Simple linear regression I

EC 320, Set 04

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Housekeeping

PS02:

- Due Tuesday at 11:59p
- Covers remaining topics from the review and from lecture Wednesday

LA03:

Due Friday at 5:00p

Reading: (up to this point)

ItE: R, 1 MM: 1, 2

Regression logic

Regression models

Modeling is about reducing something *really complicated* into something *simple* that represents some part of the complicated reality.

- Try to tell stories that are easy to understand, and easy to learn from
- Model toy versions of reality

Economists often rely on linear regression for statistical comparisons.

- Describes the relationship between a dependent (endogenous) variable and one or more explanatory (exogenous) variable(s)
- "Linear" is more flexible than you think

Regression models

Regression analysis helps us make all else equal comparisons

Running regressions provide correlative (and even causal) information between two variables

Ex. By how much does Y change when X increases by one unit?

Regression models

Modelling forces us to be explicit about the potential sources of selection bias

- ullet Model the effect of X on Y while **controlling** for potential confounders that may muddy the water
- Failure to account for sources of selection bias, leads to biased estimates.

Ex. Not controlling for confounding variables, leads to omitted-variable bias, a close cousin of selection bias

 Why? Omitted variables that correlate with our covariate of interest, hides within the model, distorting our results.

Returns to education

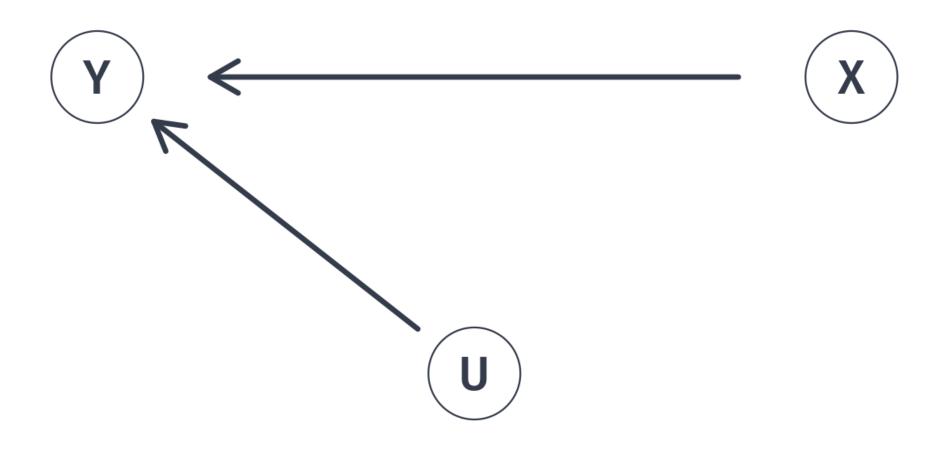
Research Question: By how much does an additional year of schooling increase future earnings?

- Dependent variable: Earnings
- Independent variable: An additional year of school
- Q. How might education increase earnings?
- Q. Why might a simple comparison between high and low educated not isolate the economic returns to education?

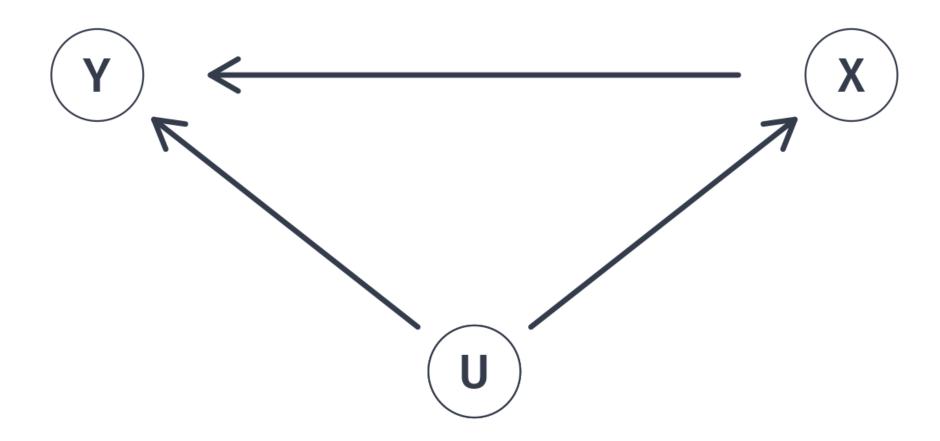
More education (**X**) increases lifetime earnings (**Y**)



More education (**X**) increases lifetime earnings (**Y**) along with a lot of other things (**U**).



More education (**X**) increases lifetime earnings (**Y**) along with a lot of other things (**U**). But a lot of other things (**U**) also impact education (**X**).



Any unobserved variable that connects a backdoor path between education (X) and earnings (Y) is called a **confounder**

Returns to education

How might we estimate the causal effect of an additional year of schooling on earnings?

Approach 1: Compare average earnings of private college graduates with those of public college graduates.

• Prone to selection bias by variety of confounding variables

Approach 2: Estimate a regression that compares the earnings of individuals with the same admissions profiles.

Try to control for confounders by including them in the model

But before taking on *confounders* and using regression to link causal relationships... let's breakdown the anatomy of the *simple regression model*

Simple regression model

The regression model

We can estimate the effect of \boldsymbol{X} on \boldsymbol{Y} by estimating a **regression model:**

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

- Y_i is the outcome variable.
- X_i is the treatment variable (continuous).
- β_0 is the **intercept** parameter. $\mathbb{E}[Y_i|X_i=0]=\beta_0$
- β_1 is the **slope** parameter, which under the correct causal setting represents marginal change in X_i 's effect on Y_i . $\frac{\partial Y_i}{\partial X_i} = \beta_1$
- ullet u_i is an error term including all other (omitted) factors affecting Y_i .

 u_i is quite special

Consider the data generating process of variable Y_i ,

ullet u_i captures all unobserved variables that explain variation in Y_i .

Some error will exist in all models, our aim is to minimize error under a set of constraints

• Error is the price we are willing to accept for simplified model

Five items contribute to the existence of the disturbance term:

1. Omission of explanatory variables

- \bullet Our description (model) of the relationship between Y and X is a simplification
- Other variables have been left out (omitted)

Five items contribute to the existence of the disturbance term:

1. Omission of explanatory variables

2. Aggregation of Variables

- Microeconomic relationships are often summarized
- ullet Ex. Housing prices (X) are described by county-level median home value data

- 1. Omission of explanatory variables
- 2. Aggregation of Variables
- 3. Model misspecificiation
- Model structure is incorrectly specified
- \bullet $\textit{Ex.}\ Y$ depends on the anticipated value of X in the previous period, not X

- 1. Omission of explanatory variables
- 2. Aggregation of Variables
- 3. Model misspecificiation
- 4. Functional misspecificiation
- The functional relationship is specified incorrectly
- True relationship is nonlinear, not linear

- 1. Omission of explanatory variables
- 2. Aggregation of Variables
- 3. Model misspecificiation
- 4. Functional misspecificiation
- **5.** Measurement error
- Measurement of the variables in the data is just wrong
- Y or X

- 1. Omission of explanatory variables
- 2. Aggregation of Variables
- 3. Model misspecificiation
- 4. Functional misspecificiation
- **5.** Measurement error

Running regressions

Using an estimator with data on X_i and Y_i , we can estimate a **fitted** regression line:

$$\hat{\mathbf{Y}}_{i} = \hat{\beta_{0}} + \hat{\beta_{1}} \mathbf{X}_{i}$$

- \hat{Y}_i is the **fitted value** of Y_i .
- $\hat{\beta_0}$ is the **estimated intercept**.
- $\hat{\beta_1}$ is the **estimated slope**.

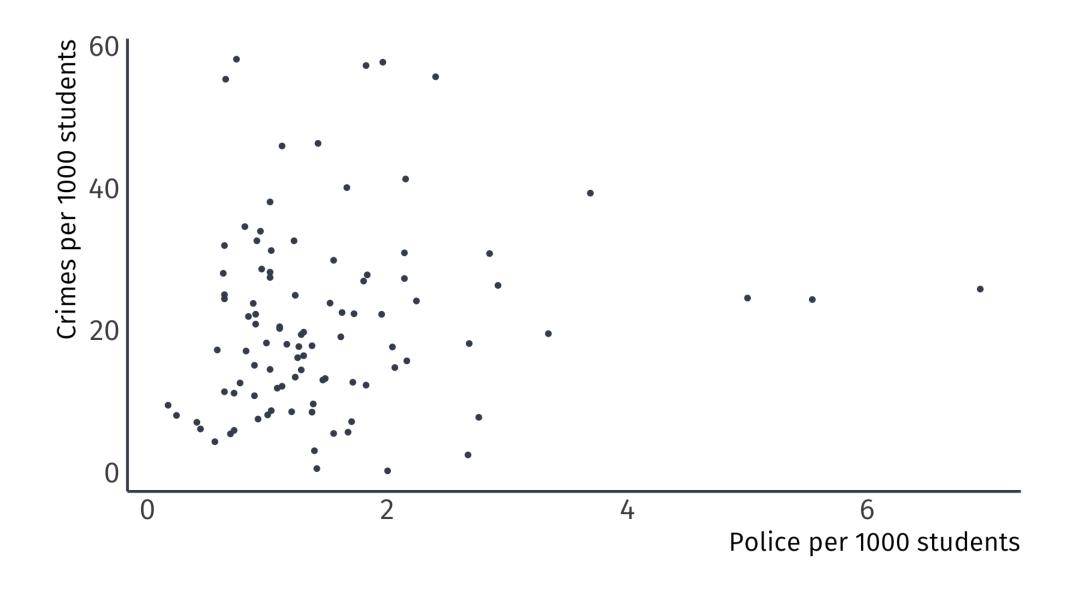
This procedure produces misses, known as **residuals**, $Y_i - \overset{\wedge}{Y_i}$

I think it would example.	be easier to think	about regression	with a concrete

Empirical question:

Does the number of on-campus police officers affect campus crime rates? If so, by how much?

Always plot your data first



The scatter plot suggest that a weak positive relationship exists

• A sample correlation of 0.14 confirms this

But correlation does not imply causation

Lets estimate a statistical model

We express the relationship between a **explained variable** and an **explanatory variable** as linear:

{\color{#81A1C1} \text{Crime}_i} = \beta_1 + \beta_2 {\color{#B48EAD} \text{Police}_i}

- β_1 is the *intercept* or constant.
- β_2 is the slope coefficient.
- \mathbf{u}_{i} is an error term or disturbance term.

The **intercept** tells us the expected value of $Crime_i$ when $Police_i = 0$.

$$Crime_i = \beta_1 + \beta_2 Police_i + u_i$$

Usually not the focus of an analysis.

The **slope coefficient** tells us the expected change in $Crime_i$ when $Police_i$ increases by one.

$$Crime_i = \beta_1 + \beta_2 Police_i + u_i$$

"A one-unit increase in $Police_i$ is associated with a β_2 -unit increase in $Crime_i$."

Interpretation of this parameter is crucial

Under certain (strong) assumptions¹, β_2 is the effect of X_i on Y_i .

• Otherwise, it's the association of X_i with Y_i .

The error term reminds us that $Police_i$ does not perfectly explain Y_i .

$$Crime_i = \beta_1 + \beta_2 Police_i + \mathbf{u}_i$$

Represents all other factors that explain Crime_i.

• Useful mnemonic: pretend that **u** stands for "unobserved" or "unexplained."

How might we apply the simple linear regression model to our question about the effect of on-campus police on campus crime?

$$Crime_i = \beta_1 + \beta_2 Police_i + u_i$$
.

- β_1 is the crime rate for colleges without police.
- β_2 is the increase in the crime rate for an additional police officer per 1000 students.

How might we apply the simple linear regression model to our question?

$$Crime_i = \beta_1 + \beta_2 Police_i + u_i$$

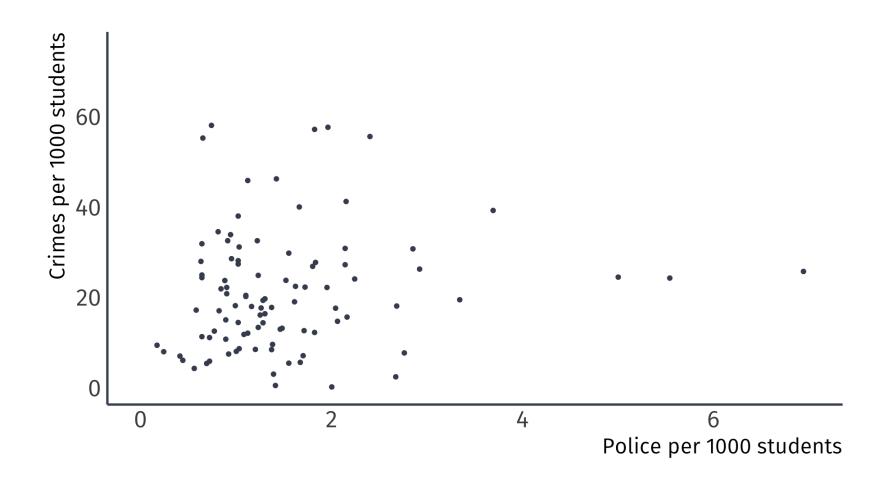
 eta_1 and eta_2 are the unobserved population parameters we want

We estimate

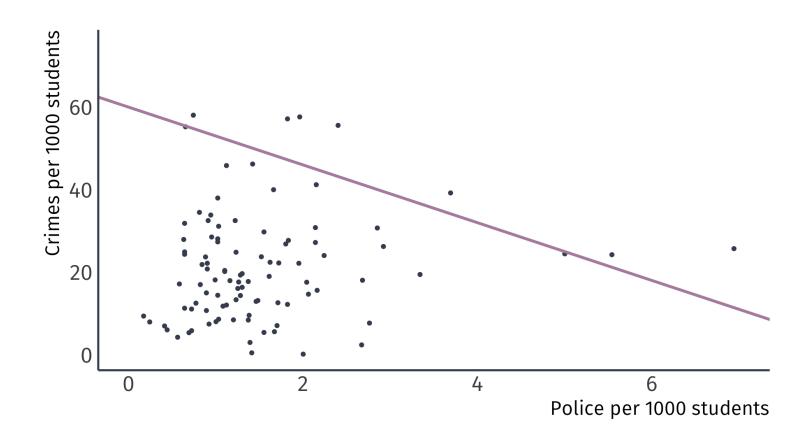
- $\hat{\beta_1}$ and $\hat{\beta_2}$ generate predictions of $Crime_i$ called $\widehat{Crime_i}$.
- We call the predictions of the dependent variable **fitted values.**
- Together, these trace a line: $\widehat{Crime}_i = \hat{\beta_1} + \hat{\beta_2} Police_i$.

So, the question becomes, how do I pick $\hat{\beta_1}$ and $\hat{\beta_2}$

Let's take some guesses:

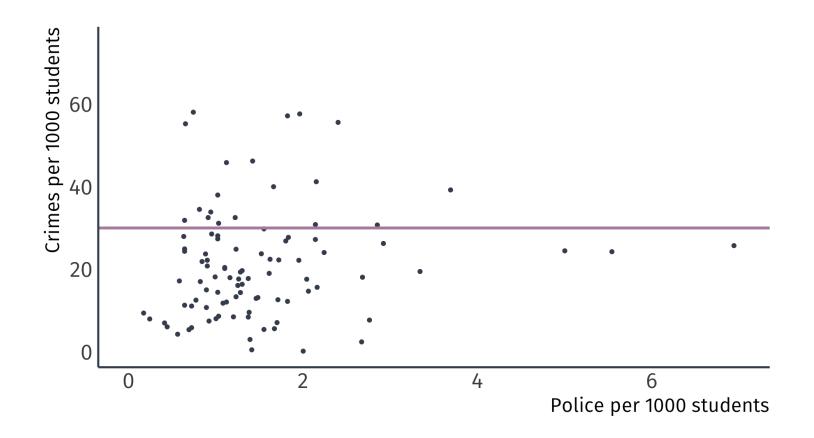


Let's take some guesses:
$$\hat{\beta_1} = 60$$
 and $\hat{\beta_2} = -7$



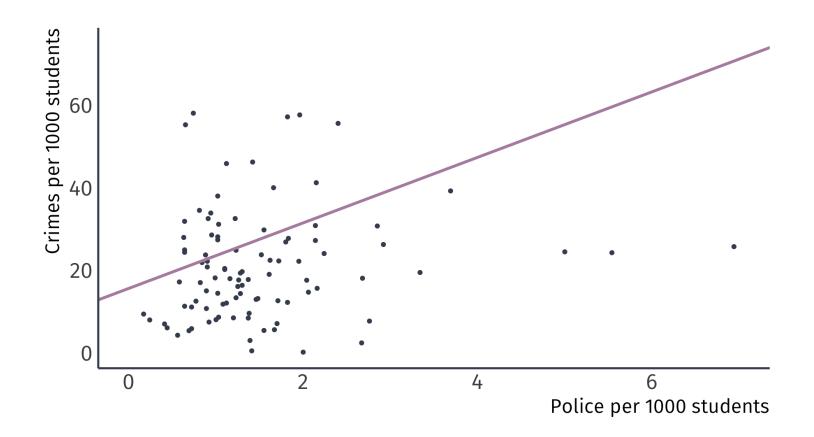
Does this line represent the data well?

Let's take some guesses: $\hat{\beta_1} = 60$ and $\hat{\beta_2} = -7$



What about this one?

Let's take some guesses: $\hat{\beta_1}=60$ and $\hat{\beta_2}=-7$



Or this one?

Residuals

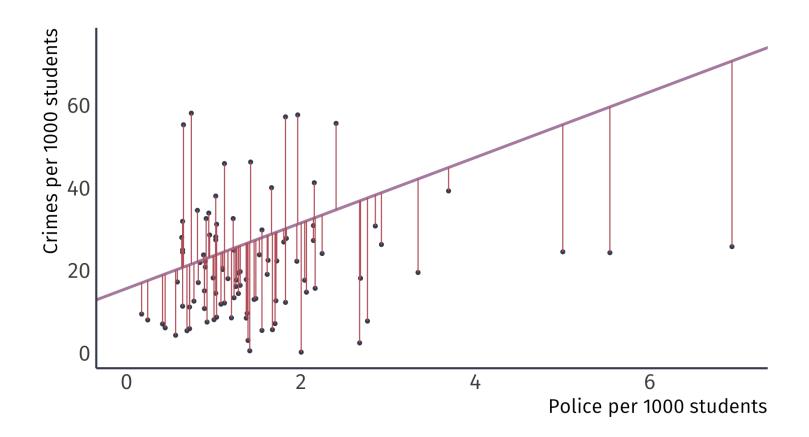
Using $\hat{\beta_1}$ and $\hat{\beta_2}$ to make $\hat{Y_i}$ generates misses.

We call these misses residuals:

{\color{#BF616A} \hat{u}_i} = {\color{#BF616A}Y_i - \hat{Y_i}}.

AKA ei.

$$\hat{\beta_1}=60$$
 and $\hat{\beta_2}=-7$



Does this line represent the data well?

Residuals

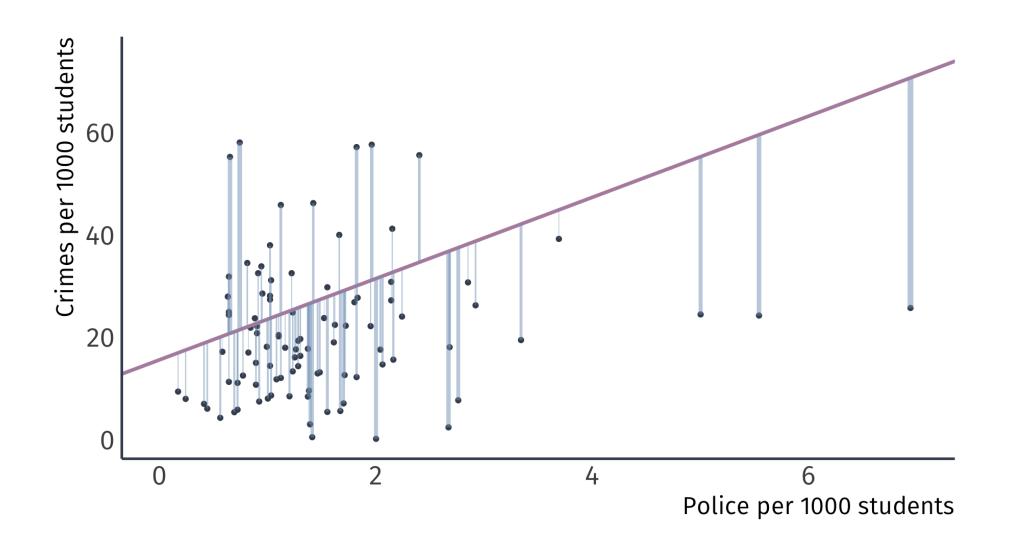
What is we picked an estimator that minimizes the residuals? Why not minimize

$$\sum_{i=1}^{n} \hat{u_i^2}$$

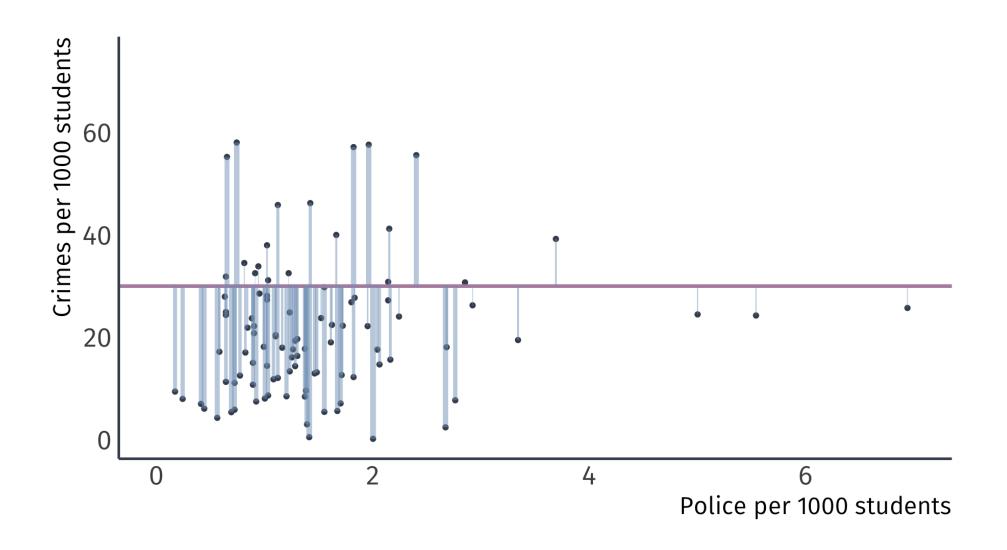
so that the estimator makes fewer big misses?

This estimator, the residual sum of squares (RSS), is convenient because squared numbers are never negative so we can minimize an absolute sum of the residuals

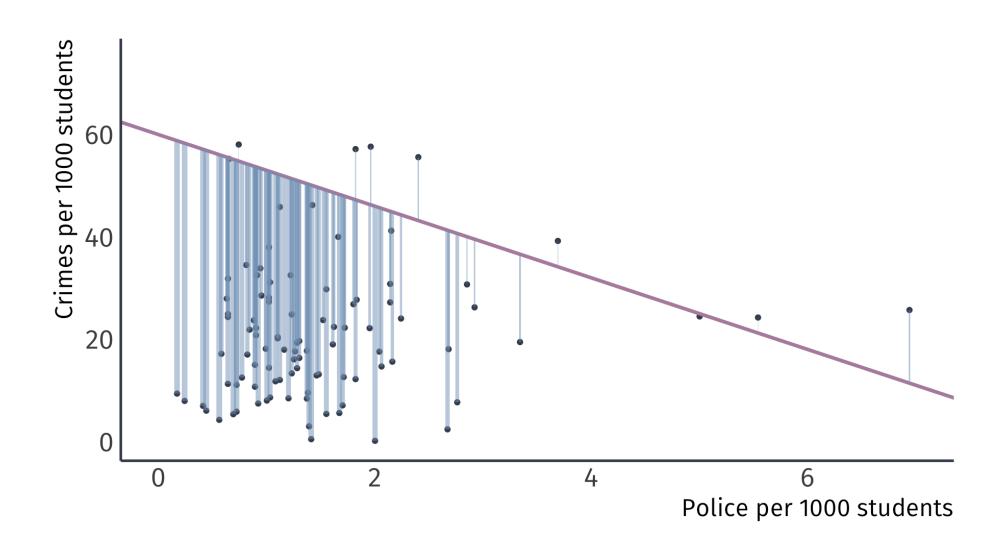
RSS gives bigger penalties to bigger residuals.



RSS gives bigger penalties to bigger residuals.



RSS gives bigger penalties to bigger residuals.



Minimizing RSS

We could test thousands of guesses of β_1 and β_2 and pick the pair the has the smallest RSS

Or... We could just do a little math

Ordinary least squares

OLS

The OLS estimator chooses the parameters $\hat{\beta_1}$ and $\hat{\beta_2}$ that minimize the residual sum of squares (RSS):

$$\min_{\hat{\beta_1}, \, \hat{\beta_2}} \quad \sum_{i=1}^{n} \hat{u_i^2}$$

This is why we call the estimator ordinary **least squares**.

OLS Formulas

For details, see the handout posted on Canvas.

Slope coefficient

$$\hat{\beta_2} = \frac{\sum_{i=1}^{n} (Y_i - Y)(X_i - X)}{\sum_{i=1}^{n} (X_i - X)^2}$$

Intercept

$$\hat{\beta_1} = Y - \hat{\beta_2} X$$

Slope coefficient

The slope estimator is equal to the sample covariance divided by the sample variance of \mathbf{X} :

$$\hat{\beta_2} = \frac{\sum_{i=1}^{n} (Y_i - Y)(X_i - X)}{\sum_{i=1}^{n} (X_i - X)^2}$$

$$= \frac{\frac{1}{n-1} \sum_{i=1}^{n} (Y_i - Y)(X_i - X)}{\frac{1}{n-1} \sum_{i=1}^{n} (X_i - X)^2}$$

$$=\frac{S_{XY}}{S_X^2}.$$

Continue with OLS next time