k=0}^{\infty} k \frac{\lambda^k}{k!} e^{-\lambda} \\ &= \sum_{k=1}^{\infty} k \frac{\lambda^k}{k!} e^{-\lambda} \\ &= \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} e^{-\lambda} \\ &= \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} e^{-\lambda} \\ &= \lambda \sum_{\ell=0}^{\infty} \frac{\lambda^{\ell}}{\ell!} e^{-\lambda} \qquad \qquad \ell = k-1 \\ &= \lambda \end{split}$$

$$\begin{split} E[X^2] &= \sum_{k=0}^{\infty} k^2 \frac{\lambda^k}{k!} e^{-\lambda} \\ &= \sum_{k=1}^{\infty} k^2 \frac{\lambda^k}{k!} e^{-\lambda} \\ &= \sum_{k=1}^{\infty} \frac{k \lambda^k}{(k-1)!} e^{-\lambda} \\ &= \sum_{\ell=0}^{\infty} \frac{(\ell+1)\lambda^{\ell+1}}{\ell!} e^{-\lambda} \\ &= \lambda \left( \sum_{\ell=0}^{\infty} \frac{\ell \lambda^{\ell}}{\ell!} e^{-\lambda} + \sum_{\ell=0}^{\infty} \frac{\lambda^{\ell}}{\ell!} e^{-\lambda} \right) \\ &= \lambda (1+\lambda) \end{split}$$

So 
$$Var[X] = E[X^2] - (E[X])^2$$
  
=  $\lambda (1 + \lambda) - (\lambda)^2$   
=  $\lambda$ 

**Example 13.1:** A radioactive substance with a large # of atoms emits 3.2 alpha particles per second on average. What is the probability that no more than 2 alpha particles are emitted in a 1 second interval?

Solution: If X = # of emitted particles in 1 second, then X is Poisson with  $E[X] = 3.2 = \lambda$ 

$$P[X \le 2] = P[X = 0] + P[X = 1] + P[X = 2]$$
$$= e^{-3.2} + 3.2e^{-3.2} + \frac{(3.2)^2}{2!}e^{-3.2}$$

## **D)** The geometric random variable [Ross 4.8.1]

Consider an infinite sequence of independent Bernoulli(p) trials.

Let X be trial # of first outcome that is a 1.

X is called **geometric** with parameter p, denoted  $X \sim \mathsf{Geometric}(p)$ 

$$p_X(k) = P[(k-1) \text{ zeros followed by a one}]$$
 for  $k=1,2,\ldots$  
$$= \begin{cases} (1-p)^{k-1}p & k \geq 1 \\ 0 & \text{else} \end{cases}$$

**Example 13.2:** A bag contains 2 white balls and 3 black balls. Balls are randomly drawn until a black ball is drawn. The selected ball is replaced before the next draw.

- a) What is the probability that exactly n draws are needed?
- b) What is the probability that at least k draws are needed?

Solution: In each draw, the probability of getting a black ball is 3/5 = 0.6.

If X = # of draws until a black ball, then  $X \sim \mathsf{Geometric}(p)$  with p = 0.6.

a)

$$P[X = n] = \left(1 - \frac{3}{5}\right)^{n-1} \times \frac{3}{5}$$
$$= \left(\frac{2}{5}\right)^{n-1} \times \frac{3}{5}$$

b)

$$\begin{split} P[X \geq k] &= \sum_{n=k}^{\infty} P[X = n] \\ &= \frac{3}{5} \times \sum_{n=k}^{\infty} \left(\frac{2}{5}\right)^{n-1} \\ &= \frac{3}{5} \times \left(\frac{2}{5}\right)^{k-1} \sum_{n=0}^{\infty} \left(\frac{2}{5}\right)^{n} \\ &= \frac{3}{5} \times \left(\frac{2}{5}\right)^{k-1} \frac{1}{1-2/5} \\ &= \left(\frac{2}{5}\right)^{k-1} \qquad (=P[\text{first } k-1 \text{ draws are white}]) \end{split}$$

## Mean and Variance

If  $X \sim \mathsf{Geometric}(p)$ , then:

$$E[X] = \sum_{k=1}^{\infty} k(1-p)^{k-1}p$$

$$= \cdots$$
 [see Ross example 4.8b]
$$= \frac{1}{p}$$

$$E[X^2] = \sum_{k=1}^{\infty} k^2 (1-p)^{k-1} p$$
 
$$= \cdots$$
 [see Ross example 4.8c] 
$$= \frac{2-p}{p^2}$$

$$\Rightarrow Var[X] = E[X^2] - (E[X])^2$$

$$= \frac{2-p}{p^2} - \left(\frac{1}{p}\right)^2$$

$$= \frac{1-p}{p^2}$$