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`presentation{ Outline`

Convergence Concepts

`\allowframebreaks`[Introduction](#)

This material comes primarily from [\[Chapter 5\]rice07](#), but will be supplemented with material from [\[Chapter 5\]casella](#).  
In this chapter, we are interested in the convergence of sequences of random variables.

In particular, we are interested in the convergence of the sample mean,  $\bar{X}_n = (X_1 + X_2 + \dots + X_n)/n$ , as the number of observations  $n$  increases.

Because  $\bar{X}_n$  is itself a random variable, we have to carefully define what it means for the convergence of a random variable.

In this class, we are mainly concerned with three types of convergence.

Because convergence of random variables is a tricky topic, we will treat them in varying amounts of detail.

`\allowframebreaks`[Convergence in Probability](#)

The first type of convergence is one of the weaker types, and is usually easy(ish) to verify.

Definition: Convergence in Probability A sequence of random variables  $X_1, X_2, \dots$  converges in probability to a random variable  $X$  if

particle<sub>i</sub>. The idea for this proof is to show that, for some neighborhood around  $t = 0$ , the mgf of  $\sqrt{n}(\bar{X}_n - \mu)/\sigma$  converges to the mgf of  $Z$ . Let  $Z_i = (X_i - \mu)/\sigma$ . Using the properties of mgf, the mgf of  $Z_i$ , denoted  $M_{Z_i}$  exists for  $|t| < \sigma t$ . The exact form of  $M_{Z_i}(t)$  is given by:

$$\begin{aligned} \text{Thus, using the properties of the mgf, we have } M_{\sqrt{n}(\bar{X}_n - \mu)/\sigma}(t) &= M_{\sum_{i=1}^n Z_i / \sqrt{n}}(t) \\ &= M_{\sum_{i=1}^n Z_i}(t/\sqrt{n}) \\ &= \left(M_Z(t/\sqrt{n})\right)^n. \text{ We now do a second order Taylor-Expansion of } M_Z(t/\sqrt{n}) \text{ about 0: } M_Z(t/\sqrt{n}) = M_Z(0) + M_Z^{(1)}(0) \frac{t}{\sqrt{n}} + \frac{t^2}{2n} + R_2(t/\sqrt{n}) \text{ Now by Taylor's theorem, there exists some function } h_2(x) \text{ such that} \\ &= 1 + \frac{t^2}{2n} + R_2(t/\sqrt{n}) \end{aligned}$$

where

or equivalently,

Since  $t$  is fixed, (and because the below equality holds when  $t = 0$ ), this implies:

$$\begin{aligned} \text{Finally, returning to the mgf of } \sqrt{n}(\bar{X}_n - \mu), \text{ we take the limit as } n \rightarrow \infty: \lim_{n \rightarrow \infty} M_{\sqrt{n}(\bar{X}_n - \mu)/\sigma}(t) &= \lim_{n \rightarrow \infty} \left(1 + \frac{t^2}{2n} + R_2(t/\sqrt{n})\right)^n \\ &= \lim_{n \rightarrow \infty} \left(1 + \frac{1}{n} \left[ \frac{t^2}{2} + nR_2(t/\sqrt{n}) \right]\right)^n \\ &= \lim_{n \rightarrow \infty} \left(1 + \frac{a_n}{n}\right)^n, \text{ where } a_n = \frac{t^2}{2} + nR_2(t/\sqrt{n}). \text{ Here, we note that } a_n \rightarrow \frac{t^2}{2} \text{ due to the convergence of } nR_2(t/\sqrt{n}) \text{ to 0.} \end{aligned}$$

Therefore, we get

Thus, by the continuity theorem,

One practical implication of the CLT is that, for large  $n$ , we can approximate

if  $X_i$  are independent and identically distributed with finite mean and variance.

In practice,  $n \approx 30$  has been found to lead to good approximations, but it depends heavily on the distribution of  $X_i$ .

A further investigation of the CLT proof shows that the convergence towards the normal distribution happens at a rate.

If we used a Taylor-series approximation with one additional order, we could derive a more accurate approximation under certain conditions.

Example: Binomial-Normal Approximation Let  $X \sim \text{Binomial}(n, p)$ . For any  $k \in \{0, 1, \dots\}$ , approximate  $P(X \leq k)$ .

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Note that if  $X$  is binomial distributed, it has the same distribution as the sum of Bernoulli random variables.

Let  $X_1, X_2, \dots, X_n$  be Bernoulli( $p$ ) random variables, and then  $X \equiv \sum_i X_i$ .

Recall  $E[X_i] = p$ ,  $\text{Var}(X_i) = p(1-p)$ .

By considering  $\bar{X}_n = \frac{1}{n} \sum_i X_i$ , we can use the CLT to approximate:

Therefore,

and

Using this approximation, we have

One thing to note is we can make a fairly simple improvement to this approximation by doing a continuity correction. Specifically,  $X$  only takes on real values, so if we approximate it with a continuous distribution  $X \rightarrow Y$ , we generally get

allowframebreaksSlutsky's Theorem

The following theorem is useful for our notes and supporting other ideas we will cover. However, we won't discuss the proof.

Theorem: Slutsky's Theorem If  $X_n \xrightarrow{d} X$  and  $Y_n \xrightarrow{p} a$  for some random variable  $X$  and constant  $a$ , then

$Y_n X_n \xrightarrow{d} aX$ .

$X_n + Y_n \xrightarrow{d} X + a$ .

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