## **Continuous Random Variables**

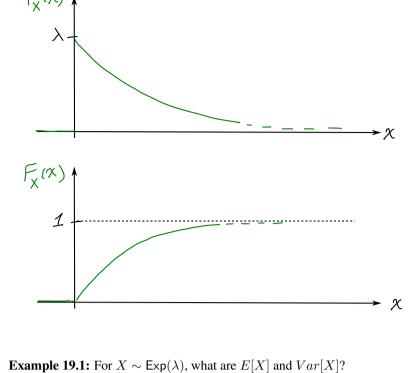
## C) Exponential Random Variable [Ross S5.5]

A random variable X with pdf

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \ge 0\\ 0 & \text{else} \end{cases}$$

is called **exponential** with **rate parameter**  $\lambda > 0$  and denoted  $X \sim \mathsf{Exp}(\lambda)$ . *Note:* If X has units of min then  $\lambda$  has units min<sup>-1</sup>.

$$F_X(a) = \int_{-\infty}^a f_X(u) du$$
$$= \begin{cases} 1 - e^{-\lambda a} & a \ge 0\\ 0 & a < 0 \end{cases}$$



 $E[X^n] = \int_{-\infty}^{\infty} x^n f_X(x) dx$ 

Solution: We compute  $E[X^n]$  first.

$$=\int_{0}^{\infty}\underbrace{x^{n}}_{u}\underbrace{\lambda e^{-\lambda x}dx}_{dv} \qquad u=x^{n} \qquad dv=\lambda e^{-\lambda x}dx$$

$$=\left[uv\Big|_{0}^{\infty}-\int_{0}^{\infty}vdu\right] \qquad du=nx^{n-1}dx \quad v=-e^{-\lambda x}$$

$$=\left[-x^{n}e^{-\lambda x}\Big|_{0}^{\infty}-\int_{0}^{\infty}-e^{-\lambda x}nx^{n-1}dx\right]$$

$$=\frac{n}{\lambda}\int_{0}^{\infty}x^{n-1}\lambda e^{-\lambda x}dx$$

$$=\frac{n}{\lambda}E[X^{n-1}]$$
Since  $E[X^{0}]=E[1]=1$ , then
$$E[X]=\frac{1}{\lambda}E[X^{0}]=\frac{1}{\lambda}$$

Hence

Hence 
$$Var[X] = E[X^2] - (E[X])^2$$
$$= \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2$$

 $E[X^2] = \frac{2}{\lambda} E[X^1] = \frac{2}{\lambda^2}$ 

**Example 19.2:** The time someone uses an ATM machine is an exponential random variable with 
$$\lambda = 1/3 \text{ min}^{-1}$$
. Someone arrives at the ATM just before you. What is the probability that you wait a) more than 3 min,

b) between 3 and 6 min? Solution:  $X \sim \text{Exp}(1/3)$ .

for all s > 0 and all t > 0

distribution.

a)  $P[X > 3] = 1 - F_X(3) = \exp(-3\lambda) = \exp(-1) \approx 0.36788$ b)  $P[3 < X < 6] = F_X(6) - F_X(3) = [1 - \exp(-6\lambda)] - [1 - \exp(-3\lambda)]$  $= \exp(-1) - \exp(-2) \approx 0.23254$ 

$$P[X>s+t\mid X>t]=P[X>s]$$

In words: The probability of waiting s seconds more given you have already waited t seconds is the same as waiting s seconds from the start. In other words, no matter how long you have waited, time to wait still has the same

**Definition 19.1:** A non-negative random variable X is called **memoryless** if

**Example 19.3:** Does  $Exp(\lambda)$  have the memoryless property? Solution: Let  $X \sim \mathsf{Exp}(\lambda)$ . Then

Yes, 
$$\mathsf{Exp}(\lambda)$$
 has the memoryless property.   
**Example 19.4:** Persons A and B are each being served by a teller. Person C arrives, and waits for one of the two tellers. All service times are exponential with the same parameter  $\lambda$ . What is the probability that C is the last to leave?   
*Solution:* C starts being served as soon as one of A or B is finished. Once this happens, the time to go for remaining person and C has the same distribution  $\mathsf{Exp}(\lambda)$  due to memoryless property. By symmetry, each has a probability 1/2 of finishing last.

with mean 10,000 km.

$$\begin{split} P[X>s+t\mid X>t] &= \frac{P[X>s+t,X>t]}{P[X>t]} \\ &= \frac{P[X>s+t]}{P[X>t]} \\ &= \frac{e^{-\lambda(s+t)}}{e^{-\lambda t}} \end{split}$$

=P[X>s]

a) What is the probability of completing a 5000 km trip without replacing the battery? b) What can we say if lifetime is not exponential? Solution: Let  $X \sim \mathsf{Exp}(\lambda)$  with  $\lambda = 1/10000$ .

> $=1-F_X(5000)$  $=\exp(-5000/10000)$

Example 19.5: A car battery has a lifetime that is exponentially distributed

a) Since battery has operated for d km so far,  $P[X > 5000 + d \mid X > d] = P[X > 5000]$ [by memoryless property]

Let d = # km that battery has been for operating so far.

- $\approx 0.607$
- b)

 $P[X > d + 5000 \mid X > d] = \frac{P[X > d + 5000, X > d]}{P[X > d]}$   $= \frac{P[X > d + 5000]}{P[X > d]}$   $= \frac{1 - F_X(d + 5000)}{1 - F_X(d)}.$ The exponential distribution can be used:

to model service times in queuing systems

· time between radioactive decays · credit risk modeling in finance

- is maximum entropy distribution on  $[0, \infty]$  subject to a specified mean.