

Markups, Firm Scale, and Distorted Economic Growth*

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Abstract

We study the consequences of markups for long-run economic growth in a model of firm-driven endogenous technological change. In this framework, differentiated firms engage in monopolistic competition, charge heterogeneous markups and invest in R&D to improve their process efficiency. Markups restrict the scale at which these firms operate and therefore distort their incentives to invest in R&D. With dispersion in markups, both the aggregate and cross-firm allocations of such investments are inefficient. Using firm-level administrative data from France to discipline our model, we find that size-dependent subsidies inducing firms to operate at their efficient scale increase the long-run growth rate of productivity by 1.2 percentage points. Nearly 75% of this faster productivity growth can be achieved by simply reallocating R&D resources across firms, revealing that it is the dispersion in markups, rather than their average level, that is more distortionary to economic growth.

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1 Introduction

The widespread rise in market concentration in recent decades has raised concerns about the aggregate consequences of product market power.¹ Concurrently, a series of studies have concluded that the macroeconomic costs of markups can be substantial, providing grounds for these concerns.² Yet, their calculations neglect the possibility that markups may not only distort the scale at which firms operate but also their incentives to bring better or cheaper products to the market. When such incentives serve as the engine of technological progress, markups may further distort the trajectory of economic growth.

The objective of this paper is therefore to quantify the consequences of markups for long-run economic growth. We do so through the lens of a dynamic general equilibrium model of firm-led endogenous technological change. In this framework, imperfectly competitive firms differ in the markups they command and make forward-looking investments in R&D to improve their technology. Under these premises, markups distort private incentives for such investments through three channels.

First, markups can induce *under-investment* in R&D. In fact, since the idea behind an improvement in technology is nonrival, once it is developed, it can be used by the firm to produce infinitely many units. Consequently, the return on an investment in R&D increases with the scale at which a firm operates. However, due to the markup it charges, an imperfectly competitive firm produces inefficiently too few units. This restricts the scale at which its improved technology can be deployed and, in turn, depresses its incentives to invest in R&D.

However, in general equilibrium, markups can instead lead to the opposite outcome, that is, *over-investment* in R&D. Through imperfect competition on the product market, firms limit the scale at which they produce, resulting in an inefficiently low aggregate demand for factors of production. Therefore, if those factors are shared between the purposes of production and R&D, markups effectively “free up” resources, which are underused in production and thus available for R&D.³

Lastly, dispersion in markup results in a *misallocation* of R&D investments across firms. Indeed, when firms differ in the markups they command, their operational scale is differentially suboptimal such that their incentives to invest in R&D are differentially distorted. Hence, whether the *level* of markups induces under- or over-investment in

¹See Grullon, Larkin and Michaely (2019), Autor, Dorn, Katz, Patterson and Van Reenen (2020), De Loecker, Eeckhout and Unger (2020), and Kehrig and Vincent (2021).

²See Baqaee and Farhi (2019), Bilbiie, Ghironi and Melitz (2019), Behrens, Mion, Murata and Suedekum (2020), Edmond, Midrigan and Xu (2023) and Afrouzi, Drenik and Kim (2023).

³This general equilibrium pecuniary externality is distinct and not to be confused with the “business stealing” externality in models of creative destruction.

R&D is ambiguous, but *dispersion* in markups invariably leads to an inefficient cross-firm allocation of such investments.

We quantify these two margins of inefficiency using a model where such dispersion in markups is endogenous. In the model we propose, differentiated firms engage in monopolistic competition and face non-isoelastic demand curves that satisfy [Marshall \(1890\)](#)’s second law of demand. This property of demand states that lower prices are met with less elastic demand, which implies that larger and more productive firms command higher markups—a fact supported by cumulating empirical evidence.⁴

Firms endogenously differ in their process efficiency, as determined by their forward-looking and risky investments in R&D.⁵ As these firms become more productive and lower their price to attract more demand, they are met with a progressively less elastic demand schedule, enabling them to charge higher markups. Hence, larger firms operate at a more suboptimal scale, which differentially depresses the return on their investments in R&D. Yet, as these investments are denominated in units of labor, for which aggregate demand is inefficiently low, the price of R&D falls short of its social (shadow) value.

The quantity of firms is endogenous and determined by the entry of newcomers and the selective survival of incumbents. New firms incur a labor-denominated entry cost to imperfectly imitate the existing technology of a randomly selected incumbent. Meanwhile, incumbent firms must bear a fixed overhead labor cost per unit of time to remain in business, and failure to do so results in endogenous exit. To meet these costs, as well as demand from production and innovation, labor is elastically supplied by a representative household.

In this model, economic growth is sustained by the profit-seeking investments of incumbent firms and the selective replacement of their unsuccessful competitors by more productive new entrants. Yet, both the level and dispersion of markups hinder this process. The former distorts economic growth by inducing under- or over-investment in R&D while the latter impedes it via two manifestations of R&D misallocation. First, since large firms differentially under-invest in R&D, the forgone improvements in their process efficiency would otherwise be rolled out over more units produced. Second, as firms restrict the scale at which they produce, their smaller, unproductive rivals face less competitive pressures and thus exit at an inefficiently low rate, obstructing their replacement by more productive new entrants.

We estimate the structural parameters of our model through a generalized method of

⁴See [De Loecker and Warzynski \(2012\)](#), [Amiti, Itskhoki and Konings \(2014\)](#), [De Loecker, Goldberg, Khandelwal and Pavcnik \(2016\)](#), [Amiti, Itskhoki and Konings \(2019\)](#), [De Loecker et al. \(2020\)](#) and [De Ridder, Grassi and Morzenti \(2023\)](#).

⁵In [Appendix A.3](#), we show that this theoretical framework is isomorphic to one in which firms achieve improvements in the quality of their product when quality and quantity are perfect substitutes.

moments (GMM) strategy. The targeted moments are calculated from an administrative panel dataset covering the near universe of French firms between 2009 and 2019.⁶ The first objective of this estimation exercise is to discipline the degree of markup dispersion, upon which the extent of R&D misallocation is contingent. Since the sole source of such dispersion in our model derives from size differences across firms, we replicate both (1) the empirical relationship between firm-level markups and market shares and (2) the extent of firm size heterogeneity in data.

The second objective of this exercise is to discipline the model’s distinct sources of productivity growth. To do so, we harness the panel dimension of our data to ensure that the model replicates the growth trajectories of continuing firms. When compared with the growth rate of the aggregate economy, this moment informs the contribution of incumbent innovation to the latter. Our model further replicates the exit rate and relative size of new entrants, thereby capturing the role of selective churn from entry and exit in propelling economic growth.

In our counterfactual analysis, we consider the implementation of size-dependent value-added subsidies to firms inducing each to operate at its efficient scale. Notably, this intervention does not decentralize the optimal allocation of the model, which features other inefficiencies such as technology spillovers. These externalities are not inherently tied to markups nor the main focus of this paper. Neither does this intervention obviate the role of markups in incentivizing R&D, as rents are required to recoup the sunk cost of an investment therein.⁷ Instead, the subsidy scheme we explore transfers the entire consumer surplus to firms, thus correcting for the *distortions* induced by markups.

We find that this intervention achieves a 1.2 percentage points increase in the long-run growth rate of total-factor productivity (TFP). This faster growth is traced back to three factors: (1) an increase in aggregate R&D employment, (2) its reallocation across firms and (3) a higher rate of entry and exit. As the policy takes hold, firms scale up operations, further invest in R&D and achieve faster productivity growth. This dynamic is most pronounced for the largest, most productive firms who outpace their competitors as R&D employment is reallocated their way. In contrast, the smallest, least productive firms grapple with dwindling R&D resources, trail behind and endogenously exit at a higher rate. As these unsuccessful firms are replaced by more productive new entrants, the productivity pool undergoes more frequent improvements, bolstering TFP growth.

To gain insight into the catalysts behind this growth acceleration, we conduct two alternative exercises. We first consider a “constrained” intervention to assess the role of

⁶This dataset excludes the financial and farming sectors.

⁷In fact, [Schumpeter \(1934\)](#) prominently posited that private incentives for such investments were fated to “perish in the vortex of competition”.

a greater aggregate allocation of labor to innovation relative to its reallocation across firms. Here, firm-level subsidies are upheld, but a uniform tax is concomitantly levied on their R&D expenditures, as to leave the aggregate allocation of labor to innovation at a level commensurate with the pre-policy equilibrium. This intervention achieves an increase in productivity growth of almost 75% of the increment recorded under its “unconstrained” counterpart. This reveals that a more efficient, rather than expanded, allocation of labor to R&D is most prominently hastening economic growth.

Finally, we explore more flexible tax and subsidy schemes to disentangle the role of the aggregate markup from that of markup dispersion. A uniform subsidy that rectifies the level of markups, while leaving their dispersion unchanged, *reduces* the long-run growth rate of TFP by a muted 4 basis points. On one hand, as firms expand in scale and distribute their R&D costs over more units sold, the return on those investments rises. On the other, these firms demand more production labor and inadvertently bid up the cost of R&D through a higher wage. These competing forces stand in contrast with the findings of [Edmond et al. \(2023\)](#): while the level of markups significantly restricts the scale of the economy, it has nearly no bearing on the rate at which it grows.

Conversely, a size-dependent tax and subsidy scheme that addressed the dispersion in markups, while holding their average level fixed, increases long-run productivity growth by 1.3 percentage points, a slight uptick from the baseline subsidy scheme.⁸ The resulting reallocation of R&D investments from small, unproductive firms towards their larger and further expanding competitors ensures that (1) marginal cost reductions are rolled out over more units sold, and (2) inefficient firms are prevented from “polluting” the pool of productivity by lingering excessively. As these firms exit, they pave the way for new entrants to replicate the existing technologies of more productive incumbents.

The rest of the paper is outlined as follows. In the remainder of this section, we discuss the relevant literature. Section 2 provides partial equilibrium intuition on the distortionary consequences of markups. Section 3 presents our general equilibrium theoretical framework. Section 4 describes the quantification of our theory. Section 5 presents the results of our counterfactual analysis and Section 6 concludes.

Related Literature

Our paper is primarily related to a longstanding literature on the macroeconomic costs of product market power. Classic analyses can be traced back to [Smith \(1776\)](#), [Lerner \(1934\)](#), [Harberger \(1954\)](#) and [Dixit and Stiglitz \(1977\)](#). Quantitatively, [Baqaee and Farhi](#)

⁸This larger increase in productivity growth is, however, not necessarily indicative of an improvement in welfare. A related discussion is provided in Section 5.4.

(2019), Bilbiie et al. (2019), Behrens et al. (2020), Edmond et al. (2023) and Afrouzi et al. (2023) are recent examples of papers concluding that the aggregate efficiency losses from markups can be large. The mechanisms stressed in this literature revolve around markups restricting the economy’s scale of production, distorting the allocation of factors of production across firms and inducing inefficient entry (Dhingra and Morrow, 2019).

Yet, by restricting the scale at which firms operate, the markups they command further distort their incentives to invest in R&D. This is akin to a firm-specific “market size” effect. The notion that incentives for R&D are dictated by the extent of the market is extensively discussed in Schmookler (1966) and succinctly captured by a quote from Matthew Boulton, a mechanical engineer, and business partner of James Watt:

“It would not be worth my while to make [steam engines] for three countries only; but I find it very well worth my while to make [them] for all the world.”
– Matthew Boulton (Scherer, 1965)

This distortion is explored in the analyses of Arrow (1962) and Dasgupta and Stiglitz (1980), and appears in the lab equipment model of Rivera-Batiz and Romer (1991).⁹ Our contribution is to quantify its consequences for economic growth in a framework that features heterogeneity in markups. This is motivated by the mounting evidence of substantial markup dispersion across firms (De Loecker and Warzynski, 2012; Amiti et al., 2014; De Loecker et al., 2016; Amiti et al., 2019; De Loecker et al., 2020; De Ridder et al., 2023). To tractably, yet endogenously replicate such dispersion, our model features a non-isoelastic demand system à la Kimball (1995), with the functional form proposed by Klenow and Willis (2016).

Under such a demand system, markup dispersion derives from price differences across firms, which, in our model, reflect heterogeneity in their process efficiency. To endogenously deliver such heterogeneity, our model borrows from the firm dynamics literature and extends the frameworks of Hopenhayn (1992) and Luttmer (2007) to allow for forward-looking and risky investments in R&D. This is in line with the findings of Foster, Haltiwanger and Krizan (2001) and Garcia-Macia, Hsieh and Klenow (2019) who infer large contributions to productivity growth from entry and exit as well as incumbent R&D on existing products. Following Ericson and Pakes (1995), Benhabib, Perla and Tonetti (2021) and Lashkari (2023), the firm’s investment choice is formulated as a stochastic optimal control problem, while its endogenous exit decision takes the form of an optimal stopping time problem.

⁹Unlike the canonical Romer (1990) model, the lab equipment model features no technology spillovers. Nevertheless, economic growth is inefficiently low in the decentralized equilibrium as private incentives to create new intermediate varieties are weakened by final good producers who substitute marked-up intermediates for labor.

The counterfactual analysis we conduct in this framework is conceptually different from those considered in Peters (2020), Cavenaile, Celik and Tian (2021) and Voronina (2021), which are otherwise closely related to our paper. The former two propose theories of Schumpeterian growth in which heterogeneous markups arise endogenously as the outcome of firms' investments in R&D. Peters (2020) quantifies the aggregate static efficiency losses from markups whereas our focus is on their dynamic consequences for long-run economic growth. Cavenaile et al. (2021) study an economy whose structural parameters change over time to replicate the observed trend in markups, and quantify the extent to which the resulting static efficiency costs are mitigated or amplified by the endogenous response of firms' investments in R&D. We adopt a different approach and instead quantify the dynamic costs of markups within a fixed economic environment.

Our counterfactual analysis is closest to Voronina (2021) who puts forth a theory of firm-driven endogenous growth in which markups are heterogeneous but *exogenous*. Through the lens of this model, she quantifies the improvement in welfare that a social planner can achieve by choosing flexible transfers to firms. However, the planner can design such flexible transfers to fix other market failures (e.g. technology spillovers) and in that sense, this counterfactual exercise does not *isolate* the costs of markups. In contrast, the subsidy schedule we consider is constrained to a structure that induces all firms to price at marginal cost, thus directly and strictly addressing the distortions caused by markups.

The misallocation of R&D investments featured in our theory relates our work to a literature focused on this market failure, including Akcigit, Celik and Greenwood (2016), Acemoglu, Akcigit, Alp, Bloom and Kerr (2018), Akcigit, Hanley and Serrano-Velarde (2020), Liu and Ma (2021), Hopenhayn and Squintani (2021), Chen, Liu, Suárez Serrato and Xu (2021), Akcigit, Hanley and Stantcheva (2022), König, Storesletten, Song and Zilibotti (2022), De Ridder (2023), Ayerst (2023) and Lehr (2023). We contribute to this literature by introducing and quantifying a novel source of such misallocation, which results from the dispersion of markups across firms within an industry.

Acemoglu (2023) and Aghion, Bergeaud, Boppart, Klenow and Li (2023) are closely related to our study, as they investigate the consequences of markups on the allocation of R&D resources. Acemoglu (2023) studies their allocation across sectors (rather than firms) in a setting with heterogeneous (yet exogenous) markups. Aghion et al. (2023) distinguish between “good” and “bad” markups. The former reflects a firm's quality advantage over its competitors, which confers positive technology spillovers onto other firms, while the latter reflects its higher process efficiency, with no associated spillovers. Hence, they find that the allocation of R&D resources is inefficiently distorted away from high markup firms in the former but not the latter case.

2 Partial Equilibrium Intuition

To form intuition on how markups distort private incentives for productivity-enhancing investments, we consider a simple two-period partial equilibrium model. In this setting, either a profit-maximizing monopolist or a welfare-maximizing agent operates a firm and allocates resources to achieve an endogenously chosen reduction in its marginal cost (Arrow, 1962; Dasgupta and Stiglitz, 1980; Tirole, 1988; Garella, 2012).

The setup is as follows. In both periods, a household inelastically supplies a factor whose price is exogenous and normalized to unity. This household has preferences over the consumption of a commodity whose price is denoted by p . Assume that these preferences imply a twice differentiable demand function $y(p)$ that satisfies:

$$\frac{\partial y(p)}{\partial p} < 0, \quad \vartheta(p) \equiv -\frac{\partial \ln(y(p))}{\partial \ln(p)} > 1 \quad \text{and} \quad \varepsilon(p) \equiv \frac{\partial \ln(\vartheta(p))}{\partial \ln(p)} \in \mathbb{R}$$

where $\vartheta(p)$ denotes the price elasticity of demand at price p , and $\varepsilon(p)$ denotes the “super-elasticity” of demand at that price. In the post-period, the commodity is produced by a firm using the factor supplied by the household according to a technology with constant returns to scale described by the marginal cost function:

$$c(z') = \exp(-z').$$

Here, z' denotes the firm’s productivity in the post-period, which can be controlled by the agent operating the firm in the pre-period. More specifically, $i(g)$ units of the factor can be invested in the pre-period to achieve a $g\%$ improvement in the firm’s post-period process efficiency:

$$z' = g + z$$

where z denotes the firm’s pre-period process efficiency. The twice differentiable function i is assumed to be strictly increasing and strictly convex and satisfies $i(0) = 0$ and $\lim_{g \rightarrow \infty} i(g) = \infty$.

In this environment, we now compare the decision problems of a profit-maximizing monopolist and a welfare-maximizing agent operating the firm. In the post-period, both agents face a *static* problem. Taking as given the household’s demand function, they must choose a unit price at which to sell the commodity to respectively maximize profits (producer surplus) or social surplus (the sum of producer and consumer surplus). The

two objectives are respectively denoted by $\pi(z', p)$ and $S(z', p)$:¹⁰

$$\pi(z', p) \equiv [p - \exp(-z')]y(p) \quad \text{and} \quad S(z', p) \equiv \pi(z', p) + \int_p^{\bar{p}} y(p')dp'.$$

Assuming demand is positive at optimally chosen prices, it is straightforward to show that maximized producer and social surpluses are given by:

$$\pi(z') \equiv \frac{p(z')y(p(z'))}{\vartheta(p(z'))} \quad \text{and} \quad S(z') \equiv \int_{c(z')}^{\bar{p}} y(p)dp$$

where $p(z')$ denotes the usual profit-maximizing price implicitly defined as:

$$p(z') \equiv \frac{\vartheta(p(z'))}{\vartheta(p(z')) - 1} \times c(z').$$

Let us now consider the agents' *dynamic* problem. In the pre-period, they must choose a factor allocation to investments in R&D to respectively maximize post-period producer or social surplus. Up to a first-order approximation of these objectives and assuming no time discounting, these dynamic problems are described by:

$$\max_g \{ \pi(z) + \pi'(z)g - i(g) \} \quad \text{and} \quad \max_g \{ S(z) + S'(z)g - i(g) \}$$

where $\pi'(z)$ and $S'(z)$ denote the partial derivatives of producer and social surplus with respect to the firm's initial productivity. The first-order conditions of each problem are:

$$\pi'(z) = i'(g) \quad \text{and} \quad S'(z) = i'(g).$$

Therefore, private and social incentives for marginal cost reductions may not coincide if *marginal* producer and social surpluses differ. The proposition that follows characterizes the ratio $R(z) \equiv \pi'(z)/S'(z)$ of these objects.

Proposition 1. *The ratio $R(z)$ of marginal producer to social surplus from an infinitesimal reduction in marginal cost is characterized by:*

$$R(z) \equiv \frac{\pi'(z)}{S'(z)} = \frac{y(p(z))}{y(\exp(-z))} < 1.$$

When the welfare-maximizing agent is instead constrained to produce at the same scale as the monopolist, the ratio $R^c(z)$ of marginal producer surplus to “constrained” marginal social

¹⁰Here, \bar{p} denotes the choke price.

surplus is characterized by:

$$R^c(z) \equiv \frac{\pi'(z)}{\pi'(z) + C'(z)} = \frac{\vartheta(p(z)) + \varepsilon(p(z)) - 1}{2\vartheta(p(z)) + \varepsilon(p(z)) - 1} < 1$$

where $C(z) \equiv \int_{p(z)}^{\bar{p}} y(p)dp$ denotes consumer surplus.

Proposition 1 shows that in this setting, the welfare-maximizing agent always faces stronger incentives to achieve a marginal cost reduction than the monopolist, even when the two are constrained to operate at the same scale. The intuition behind this proposition is twofold. First, since the welfare-maximizing agent optimally operates at a larger scale than the monopolist, the same reduction in marginal cost applies to more units produced, thus begetting larger total cost savings. This is evident from the expression of the ratio $R(z)$, which is simply equal to the ratio of the quantities produced by the monopolist and the welfare-maximizing agent.

Second, at a given scale of operation, the welfare-maximizing agent internalizes that a reduction in marginal cost may achieve additional consumer surplus whereas the monopolist does not. This is illustrated in the second part of the proposition where it is clear that the monopolist only “appropriates” a fraction of the marginal surplus achieved by the improvement in process efficiency. We can now take this first proposition further to characterize how the distance between private and social incentives for such investments depends on the firm’s initial productivity.

Proposition 2. *The elasticity of the ratio $R(z)$ with respect to the firm’s initial productivity is characterized by:*

$$\frac{\partial \ln(R(z))}{\partial z} = \vartheta(p(z))\varrho(z) - \vartheta(\exp(-z)) \quad \text{where} \quad \varrho(z) \equiv -\frac{\partial \ln(p(z))}{\partial z}.$$

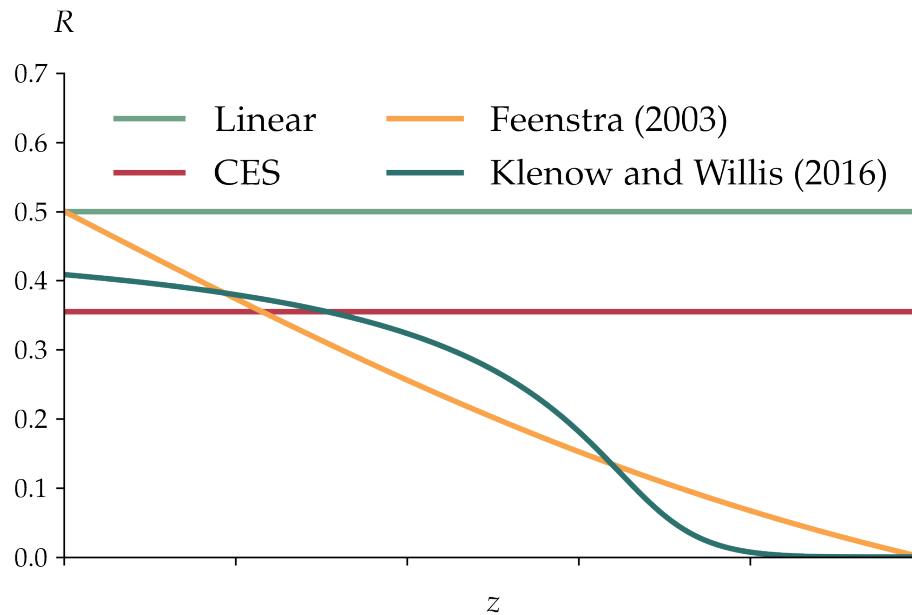
Here, $\varrho(z)$ denotes the monopolist’s productivity “pass-through” (i.e. the percent change in the monopolist’s price following a one percent improvement in its productivity) which is a function of the price elasticity and super-elasticity of demand:

$$\varrho(z) = \frac{\vartheta(p(z)) - 1}{\vartheta(p(z)) + \varepsilon(p(z)) - 1}.$$

Proposition 2 shows that the distance between private and social incentives depends on the price elasticity and super-elasticity of demand. Therefore, when the price elasticity of demand is not constant, a market equilibrium may not only suffer from an inefficient allocation of resources to productivity improvements in the aggregate but also from a misallocation of those resources across firms that differ in their process efficiency.

To illustrate these two propositions, Figure 1 plots the ratio $R(z)$ implied by four common demand functions: the isoelastic (CES) demand function, the linear demand function, the Klenow and Willis (2016) specification of the Kimball (1995) demand function and the Translog demand function proposed by Feenstra (2003). Under the CES demand function, social incentives for marginal cost reductions exceed private incentives by the same proportion, regardless of the firm's productivity. In contrast, for the non-isoelastic Kimball and Translog demand functions, the distance between social and private incentives increases with firm productivity. However, this property does not generically hold for all non-isoelastic demand functions, as the linear demand function features a constant ratio $R(z) = 1/2$.

Figure 1: Private vs. Social Incentives for Productivity Improvements



Note: The vertical axis measures the ratio $R(z)$ of marginal producer to social surplus from an infinitesimal reduction in marginal cost.

Since this section presented a stylized partial equilibrium environment, we have so far abstracted from the possibility that the private and social costs of investments in R&D may not coincide. This motivates the section that follows, which presents a more involved theory where market prices may fail to reflect the social value of resources. As discussed in Section 1, this possibility must be crucially accounted for when quantifying the dynamic consequences of markups in general equilibrium.

3 Theory

In this section, we propose a general equilibrium theory of endogenous economic growth that builds on the partial equilibrium intuition presented in the previous section. We extend the model of [Luttmer \(2007\)](#) to allow for endogenous productivity improvements by incumbent firms who face non-isoelastic demand curves.

3.1 Economic Environment

Preferences

Consider an economy populated by an infinitely-lived representative household of unit measure with separable preferences over consumption C_t and hours worked H_t such that lifetime utility is defined as:

$$U_0 = \int_0^\infty e^{-\rho t} [\ln(C_t) - v(H_t)] dt. \quad (1)$$

Here, $\rho > 0$ is the household's rate of time preference, the function v is strictly increasing and strictly convex, and time is continuous and indexed by $t \in \mathbb{R}_0^+$.

Technology

The economy is composed of two sectors: the final and intermediate sectors. The final sector produces a final good using a continuum of differentiated varieties indexed by j from the intermediate sector. The final sector's production technology has constant returns to scale and is defined implicitly by the following [Kimball \(1995\)](#) aggregator:

$$\int_{j \in \mathcal{J}_t} \Upsilon(\hat{y}_{jt}) dj = \kappa \quad \text{where} \quad \hat{y}_{jt} \equiv \frac{y_{jt}}{Y_t}. \quad (2)$$

Here, Y_t denotes aggregate output, y_{jt} is the quantity of variety j used in production and $\kappa > 0$ is a constant. The function Υ is strictly increasing, strictly concave and satisfies $\Upsilon(1) = 1$. In what will follow, we denote the measure of varieties at time t by $M_t \equiv |\mathcal{J}_t|$. This production function belongs to the family of homothetic aggregators with direct implicit additivity (HDIA) as defined in [Matsuyama and Ushchev \(2017\)](#). In particular, it nests the [Dixit and Stiglitz \(1977\)](#) aggregator when $\Upsilon(\hat{y}) = \hat{y}^{\frac{\theta-1}{\theta}}$, where $\theta > 1$ would denote the constant elasticity of substitution across varieties.

Each variety is produced by a single firm from the intermediate sector using physical

capital and production labor with Hicks-neutral productivity z_{jt} according to a Cobb-Douglas production technology:

$$y_{jt} = \exp(z_{jt})k_{jt}^\alpha l_{jt}^{1-\alpha}. \quad (3)$$

Here, k_{jt} and l_{jt} respectively denote the quantities of capital and labor used in production and $\alpha \in [0, 1]$ denotes the output elasticity of capital. As in [Hopenhayn \(1992\)](#) and [Luttmer \(2007\)](#), firms must pay an overhead of $c_O > 0$ units of labor per unit of time to remain active. If this cost is unpaid, a firm must irreversibly exit. Firms may also exit exogenously at Poisson rate $\chi > 0$.

At any point in time, a firm is fully described by its productivity $z_t \in \mathbb{R}$ such that, from this point on, we abandon the j -index notation. Over time, firms can improve their process efficiency by allocating labor to R&D. More precisely, productivity evolves according to a controlled diffusion process of the form:

$$dz_t = \gamma_t dt + \sigma dB_t$$

where $\gamma_t > 0$ is the controlled drift, dB_t is the standard normal increment of a Brownian motion and $\sigma > 0$ is its standard deviation.¹¹ Defining a firm's productivity relative to the least productive firm in the economy as $\hat{z}_t \equiv z_t - \underline{z}_t$, we obtain the following law of motion by Itô's lemma:

$$d\hat{z}_t = (\gamma_t - g_t)dt + \sigma dB_t. \quad (4)$$

Here, \underline{z}_t and g_t respectively denote the productivity lower bound and its instantaneous rate of change. The labor requirement to achieve a drift of γ for a firm with relative productivity \hat{z} is $i(\gamma, \hat{z}) : \mathbb{R}_0^+ \times \mathbb{R}_0^+ \rightarrow \mathbb{R}_0^+$, where the function i is strictly increasing and convex in its first argument and satisfies $i(0, \hat{z}) = 0$ and $\gamma(\hat{z}) < \infty$ for all $\hat{z} \in [0, \infty)$.

In every period, a measure of potential entrants can allocate $c_E > 0$ units of labor to achieve a unit flow of entry and start producing with a relative productivity draw from the cumulative density function (CDF) $F_t^E(\hat{z})$. This function is a transformation of the relative productivity CDF of incumbent firms $F_t(\hat{z})$, as defined by the non-decreasing function $T : [0, 1] \rightarrow [0, 1]$ such that $T(0) = 0$ and $T(1) = 1$:¹²

$$F_t^E(\hat{z}) = 1 - T[1 - F_t(\hat{z})]. \quad (5)$$

In particular, it is assumed that this transformation is such that the right tail of the

¹¹Productivity shocks are independent and identically distributed across firms.

¹²This type of transformation is often referred to as a dual distortion function.

relative productivity distribution decays faster for entrants than incumbents:

$$\lim_{\hat{z} \rightarrow \infty} \frac{1 - F_t^E(\hat{z})}{1 - F_t(\hat{z})} = 0.$$

As discussed in Section 3.3, this condition is imposed to achieve a unique stationary distribution of relative productivity on a balanced growth path (e.g. a simple power function for T satisfies this condition if the exponent greater than one).

Resource Constraints

The final good can either be consumed or invested in physical capital, which delivers the following resource constraint:

$$\dot{K}_t + \delta K_t + C_t \leq Y_t \quad \text{where} \quad K_t \equiv M_t \int_0^\infty k_t(\hat{z}) dF_t(\hat{z}) \quad (6)$$

and where $\delta > 0$ is the rate at which capital depreciates. The labor supplied by the household can be allocated to either production, innovation, entry or overhead, which delivers the following resource constraint:

$$L_t + I_t + c_E E_t + c_O M_t \leq H_t. \quad (7)$$

Here, E_t denotes the aggregate flow of entry and the aggregate allocations of labor to production and innovation are defined as:

$$L_t \equiv M_t \int_0^\infty l_t(\hat{z}) dF_t(\hat{z}) \quad \text{and} \quad I_t \equiv M_t \int_0^\infty i(\gamma_t(\hat{z}), \hat{z}) dF_t(\hat{z}).$$

Laws of Motion

The Kolmogorov forward equation (KFE) describing the evolution of $F_t(\hat{z})$ over time is:

$$\dot{F}_t(\hat{z}) = \mathcal{A}_t F_t(\hat{z}) - (\sigma^2/2) F_t''(0) + e_t F_t^E(\hat{z}) \quad \forall \hat{z} > 0. \quad (8)$$

Here, $e_t \equiv E_t/M_t$ denotes the entry rate and the operator \mathcal{A}_t is defined as:

$$\mathcal{A}_t \equiv -[\gamma_t(\hat{z}) - g_t] \partial_{\hat{z}} + (\sigma^2/2) \partial_{\hat{z}\hat{z}} - \chi - \dot{M}_t/M_t$$

where $\partial_{\hat{z}}$ and $\partial_{\hat{z}\hat{z}}$ denote the first and second partial derivative operators with respect to \hat{z} . The measure of varieties then follows the law of motion:

$$\dot{M}_t = [e_t - \chi - (\sigma^2/2)F_t''(0)]M_t. \quad (9)$$

The economic environment is summarized in Table 1.

Table 1: The economic environment

(1)	$U_0 = \int_0^\infty e^{-\rho t} [\ln(C_t) - v(H_t)] dt$	Preferences
(2)	$M_t \int_0^\infty \Upsilon(\hat{y}_t(\hat{z})) dF_t(\hat{z}) = \kappa$	Final good production technology
(3)	$y_t(\hat{z}) = \exp(\hat{z} + \underline{z}_t) k_t(\hat{z})^\alpha l_t(\hat{z})^{1-\alpha}$	Variety production technology
(4)	$d\hat{z}_t = (\gamma_t - g_t)dt + \sigma dB_t$	Innovation technology
(5)	$F_t^E(\hat{z}) = 1 - T[1 - F_t(\hat{z})]$	Entrants' productivity distribution
(6)	$\dot{K}_t + \delta K_t + C_t \leq Y_t$	Final good resource constraint
(7)	$L_t + I_t + c_E E_t + c_O M_t \leq H_t$	Labor resource constraint
(8)	$\dot{F}_t(\hat{z}) = \mathcal{A}_t F_t(\hat{z}) - (\sigma^2/2)F_t''(0) + e_t F_t^E(\hat{z})$	Incumbents' productivity distribution
(9)	$\dot{M}_t = [e_t - \chi - (\sigma^2/2)F_t''(0)]M_t$	Measure of varieties

3.2 Decision Problems

We now define the decision problems of economic agents, which determine equilibrium prices and quantities on the final good, varieties, labor and asset markets. In terms of market structure, it is assumed that all agents partake in perfect competition in all markets besides intermediate firms who engage in monopolistic competition and choose the price at which to sell their variety.

The Household's Problem

Taking prices as given, the household's problem is to choose its consumption and hours worked to maximize lifetime utility subject to a flow budget constraint:

$$\max_{\{C_t, H_t\}_{t \geq 0}} \int_0^\infty e^{-\rho t} [\ln(C_t) - v(H_t)] dt \quad \text{s.t.} \quad \dot{A}_t = r_t A_t + w_t H_t - C_t$$

where w_t denotes the wage rate, A_t is the value of physical capital and corporate assets, and r_t is the rate of return on those assets:

$$A_t = K_t + M_t \int_0^\infty V_t(\hat{z}) dF_t(\hat{z}) \quad \text{where} \quad \lim_{t \rightarrow \infty} e^{-\int_0^t r_{t'} dt'} A_t = 0.$$

Here, $V_t(\hat{z})$ denotes the value of a firm with relative productivity \hat{z} which is yet to be defined. The household's problem thus delivers the usual intertemporal Euler equation and static first-order condition:

$$\frac{\dot{C}_t}{C_t} = r_t - \rho \quad \text{and} \quad v'(H_t)C_t = w_t.$$

The Final Sector's Problem

Taking prices as given, the final sector's problem is to choose its relative demand for each variety to maximize profits in each period:

$$\max_{\{\hat{y}_t(\hat{z})\}_{\hat{z}=0}^\infty} \left\{ P_t - M_t \int_0^\infty p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}) \right\} Y_t \quad \text{s.t.} \quad M_t \int_0^\infty \Upsilon(\hat{y}_t(\hat{z})) dF_t(\hat{z}) = \kappa$$

where P_t and $p_t(\hat{z})$ respectively denote the price of the final good and the price charged by a firm with relative productivity \hat{z} . Therefore, this problem delivers the following inverse demand functions:

$$p_t(\hat{z}) = \Upsilon'(\hat{y}_t(\hat{z})) P_t D_t \quad \text{where} \quad P_t \equiv M_t \int_0^\infty p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}).$$

Here, the final good is chosen as the numéraire such that $P_t = 1$ for all t and D_t is a demand index defined as:

$$D_t \equiv \left(M_t \int_0^\infty \Upsilon'(\hat{y}_t(\hat{z})) \hat{y}_t(\hat{z}) dF_t(\hat{z}) \right)^{-1}.$$

The Firm's Static Problem

Firms engage in monopolistic competition in the product market but perfect competition in the input markets. That is, a firm chooses the price at which to sell its variety as well as its demand for physical capital and production labor to maximize profits in each period. The firm takes as given the demand for its variety, the rental rate of capital r_t and the wage rate w_t , which delivers the following problem:

$$\begin{aligned}\pi_t(\hat{z}) &= \max_{p_t(\hat{z}), k_t(\hat{z}), l_t(\hat{z})} \{p_t(\hat{z})y_t(\hat{z}) - (r_t + \delta)k_t(\hat{z}) - w_t l_t(\hat{z})\} - w_t c_O \\ \text{s.t. } p_t(\hat{z}) &= \Upsilon'(\hat{y}_t(\hat{z}))D_t.\end{aligned}$$

The firm's optimal choices of physical capital and production labor imply that we can rewrite the problem as:

$$\pi_t(\hat{z}) = \max_{p_t(\hat{z})} \{[p_t(\hat{z}) - \varsigma_t \exp(-\hat{z} - \underline{z}_t)]\hat{y}_t(\hat{z})\}Y_t - w_t c_O$$

where ς_t denotes the producer price index of inputs:

$$\varsigma_t \equiv \left(\frac{r_t + \delta}{\alpha}\right)^\alpha \left(\frac{w_t}{1 - \alpha}\right)^{1 - \alpha}.$$

This in turn implies that the firm sets its price to a markup $\mu_t(\hat{z})$ above marginal cost:

$$p_t(\hat{z}) = \mu_t(\hat{z}) \times \frac{\varsigma_t}{\exp(\hat{z} + \underline{z}_t)} \quad \text{where} \quad \mu_t(\hat{z}) \equiv \frac{\vartheta_t(\hat{z})}{\vartheta_t(\hat{z}) - 1}$$

and where $\vartheta_t(\hat{z})$ denotes the price elasticity of demand:

$$\vartheta_t(\hat{z}) \equiv -\frac{\Upsilon'(\hat{y}_t(\hat{z}))}{\Upsilon''(\hat{y}_t(\hat{z}))\hat{y}_t(\hat{z})} \in (1, \infty).$$

Since the markup function is monotonic in relative demand, which is itself monotonic in the firm's price, there exists a unique solution for the latter as an implicit function of productivity and calendar time. As such, firm-level profits can be expressed as:

$$\pi_t(\hat{z}) = \frac{p_t(\hat{z})\hat{y}_t(\hat{z})Y_t}{\vartheta_t(\hat{z})} - w_t c_O.$$

Note that active firms whose productivity is too low to set a price below the choke price $\bar{p}_t \equiv \Upsilon'(0)D_t$ face no demand and therefore make negative profits.

The Firm's Dynamic Problem

Given the above static profit function and taking the wage rate as given, firms control the drift of their productivity and choose an optimal exit time τ at which to shut down operations:

$$V_t(\hat{z}) = \max_{\tau, \{\gamma_s\}_{s \geq t}} \mathbb{E}_{\hat{z}} \left\{ \int_t^{t+\tau} e^{-\int_t^s (r_{t'} + \chi) dt'} [\pi_s(\hat{z}_s) - w_s i(\gamma_s, \hat{z}_s)] ds \right\}$$

where $\mathbb{E}_{\hat{z}}$ denotes the expectation operator with respect to the diffusion process $\{\hat{z}_s\}_{s \geq t}$ when its initial value is $\hat{z}_t = \hat{z}$. Within the continuation region of productivity (i.e. where it is not optimal to exit), the firm's value function satisfies the standard Hamilton-Jacobi-Bellman equation (HJBE):

$$(r_t + \chi)V_t(\hat{z}) = \pi_t(\hat{z}) + \max_{\gamma} \{(\gamma - g_t)V'_t(\hat{z}) - w_t i(\gamma, \hat{z})\} + \sigma^2 V''_t(\hat{z})/2 + \dot{V}_t(\hat{z})$$

with value matching, smooth pasting and first-order conditions:

$$V_t(0) = V'_t(0) = 0 \quad \text{and} \quad V'_t(\hat{z}) = w_t \times \frac{\partial i(\gamma, \hat{z})}{\partial \gamma}.$$

The Entrant's Problem

Entrants engage in perfect competition on the labor market, and therefore choose a flow of entry to maximize future expected profits while taking the wage rate as given:

$$V_t^E = \max_{E_t} \left\{ E_t \int_0^\infty V_t(\hat{z}) dF_t^E(\hat{z}) - w_t c_E E_t \right\}.$$

The first-order condition of the entrant's problem delivers what will be referred to as the free-entry condition, which is here written in complementary-slackness form:

$$\left(\int_0^\infty V_t(\hat{z}) dF_t^E(\hat{z}) - w_t c_E \right) E_t = 0.$$

The derivations of the optimality conditions are presented in [Appendix A.1](#).

3.3 Equilibrium Allocation

Now that all decision problems have been described, we can define the concept of an equilibrium allocation.

Definition 1. Given initial conditions $\{\underline{z}_0, K_0, F_0(\hat{z}), M_0\}$, an equilibrium allocation consists of time paths for quantities, prices and policy functions such that the following conditions hold:

1. $\{C_t, H_t\}_{t \geq 0}$ solve the household's problem.
2. $\{\hat{y}_t(\hat{z})\}_{t \geq 0}$ solve the final sector's problem.
3. $\{p_t(\hat{z}), k_t(\hat{z}), l_t(\hat{z})\}_{t \geq 0}$ solve the firm's static problem.
4. $\{\gamma_t(\hat{z}), \underline{z}_t\}_{t \geq 0}$ solve the firm's dynamic problem.
5. $\{E_t\}_{t \geq 0}$ solves the entrant's problem.
6. $\{Y_t\}_{t \geq 0}$ satisfies the **Kimball (1995)** aggregator.
7. $\{p_t(\hat{z})\}_{t \geq 0}$ clear the variety markets.
8. $\{w_t\}_{t \geq 0}$ clears the labor market.
9. $\{r_t\}_{t \geq 0}$ clears the asset market.
10. The capital stock evolves according to equation (6).
11. The cumulative density of firms evolves according to equation (8).
12. The measure of varieties evolves according to equation (9).

Aggregation

Despite its complex structure, our theory admits tractable aggregation of the equilibrium allocation. In particular, aggregate output can be expressed as:

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha} \quad \text{where} \quad Z_t \equiv \left(M_t \int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z} - \underline{z}_t) dF_t(\hat{z}) \right)^{-1}.$$

Here, Z_t denotes the economy's TFP, which is a quantity-weighted harmonic aggregate of firm-level productivity. The aggregate demand for physical capital and production labor is given by:

$$K_t = \frac{\alpha Y_t}{(r_t + \delta) \mathcal{M}_t} \quad \text{and} \quad L_t = \frac{(1 - \alpha) Y_t}{w_t \mathcal{M}_t}$$

where \mathcal{M}_t denotes the cost-weighted average of firm-level markups:

$$\mathcal{M}_t \equiv \frac{\int_0^\infty \mu_t(\hat{z}) \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}. \quad (10)$$

Therefore, we recover the result from [Edmond et al. \(2023\)](#) that the “aggregate” markup reduces the quantity of variable inputs used in production.

Balanced Growth Path

With these aggregation results, we can now define the concept of a balanced growth path equilibrium allocation. The propositions that follow characterize the growth rate of TFP as well as the asymptotic behavior of the relative productivity distribution on this balanced growth path.

Definition 2. *A balanced growth path equilibrium allocation is an equilibrium allocation as defined in Definition 1 such that all quantities, prices and policy functions are either stationary or grow at a constant rate, and the distribution of relative productivity is stationary.*

There exists a continuum of such allocations, indexed by the initial condition for the exit threshold.¹³ Since the economy is growing over time, the distribution of firm-level productivity behaves as a “traveling wave”. Hence, for this distribution to be stationary, it must be normalized by a variable that travels at the same speed on a balanced growth path. Here, we choose this variable to be the endogenous exit threshold such that the scale of the productivity distribution (and of the economy more generally) is “pinned down” by the initial condition for that threshold.¹⁴ On a balanced growth path, the growth rate of TFP is characterized by the following proposition.

Proposition 3. *Letting the price elasticity of demand $\vartheta(\hat{z})$ as well as the firm’s productivity pass-through $q(\hat{z})$ be defined as in Section 2, the stationary growth rate of TFP can be decomposed into the contribution of (1) incumbent firms’ productivity drift, (2) incumbent firms’ productivity*

¹³This continuum of balanced growth path equilibrium allocations is also found in the “AK” model (indexed by the initial condition for the physical capital stock) or the neoclassical growth model (indexed by the initial condition for TFP).

¹⁴This parallels the choices of [Perla, Tonetti and Waugh \(2021\)](#) and [Benhabib et al. \(2021\)](#) who normalize the distribution of productivity by the endogenous adoption threshold in theories of technology adoption.

volatility, (3) endogenous exit and (4) exogenous exit:

$$\begin{aligned}
g = & \frac{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})\gamma(\hat{z})dF(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} && \text{Incumbents' drift} \\
& - \frac{(\sigma^2/2) \int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF'(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} && \text{Incumbents' volatility} \\
& + \frac{(\sigma^2/2)F''(0)[\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF^E(\hat{z}) - \hat{y}(0)]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} && \text{Endogenous exit} \\
& - \frac{\chi[\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z}) - \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF^E(\hat{z})]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} && \text{Exogenous exit.}
\end{aligned}$$

Here, the growth rate of TFP is also denoted by g since it must be equated to the growth rate of the endogenous exit threshold on a balanced growth path.

Despite its complexity, the growth rate derived in Proposition 3 is similar to those obtained in other prominent theories of endogenous growth.¹⁵ To see this, suppose for simplicity that the price elasticity of demand is constant $\vartheta(\hat{z}) = \theta > 1$, the productivity pass-through is constant and complete $\varrho(\hat{z}) = 1$ and the productivity drifts are constant $\gamma(\hat{z}) = \gamma > 0$.¹⁶ Suppose further that entrants draw their productivity as to capture a fraction $s_E < 1$ of the average market share of incumbent firms while endogenously exiting firms' market share is equal to a fraction $s_X < s_E$ of that average. Under those assumptions, the above formula boils down to:

$$g = \gamma + \frac{(\theta - 1)\sigma^2}{2} + \frac{\sigma^2 F''(0)(s_E - s_X)}{2(\theta - 1)} - \frac{\chi(1 - s_E)}{\theta - 1}.$$

The first term reflects the positive contribution of incumbent firms' productivity drift to economic growth. The second term reflects how the volatility of incumbent firms' productivity contributes positively to growth. Indeed, since varieties are substitutes ($\theta > 1$), independent productivity shocks allow the final sector to reallocate expenditures from varieties that receive bad productivity shocks to those that receive good ones. The third and fourth terms reflect the contribution of entry and exit, by which entrants replace two types of firms: (1) the least productive firms who are swept below the endogenous exit threshold at rate $(\sigma^2/2)F''(0)$ and (2) randomly selected firms who exit exogenously at rate χ . Since the measure of varieties is constant on a balanced growth

¹⁵See Lashkari (2023) for a discussion of the growth rate that arises in different theories of endogenous technological change.

¹⁶The assumptions of a constant price elasticity of demand and a complete pass-through are obtained with a CES demand function.

path, entry does not contribute positively to growth through a “love for variety”.¹⁷

The general formula provided in Proposition 3 further accounts for heterogeneity in demand elasticities, pass-throughs and productivity drifts. This heterogeneity matters in that productivity improvements can be translated into *further* economic growth if (1) they are rolled out over more units produced, (2) they are passed on through lower prices and (3) these lower prices are rewarded with additional demand. To see this, notice that the term reflecting the growth contribution of incumbents’ investments in R&D is a weighted average of their productivity drifts:

$$\int_0^\infty \omega(\hat{z})\gamma(\hat{z})d\hat{z} \quad \text{where} \quad \omega(\hat{z}) \equiv \frac{[\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z})\exp(-\hat{z})F'(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z})\exp(-\hat{z})F'(\hat{z})d\hat{z}}.$$

In particular, the term $\hat{y}(\hat{z})\exp(-\hat{z})$ in the “weights” $\omega(\hat{z})$ is proportional to a firm’s total expenditures on inputs, whereas the term $\vartheta(\hat{z})\varrho(\hat{z}) - 1$ denotes the elasticity of those expenditures with respect to the firm’s productivity. Hence, these weights can be thought of as reflecting the extent to which a firm *differentially* expands in scale following an improvement in its process efficiency.

To elucidate this, consider the special case where firms’ pass-throughs are uniformly equal to zero, that is, not a single firm reduces its price following an improvement in their process efficiency. Under this assumption, the term above reduces to a cost-weighted average of firm-level productivity drifts. Even in this case, dispersion in markups can act as a drag on economic growth: since the largest firms under-invest in R&D to the greatest extent, the correlation between firms’ productivity drifts and their cost weights is inefficiently low, which reduces the value of the corresponding weighted average.

Conversely, in a more realistic case where firms’ pass-throughs are not uniformly equal to zero, but instead declining in size as dictated by Marshall (1890)’s third law of demand, that drag on economic growth is amplified. In this case, larger firms not only disproportionately under-invest in R&D but also expand the least in scale following improvements in their process efficiency. This is due to their unwillingness to reduce their price through which they fail to attract more demand.

Intuitively, this implies that the rate of TFP growth will be higher if the firms that achieve the largest drift in productivity are not only larger and more numerous, but also the firms towards which demand is mostly reallocated as a result. The magnitude of this correlation depends on the stationary distribution of relative productivity, whose asymptotic behavior is characterized by the following proposition.

Proposition 4. *On a balanced growth path, if $\chi > 0$, $\lim_{\hat{z} \rightarrow \infty} \gamma(\hat{z}) = \bar{\gamma} < \infty$ and the*

¹⁷This is a result of our assumption of a constant population.

transformation function T satisfies the assumptions described in Section 3, the distribution of relative productivity asymptotes to an exponential distribution as $\hat{z} \rightarrow \infty$:

$$\lim_{\hat{z} \rightarrow \infty} F(\hat{z}) = 1 - \exp(-\lambda \hat{z}) \quad \text{where} \quad \lambda \equiv \frac{g - \bar{\gamma} + \sqrt{(g - \bar{\gamma})^2 + 2\chi\sigma^2}}{\sigma^2}$$

and where g is the stationary growth rate of TFP. If $g > \bar{\gamma} + \sigma^2/2 - \chi$, the stationary distribution of $\exp(\hat{z})$ is Pareto with shape parameter $\lambda > 1$ and has a finite mean.

The rate parameter λ of the exponential tail is inversely related to the dispersion in relative productivity such that more churning from a higher growth rate g or a higher exit rate χ implies a thinner right tail, whereas a higher instantaneous productivity volatility σ implies a fatter right tail. Since TFP growth is endogenous, Proposition 4 illustrates how firm-level heterogeneity determines aggregate economic growth, which in turn determines the extent of this heterogeneity.

3.4 Characterization

For our theory to deliver quantifiable predictions, we impose additional parametric functional form assumptions on preferences and technologies. This subsection describes those choices, which are both standard in the literature and consistent with relevant empirical regularities.

In terms of preferences, we assume that the household has standard [MaCurdy \(1981\)](#) flow disutility from hours worked:

$$v(H) = \beta \times \frac{H^{1+\eta}}{1+\eta}$$

where $\eta > 0$ is the inverse of the Frisch elasticity of labor supply and $\beta > 0$ is the utility weight on hours worked.

The final sector's [Kimball \(1995\)](#) production technology is defined according to the functional form introduced by [Klenow and Willis \(2016\)](#):¹⁸

$$\Upsilon(\hat{y}) = 1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon - 1} \left[\Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) - \Gamma\left(\frac{\theta}{\epsilon}, \frac{\hat{y}^{\epsilon/\theta}}{\epsilon}\right) \right]$$

where $\theta > 1$ and $\epsilon > 0$. Note that as $\epsilon \rightarrow 0$, this functional form converges to the [Dixit and Stiglitz \(1977\)](#) aggregator with constant elasticity of substitution across varieties.

¹⁸Here, $\Gamma(s, x) = \int_x^\infty t^{s-1} e^{-t} dt$ denotes the upper incomplete gamma function.

This functional form is chosen as it is flexible enough to capture two important empirical regularities, which are sometimes referred to as [Marshall \(1890\)](#)’s second and third laws of demand ([Matsuyama and Ushchev, 2022](#)). Namely, that the price elasticity of demand increases in the price charged whereas its rate of change (the “super-elasticity” of demand) decreases therein. These “laws” of demand imply that larger firms command both higher markups and lower pass-throughs, which is empirically documented in [Amiti et al. \(2014\)](#) and [Amiti et al. \(2019\)](#). Specifically, these markups and pass-throughs take the following form as functions of relative demand:

$$\mu(\hat{y}) = \frac{\theta}{\theta - \hat{y}^{\epsilon/\theta}} \quad \text{and} \quad \varrho(\hat{y}) = \frac{\theta - \hat{y}^{\epsilon/\theta}}{\theta + \epsilon - \hat{y}^{\epsilon/\theta}}.$$

As in [Acemoglu et al. \(2018\)](#) and [Akcigit and Kerr \(2018\)](#), the firm’s innovation technology is characterized by an isoelastic cost function:

$$i(\gamma, \hat{z}) = \frac{\exp[c_I + (1 + \zeta)\hat{z}]\gamma^{1+\zeta}}{1 + \zeta}$$

where $c_I > 0$ measures the scale of that cost function and $\zeta > 0$ disciplines its elasticity. In particular, this functional form implies that all firms must allocate the same quantity of labor to achieve a given *absolute* drift of relative productivity. Therefore, it becomes more and more costly to achieve a *proportional* drift as a firm becomes more productive.

Finally, we follow [Benhabib et al. \(2021\)](#) and choose the transformation function $T(x) = x^{\tilde{\zeta}}$ where $\tilde{\zeta} > 1$ such that:

$$F_t^E(\hat{z}) = 1 - [1 - F_t(\hat{z})]^{\tilde{\zeta}}.$$

This functional form is both parsimonious and satisfies the required assumptions to achieve a stationary distribution of relative productivity on a balanced growth path. Specifically, for $\tilde{\zeta} > 1$, entrants start producing with lower relative productivity than incumbents on average. The functional form assumptions are summarized in [Table 2](#).

4 Quantification

In this section, we present the estimation of our theory’s structural parameters, which is performed via a GMM strategy, targeting aggregate and firm-level moments from France. We first describe the data from which these moments are calculated, after which we discuss the identification of each parameter.

Table 2: Functional forms

Function	Source
$v(H) = \beta \times \frac{H^{1+\eta}}{1+\eta}$	MaCurdy (1981)
$\Upsilon(\hat{y}) = 1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon-1} \left[\Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) - \Gamma\left(\frac{\theta}{\epsilon}, \frac{\hat{y}^{\epsilon/\theta}}{\epsilon}\right) \right]$	Klenow and Willis (2016)
$i(\gamma, \hat{z}) = \exp[c_I + (1 + \zeta)\hat{z}] \gamma^{1+\zeta} / (1 + \zeta)$	Acemoglu et al. (2018)
$F_t^E(\hat{z}) = 1 - [1 - F_t(\hat{z})]^\zeta$	Benhabib et al. (2021)

4.1 Data

Our main source of data is the *Fichier Approché des Résultats d'Esane* (FARE) which is an annual panel dataset with the balance sheet and income statements of all firms in France that are subject to the standard corporate tax (excluding the financial and farming sectors). Our sample consists of 5.4 million (firm-year) observations between 2009 and 2019, with around 830 thousand unique firms overall and 460 thousand firms each year.¹⁹ The main variables of interest are the firm's main industry of operation, value-added, wage bill, and its stock of capital.

From this data, we measure the markup of firm j from industry i in year t as:²⁰

$$\mu_{jit} = \frac{p_{jit} y_{jit}}{[(r_t + \delta) k_{jit}]^{\alpha_{it}} (w_t l_{jit})^{1-\alpha_{it}}}$$

where $p_{jit} y_{jit}$ is its value-added, k_{jit} is its stock of capital (in current value), $w_t l_{jit}$ is its total expenditures on labor and α_{it} is the output elasticity of physical capital, which we assume is common to all firms in the same 2-digit NACE industry.²¹ Given the Cobb-Douglas production function, we calculate this elasticity as the cost-weighted

¹⁹Refer to Appendix C for details on criteria we set for inclusion of an observation in our sample.

²⁰This uses Euler's theorem and abstracts from the proportionality term $\alpha^\alpha (1 - \alpha)^{1-\alpha}$. We verified that including that term affects the level of the implied markups, but not their correlation with market shares.

²¹See Appendix C for further details on the construction of each variable used in the measurement of firm-level markups.

average of each firm's capital cost share in industry i and year t :²²

$$\alpha_{it} = \sum_{j \in i} \frac{\omega_{jit}(r_t + \delta)k_{jit}}{[(r_t + \delta)k_{jit} + w_t l_{jit}]} \quad \text{where} \quad \omega_{jit} \equiv \frac{(r_t + \delta)k_{jit} + w_t l_{jit}}{\sum_{j \in i} [(r_t + \delta)k_{jit} + w_t l_{jit}]}.$$

Finally, we define a firm's market share as the share of value-added it captures in a given year within its 5-digit NACE industry.

4.2 Structural Estimation

Our theory features 15 parameters to be determined, collected in the set Ω :

$$\Omega = \{\rho, \beta, \eta, \theta, \epsilon, \kappa, \alpha, \delta, c_O, \sigma, c_I, \zeta, c_E, \xi, \chi\}.$$

To identify each of these, we assign conventional values to $\{\rho, \eta, \alpha, \delta\}$ and estimate the remaining parameters. The representative household's rate of time preference ρ is set equal to 0.04, the parameter η is set to unity as to deliver a unit Frisch elasticity of labor supply, the capital share is set to 1/3 and the rate of depreciation of physical capital to 0.06.²³ The overhead cost parameter c_O is normalized without loss of generality.²⁴

We then independently identify the following three parameters $\{\beta, \chi, \epsilon/\theta\}$ with three moments. First, the utility weight on labor supply β is chosen to match average hours worked per year per person in France between 1995 and 2019.²⁵ Second, the exogenous exit rate χ is set equal to the average exit rate of 1.34% for firms with more than 10 employees between 2009 and 2019 in France.²⁶ Third, the ratio ϵ/θ of the [Klenow and Willis \(2016\)](#) elasticity and super-elasticity parameters is estimated from the relationship between firm-level markups and market shares in our data. Figure 2 illustrates how that relationship depends on the value ascribed to the ratio ϵ/θ .

Following [Edmond et al. \(2023\)](#), we show in Appendix C that in a generalization our theory (with time-varying industry-level demand shifters and time-invariant firm-level

²²The chosen value for $r_t + \delta$ is consistent with the equilibrium interest rate of our model and the rate of physical capital depreciation we assumed.

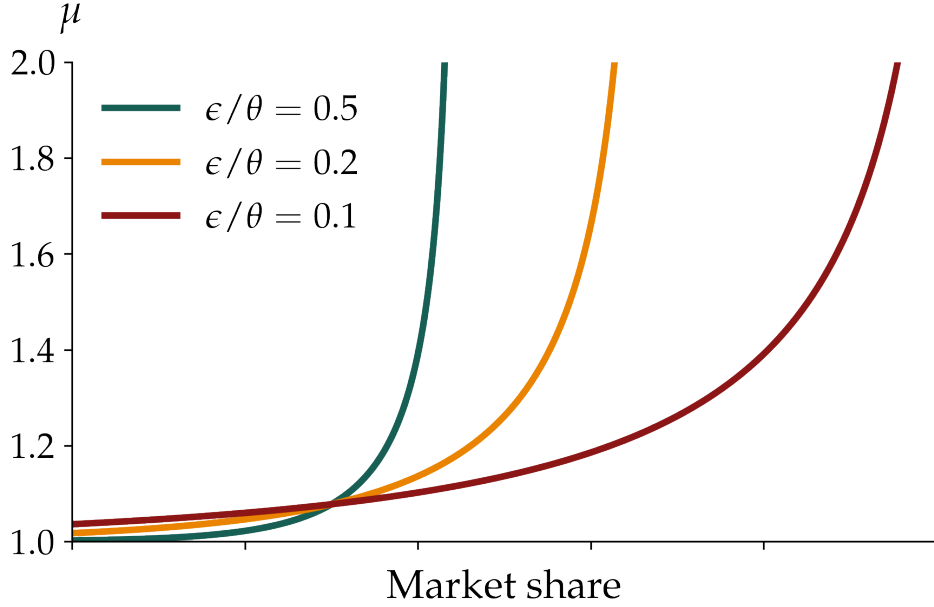
²³The cross-year average of the industry capital shares α_{it} we measure in the FARE data is equal to 0.22. To deal with this discrepancy, we consider a robustness check in Appendix C.3 where we inflate capital shares by a constant such that they aggregate to our model's capital share of 1/3.

²⁴Given an initial condition for the exit threshold, we choose its value such that aggregate output is normalized to unity in our model's initial stationary equilibrium.

²⁵For a time endowment of 16 hours per day and 365 days per year, the average of 1540.3 hours worked per person per year (calculated from [EU-KLEMS](#)) implies a value of $1540.3/(16 \times 365) \approx 0.26$ for H_t .

²⁶This moment is calculated from [Eurostat's](#) Business Demography Statistics, which goes back to 2008.

Figure 2: Markups and Market Shares



Note: This figure plots the relationship between firm-level markups and market shares, where the horizontal axis is defined on a logarithmic scale. Units are omitted for that axis since they provide no information.

demand shifters), this relationship is nonlinear and given by the following equation:

$$\mu_{jit}^{-1} + \ln(1 - \mu_{jit}^{-1}) = b + b_t + b_i + b_{it} + b_j + (\epsilon/\theta) \ln(s_{jit}) \quad \text{where} \quad s_{jit} \equiv \frac{p_{jit}y_{jit}}{P_{it}Y_{it}}.$$

Here, s_{jit} denotes the market share of firm j operating within industry i in year t , b is a constant, b_t is a time fixed effect, b_i is an industry fixed effect, b_{it} is a time-industry fixed effect and b_j is a firm fixed effect. We estimate this relationship in the FARE data using our firm-level markup and market share measurements described in Section 4.1. The results of this estimation exercise are reported in Table 3 for the entire sample and for the manufacturing sector only. We therefore obtain an estimate of $\epsilon/\theta = 0.243$ with a standard error of 0.001 clustered at the firm level. This estimate is in line with those reported in Edmond et al. (2023) and Amiti et al. (2019) who find values of 0.16 and 0.32, respectively. The latter identify this ratio by matching the variability in markups and resulting pass-throughs among Belgian manufacturing firms.

The demand elasticity parameter θ itself and the Kimball (1995) aggregation scalar κ only affect the measure of firms, for which there is no obvious empirical counterpart. We therefore choose their value such that the over-determining equilibrium condition

Table 3: Markups and Market Shares

Dependent variable: $\mu_{jit}^{-1} + \ln(1 - \mu_{jit}^{-1})$						
	Full sample			Manufacturing		
$\ln(s_{jit})$	0.047 (0.000)	0.234 (0.001)	0.243 (0.001)	0.049 (0.001)	0.321 (0.004)	0.331 (0.004)
Firm fixed effects		Y	Y		Y	Y
Industry \times year fixed effects	Y	Y	Y	Y	Y	Y
Industry fixed effects			Y			Y
Year fixed effects			Y			Y
Age group fixed effects	Y		Y	Y		Y
R^2	0.090	0.505	0.507	0.056	0.489	0.490
Observations	4.9M	4.9M	4.9M	0.5M	0.5M	0.5M

Note: Firm-level markups and market shares are constructed from the FARE dataset as described in Section 4.1. This table presents different regression specifications with firm fixed effects, 5-digit NACE industry fixed effects and age group fixed effects (for a total of 20 evenly-spaced age groups). Standard errors (in parentheses) are clustered at the firm level. The total number of observations is below the total sample size of 5.4M because negative markups were estimated for some firms.

of our model is met both in the initial and terminal stationary equilibria we analyze. More specifically, since the firm's problem is only a function of its productivity *relative* to the exit threshold, the scale of the latter is not pinned down by the smooth pasting condition.²⁷ Yet, that optimality condition must be satisfied in equilibrium. As we will analyze two equilibria (pre- and post-policy), both feature one more equation than the number of unknowns, which we satisfy by treating θ and κ as unknowns. The resulting estimate of 14.9 for θ is well within the range of values considered in the literature.²⁸

Finally, the remaining five parameters $\{\sigma, c_I, \zeta, c_E, \xi\}$ are jointly (over)identified by the following six moments via a GMM estimation strategy:

1. An aggregate (cost-weighted average) markup of 1.3, which averages the estimates of 1.1 and 1.5 from [De Ridder et al. \(2023\)](#) using the FARE (manufacturing) data.
2. The average annual growth rate of 1.16% of real GDP per hour worked in France between 1995 and 2019 calculated from [EU-KLEMS](#)'s national growth accounts.

²⁷The initial condition for that threshold is thus taken as given, as discussed in Section 3.3.

²⁸For example, [Edmond et al. \(2023\)](#) consider values between $\theta \in [5.66, 29.1]$.

3. The average annual growth rate of (deflated) firm-level value added of 1.24% calculated from the FARE data.
4. The average annual exit rate of 5.61% among all French firms between 2009 and 2019 calculated from Eurostat’s Business Demography Statistics.
5. The average size (value added) of entrants relative to incumbents of 31% calculated from the FARE data.
6. The within-industry standard deviation of log value added of 1.54 calculated from the FARE data.

The objective we minimize is the squared percent deviation between these moments and their counterpart in our theory’s stationary equilibrium allocation. We pose this estimation exercise as a mathematical program with equilibrium constraints (MPEC) (Su and Judd, 2012; Dubé, Fox and Su, 2012). Doing so allows us to perform the parameter search without repeatedly solving the model’s equilibrium conditions at each guess of parameters.²⁹ Table 4 reports the resulting values of our structural parameters.

Table 5 compares the empirical and theoretical moments listed above evaluated at the estimated parameter values. All of our targeted moments are matched with relatively high accuracy. Our model also replicates several untargeted moments in the FARE data such as the Gini coefficient of value-added, the share of total value added captured by the largest firms, the relative size of entrants by employment as well as the average and median age of a firm. It is also consistent with the empirically evidenced decreasing relationships between (1) the rate of exit and firm age (Caves, 1998), and (2) firm-level sales volatility and firm size (Yeh, 2021).

Figure 3 plots several static firm-level outcomes against relative productivity, where the stationary distribution of the latter is plotted in transparency to emphasize the “relevant” domain of each of those functions. Panel 3(a) plots the downward-sloping elasticity of the demand schedule faced by the firm, illustrating Marshall (1890)’s second law of demand. Panel 3(b) plots the price chosen by the firm (relative to the choke price) while Panel 3(c) plots the implied markup over marginal cost, which is increasing in size. Finally, Panel 3(d) plots the market share captured by more or less productive firms.

In Figure 4, we present the probability distribution function of firm-level markups, including various percentiles. Although our target is a cost-weighted average markup of 1.3, the *unweighted* median markup stands notably lower at 1.1. The implied dispersion in markups also appears more modest than its empirical counterpart. While we infer an interquartile range of 0.05 for the logarithm of markups, De Ridder et al. (2023) estimate

²⁹Appendix B.2 describes this exercise in detail and provides a formal discussion on identification.

Table 4: Structural Parameters

Parameter	Symbol	Value
<i>Household preferences:</i>		
Rate of time preference	ρ	0.04
Labor supply utility weight	β	10.3
Frisch elasticity of labor supply reciprocal	η	1
<i>Final sector technology:</i>		
Klenow and Willis (2016) elasticity parameter	θ	14.9
Klenow and Willis (2016) super-elasticity parameter	ϵ	3.62
Kimball (1995) aggregation constant	κ	0.75
<i>Firm production technology:</i>		
Output elasticity of physical capital	α	0.33
Depreciation rate of physical capital	δ	0.06
Overhead cost parameter	c_O	0.03
<i>Firm innovation technology:</i>		
Brownian motion standard deviation	σ	0.03
Innovation cost scale parameter	c_I	9.23
Innovation cost elasticity parameter	ζ	0.99
<i>Entry and exit:</i>		
Entry cost parameter	c_E	6.61
Entry distribution parameter	ξ	1.71
Exogenous exit rate	χ	1.34%

Note: This table presents the assigned/estimated structural parameters of our theory.

values of 0.48 and 0.2, respectively, using information on quantities or revenues from the FARE data (for the manufacturing sector). This serves to illustrate that our model strictly captures the markup dispersion that is systematically related to firm size. In that sense, our estimation strategy leans somewhat conservatively on the extent of heterogeneity in markups.

Table 5: Moments

Moment	Source	Model	Data
<i>Targeted:</i>			
Aggregate markup	FARE	1.30	1.30
GDP per hour worked growth rate	EU-KLEMS	1.23%	1.16%
Incumbent value added growth rate	FARE	1.18%	1.24%
Exit rate of all firms	Eurostat	5.32%	5.61%
Relative size of entrants by value added	FARE	0.30	0.31
Standard deviation of log value added	FARE	1.51	1.54
<i>Untargeted:</i>			
Gini coefficient of value added	FARE	0.78	0.73
Top 5% value added share	FARE	52%	51%
Top 10% value added share	FARE	70%	64%
Relative size of entrants by employment	FARE	0.33	0.32
Average firm age	FARE	27.8	19.3
Median firm age	FARE	15.1	15.7

Note: This table presents moments (targeted or not in our GMM estimation exercise) and their resulting value in our model. Moments measured in the FARE data are first calculated within 2-digit NACE industries and then aggregated with each industry's share of total value added.

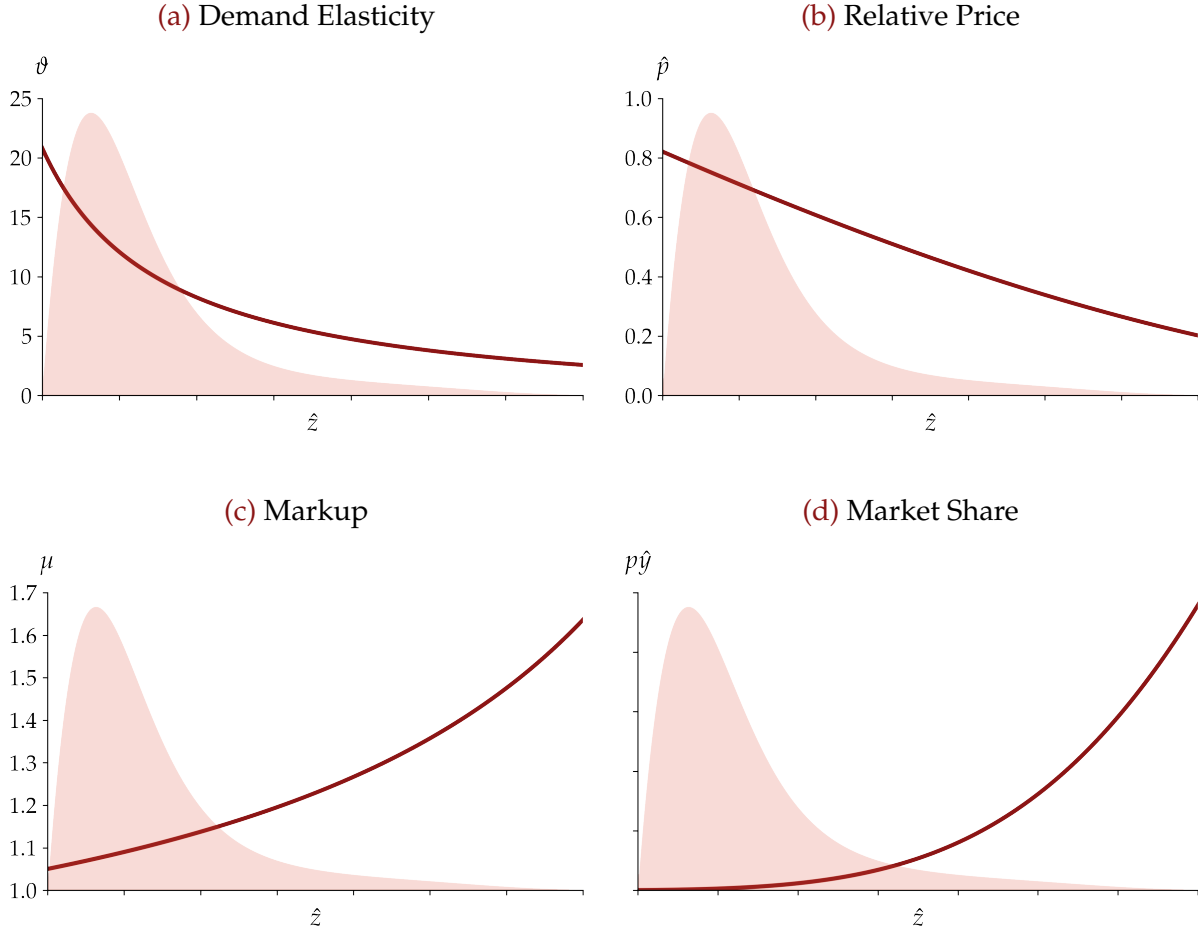
Parameter Identification

While the five parameters $\{\sigma, c_I, \zeta, c_E, \tilde{\zeta}\}$ are jointly (over)identified by the six moment conditions above, we see it as informative to intuitively discuss the reasoning behind our selection of these six moments. Appendix B.2 provides a more formal analysis of the identification of these parameters.

We include the aggregate markup as a target in our estimation exercise since it has a direct bearing on the extent to which the aggregate scale of the economy is inefficiently restricted.³⁰ While none of the five aforementioned parameters explicitly appear in the expression for the aggregate markup, several of them influence the shape of the relative productivity distribution over which firm-level markups are aggregated and are, in that sense, identified by that aggregate. This is evident from the KFE in equation (8), in which the parameters $\{\sigma, c_I, \zeta, \tilde{\zeta}\}$ appear either through the firm-level productivity

³⁰Edmond et al. (2023) find that the “static” welfare losses from markups exhibit significant convexity in the aggregate markup target.

Figure 3: Static Firm-Level Outcomes



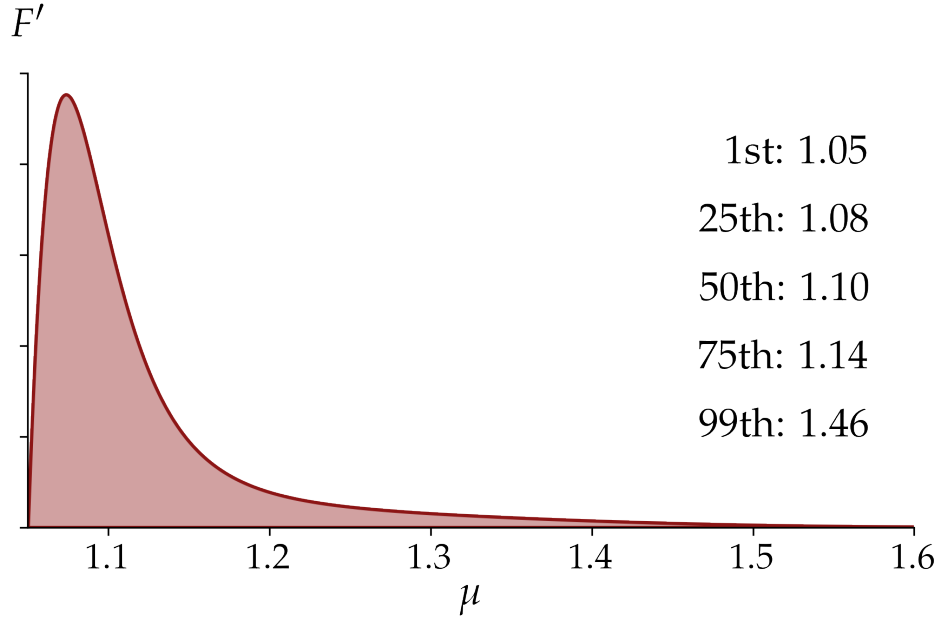
Note: Here, all variables are defined according to the estimated values for θ and ϵ , which are reported in Table 4. The distribution plotted in transparency is the stationary distribution of relative productivity. Units are omitted on axes where they are not informative.

process or the dynamics of entry and exit.

We aim for a value of 1.3 for the aggregate markup, which aligns with the range of values estimated by De Ridder et al. (2023) using the FARE data for the manufacturing sector. More precisely, they calculate the sales-weighted harmonic average markup to be approximately equal to 1.1 or 1.5 with revenue and quantity data, respectively. Our choice of targeting the average of these two estimates is motivated by recent evidence that markups are difficult to identify with either revenue or production data (Bond, Hashemi, Kaplan and Zoch, 2021; De Ridder et al., 2023; Flynn, Traina and Gandhi, 2019; Raval, 2023).³¹ In Section 5.5, we consider how our results change as we aim for lower

³¹For different empirical counterparts to the markup in our theory, we have estimated cost-weighted

Figure 4: Distribution of Firm-Level Markups



Note: This figure plots the unweighted distribution of firm-level markups with its 1st, 25th, 50th, 75th and 99th percentiles. The interquartile range of the logarithm of markups is equal to 0.05, which is more modest than the values of 0.48 and 0.2 estimated by De Ridder et al. (2023) on quantity and revenue data, respectively.

or higher values.

We target the *aggregate* and *firm-level* value-added growth rate moments with two objectives in mind. First, since we propose a model of endogenous economic growth, it appears imperative that our parameterization is consistent with the growth rate of the French economy. Second, since our paper explores the distortionary consequences of markups on incumbent firms' investment decisions, it seems important to discipline their contribution to TFP growth, despite the absence of a clear, direct empirical counterpart.³² As such, the aggregate and firm-level growth rate moments are intended to identify the parameters of the firm's innovation cost function, c_I and ζ , thereby delineating the pace of economic growth and the contribution of incumbent firms to this progress.

As accentuated in the technology diffusion literature, yet another contribution to long-run productivity growth comes from the selective survival of successful firms and the adoption of existing technologies by new entrants. The scope of this contribution hinges on two factors: (1) the frequency at which underperforming firms are supplanted

average markups that ranged from 5% to almost 80%.

³²There have been substantial attempts to infer this contribution indirectly, as evidenced by Luttmer (2007) and Garcia-Macia et al. (2019).

by more efficient newcomers and (2) the productivity differential between these two groups of firms. To discipline these two factors, we target the rate at which firms exit—which must be equated to the entry rate on a balanced growth path since our model features a constant population—and the initial value-added of new entrants relative to incumbents, which positively correlates with productivity in our model.³³ Hence, these two moments partly identify the parameters ζ and σ , where the former regulates the transformation of the incumbent distribution from which entrants draw their relative productivity and the latter directly influences the rate at which unsuccessful firms are swept below the endogenous exit threshold.

Finally, we include the standard deviation of the logarithm of value added as a target in our estimation exercise, enabling us to regulate the extent of markup dispersion. Our model posits that the sole source of dispersion in markups across firms comes from their endogenous differences in process efficiency, which cause them to charge varying prices at which demand is more or less elastic. That is, the extent of dispersion in markups is intrinsically tied to the degree of heterogeneity in firm size. As contended in [Edmond et al. \(2023\)](#), we regard this approach as more conservative than the alternative of directly targeting the empirically observed dispersion in markups, which could instead reflect dispersion in other types of distortions unrelated to markups.

5 Counterfactual Analysis

To quantify how heterogeneous markups differentially distort firms’ incentives to invest in R&D, we consider the introduction of size-dependent subsidies inducing each firm to price at marginal cost and therefore produce at the efficient scale. We analyze how firm-level decisions and economic aggregates endogenously respond to this intervention.

5.1 Policy Intervention

Why do we consider such a particular policy intervention rather than simply comparing the economy’s *laissez-faire* and optimal resource allocations? The reason is that, in order to achieve a stationary distribution of relative productivity, our economic environment must feature technological spillovers across firms, which introduces externalities that are not intrinsically tied to markups.³⁴ Therefore, a comparison of the *laissez-faire* and optimal allocations would not *isolate* the costs of markups. Instead, the implementation

³³We target the exit rate rather than the entry rate, since the latter may reflect long-run growth in the number of firms, which is not a feature of our model.

³⁴These externalities are further discussed in [Section 5.4](#).

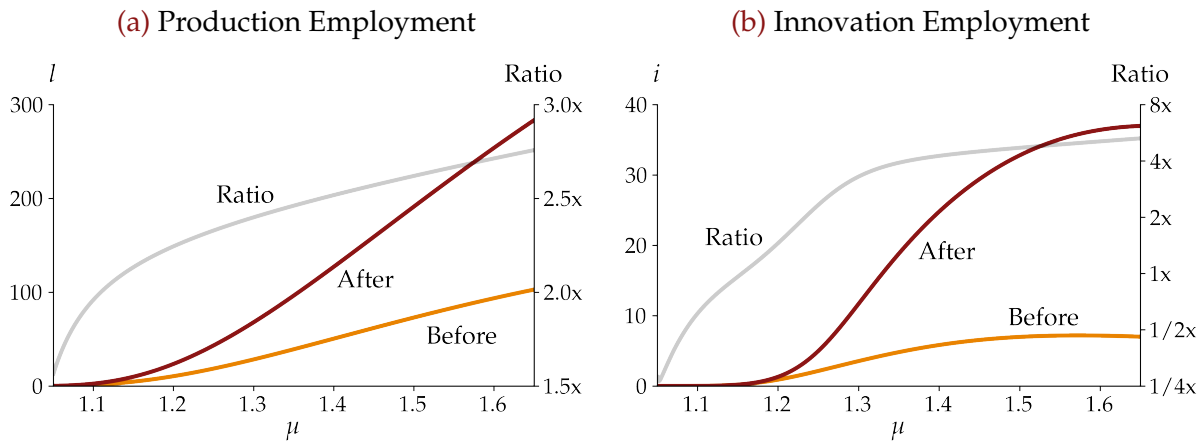
of a policy intervention devised to induce firms to operate at their efficient scale directly addresses the product market distortions caused by markups. Following [Edmond et al. \(2023\)](#), we show in Appendix [A.1.7](#) that the size-dependent subsidy scheme $T_t(\hat{y})$ that achieves this is:

$$T_t(\hat{y}) = [\Upsilon(\hat{y}) - \Upsilon'(\hat{y})\hat{y}]Y_tD_t \quad (11)$$

which we assume is financed by lump-sum taxes levied on the representative household. Under this policy, firms optimally price at marginal cost as the subsidy schedule outlined in equation (11) is such that they consider the final sector's output rather than their revenue as part of their objective. That is, these subsidies transfer the entire consumer surplus to firms.³⁵

As an intended result of this intervention, firms optimally operate at a larger scale. This is depicted in Panel 5(a), which plots firms' production employment in both the pre- and post-policy stationary equilibrium allocations. In most of the figures that follow, firms' initial markup before the intervention is plotted on the horizontal axis whose range covers more than 99.9% of the measure of firms. Notably, production employment increases for all firms but is also reallocated towards larger, more productive firms that initially commanded a higher markup. While firm-level production employment roughly doubles on average, it almost triples for the largest firms.

Figure 5: Firm-Level Labor Allocations



Note: Both functions are plotted over the same support of initial markups, covering more than 99.9% of the measure of firms both before and after the intervention. The red and orange lines (left axes) respectively plot labor allocations before and after the policy intervention. The gray line (right axis) plots their ratio (post- relative to pre-policy).

³⁵This policy intervention is large in magnitude, equivalent to about 17% of aggregate output.

The reallocation of innovation employment is even starker, as illustrated in Panel 5(b). As the size of the market (relative to the equilibrium wage) contracts for the smallest firms, but expands for the largest ones, R&D resources are reallocated towards the latter. On average, firm-level innovation employment more than triples. Yet, it shrinks by about 75% for the least productive firms, while their most productive counterparts see a more than 5-fold increase in their allocation.

Table 6 presents the aggregation of these firm-level outcomes. Aggregate demand for production, innovation, and entry labor increases considerably, matched by a 23% rise in labor supply.³⁶ However, the aggregate labor allocation to overhead contracts by 41.8% due to a proportional decrease in the measure of varieties. While more resources are allocated to entry post-policy due to the greater scale and convexity of profits, the endogenous exit rate increases disproportionately from 4% to 15.5%, thus depleting the stock of varieties. As R&D resources are redirected towards more productive firms, the smallest firms fail to keep up with the competition, trail behind and eventually exit (endogenously) at a higher rate.

This dynamic is most clearly illustrated in Figure 6. In particular, Panel 6(a) plots the firm-level productivity drifts achieved pre- and post-policy. Consistent with Figure 5, the growth trajectory of the smaller, less efficient firms decelerates post-policy and, as a result, they congregate near the exit threshold. This is depicted in Panel 6(b), where it is shown that the post-policy stationary distribution of relative productivity admits a higher density of small firms. However, it also features a higher density of fast-growing large firms, such that the distribution becomes slightly bimodal after the intervention. Overall, market concentration rises only slightly after the intervention, as presented in the last two rows of Table 6, which report the share of total value added captured by the top 5% and 10% largest firms.

A closer look at the firm's dynamic first-order condition sheds light on the disparity in growth trajectories depicted in Panel 6(a):

$$\frac{V'_t(\hat{z})}{w_t} = \frac{\partial i(\gamma, \hat{z})}{\partial \gamma}.$$

This condition implies that a firm will achieve a larger productivity drift if the resulting change in its value is large relative to the prevailing wage rate. Despite the absence of an analytical solution for the former, one can gain insight into what determines the extent of the firm's marginal value through the following asymptotic proposition:

Proposition 5. *On a balanced growth path, the firm's value function asymptotes to the present*

³⁶As a reference, Edmond et al. (2023) find, under targeted aggregate markups of 1.25 and 1.35, that the same policy intervention achieves a 30.1% and 42.1% increase in labor supply, respectively.

Table 6: Economic Aggregates

Aggregate	Before	After	Change
<i>Labor allocations:</i>			
Labor supply	0.264	0.325	+23.0%
Production labor	0.227	0.260	+14.5%
Innovation labor	0.020	0.036	+76.9%
Entry labor	0.015	0.028	+84.7%
Overhead labor	0.002	0.001	-41.8%
<i>Firms, entry and exit:</i>			
Measure of varieties	0.044	0.025	-41.8%
Entry rate	5.32%	16.87%	+11.6p.p.
Endogenous exit rate	3.97%	15.52%	+11.6p.p.
<i>Market concentration:</i>			
Top 5% value added share	51.6%	54.6%	+3.0p.p.
Top 10% value added share	69.8%	77.2%	+7.4p.p.

Note: This table presents the pre- and post-policy level of various economic aggregates as well as the corresponding percentage change.

discounted value of asymptotic profits:

$$\lim_{\hat{z} \rightarrow \infty} V_t(\hat{z}) = \bar{V}_t \quad \text{where} \quad \bar{V}_t \equiv \frac{\bar{\pi} Y_t D_t - w_t c_O}{\rho + \chi}$$

and where the constant $\bar{\pi}$ is given by:

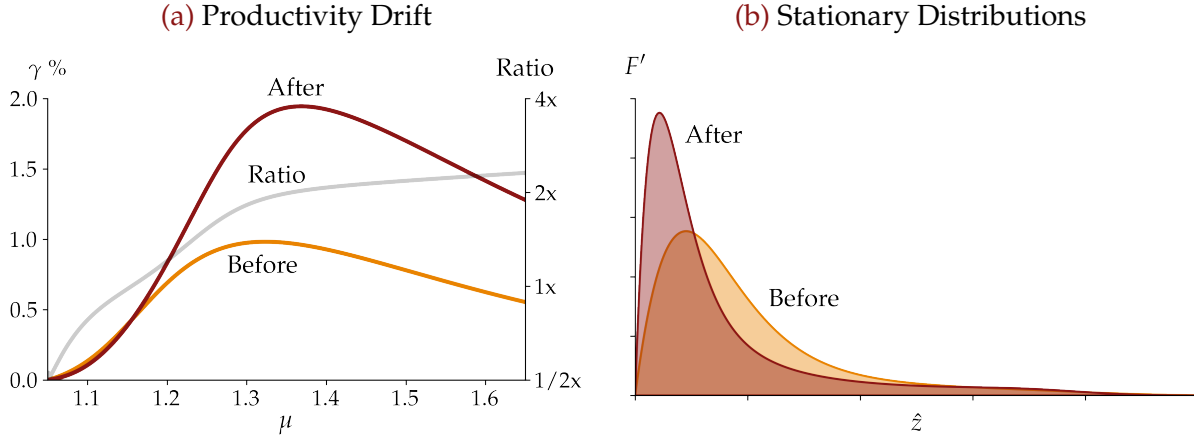
$$\bar{\pi} = \begin{cases} (\theta - 1) \exp[(1 - \theta)/\epsilon] \theta^{\theta/\epsilon - 1} & \text{Pre-policy,} \\ 1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon - 1} \Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) & \text{Post-policy.} \end{cases}$$

Using Proposition 5 together with the expression for aggregate labor demand, let us denote the firm's value function relative to its asymptote by $\hat{V}_t(\hat{z}) \equiv V_t(\hat{z})/\bar{V}_t$. This allows us to recast the firm's dynamic first-order condition as:

$$\frac{\bar{\pi} \mathcal{M}_t L_t D_t - (1 - \alpha) c_O}{(1 - \alpha)(\rho + \chi)} \times \hat{V}'_t(\hat{z}) = \frac{\partial i(\gamma, \hat{z})}{\partial \gamma}.$$

Given that the term $\bar{\pi} \mathcal{M}_t L_t D_t$ increases by a factor of almost 5 after the intervention, all

Figure 6: Productivity Drifts and Stationary Distributions



Note: For consistency, both functions are plotted over the same relative productivity support as in Figure 5. In Panel 6(a), the red and orange lines (left axes) respectively plot productivity drifts before and after the policy intervention. The gray line (right axis) plots their ratio (post- relative to pre-policy). It is worth noting that the distributions' support in Panel 6(b) are defined *relative* to the least productive firm in each allocation. The *level* of the exit threshold may change due to the policy intervention.

firms would be set on a trajectory of accelerated growth were it not for changes in the curvature of the firm's value function. Consequently, it is the curvature in the subsidy scheme that effectively reallocates innovation employment away from smaller firms towards their larger counterparts.

5.2 Long-Run Economic Growth

What are the implications of these findings for long-run economic growth? Our analysis reveals that the stationary growth rate of TFP increases by 1.2 percentage points, from 0.82% to 2.05%. Due to physical capital accumulation, this leads to a corresponding increase in the growth rate of aggregate output, from an initially targeted value of 1.23% to a counterfactual value of 3.08%.

Table 7 provides a breakdown of the pre- and post-policy growth rates, along with the corresponding change, according to Proposition 3. In particular, this decomposition presents the contributions of incumbent firms' productivity drift and volatility, as well as that of entry and exit. Of noteworthy importance is the role of incumbents' productivity drift, which constitutes the largest share of both growth levels (before and after the intervention): it respectively accounts for 65.5% and 50% of TFP growth in the pre- and post-policy stationary equilibria. Nonetheless, it contributes slightly less to the growth

rate differential between the two, comprising 39.8% of it, while the remaining share of the growth rate change is attributed to the contribution from entry and exit.

Table 7: Growth Rate Decomposition

Contribution	Before	After	Change (p.p.)
Incumbent drift	0.54%	1.02%	+0.49%
Incumbent volatility	0.18%	0.19%	+0.01%
Entry and exit	0.10%	0.84%	+0.74%
Total	0.82%	2.05%	+1.23%

Note: This table presents the contributions to the level and change of TFP growth.

The contribution is an indirect consequence of the R&D reallocation prompted by the policy. As R&D expenditures are redirected from small, unproductive firms to their larger, more productive competitors, the former grapple with dwindling resources, trail behind and exit at a higher rate. However, these departing firms are replaced by new entrants who seize the opportunity to replicate the existing technologies of more productive incumbents. Consequently, the higher churn rate brought about by the R&D reallocation leads to more frequent improvements in the pool of productivity.

As discussed in Section 3.3, the growth contribution from incumbent innovation takes the form of a weighted average of firm-level productivity drifts. Denoting changes between the pre- and post-policy stationary equilibria by the operator Δ , the change in this weighted average can be decomposed as:

$$\begin{aligned}
\Delta \int_0^\infty \omega(\hat{z})\gamma(\hat{z})d\hat{z} &= \int_0^\infty \Delta\gamma(\hat{z}) \times \omega(\hat{z})d\hat{z} && \text{Productivity growth} && (62.7\%) \\
&+ \int_0^\infty \Delta\omega(\hat{z}) \times \gamma(\hat{z})d\hat{z} && \text{Market expansion} && (14.9\%) \\
&+ \int_0^\infty \Delta\omega(\hat{z}) \times \Delta\gamma(\hat{z})d\hat{z} && \text{Covariance} && (22.4\%).
\end{aligned} \tag{12}$$

The first term reflects the average change in firm-level productivity drifts, evaluated over firms' initial expansion responses (and their initial composition). The second term instead reflects the change in these expansion responses, keeping firms' productivity drifts constant. Finally, the third term captures the covariance between these changes. The "productivity growth" term accounts for 62.7% of the total, compared to 14.9% and

22.4% for the “market expansion” and “covariance” terms, respectively.

Therefore, the larger contribution from incumbent innovation is mostly due to (1) firms simply growing faster on average and (2) initially larger firms achieving the largest increases in their productivity drift. However, a noteworthy contribution comes from the heightened correlation between firms’ productivity growth, their size and their density. Larger firms, who initially commanded higher markups and whose scale was consequently most restricted, achieve disproportionately faster productivity growth, differentially expand in scale and become more numerous following the intervention. Accordingly, process efficiency improvements are rolled out over more units sold. This contrasts with models of CES demand, in which firms’ scale are uniformly distorted, independently of their size.

5.3 Aggregate and Cross-Firm Allocations of R&D

In light of the preceding analysis, a question arises: is the accelerated growth in TFP the result of a greater aggregate allocation of labor to innovation, or is it attributable to its reallocation across firms? To elucidate this, we consider an alternative intervention wherein the baseline subsidies are upheld, but concomitantly, a uniform tax is levied on firms’ R&D expenditures. This tax is precisely chosen to fix the aggregate allocation of labor to innovation at a level commensurate with the initial stationary equilibrium, which isolates the role of R&D misallocation.

Table 8: Growth Rate Decomposition for Fixed Innovation Labor

Contribution	Before	Baseline		Fixed R&D	
		After	Change	After	Change
Incumbent drift	0.54%	1.02%	+0.49%	0.71%	+0.17%
Incumbent volatility	0.18%	0.19%	+0.01%	0.22%	+0.04%
Entry and exit	0.10%	0.84%	+0.74%	0.82%	+0.72%
Total	0.82%	2.05%	+1.23%	1.74%	+0.92%

Note: This table presents the contributions to the level and change of TFP growth when fixing or not the aggregate allocation of labor to innovation to its initial level before the implementation of the policy intervention. Doing so requires imposing a uniform tax of 44.5% on firms’ expenditures on R&D.

The outcomes of this alternate policy are detailed in Table 8. It achieves an increase in TFP growth equal to 74.8% of the increment observed under the baseline policy, where variations in the aggregate allocation of labor to innovation were admissible. A little over three-quarters of this accelerated growth results from an intensified selection of firms. The redistribution of R&D resources from less productive, smaller firms to their larger, more efficient competitors induces the exit of the former, which are replaced by more productive newcomers. Furthermore, nearly one-fifth of the uplift in TFP growth is derived from an expanded contribution from the productivity drift of incumbent firms. Although the aggregate allocation of labor to innovation is kept fixed, the per-firm average allocation escalates by 49.7% owing to the diminishing number of firms.

5.4 Aggregate Markup and Markup Dispersion

To get a deeper sense of the driving forces behind the acceleration in TFP growth, we disentangle the role of the *aggregate* markup from that of markup *dispersion*. To do so, we consider a slightly more general tax and subsidy schedule. As described in Edmond et al. (2023), the transfers of equation (11) can be generalized with the following parameterization:

$$T_t(\hat{y}) = [\tau_0 \Upsilon(\hat{y}) + \tau_1 \Upsilon'(\hat{y})\hat{y}]Y_t D_t \quad (13)$$

where τ_0 and τ_1 can be appropriately chosen to either mitigate the level or dispersion in markups. In particular, we show in Appendix A.1.7 that setting $\tau_0 = 0$ and $\tau_1 = \mathcal{M}_t - 1$ delivers a uniform subsidy scheme, which leaves the dispersion in markups unchanged from the initial equilibrium, but eliminates the aggregate markup from equation (10). Instead, setting $\tau_0 = \mathcal{M}_t^{-1}$ and $\tau_1 = -1$ delivers a size-dependent tax/subsidy scheme, which eliminates markup dispersion while holding the aggregate markup fixed.

Table 9 replicates Table 7 but for those alternative schemes. In particular, the two columns labeled “Level fix” and “Dispersion fix”, respectively refer to the scheme that either rectifies the level or dispersion in markups. The “Level fix” has a comparatively muted impact on long-run TFP growth, decreasing it by a slight 4 basis points. This subdued response is largely due to (1) a nearly unchanged growth contribution from incumbent firms’ investments in R&D, and (2) a weak reallocation of R&D across firms. The former reflects a pecuniary externality: as firms expand in scale and demand more production labor, they bid up the cost of R&D through a higher wage. This conclusion contrasts with the findings of Edmond et al. (2023) who infer an important role for the aggregate markup in distorting the scale of the economy at a point in time. Instead, we find the level of markups to have little to no bearing on the rate at which the economy

grows in the long run.

Table 9: Growth Rate Decomposition for Alternative Transfer Schedules

Contribution	Before	Baseline		Level fix		Dispersion fix	
		After	Change	After	Change	After	Change
Incumbent drift	0.54%	1.02%	+0.49%	0.51%	-0.03%	1.06%	+0.53%
Incumbent vol.	0.18%	0.19%	+0.01%	0.18%	-0.00%	0.19%	+0.01%
Entry and exit	0.10%	0.84%	+0.74%	0.09%	-0.01%	0.89%	+0.79%
Total	0.82%	2.05%	+1.23%	0.78%	-0.04%	2.15%	+1.33%

Note: This table presents the contributions to the level and change (p.p.) of TFP growth under alternative policy interventions. Specifically, the columns labeled “Baseline”, “Level fix” and “Dispersion fix” refer to the transfers that rectify both, and either the level or dispersion in markups, respectively.

Compared with the “Level fix” policy, the tax/subsidy schedule that eliminates markup dispersion while leaving their average level unchanged achieves an even larger increase in long-run TFP growth than the baseline intervention. The largest contributors are, here again, the terms reflecting incumbents’ productivity drift and entry and exit. Nonetheless, it is worth noting that this faster productivity growth is not necessarily indicative of an improvement in welfare. On the one hand, too few resources may be directed towards production under this allocation, thus reducing aggregate output. On the other hand, this tax/subsidy scheme might strike a more optimal balance between rectifying product market distortions and addressing other market failures.

As mentioned earlier, these market failures take the form of technological externalities across firms. As entering firms draw their relative productivity from a transformation of the incumbent distribution, the latter do not internalize that their R&D investments benefit future cohorts of firms. Further, as emphasized in [Lashkari \(2023\)](#), since the lower bound of the productivity support is endogenous, an unproductive firm may choose to stay in business to extract rents, but in doing so, it “pollutes” the pool from which entrants draw their relative productivity.

All else equal, these inefficiencies imply that the market would (1) allocate too few resources to R&D and (2) harbor an inefficiently large density of small unproductive firms. The “Dispersion fix” tax/subsidy scheme inadvertently addresses both of these market failures, albeit imperfectly. In preserving the level of markups, production labor

demand remains inefficiently low, thus “freeing up” resources for R&D. Further, this scheme takes the form of a tax for the smallest firms, which is passed on through higher prices, in turn, reducing the demand they face and edging them out of the market. Yet, without a comprehensive welfare analysis, it remains unclear which policy achieves the largest improvement in welfare.

5.5 Robustness

An important assumption entertained in the quantification of our model is to target an aggregate markup of 30%. [De Ridder et al. \(2023\)](#) measure a sales-weighted *harmonic* average markup of 10% and 50% using French firm-level revenue and quantity data, respectively.³⁷ To assess the implications of this assumption, Table 10 replicates Table 7, albeit with a targeted aggregate markup of 10% or 50%.³⁸ Structural parameters are re-estimated under these alternative assumptions. With a lower target of 10%, the implications of the intervention are more tempered. The increase in the long-run growth rate of TFP is significantly muted at 29 basis points. Yet, we see a nontrivial contribution of 13 basis points from the faster productivity growth achieved by incumbent firms.

Table 10: Growth Rate Decomposition for Alternative Aggregate Markup Targets

Contribution	$\mathcal{M}=1.1$			$\mathcal{M}=1.5$		
	Before	After	Change	Before	After	Change
Incumbent drift	0.15%	0.28%	+0.13%	0.83%	2.12%	+1.29%
Incumbent vol.	0.33%	0.47%	+0.14%	0.14%	0.15%	+0.01%
Entry and exit	0.28%	0.30%	+0.02%	-0.20%	0.77%	+0.97%
Total	0.76%	1.05%	+0.29%	0.77%	3.03%	+2.26%

Note: This table presents the contributions to the level and change of TFP growth under alternative aggregate markup targets. Specifically, the columns labeled “ $\mathcal{M}=1.1$ ” and “ $\mathcal{M}=1.5$ ” respectively refer to parameterizations that target a cost-weighted average markup of 1.1 and 1.5.

With a target of 50% for the aggregate markup, the repercussions of the intervention are magnified. TFP growth increases by 2.26 percentage points, with a considerably

³⁷As a reference point, [Aghion et al. \(2023\)](#) entertain an aggregate markup of 50%.

³⁸Appendix C.3 further replicates Tables 4, 5 and 6 for these cases.

larger contribution from incumbent firms' productivity drift. Interestingly, decomposing the change in this term according to equation (12) reveals a larger contribution (46.1% of the total) from a greater correlation between firm-level drifts and their expansion responses. Meanwhile, the contribution of larger productivity drifts on average (48.2% of the total) is slightly subdued relative to the baseline.

5.6 Discussion

In this subsection, we delve into the nuances of our theoretical framework by exploring a range of possible extensions and alternative assumptions. A forthcoming extension is to conduct a comprehensive welfare analysis incorporating transition dynamics. As emphasized in [Atkeson, Burstein and Chatzikonstantinou \(2019\)](#), transition dynamics tend to be slow in models of endogenous economic growth. Such inertia might curtail the intervention's welfare implications if the accelerated pace of productivity growth mostly materializes in the distant future.

Process vs. Product Innovation

In our framework, firms invest in R&D to achieve improvements in their *process efficiency*. An alternative assumption is to consider *product quality* improvements as the objective for those investments. We show in Appendix A.3 that when the quality and quantity of a product are perfect substitutes, this alternative is isomorphic to our model. While the assumption of perfect substitutability between quantity and quality is surely stylized, the economic modeling of the latter lacks a cohesive consensus. Hence, the analysis of endogenous product quality improvements under non-isoelastic demand systems presents a promising frontier for exploration.

Production vs. Innovation Resource Substitution

Our analysis yields a perhaps unexpected result: a uniform subsidy that neutralizes the distortionary consequences of the aggregate markup—while preserving the dispersion in markups—mildly diminishes the long-run growth rate of TFP. It is worth noting that this result is attributable to our theoretical choices. As posited in Section 3.1, labor can be seamlessly reallocated between production and innovation, such that the level of markups can spur over-investment in R&D. Thus, the extent of substitutability between production and innovation resources determines the extent to which the aggregate markup can skew economic growth.

To elucidate, consider an alternative economy wherein an elastic supply of ‘skilled’ and ‘unskilled’ labor can only be allocated to innovation and production, *respectively*. Under these circumstances, the level of markups can only catalyze under-investment in R&D as ‘unskilled’ labor is non-transferable to it. An even more contrasting alternative is that of an economy in which the final sector (instead of intermediate firms) uses labor along with intermediate inputs to produce the final good. In this context, production labor demand would instead be inefficiently high, as the final sector would substitute away from marked-up intermediates. Hence, rectifying the level of markups would induce the final sector to reallocate expenditures toward intermediates (away from production labor), thus freeing up resources for R&D instead of restricting them. In that sense, we consider our choices to be conservative.

Monopolistic vs. Oligopolistic Competition

Another assumption of our model is the tractable premise of *monopolistic* competition. Notwithstanding, [Edmond et al. \(2023\)](#) find that the static efficiency losses from markups are greater under *oligopolistic* competition, through which they infer greater dispersion therein. It remains to be shown whether this conclusion extends to the dynamic losses from markups, but one might conjecture that it could, based on the following argument. The [Klenow and Willis \(2016\)](#) demand system makes two counterfactual predictions under monopolistic competition. First, in order to accommodate a reasonable distribution of markups, it requires unreasonably little heterogeneity in pass-throughs, which grates against the evidence documented in [Amiti et al. \(2019\)](#). Second, the sharp concavity in demand at lower prices appears empirically inconsistent when juxtaposed against the heavy observed tails of firm-level employment or sales.

What are the implications of these counterfactual predictions for economic growth? With little dispersion in pass-throughs prior to the intervention, the transition to marginal cost pricing (complete pass-throughs for all firms) barely improves the reallocation of demand towards the most productive firms. Further, the pronounced decline in the elasticity of demand implies that a streamlined transmission of cost reductions through lower prices is not met with substantially higher demand. These two properties of our model limit the extent to which markups can deliver an inefficiently low correlation between firm-level productivity growth and their resulting expansion responses. [Amiti et al. \(2019\)](#) argue that models of oligopolistic competition are better suited to replicate the joint distribution of markups and pass-throughs. This suggests that the consequences of markups might be amplified under a market structure that (1) aligns with lower pass-throughs for large firms and (2) obviates the requirement for a steeply declining demand elasticity at lower prices.

Customer Acquisition

Neglecting the role of customer acquisition might also understate such consequences. [Afrouzi et al. \(2023\)](#) find that in a model parallel to [Edmond et al. \(2023\)](#), considering the endogenous accumulation of a customer base unveils greater dispersion in markups, thereby intensifying the implied efficiency losses from markups.³⁹ In addition, [Einav, Klenow, Levin and Murciano-Goroff \(2021\)](#) document that, while new entrants accrue lower total sales compared to incumbents, their average sales per customer are nearly equivalent. Since a firm's process efficiency is related to the intensive rather than the extensive margin of demand, these findings suggest that new entrants might be more productive than inferred in models omitting a customer margin, such as ours. Hence, accounting for this extensive margin of demand could amplify the role played by the selective displacement of unsuccessful firms by more productive newcomers.

Firm Ownership, Nonlinear Pricing and Semi-Endogenous Growth

However, other forces not encompassed in our analysis—such as concentrated firm ownership, nonlinear pricing and semi-endogenous growth—might instead temper our conclusions. [Boar and Midrigan \(2019\)](#) find that when firm ownership is concentrated, tensions arise between concerns for efficiency and inequality. Whether these trade-offs are mitigated or amplified in models of endogenous economic growth remains an open question. When firms engage in second-degree price discrimination, as in [Bornstein and Peter \(2023\)](#), their ability to appropriate a larger fraction of the consumer surplus might narrow the gap between private and social incentives for innovation. The work of [Jones \(1995\)](#) adds yet another layer of nuance. Should economic growth be intrinsically tied to demographic dynamics, the acceleration in TFP growth we identified might only be transient. Combining the insights from semi-endogenous growth theory with models of firm-led technological change, such as those proposed by [Peretto \(1998\)](#), [Dinopoulos and Thompson \(1998\)](#), [Young \(1998\)](#) and [Howitt \(1999\)](#), offers a promising route for future exploration.

6 Conclusion

We studied the dynamic consequences of markups for long-run economic growth in a general equilibrium theory of firm-led endogenous technological change. In our model,

³⁹These findings may extend to frictions in accumulating factors of production such as in [Bilal, Engbom, Mongey and Violante \(2021\)](#).

firms engage in monopolistic competition, charge heterogeneous markups and economic growth is propelled by their investments in R&D, in the pursuit of profit opportunities.

Two economic insights underpinned the arguments of this paper. The first posits that the level of markups either induces under- or over-investment in R&D. In restricting firms' scale, markups (1) curb their incentives to invest in cost reductions that apply to fewer units sold, but also (2) inefficiently free up resources for such investments via general equilibrium dynamics. Second, with dispersion in markups, both the aggregate and cross-firm allocations of R&D spending can be inefficient.

To quantify these implications, we estimated the structural parameters of our model using French macroeconomic and firm-level administrative data. We found that the introduction of size-dependent subsidies that induce firms to operate at their efficient scale raises long-run productivity growth by 1.2 percentage points. This faster growth is the result of an increase in aggregate R&D spending, a reallocation of those expenditures towards firms with a broader market reach and higher churn from entry and exit.

We conducted two alternative exercises to further elucidate these findings. We first explored a "constrained" intervention, which kept the aggregate allocation of labor to innovation fixed, to determine whether the uptick in TFP growth is predominantly due to the allocation of additional resources to R&D or a more efficient reallocation of these resources across firms. Our findings suggest that nearly 75% of the baseline acceleration in growth can be attained by simply reallocating a fixed quantity of innovation labor from small, unproductive firms to their larger, more productive competitors.

Our final counterfactual exercise disentangled the role of the aggregate markup from that of markup dispersion. We found that a uniform subsidy that strictly rectifies the level of markups has a relatively muted impact on long-run TFP growth as firms bid up the cost of R&D investments via a higher wage. In contrast, a size-dependent tax/subsidy scheme that only neutralizes the dispersion in markups achieves slightly faster TFP growth than the baseline intervention. This exercise revealed that it is the dispersion in markups, rather than their average level, that is more distortionary for long-run economic growth.

To conclude, we emphasize that heterogeneous markups serve as just one illustration of a distortion that differentially restricts firms' operational scale and therefore their incentives to improve their technology. [Hsieh and Klenow \(2009\)](#) document that such distortions (e.g. size-dependent taxes and regulations, financial frictions or political connections) are plausibly large and pervasive, and [Baqae and Farhi \(2019\)](#) show that input-output linkages substantially amplify their consequences on allocative efficiency. In light of this, one can't help but surmise that the real world might be riddled with such impediments to long-run economic growth that await further study.

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A Theoretical Appendix

This section of the Appendix provides derivations, proofs and extensions for the results presented in Section 3. Appendix A.1 provides derivations, Appendix A.2 provides proofs and Appendix A.3 discusses several extensions to our framework.

A.1 Derivations

A.1.1 The Kolmogorov Forward Equation

Denoting a firm's age at time t by a_t , the aging process of the firm is simply given by:

$$da_t = dt.$$

Denote the cumulative measure of firms with relative productivity below \hat{z} and age below a at time t by $M_t(\hat{z}, a)$ such that we have the following definitions:

$$M_t(\hat{z}) \equiv \lim_{a \rightarrow \infty} M_t(\hat{z}, a), \quad M_t(a) \equiv \lim_{\hat{z} \rightarrow \infty} M_t(\hat{z}, a) \quad \text{and} \quad \lim_{\hat{z} \rightarrow \infty} M_t(\hat{z}) = \lim_{a \rightarrow \infty} M_t(a) = M_t$$

as well as the following boundary conditions:

$$\begin{aligned} M_t(0, a) &= 0 \quad \forall a, \\ \partial_{\hat{z}} M_t(0, a) &= 0 \quad \forall a, \\ \lim_{\hat{z} \rightarrow \infty} \partial_{\hat{z}} M_t(\hat{z}, a) &= 0 \quad \forall a, \\ \lim_{\hat{z} \rightarrow \infty} \partial_{\hat{z}\hat{z}} M_t(\hat{z}, a) &= 0 \quad \forall a, \\ M_t(\hat{z}, 0) &= 0 \quad \forall \hat{z}, \\ \partial_a M_t(\hat{z}, 0) &= E_t F_t^E(\hat{z}) \quad \forall \hat{z}, \\ \lim_{a \rightarrow \infty} \partial_a M_t(\hat{z}, a) &= 0 \quad \forall \hat{z} \end{aligned}$$

where ∂_x denotes the partial derivative operator with respect to a variable x while ∂_{xy} denotes the cross partial derivative operator with respect to variables x and y . Then, we can extend the KFE in Section 3.1 to:

$$\begin{aligned} \dot{M}_t(\hat{z}, a) &= -[\gamma_t(\hat{z}) - g_t] \partial_{\hat{z}} M_t(\hat{z}, a) - \partial_a M_t(\hat{z}, a) + (\sigma^2/2) [\partial_{\hat{z}\hat{z}} M_t(\hat{z}, a) - \partial_{\hat{z}\hat{z}} M_t(0, a)] \\ &\quad + E_t F_t^E(\hat{z}) - \chi M_t(\hat{z}, a) \end{aligned}$$

for all $\hat{z} > 0$ and $a > 0$. By the same argument as in Section 3.1, the KFE of the CDF $F_t(\hat{z}, a) \equiv M_t(\hat{z}, a)/M_t$ is given by:

$$\begin{aligned}\dot{F}_t(\hat{z}, a) = & -[\gamma_t(\hat{z}) - g_t]\partial_{\hat{z}}F_t(\hat{z}, a) - \partial_a F_t(\hat{z}, a) + (\sigma^2/2)[\partial_{\hat{z}\hat{z}}F_t(\hat{z}, a) - \partial_{\hat{z}\hat{z}}F_t(0, a)] \\ & + e_t[F_t^E(\hat{z}) - F_t(\hat{z}, a)] + (\sigma^2/2)F_t(\hat{z}, a)F_t''(0)\end{aligned}$$

where $F_t(\hat{z})$ and $F_t(a)$ are defined analogously as $M_t(\hat{z})$ and $M_t(a)$ and where $F_t''(0)$ is defined over relative productivity. It is straightforward to see that by taking the limit of this equation as $a \rightarrow \infty$, we recover equation (8). Taking its limit as $\hat{z} \rightarrow \infty$, we instead obtain the law of motion for the cumulative density of firms below age a :

$$\dot{F}_t(a) = -\partial_a F_t(a) - (\sigma^2/2)\partial_{\hat{z}\hat{z}}F_t(0, a) + e_t[1 - F_t(a)] + (\sigma^2/2)F_t(a)F_t''(0)$$

where $F_t''(0)$ is here again defined over relative productivity

A.1.2 The Household's Problem

Taking prices as given, the household's problem is to choose its consumption and hours worked to maximize lifetime utility subject to a flow budget constraint:

$$\max_{\{C_t, H_t\}_{t \geq 0}} \int_0^\infty e^{-\rho t} [\ln(C_t) - v(H_t)] dt \quad \text{s.t.} \quad \dot{A}_t = r_t A_t + w_t H_t - C_t.$$

Reformulating the Household's problem using the current-value Hamiltonian, we have:

$$\mathcal{H}_t(C_t, H_t, A_t, \nu_t) = \ln(C_t) - v(H_t) + \nu_t(r_t A_t + w_t H_t - C_t)$$

where ν_t denotes the costate variable. The first-order conditions are:

$$\begin{aligned}\frac{\partial \mathcal{H}_t}{\partial C_t} &= 1/C_t - \nu_t = 0, \\ \frac{\partial \mathcal{H}_t}{\partial H_t} &= -v'(H_t) + \nu_t w_t = 0, \\ \frac{\partial \mathcal{H}_t}{\partial A_t} &= \nu_t r_t = \nu_t \rho - \dot{\nu}_t\end{aligned}$$

together with the No-Ponzi and transversality conditions, which jointly imply:

$$\lim_{t \rightarrow \infty} e^{-\int_0^t r_{t'} dt'} A_t = 0$$

Combining these equations, we obtain the usual intertemporal Euler equation and static first-order condition:

$$\frac{\dot{C}_t}{C_t} = r_t - \rho \quad \text{and} \quad v'(H_t)C_t = w_t.$$

A.1.3 The Final Sector's Problem

Taking prices as given, the final sector's problem is to choose its relative demand for each variety to maximize profits in each period:

$$\max_{\{\hat{y}_t(\hat{z})\}_{\hat{z}=0}^{\infty}} \left\{ P_t - M_t \int_0^{\infty} p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}) \right\} Y_t \quad \text{s.t.} \quad M_t \int_0^{\infty} \Upsilon(\hat{y}_t(\hat{z})) dF_t(\hat{z}) = \kappa.$$

Reformulating the final sector's problem as a cost-minimization problem subject to the [Kimball \(1995\)](#) aggregator constraint using the Lagrangian, we have:

$$\mathcal{L}_t(\{\hat{y}_t(\hat{z})\}_{\hat{z}=0}^{\infty}, \nu_t) = M_t Y_t \int_0^{\infty} p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}) + \nu_t \left(M_t \int_0^{\infty} \Upsilon(\hat{y}_t(\hat{z})) dF_t(\hat{z}) - 1 \right)$$

where ν_t now denotes the Lagrange multiplier. The first-order conditions are:

$$p_t(\hat{z}) = \nu_t \Upsilon'(\hat{y}_t(\hat{z})) / Y_t \quad \text{and} \quad M_t \int_0^{\infty} \Upsilon(\hat{y}_t(\hat{z})) dF_t(\hat{z}) = \kappa.$$

Since the final sector is perfectly competitive and makes no profit, we have:

$$P_t = M_t \int_0^{\infty} p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}).$$

Substituting in the first-order conditions, we obtain a solution for ν_t :

$$\nu_t = P_t Y_t D_t \quad \text{where} \quad D_t \equiv \left(M_t \int_0^{\infty} \Upsilon'(\hat{y}_t(\hat{z})) \hat{y}_t(\hat{z}) dF_t(\hat{z}) \right)^{-1}.$$

This delivers the following inverse demand functions:

$$p_t(\hat{z}) = \Upsilon'(\hat{y}_t(\hat{z})) P_t D_t.$$

A.1.4 The Firm's Static Problem

Firms engage in monopolistic competition in the product market but perfect competition in the input markets. That is, a firm chooses the price at which to sell its variety as well

as its demand for physical capital and production labor to maximize profits in each period. The firm takes as given the demand for its variety, the rental rate of capital r_t , the wage rate w_t and any transfer $T_t(\hat{y}_t(\hat{z}))$, which delivers the following problem:

$$\pi_t(\hat{z}) = \max_{p_t(\hat{z}), k_t(\hat{z}), l_t(\hat{z})} \{p_t(\hat{z})y_t(\hat{z}) - (r_t + \delta)k_t(\hat{z}) - w_t l_t(\hat{z}) + T_t(\hat{y}_t(\hat{z}))\} - w_t c_O$$

subject to the inverse demand function $p_t(\hat{z}) = \Upsilon'(\hat{y}_t(\hat{z}))D_t$. Let us first consider the sub-problem of optimally choosing the demand for capital and labor, which can be reformulated as a cost-minimization problem. Using the Lagrangian, we have:

$$\mathcal{L}_t(k_t(\hat{z}), l_t(\hat{z}), v_t(\hat{z})) = (r_t + \delta)k_t(\hat{z}) + w_t l_t(\hat{z}) + v_t(\hat{z})[y_t(\hat{z}) - \exp(\hat{z} + \underline{z}_t)k_t(\hat{z})^\alpha l_t(\hat{z})^{1-\alpha}]$$

where $v_t(\hat{z})$ denotes the Lagrange multiplier. The first-order conditions are:

$$\begin{aligned} k_t(\hat{z}) &= \frac{\alpha v_t(\hat{z}) y_t(\hat{z})}{r_t + \delta}, \\ l_t(\hat{z}) &= \frac{(1 - \alpha) v_t(\hat{z}) y_t(\hat{z})}{w_t}, \\ y_t(\hat{z}) &= \exp(\hat{z} + \underline{z}_t) k_t(\hat{z})^\alpha l_t(\hat{z})^{1-\alpha}. \end{aligned}$$

Solving for the Lagrange multiplier, we have:

$$v_t(\hat{z}) = \varsigma_t \exp(-\hat{z} - \underline{z}_t) \quad \text{where} \quad \varsigma_t \equiv \left(\frac{r_t + \delta}{\alpha} \right)^\alpha \left(\frac{w_t}{1 - \alpha} \right)^{1-\alpha}.$$

Therefore, we can rewrite the firm's static problem as:

$$\begin{aligned} \pi_t(\hat{z}) &= \max_{p_t(\hat{z})} \{[p_t(\hat{z}) - \varsigma_t \exp(-\hat{z} - \underline{z}_t)] \hat{y}_t(\hat{z}) Y_t + T_t(\hat{y}_t(\hat{z}))\} - w_t c_O \\ \text{s.t.} \quad p_t(\hat{z}) &= \Upsilon'(\hat{y}_t(\hat{z})) D_t. \end{aligned}$$

Reformulating it as a choice of $\hat{y}_t(\hat{z})$ given the inverse demand function $p_t(\hat{z})$, we have:

$$\pi_t(\hat{z}) = \max_{\hat{y}_t(\hat{z})} \{[\Upsilon'(\hat{y}_t(\hat{z})) D_t - \varsigma_t \exp(-\hat{z} - \underline{z}_t)] \hat{y}_t(\hat{z}) Y_t + T_t(\hat{y}_t(\hat{z}))\} - w_t c_O.$$

The first-order condition is:

$$[\Upsilon''(\hat{y}_t(\hat{z})) \hat{y}_t(\hat{z}) + \Upsilon'(\hat{y}_t(\hat{z}))] D_t + T'_t(\hat{y}_t(\hat{z})) / Y_t = \varsigma_t \exp(-\hat{z} - \underline{z}_t).$$

With $T_t(\hat{y}_t(\hat{z})) = 0$ for all \hat{z} , we can rearrange this expression as:

$$p_t(\hat{z}) = \frac{\mu_t(\hat{z})\zeta_t}{\exp(\hat{z} + \underline{z}_t)} \quad \text{where} \quad \mu_t(\hat{z}) \equiv \frac{\vartheta_t(\hat{z})}{\vartheta_t(\hat{z}) - 1}$$

and where $\vartheta_t(\hat{z})$ denotes the price elasticity of demand:

$$\vartheta_t(\hat{z}) \equiv -\frac{\Upsilon'(\hat{y}_t(\hat{z}))}{\Upsilon''(\hat{y}_t(\hat{z}))\hat{y}_t(\hat{z})} \in (1, \infty).$$

Substituting the monopoly pricing function in the profit function, we have:

$$\pi_t(\hat{z}) = \frac{p_t(\hat{z})\hat{y}_t(\hat{z})Y_t}{\vartheta_t(\hat{z})} - w_t c_O.$$

Denoting the firm's price relative to the choke price as $\hat{p}_t(\hat{z}) \equiv p_t(\hat{z})/\bar{p}_t$, we can rewrite:

$$\hat{p}_t(\hat{z}) = \frac{\mu(\hat{z})\zeta_t/\bar{p}_t}{\exp(\hat{z} + \underline{z}_t)} \quad \text{and} \quad \pi_t(\hat{z}) = \frac{\hat{p}_t(\hat{z})\hat{y}_t(\hat{z})\bar{p}_t Y_t}{\vartheta_t(\hat{z})} - w_t c_O.$$

With the transfers from equation (11), the first-order condition is instead:

$$\hat{p}_t(\hat{z}) = \frac{\zeta_t/\bar{p}_t}{\exp(\hat{z} + \underline{z}_t)}$$

such that all firms price at marginal cost. This implies that profits are simply given by:

$$\pi_t(\hat{z}) = [\Upsilon(\hat{y}_t(\hat{z})) - \Upsilon'(\hat{y}_t(\hat{z}))\hat{y}_t(\hat{z})]Y_t D_t - w_t c_O.$$

A.1.5 The Firm's Dynamic Problem

Given the above static profit function and taking the wage rate as given, firms control the drift of their productivity and choose an exit stopping time τ :

$$V_t(\hat{z}) = \max_{\tau, \{\gamma_s\}_{s \geq t}} \mathbb{E}_{\hat{z}} \left\{ \int_t^{t+\tau} e^{-\int_t^s (r_{t'} + \chi) dt'} [\pi_s(\hat{z}_s) - w_s i(\gamma_s, \hat{z}_s)] ds \right\}$$

where $\mathbb{E}_{\hat{z}}$ denotes the expectation operator with respect to the diffusion process $\{\hat{z}_s\}_{s \geq t}$ when its initial value is $\hat{z}_t = \hat{z}$. Consider a random walk approximation to this diffusion process. That is, within a discrete time interval of $\Delta > 0$, the firm's productivity may increase by $\sigma\sqrt{\Delta}$ with probability $p_t(\gamma) = [1 + (\gamma - g_t)\sqrt{\Delta}/\sigma]/2$ or decrease by the same increment with the complementary probability $q_t(\gamma) = 1 - p_t(\gamma)$. Then, the

recursive formulation of the firm's dynamic problem is described by:

$$V_t(\hat{z}) = \pi_t(\hat{z})\Delta + \max_{\gamma} \{ (1 - \chi\Delta)(1 - r_t\Delta)[p_t(\gamma) \max\{V_{t+\Delta}(\hat{z} + \sigma\sqrt{\Delta}), 0\} \\ + q_t(\gamma) \max\{V_{t+\Delta}(\hat{z} - \sigma\sqrt{\Delta}), 0\}] - w_t i(\gamma, \hat{z})\Delta \}.$$

At a productivity state such that $V_{t+\Delta}(\hat{z} - \sigma\sqrt{\Delta}) > 0$, the firm's value function satisfies:

$$V_t(\hat{z}) = \pi_t(\hat{z})\Delta + \max_{\gamma} \{ (1 - \chi\Delta)(1 - r_t\Delta)[p_t(\gamma)V_{t+\Delta}(\hat{z} + \sigma\sqrt{\Delta}) \\ + q_t(\gamma)V_{t+\Delta}(\hat{z} - \sigma\sqrt{\Delta})] - w_t i(\gamma, \hat{z})\Delta \}.$$

Up to a second-order approximation, we can rewrite:

$$V_t(\hat{z}) = \pi_t(\hat{z})\Delta + \max_{\gamma} \{ (1 - \chi\Delta)(1 - r_t\Delta)[V_t(\hat{z}) + \dot{V}_t(\hat{z})\Delta + \ddot{V}_t(\hat{z})\Delta^2/2 + \sigma^2 V_t''(\hat{z})\Delta/2 \\ + (2p_t(\gamma) - 1)\sigma V_t'(\hat{z})\sqrt{\Delta} + o(\Delta)] - w_t i(\gamma, \hat{z})\Delta \}$$

where a single and double dot above a function respectively denote its first and second partial derivatives with respect to time. Subtracting $(1 - \chi\Delta)(1 - r_t\Delta)V_t(\hat{z})$ from both sides, substituting in the expression for $p_t(\gamma)$, dividing both sides by Δ and then taking the limit as $\Delta \rightarrow 0$ delivers the following HJBE:

$$(r_t + \chi)V_t(\hat{z}) = \pi_t(\hat{z}) + \max_{\gamma} \{ (\gamma - g_t)V_t'(\hat{z}) - w_t i(\gamma, \hat{z}) \} + \sigma^2 V_t''(\hat{z})/2 + \dot{V}_t(\hat{z}).$$

As in [Stokey \(2009\)](#), the optimality conditions of the firm's dynamic problem are the value matching, smooth pasting and first-order conditions:

$$V_t(0) = 0, \quad V_t'(0) = 0 \quad \text{and} \quad V_t'(\hat{z}) = w_t \times \frac{\partial i(\gamma, \hat{z})}{\partial \gamma}$$

together with the “no bubble” condition:

$$\lim_{\hat{z} \rightarrow \infty} V_t(\hat{z}) = \lim_{\hat{z} \rightarrow \infty} \max_{\{\gamma_s\}_{s \geq t}} \mathbb{E}_{\hat{z}} \left\{ \int_t^{\infty} e^{-\int_t^s (r_{t'} + \chi) dt'} [\pi_s(\hat{z}_s) - w_s i(\gamma_s, \hat{z}_s)] ds \right\}.$$

A.1.6 Aggregation

By the definition of the final good's price index, we have:

$$1 = M_t \int_0^{\infty} p_t(\hat{z}) \hat{y}_t(\hat{z}) dF_t(\hat{z}).$$

Substituting in the monopoly pricing condition, we obtain:

$$1 = \varsigma_t M_t \int_0^\infty \mu_t(\hat{z}) \hat{y}_t(\hat{z}) \exp(-\hat{z} - \underline{z}_t) dF_t(\hat{z}).$$

Multiplying both sides by the definition of TFP Z_t :

$$Z_t = \varsigma_t \mathcal{M}_t \quad \text{where} \quad \mathcal{M}_t \equiv \frac{\int_0^\infty \mu_t(\hat{z}) \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}.$$

The definitions of aggregate physical capital and production labor demand are:

$$K_t = M_t \int_0^\infty k_t(\hat{z}) dF_t(\hat{z}) \quad \text{and} \quad L_t = M_t \int_0^\infty l_t(\hat{z}) dF_t(\hat{z}).$$

Substituting in the firm-level demand functions for physical capital and production labor derived in Appendix A.1.4, we obtain:

$$K_t = \frac{\alpha \varsigma_t Y_t}{(r_t + \delta) Z_t} = \frac{\alpha Y_t}{(r_t + \delta) \mathcal{M}_t} \quad \text{and} \quad L_t = \frac{(1 - \alpha) \varsigma_t Y_t}{w_t Z_t} = \frac{(1 - \alpha) Y_t}{w_t \mathcal{M}_t}.$$

Solving for aggregate output delivers:

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha}.$$

Under the transfers of equation (11), we simply have that $\mathcal{M}_t = 1$.

A.1.7 Transfer Schedules

In this subsection, we derive the transfers presented in Sections 5.1 and 5.4. Let us first rewrite the firm's static problem of choosing a scale at which to produce in order to maximize profits given a demand schedule:

$$\pi_t(\hat{z}) = \max_{\hat{y}_t(\hat{z})} \{ [\Upsilon'(\hat{y}_t(\hat{z})) D_t - \varsigma_t \exp(-\hat{z} - \underline{z}_t)] \hat{y}_t(\hat{z}) Y_t + T_t(\hat{y}_t(\hat{z})) \} - w_t c_O.$$

The first-order condition of this problem is:

$$[\Upsilon''(\hat{y}_t(\hat{z})) \hat{y}_t(\hat{z}) + \Upsilon'(\hat{y}_t(\hat{z}))] D_t + T'_t(\hat{y}_t(\hat{z})) / Y_t = \varsigma_t \exp(-\hat{z} - \underline{z}_t).$$

Let us now momentarily abandon the \hat{z} dependence notation for clarity. What transfer schedule would induce firms to price at marginal cost? For firms to do so, we must have:

$$p_t(\hat{y}) = \Upsilon'(\hat{y})D_t = \varsigma_t \exp(-\hat{z} - \underline{z}_t).$$

Substituting this constraint in the first-order condition, we obtain:

$$T'_t(\hat{y}) = -\Upsilon''(\hat{y})\hat{y}Y_tD_t.$$

This is an ordinary differential equation, which can be integrated by parts with the initial condition $T_t(0) = 0$:

$$T_t(\hat{y}) = -Y_tD_t \int \Upsilon''(\hat{y}')\hat{y}'d\hat{y}' + C_0 = [\Upsilon(\hat{y}) - \Upsilon'(\hat{y})\hat{y}]Y_tD_t.$$

Here C_0 is the constant of integration, which is equal to zero given the initial condition. If the transfers were instead intended to induce firms to set a markup of \mathcal{M}_t above marginal cost, we would substitute the constraint $\Upsilon'(\hat{y})D_t = \mathcal{M}_t\varsigma_t \exp(-\hat{z} - \underline{z}_t)$ in the first-order condition to obtain:

$$T'_t(\hat{y}) = [(\mathcal{M}_t^{-1} - 1)\Upsilon'(\hat{y}) - \Upsilon''(\hat{y})\hat{y}]Y_tD_t.$$

Integrating by parts with the same initial condition, we find:

$$T_t(\hat{y}) = [\mathcal{M}_t^{-1}\Upsilon(\hat{y}) - \Upsilon'(\hat{y})\hat{y}]Y_tD_t.$$

Finally, if the transfers were intended to induce firms to set a markup of $\mu(\hat{y})/\mathcal{M}_t$ above marginal cost, we would substitute the constraint $\Upsilon'(\hat{y})D_t = \mu(\hat{y})\varsigma_t \exp(-\hat{z} - \underline{z}_t)/\mathcal{M}_t$ in the first-order condition to obtain:⁴⁰

$$T'_t(\hat{y}) = (\mathcal{M}_t - 1)[\Upsilon'(\hat{y}) + \Upsilon''(\hat{y})\hat{y}]Y_tD_t.$$

Integrating by parts with the same initial condition, we find:

$$T_t(\hat{y}) = (\mathcal{M}_t - 1)\Upsilon'(\hat{y})\hat{y}Y_tD_t.$$

⁴⁰Here, $\mu(\hat{y}) = \Upsilon'(\hat{y})/[\Upsilon'(\hat{y}) + \Upsilon''(\hat{y})\hat{y}]$.

A.1.8 Relative Prices and Demand Under Klenow and Willis (2016)

From the final sector's problem and the firm's static problem, we have the following relative demand function and monopoly pricing condition for varieties:

$$\hat{y}_t(\hat{z}) = \Upsilon'^{-1}[p_t(\hat{z})/D_t] \quad \text{and} \quad p_t(\hat{z}) = \frac{\mu_t(\hat{z})\zeta_t}{\exp(\hat{z} + \bar{z}_t)}$$

where we have the following two definitions:

$$\mu_t(\hat{z}) \equiv \frac{\vartheta_t(\hat{z})}{\vartheta_t(\hat{z}) - 1} \quad \text{and} \quad \vartheta_t(\hat{z}) \equiv -\frac{\Upsilon'(\hat{y}_t(\hat{z}))}{\Upsilon''(\hat{y}_t(\hat{z}))\hat{y}_t(\hat{z})}.$$

Denoting the choke price by $\bar{p}_t \equiv \Upsilon'(0)D_t$, we can define a variety's relative price as $\hat{p}_t(\hat{z}) \equiv p_t(\hat{z})/\bar{p}_t$. This allows us to rewrite the relative demand function as:

$$\hat{y}(\hat{p}) = \Upsilon'^{-1}[\hat{p}\Upsilon'(0)]$$

which is now a stationary function of the corresponding variety's relative price. Using the [Klenow and Willis \(2016\)](#) specification of the [Kimball \(1995\)](#) aggregator, we can rewrite the monopoly pricing condition (relative to the choke price) as:

$$\hat{p}_t(\hat{z}) = \frac{\zeta_t/\bar{p}_t}{[1 + (\epsilon/\theta) \ln(\hat{p}_t(\hat{z}))] \exp(\hat{z} + \bar{z}_t)}$$

Note that we can rearrange this equation to obtain:

$$\exp\{(\theta/\epsilon)[\zeta_t[\hat{p}_t(\hat{z})\bar{p}_t \exp(\hat{z} + \bar{z}_t)]^{-1} - 1]\}\hat{p}_t(\hat{z})^{-1} = 1$$

Multiplying both sides by $(\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \bar{z}_t) \zeta_t/\bar{p}_t$, we have:

$$W^{-1}\{(\theta/\epsilon)\zeta_t[\hat{p}_t(\hat{z})\bar{p}_t \exp(\hat{z} + \bar{z}_t)]^{-1}\} = (\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \bar{z}_t) \zeta_t/\bar{p}_t$$

where W is the [Lambert W-function](#) defined by the inverse mapping $W^{-1}(x) = xe^x$. Finally, solving for $\hat{p}_t(\hat{z})$ delivers:

$$\hat{p}_t(\hat{z}) = \frac{(\theta/\epsilon) \exp(-\hat{z} - \bar{z}_t) \zeta_t/\bar{p}_t}{W[(\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \bar{z}_t) \zeta_t/\bar{p}_t]}.$$

The Lambert W -function has two useful properties:

1. $W(x) > 0$ for all $x > 0$ and $\lim_{x \rightarrow 0} W(x) = 0$.

2. $W'(x) = [x + e^{W(x)}]^{-1}$ for all $x > 0$ and $\lim_{x \rightarrow 0} W'(x) = 1$.

We can use those properties to study the limiting behavior of a variety's relative price and demand for the most productive firms:

Proposition 6. *At any given point in time, a variety's relative price and demand asymptote to constants as $\hat{z} \rightarrow \infty$:*

$$\lim_{\hat{z} \rightarrow \infty} \hat{p}_t(\hat{z}) = \exp(-\theta/\epsilon) \quad \text{and} \quad \lim_{\hat{z} \rightarrow \infty} \hat{y}(\hat{z}) = \theta^{\theta/\epsilon}.$$

Let us now consider the transfer schedule of equation (13):

$$T_t(\hat{y}) = [\tau_0 \Upsilon(\hat{y}) + \tau_1 \Upsilon'(\hat{y}) \hat{y}] Y_t D_t.$$

In particular, we will explore the three cases derived in the previous subsection of this appendix: $(\tau_0, \tau_1) \in \{(1, -1), (0, \mathcal{M}_t - 1), (\mathcal{M}_t^{-1}, -1)\}$. In the first case, the firm's optimally chosen relative price is:

$$\hat{p}_t(\hat{z}) = \frac{\varsigma_t / \bar{p}_t}{\exp(\hat{z} + \underline{z}_t)}$$

which asymptotes to zero as $\hat{z} \rightarrow \infty$. Therefore, the firm's relative demand asymptotes to infinity as $\hat{z} \rightarrow \infty$. However, firm-level profits remain finite at any given point in time as $\hat{z} \rightarrow \infty$. Indeed, we have that:

$$\lim_{\hat{z} \rightarrow \infty} \pi_t(\hat{z}) = \left[1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon - 1} \Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) \right] Y_t D_t - w_t c_O.$$

In the second case, the firm's optimally chosen relative price is:

$$\hat{p}_t(\hat{z}) = \frac{\mu(\hat{z}) \varsigma_t / \bar{p}_t}{\mathcal{M}_t \exp(\hat{z} + \underline{z}_t)}.$$

With the same derivation as above, we can express that relative price function as:

$$\hat{p}_t(\hat{z}) = \frac{(\theta/\epsilon) \exp(-\hat{z} - \underline{z}_t) \varsigma_t / (\bar{p}_t \mathcal{M}_t)}{W[(\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \underline{z}_t) \varsigma_t / (\bar{p}_t \mathcal{M}_t)]}$$

which holds the same asymptotic properties as in Proposition 6. In the third case, the firm's optimally chosen relative price is:

$$\hat{p}_t(\hat{z}) = \frac{\mathcal{M}_t \varsigma_t / \bar{p}_t}{\exp(\hat{z} + \underline{z}_t)}$$

which asymptotes to zero as $\hat{z} \rightarrow \infty$. Therefore, the firm's relative demand asymptotes to infinity as $\hat{z} \rightarrow \infty$. However, firm-level profits remain finite at any given point in time as $\hat{z} \rightarrow \infty$. Indeed, we have that:

$$\lim_{\hat{z} \rightarrow \infty} \pi_t(\hat{z}) = \left[1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon - 1} \Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) \right] Y_t D_t / \mathcal{M}_t - w_t c_O.$$

A.1.9 Computing a Balanced Growth Path Equilibrium Allocation

Consider a balanced growth path equilibrium allocation as defined in Definition 2 with initial condition $\underline{z}_0 = 0$. Using the household's Euler equation together with the restriction that the value of a firm must grow at the same rate as aggregate consumption, we have the following expression for the firm's HJBE in the continuation region:

$$(\rho + \chi)V_t(\hat{z}) = \pi_t(\hat{z}) + \max_{\gamma} \{(\gamma - g)V'_t(\hat{z}) - w_t i(\gamma, \hat{z})\} + \sigma^2 V''_t(\hat{z})/2.$$

Here, g still denotes the instantaneous growth rate of TFP, which is constant on a balanced growth path. Profits and the wage rate both grow at constant rate $g/(1 - \alpha)$ such that we can define the stationary function $V(\hat{z}) \equiv V_t(\hat{z}) \exp[-gt/(1 - \alpha)]$ and rewrite the firm's HJBE as:

$$(\rho + \chi)V(\hat{z}) = \pi_0(\hat{z}) + \max_{\gamma} \{(\gamma - g)V'(\hat{z}) - w_0 i(\gamma, \hat{z})\} + \sigma^2 V''(\hat{z})/2$$

where X_0 denotes the detrended value of a variable X_t that grows at a constant rate on a balanced growth path. As in Appendix A.1.5, the firm's dynamic problem delivers the value matching, smooth pasting and first-order conditions:

$$V(0) = 0, \quad V'(0) = 0 \quad \text{and} \quad \gamma(\hat{z}) = \left[\frac{V'(\hat{z})}{w_0 \exp(c_I + (1 + \zeta)\hat{z})} \right]^{\frac{1}{\zeta}}.$$

Substituting this first-order condition in the firm's stationary HJBE, we obtain a second-order nonlinear ordinary differential equation:

$$(\rho + \chi)V(\hat{z}) = \pi_0(\hat{z}) + \frac{\zeta \gamma(\hat{z}) V'(\hat{z})}{1 + \zeta} - g V'(\hat{z}) + \frac{\sigma^2 V''(\hat{z})}{2}$$

in which the stationary profit function is given by:

$$\pi_0(\hat{z}) = \begin{cases} \hat{p}(\hat{z})\hat{y}(\hat{z})^{1+\epsilon/\theta}\Upsilon'(0)Y_0D/\theta - w_0c_O & \text{Pre-policy,} \\ [\Upsilon(\hat{y}(\hat{z})) - \Upsilon'(\hat{y}(\hat{z}))\hat{y}(\hat{z})]Y_0D - w_0c_O & \text{Post-policy,} \end{cases}$$

and the relative demand and relative price functions are in turn given by:

$$\begin{aligned} \hat{p}(\hat{z}) &= \begin{cases} \frac{(\theta/\epsilon)\exp(-\hat{z}-\underline{z}_0)\zeta_0/\bar{p}}{W[(\theta/\epsilon)\exp(\theta/\epsilon-\hat{z}-\underline{z}_0)\zeta_0/\bar{p}]} & \text{Pre-policy,} \\ \exp(-\hat{z}-\underline{z}_0)\zeta_0/\bar{p} & \text{Post-policy,} \end{cases} \\ \hat{y}(\hat{z}) &= [-\epsilon \max\{\ln(\hat{p}(\hat{z})), 0\}]^{\theta/\epsilon}. \end{aligned}$$

It is straightforward to verify that Y_t must be growing at the same rate as w_t , and ζ_t must be growing at the same rate as \underline{z}_t on a balanced growth path. From Proposition 5, we know that the stationary value function asymptotes to an endogenous constant \bar{V} :

$$\bar{V} = \frac{w_0c_O(1-x)}{(\rho+\chi)x} \quad \text{where} \quad x \equiv \frac{w_0c_O}{\bar{\pi}Y_0D} \in (0,1)$$

where the constant $\bar{\pi}$ is given by:

$$\bar{\pi} = \begin{cases} (\theta-1)\exp[(1-\theta)/\epsilon]\theta^{\theta/\epsilon-1} & \text{Pre-policy,} \\ 1 + (\theta-1)\exp(1/\epsilon)\epsilon^{\theta/\epsilon-1}\Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) & \text{Post-policy.} \end{cases}$$

Therefore, let us respectively define the firm's *normalized* value and profit functions as $\hat{V}(\hat{z}) \equiv V(\hat{z})/\bar{V}$ and $\hat{\pi}_0(\hat{z}) \equiv \pi_0(\hat{z})/\bar{V}$ such that we can rewrite the firm's HJBE as:

$$(\rho+\chi)\hat{V}(\hat{z}) = \hat{\pi}_0(\hat{z}) + \frac{\zeta\gamma(\hat{z})\hat{V}'(\hat{z})}{1+\zeta} - g\hat{V}'(\hat{z}) + \frac{\sigma^2\hat{V}''(\hat{z})}{2}$$

with boundary conditions $\hat{V}(0) = 0$ and $\lim_{\hat{z} \rightarrow \infty} \hat{V}(\hat{z}) = 1$, and where the first-order condition of the firm's dynamic problem becomes:

$$\gamma(\hat{z}) = \left[\frac{c_O(1-x)\hat{V}'(\hat{z})}{(\rho+\chi)\exp(c_I + (1+\zeta)\hat{z})x} \right]^{\frac{1}{\zeta}}.$$

Since the measure of varieties is constant on a balanced growth path, the entry rate must be equated to the sum of the exogenous and endogenous exit rates:

$$e = \chi + (\sigma^2/2)F''(0).$$

The stationary Kolmogorov forward equation of the CDF $F(\hat{z})$ therefore delivers the following second-order nonlinear ordinary differential equation:

$$0 = -[\gamma(\hat{z}) - g]F'(\hat{z}) + (\sigma^2/2)\{F''(\hat{z}) - F''(0)[1 - F(\hat{z})]\} + e[F^E(\hat{z}) - F(\hat{z})]$$

with boundary conditions $F(0) = 0$ and $\lim_{\hat{z} \rightarrow \infty} F(\hat{z}) = 1$, and where the growth rate g (derived in Appendix A.2) is given by:

$$\begin{aligned} g = & \frac{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})\gamma(\hat{z})dF(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} \\ & - \frac{(\sigma^2/2) \int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF'(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} \\ & + \frac{(\sigma^2/2)F''(0)[\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF^E(\hat{z}) - \hat{y}(0)]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})} \\ & - \frac{\chi[\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z}) - \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z})dF^E(\hat{z})]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1]\hat{y}(\hat{z}) \exp(-\hat{z})dF(\hat{z})}. \end{aligned}$$

Here, $\vartheta(\hat{z})$ and $\varrho(\hat{z})$ are both functions of $\hat{p}(\hat{z})$ and parameters only.

To solve for the two second-order nonlinear ordinary differential equations above (the firm's HJBE and the KFE), we need equilibrium conditions that pin down the aggregate variables upon which they depend. More precisely, given an initial condition \underline{z}_0 , we are looking for ten equations to identify the unknowns $\{Y_0, C_0, Z_0, w_0, \varsigma_0, r, \bar{p}, D, M, H\}$:

1. The **Kimball (1995)** aggregation condition:

$$M \int_0^\infty \Upsilon(\hat{y}(\hat{z}))dF(\hat{z}) = \kappa.$$

2. The demand index:

$$D = \left(M \int_0^\infty \Upsilon'(\hat{y}(\hat{z}))\hat{y}(\hat{z})dF(\hat{z}) \right)^{-1}.$$

3. The free-entry condition:

$$(1 - x) \int_0^\infty \hat{V}(\hat{z})dF^E(\hat{z}) = (c_E/c_O)(\rho + \chi)x.$$

4. The producer price index:

$$\zeta_0 = \left(\frac{r + \delta}{\alpha} \right)^\alpha \left(\frac{w_0}{1 - \alpha} \right)^{1 - \alpha}.$$

5. The household's Euler equation:

$$\frac{g}{1 - \alpha} = r - \rho.$$

6. The household's static first-order condition:

$$\beta H^\eta C_0 = w_0.$$

7. The final good market clearing condition:

$$C_0 = Y_0 \left[1 - \frac{\alpha \zeta_0 (r + \delta - \rho)}{Z_0 (r + \delta)} \right].$$

8. The labor market clearing condition:

$$\frac{(1 - \alpha) \zeta_0 Y_0}{w_0 Z_0} + \frac{M c_O (1 - x) \int_0^\infty \gamma(\hat{z}) \hat{V}'(\hat{z}) dF(\hat{z})}{(1 + \zeta)(\rho + \chi)x} + c_E e M + c_O M = H.$$

9. Total-factor productivity:

$$Z_0 = \exp(\underline{z}_0) \left(M \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z}) \right)^{-1}.$$

10. The choke price:

$$\bar{p} = \Upsilon'(0)D.$$

A.2 Proofs

Proof of Propositions 1 and 2. The monopolist's profits are given by:

$$\pi(z) = \frac{p(z)y(p(z))}{\vartheta(p(z))}.$$

The elasticity of profits with respect to productivity is given by:

$$\frac{\partial \ln(\pi(z))}{\partial z} = [\vartheta(p(z)) + \varepsilon(p(z)) - 1]\varrho(z) \quad \text{where} \quad \varrho(z) \equiv -\frac{\partial \ln(p(z))}{\partial z}$$

where $\varrho(z)$ denotes the productivity “pass-through”. Using the expression provided in Section 2 for the profit-maximizing price, we can rewrite $\varrho(z)$ as:

$$\varrho(z) = \frac{\vartheta(p(z)) - 1}{\vartheta(p(z)) + \varepsilon(p(z)) - 1}.$$

Substituting this result in the previous expression, we have:

$$\frac{\partial \ln(\pi(z))}{\partial z} = \vartheta(p(z)) - 1$$

such that the partial derivative of profits with respect to productivity is given by:

$$\pi'(z) = y(p(z)) \exp(-z).$$

The social surplus is instead given by:

$$S(z) = \int_{c(z)}^{\bar{p}} y(p) dp.$$

Therefore, its partial derivative with respect to productivity is given by:

$$S'(z) = y(\exp(-z)) \exp(-z).$$

This shows that the ratio $R(z) \equiv \pi'(z)/S'(z)$ is simply given by the output ratio of the monopolist and the welfare-maximizing agent, which completes the proof for the first part of Proposition 1. Taking the elasticity of this ratio with respect to productivity and using the expression derived above for the monopolist’s productivity pass-through completes the proof of Proposition 2. To prove the second part of Proposition 1, define consumer surplus as:

$$C(z) \equiv \int_{p(z)}^{\bar{p}} y(p) dp.$$

Its partial derivative with respect to productivity is given by:

$$\begin{aligned} C'(z) &= -\frac{\partial p(z)}{\partial z} \times y(p(z)) \\ &= \varrho(z)p(z)y(p(z)) \\ &= \frac{\vartheta(p(z))}{\vartheta(p(z)) + \varepsilon(p(z)) - 1} \times \pi'(z). \end{aligned}$$

Therefore, the ratio of marginal producer surplus to the sum of marginal consumer and producer surplus is:

$$\frac{\pi'(z)}{C'(z) + \pi'(z)} = \frac{\vartheta(p(z)) + \varepsilon(p(z)) - 1}{2\vartheta(p(z)) + \varepsilon(p(z)) - 1}$$

which completes the proof for the second part of Proposition 1. \square

Proof of Proposition 3. Let us start with the expression derived in Appendix A.1.6 for aggregate output:

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha} \quad \text{where} \quad Z_t \equiv \left(M_t \int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z} - \underline{z}_t) dF_t(\hat{z}) \right)^{-1}.$$

On a balanced growth path, L_t and M_t are constant, and Y_t and K_t grow at the same rate. This implies that aggregate output grows at rate $g/(1-\alpha)$ where g is the constant growth rate of the endogenous exit threshold, which must grow at the same rate as TFP on a balanced growth path. To derive an expression for g , let us differentiate the logarithm of Z_t with respect to time, to obtain:

$$\frac{\dot{Z}_t}{Z_t} = g_t - \frac{\dot{M}_t}{M_t} + \frac{(\dot{\zeta}_t/\zeta_t - g_t) \int_0^\infty \vartheta_t(\hat{z}) \varrho_t(\hat{z}) \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})} - \frac{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) \dot{F}_t'(\hat{z}) d\hat{z}}{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})}.$$

Here g_t and ζ_t still denote the rate of change of the endogenous exit threshold and the producer price index, respectively. On a balanced growth path, since the measure of varieties is stationary and the producer price index grows at the same constant rate as the endogenous exit threshold, we obtain:

$$\frac{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) \dot{F}_t'(\hat{z}) d\hat{z}}{\int_0^\infty \hat{y}_t(\hat{z}) \exp(-\hat{z}) dF_t(\hat{z})} = 0.$$

Substituting in the Kolmogorov forward equation of $F_t(\hat{z})$ and abandoning the time

subscripts, we have:

$$0 = -\chi \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z}) + e \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF^E(\hat{z}) - \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) \frac{\partial[\gamma(\hat{z})F'(\hat{z})]}{\partial \hat{z}} d\hat{z} \quad (\text{A.1})$$

$$+ g \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) F''(\hat{z}) d\hat{z} \quad (\text{A.2})$$

$$+ (\sigma^2/2) \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) F'''(\hat{z}) d\hat{z} \quad (\text{A.3})$$

Now, let us consider each of the three last terms above separately. First, using integration by parts, the term (A.1) can be rewritten as:

$$(\text{A.1}) = -[\hat{y}(\hat{z}) \exp(-\hat{z}) \gamma(\hat{z}) F'(\hat{z})]_{\hat{z}=0}^\infty + \int_0^\infty [\vartheta(\hat{z}) \varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) \gamma(\hat{z}) dF(\hat{z})$$

where $\varrho(\hat{z})$ denotes the productivity “pass-through”:

$$\varrho(\hat{z}) \equiv -\frac{\partial \ln(\hat{p}(\hat{z}))}{\partial \hat{z}}.$$

Similarly, the term (A.2) can be rewritten as:

$$(\text{A.2}) = g[\hat{y}(\hat{z}) \exp(-\hat{z}) F'(\hat{z})]_{\hat{z}=0}^\infty - g \int_0^\infty [\vartheta(\hat{z}) \varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z}).$$

Notice that the boundary conditions $\lim_{\hat{z} \rightarrow \infty} F'(\hat{z}) = F'(0) = 0$, the smooth pasting condition (implying that $\gamma(0) = 0$ by the firm’s dynamic first-order condition), the limit $\lim_{\hat{z} \rightarrow \infty} \hat{y}(\hat{z}) = \theta^{\theta/\epsilon}$ (proved in Proposition 6) and the limit $\lim_{\hat{z} \rightarrow \infty} \gamma(\hat{z}) = 0$ (derived in Appendix A.1.9) imply that:

$$[\hat{y}(\hat{z}) \exp(-\hat{z}) F'(\hat{z})]_{\hat{z}=0}^\infty = [\hat{y}(\hat{z}) \exp(-\hat{z}) \gamma(\hat{z}) F'(\hat{z})]_{\hat{z}=0}^\infty = 0$$

Using integration by parts once again, the term (A.3) can be rewritten as:

$$(\text{A.3}) = (\sigma^2/2)[\hat{y}(\hat{z}) \exp(-\hat{z}) F''(\hat{z})]_{\hat{z}=0}^\infty - (\sigma^2/2) \int_0^\infty [\vartheta(\hat{z}) \varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF'(\hat{z}).$$

Notice here again that the boundary condition $\lim_{\hat{z} \rightarrow \infty} F''(\hat{z}) = 0$ together with the limit $\lim_{\hat{z} \rightarrow \infty} \hat{y}(\hat{z}) = \theta^{\theta/\epsilon}$ (proved in Proposition 6) imply that:

$$[\hat{y}(\hat{z}) \exp(-\hat{z}) F''(\hat{z})]_{\hat{z}=0}^\infty = -\hat{y}(0) F''(0).$$

Collecting all terms above and using the fact that the measure of varieties is constant on a balanced growth path (implying that the entry rate is equal to the sum of the endogenous and exogenous exit rates), we can solve for g to obtain the expression:

$$\begin{aligned}
g = & \frac{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) \gamma(\hat{z}) dF(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z})} \\
& - \frac{(\sigma^2/2) \int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF'(\hat{z})}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z})} \\
& + \frac{(\sigma^2/2) F''(0) [\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF^E(\hat{z}) - \hat{y}(0)]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z})} \\
& - \frac{\chi [\int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z}) - \int_0^\infty \hat{y}(\hat{z}) \exp(-\hat{z}) dF^E(\hat{z})]}{\int_0^\infty [\vartheta(\hat{z})\varrho(\hat{z}) - 1] \hat{y}(\hat{z}) \exp(-\hat{z}) dF(\hat{z})}.
\end{aligned}$$

□

Proof of Proposition 4. We want to prove that on a balanced growth path equilibrium allocation:

$$\lim_{\hat{z} \rightarrow \infty} F(\hat{z}) = 1 - \exp(-\lambda \hat{z}).$$

To do so, let us guess that as $\hat{z} \rightarrow \infty$:

$$\begin{aligned}
F(\hat{z}) &= 1 - \exp(-\lambda \hat{z}), \\
F'(\hat{z}) &= \lambda \exp(-\lambda \hat{z}), \\
F''(\hat{z}) &= -\lambda^2 \exp(-\lambda \hat{z})
\end{aligned}$$

for $\lambda > 0$. Substituting these guesses in the stationary KFE derived in Appendix A.1.9 and cancelling terms:

$$0 = [g - \gamma(\hat{z})]\lambda + (\sigma^2 \lambda^2 / 2 - \chi) \left\{ \frac{T[\exp(-\lambda \hat{z})]}{\exp(-\lambda \hat{z})} - 1 \right\}.$$

Taking the limit of this equation and using the assumptions that $\lim_{\hat{z} \rightarrow \infty} \gamma(\hat{z}) = \bar{\gamma} < \infty$ and $\lim_{\hat{z} \rightarrow \infty} \frac{1 - F_t^E(\hat{z})}{1 - F_t(\hat{z})} = 0$, we obtain the quadratic equation:

$$0 = \sigma^2 \lambda^2 / 2 - (g - \bar{\gamma})\lambda - \chi.$$

The positive root of this quadratic equation is:

$$\lambda = \frac{g - \bar{\gamma} + \sqrt{(g - \bar{\gamma})^2 + 2\chi\sigma^2}}{\sigma^2}.$$

For that root to be greater than one, we have the additional restriction:

$$g > \bar{\gamma} + \sigma^2/2 - \chi$$

which implies that the distribution of $\exp(\hat{z})$ is Pareto with a finite mean. \square

Proof of Proposition 5. Using the household's Euler equation together with the restriction that the value of a firm must grow at the same rate as aggregate output on a balanced growth path, we have the following expression for the firm's HJBE:

$$(\rho + \chi)V_t(\hat{z}) = \pi_t(\hat{z}) + \max_{\gamma} \{(\gamma - g)V'_t(\hat{z}) - w_t i(\gamma, \hat{z})\} + \sigma^2 V''_t(\hat{z})/2$$

with value matching, smooth pasting and first-order conditions:

$$V_t(0) = 0, \quad V'_t(0) = 0 \quad \text{and} \quad \gamma_t(\hat{z}) = \left[\frac{V'_t(\hat{z})}{w_t \exp(c_I + (1 + \zeta)\hat{z})} \right]^{1/\zeta}.$$

Substituting the first-order condition in the firm's HJBE, we obtain the second-order nonlinear ordinary differential equation:

$$(\rho + \chi)V_t(\hat{z}) = \pi_t(\hat{z}) + \frac{\zeta \gamma_t(\hat{z}) V'_t(\hat{z})}{1 + \zeta} - g V'_t(\hat{z}) + \frac{\sigma^2 V''_t(\hat{z})}{2}$$

in which the profit function is given by:

$$\pi_t(\hat{z}) = \begin{cases} \hat{p}_t(\hat{z}) \hat{y}_t(\hat{z})^{1+\epsilon/\theta} \Upsilon'(0) Y_t D_t / \theta - w_t c_O & \text{Pre-policy,} \\ [\Upsilon(\hat{y}_t(\hat{z})) - \Upsilon'(\hat{y}_t(\hat{z})) \hat{y}_t(\hat{z})] Y_t D_t - w_t c_O & \text{Post-policy} \end{cases}$$

and the relative demand and relative price functions are in turn given by:

$$\begin{aligned} \hat{p}_t(\hat{z}) &= \begin{cases} \frac{(\theta/\epsilon) \exp(-\hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t}{W[(\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t]} & \text{Pre-policy,} \\ \exp(-\hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t & \text{Post-policy,} \end{cases} \\ \hat{y}_t(\hat{z}) &= [-\epsilon \max\{\ln(\hat{p}_t(\hat{z})), 0\}]^{\theta/\epsilon}. \end{aligned}$$

Let us now guess and verify that on a balanced growth path, the firm's value function

asymptotes to an endogenous constant as $\hat{z} \rightarrow \infty$:

$$\lim_{\hat{z} \rightarrow \infty} V_t(\hat{z}) = \bar{V}_t.$$

Taking the limit of the firm's HJBE and using Proposition 6, we obtain:

$$(\rho + \chi)\bar{V}_t = \lim_{\hat{z} \rightarrow \infty} \pi_t(\hat{z}) = \bar{\pi}Y_tD_t - w_t c_O$$

where the constant $\bar{\pi}$ is given by:

$$\bar{\pi} = \begin{cases} (\theta - 1) \exp[(1 - \theta)/\epsilon] \theta^{\theta/\epsilon - 1} & \text{Pre-policy,} \\ 1 + (\theta - 1) \exp(1/\epsilon) \epsilon^{\theta/\epsilon - 1} \Gamma\left(\frac{\theta}{\epsilon}, \frac{1}{\epsilon}\right) & \text{Post-policy} \end{cases}$$

which completes the proof. \square

Proof of Proposition 6. We want to prove that:

$$\lim_{\hat{z} \rightarrow \infty} \hat{p}_t(\hat{z}) = \exp(-\theta/\epsilon) \quad \text{and} \quad \lim_{\hat{z} \rightarrow \infty} \hat{y}_t(\hat{z}) = \theta^{\theta/\epsilon}.$$

Let us remember that a variety's relative price is given by:

$$\hat{p}_t(\hat{z}) = \frac{(\theta/\epsilon) \exp(-\hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t}{W[(\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t]}.$$

Defining the change of variable $x \equiv (\theta/\epsilon) \exp(\theta/\epsilon - \hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t$ and using l'Hôpital's rule together with the second property of the Lambert W-function, we find:

$$\lim_{\hat{z} \rightarrow \infty} \hat{p}_t(\hat{z}) = \lim_{x \rightarrow 0} \frac{x}{W(x) \exp(\theta/\epsilon)} = \exp(-\theta/\epsilon).$$

Clearly, since the relative demand function is given by:

$$\hat{y}(\hat{p}) = \begin{cases} [-\epsilon \ln(\hat{p})]^{\theta/\epsilon} & \text{if } \hat{p} < 1, \\ 0 & \text{otherwise} \end{cases}$$

we immediately obtain the result that $\lim_{\hat{z} \rightarrow \infty} \hat{y}_t(\hat{z}) = \theta^{\theta/\epsilon}$. \square

A.3 Extensions

Product Quality Improvements

In this section, we consider both productivity and quality improvements as the nature of an innovation. In particular, we denote a firm's quality relative to the lowest quality \underline{q}_t in the economy by $\hat{q} \in [0, \infty)$ and adopt the theoretical definition of a variety's quality proposed by [Baqae, Farhi and Sangani \(2023\)](#). That is, the quality and quantity of a product are assumed to be perfect substitutes.

Taking prices as given, the final sector's problem is to choose its relative demand for each variety to maximize profits in each period:

$$\begin{aligned} \max_{\{\hat{y}_t(\hat{z}, \hat{q})\}_{\hat{z}, \hat{q}=0}^{\infty}} & \left\{ P_t - M_t \int p_t(\hat{z}, \hat{q}) \hat{y}_t(\hat{z}, \hat{q}) dF_t(\hat{z}, \hat{q}) \right\} Y_t \\ \text{s.t.} & \quad M_t \int \Upsilon[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] dF_t(\hat{z}, \hat{q}) = \kappa. \end{aligned}$$

Reformulating the final sector's problem as a cost-minimization problem subject to the [Kimball \(1995\)](#) aggregator constraint using the Lagrangian, we have:

$$\mathcal{L}_t(\{\hat{y}_t(\hat{q})\}_{\hat{z}, \hat{q}=0}^{\infty}, \nu_t) = M_t \int p_t(\hat{z}, \hat{q}) \hat{y}_t(\hat{z}, \hat{q}) dF_t(\hat{z}, \hat{q}) + \nu_t \left(M_t \int \Upsilon[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] dF_t(\hat{z}, \hat{q}) - 1 \right)$$

where ν_t now denotes the Lagrange multiplier. The first-order conditions are:

$$p_t(\hat{z}, \hat{q}) = \nu_t \exp(\hat{q} + \underline{q}_t) \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] \quad \text{and} \quad M_t \int \Upsilon[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] dF_t(\hat{z}, \hat{q}) = \kappa.$$

Since the final sector is perfectly competitive and makes no profit, we have:

$$P_t = M_t \int p_t(\hat{z}, \hat{q}) \hat{y}_t(\hat{z}, \hat{q}) dF_t(\hat{z}, \hat{q}).$$

Substituting in the first-order conditions, we obtain a solution for ν_t :

$$\nu_t = P_t D_t \quad \text{where} \quad D_t \equiv \left(M_t \int \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] \exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q}) dF_t(\hat{z}, \hat{q}) \right)^{-1}.$$

This delivers the following inverse demand functions:

$$p_t(\hat{z}, \hat{q}) = \exp(\hat{q} + \underline{q}_t) \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] P_t D_t.$$

Firms engage in monopolistic competition in the product market but perfect competition in the input markets. That is, a firm chooses the price at which to sell its variety as

well as its demand for physical capital and production labor to maximize profits in each period. The firm takes as given the demand for its variety, the rental rate of capital r_t and the wage rate w_t , which delivers the following problem:

$$\pi_t(\hat{z}, \hat{q}) = \max_{p_t(\hat{z}, \hat{q}), k_t(\hat{z}, \hat{q}), l_t(\hat{z}, \hat{q})} \{p_t(\hat{z}, \hat{q})y_t(\hat{z}, \hat{q}) - (r_t + \delta)k_t(\hat{z}, \hat{q}) - w_t l_t(\hat{z}, \hat{q})\} - w_t c_O$$

subject to the inverse demand function $p_t(\hat{z}, \hat{q}) = \exp(\hat{q} + \underline{q}_t) \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] D_t$. Let us first consider the sub-problem of optimally choosing the demand for capital and labor, which can be reformulated as a cost-minimization problem. Using the Lagrangian:

$$\mathcal{L}_t(k_t(\hat{z}, \hat{q}), l_t(\hat{z}, \hat{q}), v_t) = (r_t + \delta)k_t(\hat{z}, \hat{q}) + w_t l_t(\hat{z}, \hat{q}) + v_t[y_t(\hat{z}, \hat{q}) - \exp(\hat{z} + \underline{z}_t) k_t(\hat{z}, \hat{q})^\alpha l_t(\hat{z}, \hat{q})^{1-\alpha}]$$

where v_t denotes the Lagrange multiplier. The first-order conditions are:

$$\begin{aligned} k_t(\hat{z}, \hat{q}) &= \frac{\alpha v_t y_t(\hat{z}, \hat{q})}{r_t + \delta}, \\ l_t(\hat{z}, \hat{q}) &= \frac{(1 - \alpha) v_t y_t(\hat{z}, \hat{q})}{w_t}, \\ y_t(\hat{z}, \hat{q}) &= \exp(\hat{z} + \underline{z}_t) k_t(\hat{z}, \hat{q})^\alpha l_t(\hat{z}, \hat{q})^{1-\alpha}. \end{aligned}$$

Solving for the Lagrange multiplier, we have:

$$v_t = \varsigma_t \exp(-\hat{z} - \underline{z}_t) \quad \text{where} \quad \varsigma_t \equiv \left(\frac{r_t + \delta}{\alpha} \right)^\alpha \left(\frac{w_t}{1 - \alpha} \right)^{1-\alpha}.$$

Therefore, we can rewrite the firm's static problem as:

$$\begin{aligned} \pi_t(\hat{z}, \hat{q}) &= \max_{p_t(\hat{z}, \hat{q})} \{[p_t(\hat{z}, \hat{q}) - \varsigma_t \exp(-\hat{z} - \underline{z}_t)] \hat{y}_t(\hat{z}, \hat{q})\} Y_t - w_t c_O \\ \text{s.t.} \quad p_t(\hat{z}, \hat{q}) &= \exp(\hat{q} + \underline{q}_t) \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] D_t. \end{aligned}$$

Reformulating it as a choice of $\hat{y}_t(\hat{z}, \hat{q})$ given the inverse demand function $p_t(\hat{z}, \hat{q})$:

$$\pi_t(\hat{z}, \hat{q}) = \max_{\hat{y}_t(\hat{z}, \hat{q})} \{[\exp(\hat{q} + \underline{q}_t) \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] D_t - \varsigma_t \exp(-\hat{z} - \underline{z}_t)] \hat{y}_t(\hat{z}, \hat{q})\} Y_t - w_t c_O.$$

The first-order condition is:

$$\{\Upsilon''[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] \exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q}) + \Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})]\} \exp(\hat{q} + \underline{q}_t) D_t = \varsigma_t \exp(-\hat{z} - \underline{z}_t).$$

We can rearrange this expression as:

$$p_t(\hat{z}, \hat{q}) = \frac{\mu_t(\hat{z}, \hat{q}) \zeta_t}{\exp(\hat{z} + \underline{z}_t)} \quad \text{where} \quad \mu_t(\hat{z}, \hat{q}) \equiv \frac{\vartheta_t(\hat{z}, \hat{q})}{\vartheta_t(\hat{z}, \hat{q}) - 1}$$

and where $\vartheta_t(\hat{z}, \hat{q})$ denotes the price elasticity of demand:

$$\vartheta_t(\hat{z}, \hat{q}) \equiv - \frac{\Upsilon'[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})]}{\Upsilon''[\exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})] \exp(\hat{q} + \underline{q}_t) \hat{y}_t(\hat{z}, \hat{q})} \in (1, \infty).$$

Substituting the monopoly pricing function in the profit function, we have:

$$\pi_t(\hat{z}, \hat{q}) = \frac{p_t(\hat{z}, \hat{q}) \hat{y}_t(\hat{z}, \hat{q}) Y_t}{\vartheta_t(\hat{z}, \hat{q})} - w_t c_O.$$

Denoting the quality-specific choke price by $\bar{p}_t(\hat{q}) \equiv \exp(\hat{q} + \underline{q}_t) \Upsilon'(0) D_t$, we can define a variety's relative price as $\hat{p}_t(\hat{z}, \hat{q}) \equiv p_t(\hat{z}, \hat{q}) / \bar{p}_t(\hat{q})$. This allows us to rewrite the relative demand function as:

$$\hat{y}_t(\hat{z}, \hat{q}) = \Upsilon'^{-1}[\hat{p}_t(\hat{z}, \hat{q}) \Upsilon'(0)] / \exp(\hat{q} + \underline{q}_t).$$

Using the **Klenow and Willis (2016)** specification of the **Kimball (1995)** aggregator, we can rewrite the monopoly pricing condition (relative to the choke price) as:

$$\hat{p}_t(\hat{z}, \hat{q}) = \frac{\zeta_t}{[1 + (\epsilon/\theta) \ln(\hat{p}_t(\hat{z}, \hat{q}))] \exp(\hat{z} + \underline{z}_t) \bar{p}_t(\hat{q})}$$

Note that we can rearrange this equation to obtain:

$$\exp\{(\theta/\epsilon)[\zeta_t[\hat{p}_t(\hat{z}, \hat{q}) \exp(\hat{z} + \underline{z}_t) \bar{p}_t(\hat{q})]^{-1} - 1]\} \hat{p}_t(\hat{z}, \hat{q})^{-1} = 1$$

Multiplying both sides by $(\theta/\epsilon) \exp(\theta/\epsilon) \zeta_t [\exp(\hat{z} + \underline{z}_t) \bar{p}_t(\hat{q})]^{-1}$, we have:

$$W^{-1}\{(\theta/\epsilon) \zeta_t [\hat{p}_t(\hat{z}, \hat{q}) \exp(\hat{z} + \underline{z}_t) \bar{p}_t(\hat{q})]^{-1}\} = (\theta/\epsilon) \exp(\theta/\epsilon) \zeta_t [\exp(\hat{z} + \underline{z}_t) \bar{p}_t(\hat{q})]^{-1}.$$

Finally, solving for $\hat{p}_t(\hat{z}, \hat{q})$ delivers:

$$\hat{p}_t(\hat{z}, \hat{q}) = \frac{(\theta/\epsilon) \exp(-\hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t(\hat{q})}{W[(\theta/\epsilon) \exp(\theta/\epsilon) \exp(-\hat{z} - \underline{z}_t) \zeta_t / \bar{p}_t(\hat{q})]}.$$

Substituting in the expression for the quality-specific choke price and defining the firm's

composite state variable as $x_t \equiv z_t + q_t$, we have:

$$\hat{p}_t(\hat{x}) = \frac{(\theta/\epsilon) \exp(-\hat{x} - \underline{x}_t) \zeta_t / \bar{p}_t}{W[(\theta/\epsilon) \exp(\theta/\epsilon) \exp(-\hat{x} - \underline{x}_t) \zeta_t / \bar{p}_t]} \quad \text{where} \quad \bar{p}_t \equiv \Upsilon'(0) D_t$$

which is isomorphic to our framework. Therefore, profits are given by:

$$\pi_t(\hat{x}) = \frac{\hat{p}_t(\hat{x}) \Upsilon'^{-1}[\hat{p}_t(\hat{x}) \Upsilon'(0)] \bar{p}_t Y_t}{\vartheta_t(\hat{x})} - w_t c_O$$

which is also isomorphic to our framework. Using the above expressions, the firm's markup is given by the following function:

$$\mu_t(\hat{x}) = \frac{\theta/\epsilon}{W[(\theta/\epsilon) \exp(\theta/\epsilon) \exp(-\hat{x} - \underline{x}_t) \zeta_t / \bar{p}_t]}.$$

Inverting this function, we have:

$$\hat{x}_t(\mu) = \ln(\mu) + (\theta/\epsilon)(1 - \mu^{-1}) + \ln(\zeta_t / \bar{p}_t) - \underline{x}_t.$$

B Numerical Appendix

This section of the Appendix provides details on the numerical strategies we use to solve and quantify the model.

B.1 Spectral Collocation and Quadrature

Following [Miranda and Fackler \(2004\)](#), we approximate the solutions of the HJBE and KFE using spectral collocation in the spatial dimension on the interval $\hat{z} \in [0, \infty)$. But first, we consider a change of variable to approximate the solutions of the HJBE and KFE on the unit line:

$$V_t(\hat{z}) = \mathcal{V}_t(\tilde{z}) \quad \text{and} \quad F_t(\hat{z}) = \mathcal{F}_t(\tilde{z}) \quad \text{where} \quad \tilde{z} \equiv \frac{\nu \hat{z}}{\nu \hat{z} + 1}.$$

Here, $\nu > 0$ is a parameter governing the curvature of the change of variable. With this change of variable, we can use the chain rule to obtain:

$$\begin{aligned} V_t'(\hat{z}) &= \nu(1 - \tilde{z})^2 \mathcal{V}'(\tilde{z}), \\ V_t''(\hat{z}) &= \nu^2(1 - \tilde{z})^3 [(1 - \tilde{z}) \mathcal{V}_t''(\tilde{z}) - 2\mathcal{V}_t'(\tilde{z})], \\ F_t'(\hat{z}) &= \nu(1 - \tilde{z})^2 \mathcal{F}_t'(\tilde{z}), \\ F_t''(\hat{z}) &= \nu^2(1 - \tilde{z})^3 [(1 - \tilde{z}) \mathcal{F}_t''(\tilde{z}) - 2\mathcal{F}_t'(\tilde{z})]. \end{aligned}$$

Under this reformulation, the boundary conditions of $\mathcal{V}_t(\tilde{z})$ and $\mathcal{F}_t(\tilde{z})$ are:

$$\mathcal{V}_t(0) = \mathcal{V}_t'(0) = \mathcal{F}_t(0) = 0 \quad \text{and} \quad \mathcal{F}_t(1) = 1.$$

We approximate the functions $\mathcal{V}_t(\hat{z})$ and $\mathcal{F}_t(\hat{z})$ over $n - 2$ Chebyshev nodes $\{\tilde{z}_i\}_{i=2}^{n-1}$ on the unit line to which we append the boundaries $\tilde{z}_1 = 0$ and $\tilde{z}_n = 1$. The approximation is a linear combination of Chebyshev basis functions $\{b_j(\tilde{z})\}_{j=1}^n$ of degree n whose time-varying coefficients $\{c_j^V(t), c_j^F(t)\}_{j=1}^n$ are to be determined:

$$\mathcal{V}_t(\tilde{z}_i) \approx \sum_{j=1}^n c_j^V(t) b_j(\tilde{z}_i) \quad \text{and} \quad \mathcal{F}_t(\tilde{z}_i) \approx \sum_{j=1}^n c_j^F(t) b_j(\tilde{z}_i).$$

In particular, the coefficients $\{c_j^V(t), c_j^F(t)\}_{j=1}^n$ are chosen to satisfy the HJBE and KFE over the nodes $\{\tilde{z}_i\}_{i=2}^{n-1}$ as well as the boundary conditions of $\mathcal{V}_t(\hat{z})$ and $\mathcal{F}_t(\hat{z})$ at $\tilde{z}_1 = 0$ and $\tilde{z}_n = 1$. In addition, the value of the free boundary \underline{z}_t is chosen to satisfy the smooth pasting condition $\mathcal{V}_t'(0) = 0$. However, the solutions of the HJBE and KFE depend on unknown endogenous economic variables that are constrained by equilibrium conditions. These equilibrium conditions involve integrals which we approximate using Chebyshev–Gauss quadrature over the same set of nodes $\{\tilde{z}_i\}_{i=2}^{n-1}$ as above. That is, for an arbitrary function f , we approximate the following integral:

$$\int_0^\infty f_t(\hat{z}) d\hat{z} = \int_0^1 \frac{f_t(\hat{z}(\tilde{z}))}{\nu(1 - \tilde{z})^2} d\tilde{z} \approx \sum_{i=2}^{n-1} \frac{f_t(\hat{z}(\tilde{z}_i))}{\nu(1 - \tilde{z}_i)^2} \omega_i \quad \text{where} \quad \hat{z}(\tilde{z}) \equiv \frac{\tilde{z}}{\nu(1 - \tilde{z})}$$

and where $\{\omega_i\}_{i=2}^{n-1}$ denote the Chebyshev–Gauss quadrature weights.

B.2 MPEC-BGP Estimation Strategy

To estimate the structural parameters of our theory, solving for a balanced growth path equilibrium at multiple parameter guesses is computationally expensive. An alternative

approach discussed in [Su and Judd \(2012\)](#) and [Dubé et al. \(2012\)](#) is to reformulate the estimation problem as a mathematical program with equilibrium constraints (MPEC).

Instead of solving for a balanced growth path equilibrium allocation at multiple guesses of these parameters, our approach is to treat the model's equilibrium conditions as constraints in the optimization problem. That is, we search for parameters as well as endogenous economic variables to minimize a GMM objective subject to the constraints that the model's equilibrium conditions are met.

More formally, let us describe the balanced growth path equilibrium allocation of our model as an $N_x \times 1$ vector of endogenous variables \mathbf{x} that depend on an $N_\Omega \times 1$ vector of parameters Ω through the $N_x \times 1$ vector of equilibrium conditions:

$$h(\mathbf{x}, \Omega) = \mathbf{0}.$$

We let $X(\Omega)$ denote the set of all \mathbf{x} such that $h(\mathbf{x}, \Omega) = \mathbf{0}$:

$$X(\Omega) := \{\mathbf{x} : h(\mathbf{x}, \Omega) = \mathbf{0}\}.$$

For some weighting $N_m \times N_m$ matrix \mathbf{W} and an $N_m \times 1$ vector of moments $m(\mathbf{x}, \Omega)$, define the GMM estimator as the vector Ω^* that solves the problem:

$$\Omega^* = \arg \min_{\Omega} \left\{ \min_{\mathbf{x} \in X(\Omega)} m(\mathbf{x}, \Omega)^\top \mathbf{W} m(\mathbf{x}, \Omega) \right\}.$$

Denote by $\mathbf{x}^*(\Omega^*)$ the optimal solution of endogenous variables for this problem. This approach is particularly powerful when the functions $h(\mathbf{x}, \Omega)$ and $m(\mathbf{x}, \Omega)$ are both twice differentiable in their arguments since we can exploit their Jacobian and Hessian to efficiently find a solution. We implement this using the commercial nonlinear solver KNITRO ([Byrd, Nocedal and Waltz, 2006](#)) as well as the open source nonlinear solver IPOPT ([Wächter and Biegler, 2006](#)) through the interface of JuMP ([Lubin, Dowson, Dias Garcia, Huchette, Legat and Vielma, 2023](#)), a modeling language for mathematical optimization embedded in [Julia](#).

B.2.1 Identification

[Andrews, Gentzkow and Shapiro \(2017\)](#) suggest that researchers report the sensitivity of their parameter estimates to moment conditions. This sensitivity matrix is denoted by Λ and defined as:

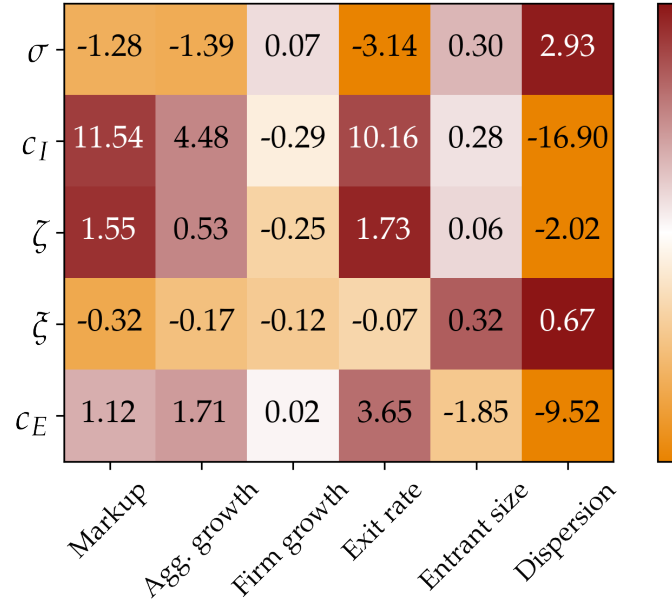
$$\Lambda = -[\mathbf{G}(\mathbf{x}^*(\Omega^*), \Omega^*)^\top \mathbf{W} \mathbf{G}(\mathbf{x}^*(\Omega^*), \Omega^*)]^{-1} \mathbf{G}(\mathbf{x}^*(\Omega^*), \Omega^*)^\top \mathbf{W}$$

where $\mathbf{G}(\mathbf{x}^*(\boldsymbol{\Omega}^*), \boldsymbol{\Omega}^*)$ is the $N_m \times N_\Omega$ Jacobian of $m(\mathbf{x}, \boldsymbol{\Omega})$ with respect to parameters, evaluated at $\mathbf{x}^*(\boldsymbol{\Omega}^*)$ and $\boldsymbol{\Omega}^*$. However, to calculate this Jacobian, we must account for the dependency between endogenous variables and parameters through the equilibrium conditions $h(\mathbf{x}^*(\boldsymbol{\Omega}^*), \boldsymbol{\Omega}^*) = \mathbf{0}$. As such, we can define that Jacobian using the implicit function theorem:

$$\mathbf{G}(\mathbf{x}^*(\boldsymbol{\Omega}^*), \boldsymbol{\Omega}^*) := \nabla_{\boldsymbol{\Omega}} m(\mathbf{x}^*, \boldsymbol{\Omega}^*) - \nabla_{\mathbf{x}} m(\mathbf{x}^*, \boldsymbol{\Omega}^*) [\nabla_{\mathbf{x}} h(\mathbf{x}^*, \boldsymbol{\Omega}^*)]^{-1} \nabla_{\boldsymbol{\Omega}} h(\mathbf{x}^*, \boldsymbol{\Omega}^*).$$

Figures B.7 and B.8 respectively plot the matrices $\boldsymbol{\Lambda}$ and $\mathbf{G}(\mathbf{x}^*(\boldsymbol{\Omega}^*), \boldsymbol{\Omega}^*)$ (in elasticity form) for the five jointly estimated parameters and the six moment conditions presented in Section 4.2.

Figure B.7: Sensitivity Matrix

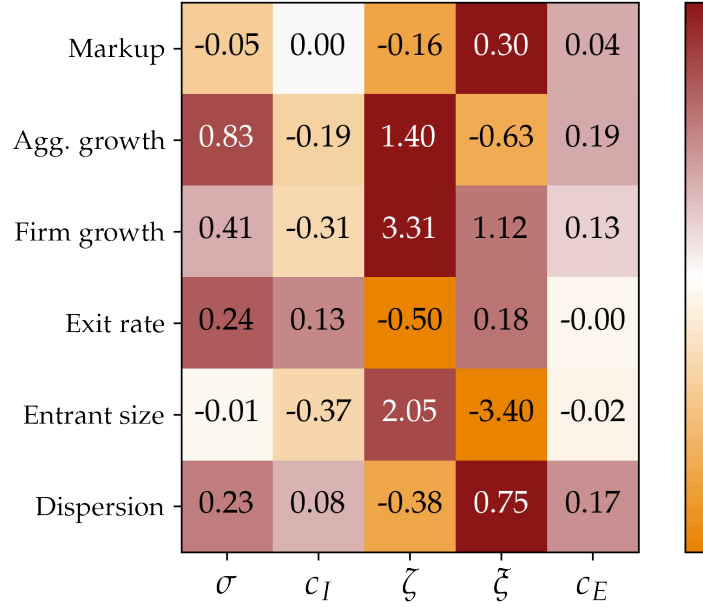


Note: This heat map plots the sensitivity matrix $\boldsymbol{\Lambda}$ of parameters with respect to moment conditions. The color palette is normalized for each row separately to ease visualization and the matrix is presented in elasticity form.

A first observation is that the standard deviation of the Brownian motion σ is mostly identified by the dispersion of value added and the overall exit rate. It is perhaps no surprise that this parameter is sensitive to the dispersion in value-added, since it determines the volatility of the productivity process. But this volatility also influences the rate at which unproductive firms are swept below the exit threshold, explaining why σ is sensitive to the exit rate. Further, this parameter is also sensitive to the aggregate markup—since it determines the shape of the relative productivity distribution over

which firm-level markups are aggregated—and to the aggregate growth rate, which is directly reliant on this parameter as noted in Proposition 3.

Figure B.8: Jacobian Matrix



Note: This heat map plots the Jacobian matrix $\mathbf{G}(\mathbf{x}^*(\mathbf{\Omega}^*), \mathbf{\Omega}^*)$ of moment conditions with respect to parameters. The color palette is normalized for each row separately to ease visualization and the matrix is presented in elasticity form.

Similarly, the innovation cost scale and elasticity parameters c_I and ζ are mainly identified by these same four moments. The intuition underlying these relationships is as follows: to decide how much to invest in R&D, the firm compares the marginal value to the marginal cost of such investments. The parameters c_I and ζ influence the scale and shape of that marginal cost function, which thus dictate the firm's productivity drift function and in turn the shape of the relative productivity distribution. Since these four moments are explicitly dependent on this stationary distribution, it is consistent with our understanding that they identify c_I and ζ .

The parameter $\tilde{\zeta}$ determines the transformation of the incumbent distribution from which entrants draw their relative productivity. Therefore, it is not surprising that this parameter is sensitive to the relative size of entrants. However, it is also sensitive to the dispersion of value-added since it acts as a compressing force on the variance of the relative productivity distribution. Therefore, and by the same intuition mentioned above, this parameter is also sensitive to the aggregate markup.

Finally, the entry cost parameter c_E is primarily identified by the dispersion in value-

added, which is intricately tied to the shape of the relative productivity distribution. But the shape of that distribution is in large part determined by firms' investments in R&D who compete against one another for labor. The intensity of that competition, in turn, is dictated by the endogenous measure of varieties, which depends on the flow of entrants and, consequently, on the entry cost parameter itself.

C Empirical Appendix

This section of the Appendix provides details on the construction of variables of interest from our main source of data and on the estimation of our theory's structural parameters.

C.1 Data

Our main source of data is the *Fichier Approché des Résultats d'Esane* (FARE). This is an annual panel dataset, covering the period 2009–2019, with the balance sheet and income statements for the universe of firms in France that are subject to the standard corporate tax (excluding the financial and farming sectors).

Given that this dataset is compiled from tax declarations, the unit of observation is a legal entity (*unité légale*), each identified by a unique Siren number. Recognizing that this does not correspond to what users of the data would call a firm, the **National Institute of Statistics and Economic Studies** (INSEE) developed definitions of consolidated firms (*entreprises profilées*): a collection of legal entities that are part of the same group identified by a unique Sirus number. We use the *Contour des Entreprises Profilées* from 2019 to define the boundaries of the firms in our sample.

The main variables of interest are the firm's industry of operation, value-added, wage bill, and capital stock.

- The main industry of operation for the firm is a 5-digit industry from the **NAF** classification. For consolidated firms, this is provided in the Contours files.
- Annual value added excludes VAT and is calculated as gross output (sum of sales and the gross value of stored production) net of expenditures on intermediate inputs, materials, and other external expenses, as well as changes in the stock of intermediate inputs and materials.
- The annual wage bill includes all labor costs for the firm and is obtained by summing expenditures on salaries (including bonuses) and social security payments.

- Capital is the sum of tangible capital, inventories, and rental and lease payments. We calculate the book value of tangible capital as the gross acquisition value net of accumulated depreciation. To convert to current prices, we multiply this book value by the ratio of the aggregate price index for gross fixed capital formation in the current year to its value in the acquisition year.⁴¹ For every firm-by-year observation, we have an estimate of the acquisition year of capital since, assuming a constant depreciation rate, the ratio of accumulated depreciation to gross acquisition value can be used to recover an average age for the firm's capital stock.⁴² Formally, denoting by k_{jit}^T the stock of tangible capital of firm j in industry i in year t , k_{jit}^I its inventories and R_{jit} its lease and rental payments, we define:

$$k_{jit} = k_{jit}^T + k_{jit}^I + \frac{R_{jit}}{r_t + \delta} \quad \text{s.t.} \quad (r_t + \delta)k_{jit} = (r_t + \delta)(k_{jit}^T + k_{jit}^I) + R_{jit}.$$

We obtain firm-level value-added, wage bill and capital stock by summing each of these variables over the different legal units that constitute a firm.

We keep in our sample private businesses with a regular taxation scheme and drop any firm with a negative value for value-added, wage bill, stock of tangible capital, inventories, or rental and lease payments. Additionally, we winsorize each of these variables at the 1% level at the 2-digit industry by year level. With these selection criteria, we end up with a sample of 5,423,743 (firm-year) observations between 2009 and 2019, with 831,297 unique firms overall. Table C.11 presents summary statistics for the main variables of interest in our sample.

Table C.11: Summary Statistics

Variable	Mean	5th %ile	25th %ile	Median	75th %ile	95th %ile
Value Added	938	26	116	265	584	2790
Wage bill	700	18	94	214	467	2149
Tangible capital + inventories	823	1	19	77	256	1987

Note: Units are in thousands of current Euros.

⁴¹This price index is provided by the INSEE.

⁴²We use a depreciation rate of 10% for the tangible capital stock, consistent with French accounting standards.

C.2 Structural Estimation

The Klenow and Willis (2016) Parameters

A key parameter to discipline in our theory is the ratio of ϵ and θ . This ratio measures how steeply markups increase in firms' productivity and therefore partly determines the degree of dispersion in markups. If both markups and market shares are observed at the firm level, one can use the demand functions for varieties to estimate this ratio.⁴³ In particular, let us consider a generalization of the [Kimball \(1995\)](#) aggregator:

$$\sum_{i=1}^I x_{it} \int_{j \in \mathcal{J}_t} q_{ji} \Upsilon(\hat{y}_{jit}) dj = 1$$

where x_{it} is an industry-time-specific demand shifter and q_{ji} denotes the unobserved time-invariant “quality” of variety j . This demand system implies the following inverse demand functions in logarithms:

$$\ln(p_{jit}) = \ln(x_{it}) + \ln(q_{ji}) + \ln(D_t) + \ln(\Upsilon'(\hat{y}_{jit}))$$

where D_t is the time-varying demand index and the function $\Upsilon'(\hat{y})$ is given by:

$$\Upsilon'(\hat{y}) = \left(\frac{\theta - 1}{\theta} \right) \exp \left(\frac{1 - \hat{y}^{\epsilon/\theta}}{\epsilon} \right).$$

Adding $\ln(\hat{y}_{jit})$ to both sides of this equation, and denoting the market share of firm j from industry i at time t by s_{jit} , we obtain:

$$\ln(s_{jit}) = \ln(x_{it}) + \ln(q_{ji}) + \ln(D_t) + \ln(\Upsilon'(\hat{y}_{jit})) + \ln(\hat{y}_{jit}).$$

Since markups are related to relative demand as:

$$\mu_{jit}^{-1} = 1 - \hat{y}_{jit}^{\epsilon/\theta} / \theta$$

we can rewrite the previous equation as:

$$\mu_{jit}^{-1} + \ln(1 - \mu_{jit}^{-1}) = \psi + (\epsilon/\theta)[\ln(s_{jit}) - \ln(x_{it}) - \ln(q_{ji}) - \ln(D_t)]$$

where the constant is given by $\psi \equiv \frac{\theta-1}{\theta} - (\epsilon/\theta) \ln \left(\frac{\theta-1}{\theta} \right)$. Hence, regressing the nonlinear transformation $\mu_{jit}^{-1} + \ln(1 - \mu_{jit}^{-1})$ of firm-level markups on firm-level market shares,

⁴³We calculate μ_{jit} and market share as described in [Section 4.1](#)

controlling for industry, industry-time, time and firm fixed effects delivers a consistent estimate of θ/ϵ .

Calculating the Targeted Moments

Our GMM estimation strategy uses three moments calculated from the FARE data:

- Within-industry standard deviation of log value added.
- Average annual growth of firm-level value-added, deflated with the GDP deflator.
- Relative size of entrants, calculated as the ratio of the average value added of an incumbent to the average value added of an entrant. In calculating this moment, we treat any firm as an entrant if it has been created within the past five years, and an incumbent otherwise.

We calculate each of these moments at the 2-digit industry-by-year level. We then aggregate by year, weighting each industry by its share of value added in that year. Finally, we take a simple average of each of these moments across the different years in our sample.

C.3 Additional Tables

In this section of the Appendix, we present tables for the alternative assumptions and policy counterfactuals discussed and considered in the main text. In particular, we replicate Tables 3, 4, 5 and 6.

Table C.12: Markups and Market Shares

Dependent variable: $\mu_{jit}^{-1} + \ln(1 - \mu_{jit}^{-1})$						
	Unadjusted capital share			Adjusted capital share		
$\ln(s_{jit})$	0.047 (0.000)	0.234 (0.001)	0.243 (0.001)	0.048 (0.000)	0.224 (0.001)	0.231 (0.001)
Firm fixed effects		Y	Y		Y	Y
Industry \times year fixed effects	Y	Y	Y	Y	Y	Y
Industry fixed effects			Y			Y
Year fixed effects			Y			Y
Age group fixed effects	Y		Y	Y		Y
R^2	0.090	0.505	0.507	0.087	0.523	0.524
Observations	4.9M	4.9M	4.9M	5M	5M	5M

Note: Firm-level markups and market shares are constructed from the FARE dataset as described in Section 4.1. This table presents different regression specifications with firm fixed effects, 5-digit NACE industry fixed effects and age group fixed effects (for a total of 20 evenly-spaced age groups). Standard errors (in parentheses) are clustered at the firm level. The total number of observations is below the total sample size of 5.4M because negative markups were estimated for some firms. In the columns labeled “adjusted capital share”, the industry-specific capital cost shares measured in the data are inflated by a constant such that the aggregate capital cost share is equal to 1/3, consistent with our calibrated model.

Table C.13: Economic Aggregates for Fixed Innovation Labor

Aggregate	Before	After	Change
<i>Labor allocations:</i>			
Labor supply	0.264	0.319	+20.8%
Production labor	0.227	0.263	+16.1%
Innovation labor	0.020	0.020	0.0%
Entry labor	0.015	0.034	+122.2%
Overhead labor	0.002	0.001	-33.2%
<i>Firms, entry and exit:</i>			
Measure of varieties	0.044	0.029	-33.2%
Entry rate	5.32%	17.68%	+12.4p.p.
Endogenous exit rate	3.97%	16.33%	+12.4p.p.

Note: This table presents the pre- and post-policy level of various economic aggregates as well as the corresponding percentage change when fixing the aggregate allocation of labor to innovation to its initial level before the implementation of the policy intervention. Doing so requires imposing a uniform tax of 44.5% on firms' expenditures on R&D.

Table C.14: Economic Aggregates for Alternative Transfer Schedules

Aggregate	Before	Level fix		Dispersion fix	
		After	Change	After	Change
<i>Labor allocations:</i>					
Labor supply	0.264	0.304	+15.3%	0.284	+7.8%
Production labor	0.227	0.271	+19.5%	0.214	-5.7%
Innovation labor	0.020	0.017	-12.9%	0.039	+93.4%
Entry labor	0.015	0.014	-8.9%	0.031	+100.2%
Overhead labor	0.002	0.001	-1.7%	0.001	-40.5%
<i>Firms, entry and exit:</i>					
Measure of varieties	0.044	0.043	-1.7%	0.026	-40.5%
Entry rate	5.32%	4.93%	-0.4p.p.	17.89%	+12.6p.p.
Endogenous exit rate	3.97%	3.59%	-0.4p.p.	16.54%	+12.6p.p.

Note: This table presents the pre- and post-policy level of various economic aggregates as well as the corresponding percentage change under alternative policy interventions. Specifically, the columns labeled "Baseline", "Level fix" and "Dispersion fix" refer to the transfers that rectify both, and either the level or dispersion in markups, respectively.

Table C.15: Structural Parameters for Alternative Aggregate Markup Targets

Parameter	Symbol	Value $\mathcal{M}=1.1$	Value $\mathcal{M}=1.5$
<i>Household preferences:</i>			
Rate of time preference	ρ	0.04	0.04
Labor supply utility weight	β	11.6	9.3
Frisch elasticity of labor supply reciprocal	η	1	1
<i>Final sector technology:</i>			
Klenow and Willis (2016) elasticity param.	θ	54.5	6.8
Klenow and Willis (2016) super-elasticity param.	ϵ	26.78	1.53
Kimball (1995) aggregation constant	κ	0.65	0.88
<i>Firm production technology:</i>			
Output elasticity of physical capital	α	0.33	0.33
Depreciation rate of physical capital	δ	0.06	0.06
Overhead cost parameter	c_O	0.05	0.03
<i>Firm innovation technology:</i>			
Brownian motion standard deviation	σ	0.03	0.05
Innovation cost scale parameter	c_I	11.04	4.83
Innovation cost elasticity parameter	ζ	1.10	0.48
<i>Entry and exit:</i>			
Entry cost parameter	c_E	1.15	3.49
Entry distribution parameter	ξ	2.03	2.39
Exogenous exit rate	χ	1.34%	1.34%

Note: This table presents the assigned/estimated structural parameters of our theory under alternative aggregate markup targets. Specifically, the columns labeled “ $\mathcal{M}=1.1$ ” and “ $\mathcal{M}=1.5$ ” respectively refer to parameterizations that target a cost-weighted average markup of 1.1 and 1.5.

Table C.16: Moments for Alternative Aggregate Markup Targets

Moment	Model	Model	Data
	$\mathcal{M}=1.1$	$\mathcal{M}=1.5$	
Cost-weighted average markup	1.10	1.50	
GDP per hour worked growth rate	1.15%	1.16%	1.16%
Incumbent value added growth rate	1.24%	1.24%	1.24%
Exit rate of all firms	10.24%	3.62%	5.61%
Relative size of entrants by value added	0.37	0.26	0.31
Standard deviation of log value added	1.54	1.42	1.54

Note: This table presents targeted moments and their resulting value in our model under alternative aggregate markup targets. Specifically, the columns labeled “ $\mathcal{M}=1.1$ ” and “ $\mathcal{M}=1.5$ ” respectively refer to parameterizations that target a cost-weighted average markup of 1.1 and 1.5.

Table C.17: Economic Aggregates for Alternative Aggregate Markup Targets

Aggregate	$\mathcal{M}=1.1$			$\mathcal{M}=1.5$		
	Before	After	Change	Before	After	Change
<i>Labor allocations:</i>						
Labor supply	0.264	0.285	+7.9%	0.264	0.368	+39.6%
Production labor	0.245	0.258	+5.3%	0.211	0.256	+21.6%
Innovation labor	0.005	0.016	+191.3%	0.034	0.079	+134.2%
Entry labor	0.053	0.047	-10.3%	0.029	0.058	+101.4%
Overhead labor	0.003	0.002	-43.5%	0.004	0.002	-45.7%
<i>Firms, entry and exit:</i>						
Measure of varieties	0.078	0.044	-43.5%	0.119	0.065	-45.7%
Entry rate	10.24%	16.27%	+6.0p.p.	3.62%	13.43%	+9.8p.p.
Endogenous exit rate	8.90%	14.93%	+6.0p.p.	2.28%	12.09%	+9.8p.p.

Note: This table presents the pre- and post-policy level of various economic aggregates as well as the corresponding percentage change under alternative aggregate markup targets. Specifically, the columns labeled “ $\mathcal{M}=1.1$ ” and “ $\mathcal{M}=1.5$ ” respectively refer to parameterizations that target a cost-weighted average markup of 1.1 and 1.5.