

# **Mathematical Statistics II**

## **The Bayesian Approach to Parameter Estimation**

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# **Introduction**

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## Bayesian Estimation

- Much of this work is based on Rice (2007, Section 8.6).
- We have already discussed the philosophy of Bayesian statistics.
- We start with a prior belief about parameter values, and update these beliefs using observed data.
- The resulting **distribution** is called the *posterior*, and it represents our updated belief after observing data.
- This is very natural idea that is closely related to the idea of likelihood: likelihood quantifies some degree of belief about a parameter value.

# **Review**

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## Some Review

- Before we begin, we will first do a bit of review.
- In the context of Bayesian inference, we treat unknown parameter vectors as random variables, which I will denote  $\Theta$ .
- Thus, our probability model can be expressed as  $f(x|\Theta = \theta)$ , which we often shorten to  $f(x|\theta)$ .

## Some Review II

### Bayes' Theorem

Let  $X$  be the random vector representing observed data, and  $\Theta$  the random parameter vector, and  $x^*$  the observed data. Bayes Theorem states:

$$\begin{aligned}\pi_{\Theta|X}(\theta|x^*) &= \frac{f_{X|\Theta}(x^*|\theta)\pi_\Theta(\theta)}{f_X(x^*)} \\ &= \frac{f_{X|\theta}(x^*|\theta)\pi_\Theta(\theta)}{\int f_{X|\Theta}(x^*|\tau)\pi_\Theta(\tau) d\tau}\end{aligned}$$

- As before,  $f$  is taken to be either a pmf or pdf, depending on the problem.

## Some Review III

### Flipping 10 coins

Our friend hands us a coin from another country, and we want to estimate  $\theta = p$ , the probability that the coin lands heads.

Suppose we flip a coin 10 times, and see  $n$  heads. Find a Bayesian estimate for  $\theta$ .

## Some Review IV

- Even in the simple problem above, we see two of the primary challenges with Bayesian parameter estimation:
  - How do we choose the prior distribution  $\pi(\theta)$ ? A generally safe and accepted approach is a uniform prior. However, this formally only exists if  $\theta$  is bounded, which is not always the case. Also, it represents a prior belief: given a new coin, do we really think all values of  $p$  are equally likely, or maybe values close to  $p = 0.5$  are more likely than extreme values  $p = 0, 1$ ? Since the prior represents our beliefs about  $\theta$ , is a uniform prior actually appropriate? If it isn't appropriate, how exactly should we specify the prior?
  - Even in this very simple model and prior, the denominator  $f(x)$  was difficult to compute. What about more complex models and priors? A large amount of Bayesian computation and theory is dedicated to solving this problem.

## Some Review V

### Proposition: the MAP and MLE

Let  $\theta$  be a parameter of interest, and  $x^*$  the observed data. If our prior distribution is proportional to 1, i.e.,  $\pi(\theta) \propto 1$  (which is effectively a uniform prior on a bounded interval), then

$$\hat{\theta}_{\text{MAP}} = \hat{\theta}_{\text{MLE}}.$$

- This is true for the Coin-tossing example; look back at the likelihood function and posterior, and use R to plot them both.

## Examples

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## Bayesian point-estimate examples

### Poisson model

Suppose we have observations  $n$  observations, which we wish to model as IID  $\text{Poisson}(\lambda)$ . Find a Bayesian estimate of  $\Lambda = \lambda$  given the observed data  $x^*$ .

## Bayesian point-estimate examples II

### Poisson posterior, uniform prior

Revisit the Poisson( $\lambda$ ) model, while taking the alternative approach of using a uniform prior.

## Real-data example: Poisson Distribution

- Now let's look at a real-data example. These data are the 23 observations from the asbestos-filter problem.

```
x <- c(  
 31, 29, 19, 18, 31, 28, 34, 27, 34, 30, 16, 18,  
 26, 27, 27, 18, 24, 22, 28, 24, 21, 17, 24  
)  
x
```

```
[1] 31 29 19 18 31 28 34 27 34 30 16 18 26 27 27 18 24 22  
[19] 28 24 21 17 24
```

- TODO: Compare Bayes 1 to Bayes 2 to MLE

## References and Acknowledgements

Rice JA (2007). *Mathematical statistics and data analysis*, volume 371. 3 edition. Thomson/Brooks/Cole Belmont, CA.

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