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ML Estimation (Ch 4)

Suppose we have $E(y_i | x_i) = x_i \beta$ and $\text{Cov}(y_i | x_i) = \Sigma_i = \Sigma_i(\theta)$

Σ_i is a function of θ
 θ = vector of covariance parameters (e.g., σ_{ij})

Using ML to estimate β , we get $\hat{\beta}_{ML} = \left[\sum_{i=1}^N (x_i' \hat{\Sigma}_i^{-1} x_i) \right]^{-1} \sum_{i=1}^N x_i' \hat{\Sigma}_i^{-1} y_i$

Σ_i is unknown so we must estimate it (meaning we must estimate all the σ_{ij})

It can be shown that in very ^{large # of clusters} large samples $E(\hat{\beta}_{ML}) = \beta$ and $\text{Cov}(\hat{\beta}) = \left[\sum_{i=1}^N x_i' \Sigma_i^{-1} x_i \right]^{-1}$ and $\hat{\beta}_{ML} \sim \text{MVN}(\beta, \text{Cov}(\hat{\beta}))$

We could use this large sample approximation to conduct inference (e.g., for $H_0: \beta_1 = 0$, we might compute $Z = \frac{\hat{\beta}_1 - 0}{\sqrt{\text{Cov}(\hat{\beta}_1)}}$)

But this falls apart in several ways:

(1) We don't know Σ_i so must estimate it with $\hat{\Sigma}_i$.

(2) Since we don't know Σ_i , we can't compute $\text{Cov}(\hat{\beta})$ and thus must estimate it with $\hat{\text{Cov}}(\hat{\beta}) = \left[\sum_{i=1}^N x_i' \hat{\Sigma}_i^{-1} x_i \right]^{-1}$

(3) The results are based on large sample approximations and do not hold for finite (smaller) samples

Specifically, $\hat{\Sigma}_i$ underestimates Σ_i in finite samples (i.e., $\hat{\Sigma}_i$ is too small $\Rightarrow \hat{\text{Cov}}(\hat{\beta})$ is also too small)

So test statistics (e.g. $t = \frac{\hat{\beta}_1 - \beta_{01}}{\text{SE}(\hat{\beta}_1)}$) are too large and CIs (e.g. $\hat{\beta} \pm t^* [\text{SE}(\hat{\beta})]$) are too narrow

and we are therefore more likely to falsely reject H_0 ("anti-conservative" inference)

Satterthwaite and KR adjustments

For a simple test of $H_0: \beta_1 = 0$:

Under H_0 , we assume $T = \frac{\hat{\beta}_1 - \beta_{01}}{\text{SE}(\hat{\beta}_1)} \sim t_{df_r}$

but $\text{SE}(\hat{\beta})$ is too small

Satterthwaite fix: inflate df of H_0 distribution

K-R fix: inflate df of H_0 distr'n and $\text{SE}(\hat{\beta})$

